



