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# Towards dynamic quantile function models for anomaly detection

PRESENTED BY

Peter Jacobs and **Lekha Patel**

*Joint work with Anirban Bhattacharya and Debdeep Pati (TAMU)*



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- 1 Introduction: Problem Definition, Motivating Applications, and Prior Work
- 2 Our Approach: General Framework, Constraints, and Current Model
- 3 Conclusions



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We observe univariate  $X_1 \sim F_1, X_2 \sim F_2, \dots$  in a stream

$$F_1, F_2, \dots$$

vary through time. At each time  $t \geq 1$ , we want to estimate the

$$0 < q_1 < q_2 < \dots < q_K < 1$$

quantiles of  $F_t$

# Motivating Applications



## 5 Motivating Applications



### Detecting Malicious Activity in a Stream of Computer Network Data

- Multivariate points in the stream of network data  $\mathbf{X}_1, \mathbf{X}_2, \dots$  are converted via feature engineering to a discriminative 1d stream  $X_1, X_2, \dots$  as in (Barata, 2021)
- Raises in the .9 quantile without changes in the .85 quantile could indicate a small group of machines with a common characteristic have become infected

Quantile tracking also

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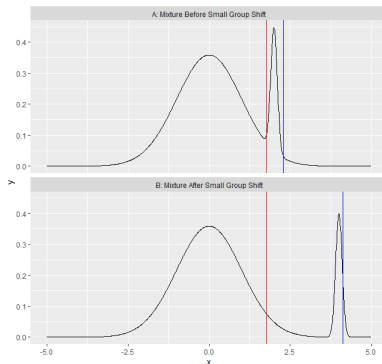


Figure: Red line is .86 quantile. Blue line is .99 quantile



# Motivating Applications







## Monitoring for Oil Price Shocks over Days

- Changes in the .9 quantile of world oil prices without changes in the .99 quantile indicates that oil producing nations that had expensive prices became more expensive but the countries with the most costly oil still had the most costly oil
- Useful for quantile quantile regression as in Barata (2021)



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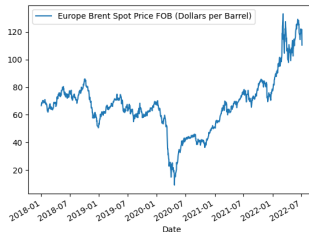


Figure: Time Series of BRENT Crude oil Dollars Per Barrel

## Prior Work



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- Consistency in iid setting is proved: If  $F_1 = F_2 = \dots = F$ , then in probability if  $Q_q = F^{-1}(q) > 0$ , and  $Q_q(0) > 0$

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- In the non-dynamic setting  $F = F_1 = F_2 = \dots$ , frequentist approaches have worked out asymptotically valid interval estimation for a wide range of distributions Barata (2021), Barata (2021), Barata (2021)
  - These methods are inherently unable to adapt to change because they weight each member of the sample equally





- Dynamic credible interval construction for a single quantile is also well studied Barata (2021)
- This work suffers from a stochastic ordering violation
  - The model can be fit to 2 quantiles separately but if estimating quantiles  $q_1$  and  $q_2$  where  $q_1 < q_2$  it is possible there will be a  $c \in \mathbb{R}$  s.t  $P(Q_{1t} > c | X_1, \dots, X_t) > P(Q_{2t} > c | X_1, \dots, X_t)$
  - This type of contradiction is likely to occur when estimating extreme low quantiles



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# General Framework: HMM



We will approach the dynamic many quantile estimation problem using a hidden markov model structure as well. Recall we seek to estimate the  $0 < q_1 < \dots < q_K < 1$  quantiles of  $F_t$  for each  $t$ . We propose the following model

## Initial Distribution

$$\mathbf{Q}_0 = (Q_{10}, \dots, Q_{K0}) \sim G_0$$

where  $G_0$  is some to be determined parametric family of distributions on  $\{(x_1, \dots, x_K) \mid -\infty < x_1 < x_2 < \dots < x_K < \infty\}$ .

And for  $t \geq 1$

## Transition Distribution

$$\mathbf{Q}_t | \mathbf{Q}_{t-1} \sim G(\mathbf{q}_{t-1}, \mathbf{V}_t)$$

where  $G$  is some to be determined parametric family of distributions on  $\{(x_1, \dots, x_K) \mid -\infty < x_1 < x_2 < \dots < x_K < \infty\}$  parameterized by the underlying quantile vector at time  $t - 1$  and a vector of parameters dictating the variance of transition at time



## Emission Distribution

$$X_t | \mathbf{Q}_t \sim \text{Hist}_{q_1, \dots, q_K}(\mathbf{Q}_t, s_t, e_t)$$

where  $\text{Hist}_{q_1, \dots, q_K}(\mathbf{Q}_t, s_t, e_t)$  is the distribution coming from the family

$$\begin{aligned} \mathcal{F}_{(q_1, \dots, q_K)} &= \{\text{Hist}_{q_1, \dots, q_K}(\mathbf{Q}, s, e) = p(x | \mathbf{Q}, s, e) \\ &= q_1(s * \exp(-(q_1 - x)s) \mathbb{I}(x < Q_1) + \\ &\quad \sum_{i=2}^{K-1} (q_i - q_{i-1}) \mathbb{I}(Q_i \leq x < Q_{i+1}) \frac{1}{Q_i - Q_{i-1}} \\ &\quad + (1 - q_K) \mathbb{I}(x \geq Q_K) (e * \exp(-(x - q_{M-1})e)) | \infty < Q_1 < \dots < Q_K, s > 0, e > 0\} \end{aligned} \quad (1)$$



The following 3 constraints allow this work to build upon what has already been done

- Order preservation
- Consistent estimation
- Adaptable



### Constraint

**Order Preserving Estimation:** For each time  $t \geq 0$  the marginal filter distributions should be stochastically ordered. In other words  $p(Q_{1t}|X_1 = x_1, \dots, X_{t-1} = x_{t-1}) \preceq \dots \preceq p(Q_{Kt}|X_1 = x_1, \dots, X_{t-1} = x_{t-1})$



### Constraint

*Consistent in iid Circumstances* For every  $F$ , if  $F_1 = F_2 = \dots = F$ , i.e. if the stream is truly an iid stream possessing true quantiles  $Q_1 = F^{-1}(q_1), \dots, Q_K = F^{-1}(q_K)$ , then for  $1 \leq k \leq K$  and every  $\epsilon > 0$

$$\lim_{t \parallel \mathbf{V}_t \parallel \rightarrow \infty, \parallel \mathbf{V}_t \parallel \rightarrow 0} p(|Q_{kt} - Q_k| > \epsilon | X_1 = x_1, \dots, X_{t-1} = x_{t-1}) = 0$$



## Constraint

**Fast Adaptation** *In practice we require  $\mathbf{V}_t$  to not tend to zero so that adaptation to distributional change is possible. At least, we must empirically verify fast adaptation to change for non-varying  $\mathbf{V}$ . Ideally we can specify a deterministic algorithm for dynamically varying  $\mathbf{V}_t$  that maintains consistency but allows adaptation as in Barata (2021).*



## Specific Model Under Analysis



This model maintains the  $(1/K, 2/K, \dots, (K-1)/K)$  quantiles in a transformed space that allows for multivariate normal transitions. Let  $\mathbf{a}_t = (a_{1t}, a_{2t}, \dots, a_{(K-1)t}, \alpha_t, \beta_t)$  for  $t \geq 1$ . Also define for  $t \geq 1$ .

$$Q_{1t}(\mathbf{a}_t) = a_{1t}$$

and for  $2 \leq j \leq K-1$

$$Q_{jt}(\mathbf{a}_t) = a_{1t} + \sum_{t=2}^j \exp(a_{jt})$$

and

$$s(\mathbf{a}_t) = \exp(\alpha_t)$$

and

$$e(\mathbf{a}_t) = \exp(\beta_t)$$

In the “a” space, the first dimension is the  $1/K$  quantile, the second dimension is the log of the difference between the  $2/K$  and  $1/K$  quantile, the third dimension is the log of the difference between  $3/K$  and  $2/K$  quantile, and so on.

## Transition Distributions: Specification of $G$

$\mathbf{a}_t | \mathbf{a}_{t-1} \sim \text{MVN}(\mathbf{a}_{t-1}, \Sigma_t)$  where  $\Sigma_t$  is  $K + 1 \times K + 1$  diagonal and

$$\text{Var}(\mathbf{a}_{1t} | \mathbf{a}_{1(t-1)}) = \sigma_L^2(t)$$

and for  $2 \leq j \leq K - 1$

$$\text{Var}(\mathbf{a}_{jt} | \mathbf{a}_{j(t-1)}) = \zeta(t, K)$$

and

$$\text{Var}(\alpha_t | \alpha_{t-1}) = \text{Var}(\beta_t | \beta_{t-1}) = \sigma_B^2(t)$$

Note in this model that  $\mathbf{V}_t := (\sigma_L^2(t), \zeta(t, K), \sigma_B^2(t))$

## Initial Distribution: Specification of $G_0$

$\mathbf{a}_1 \sim \text{MVN}([0, c, c, \dots, c, d]^T, \Sigma_0)$ .  $\Sigma_0$  is diagonal; the first entry is  $\sigma_L^2(0)$ . The last two entries are  $\sigma_B^2(0)$ . The other diagonal entries are  $\zeta(0, K)$

# Simulations



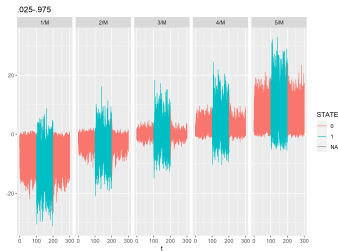


Figure: Quantiles over a standard Normal.

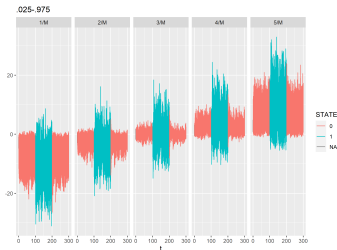


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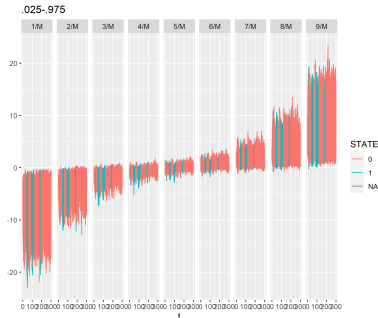


Figure: Quantiles over a t-distribution.



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- We have built a flexible framework for modeling multiple quantiles for streaming data.
- Simulated streaming data using an example model class shows promise in using this method for detecting anomalies.
- Regulatory conditions are needed to ensure the estimated quantile function is proper along the stream.
- Establishing posterior consistency for the dynamic setting requires focus on the iid setting which is currently in progress.
- Once established, will focus on credible intervals of the quantiles can be used to do anomaly detection with risk quantification.

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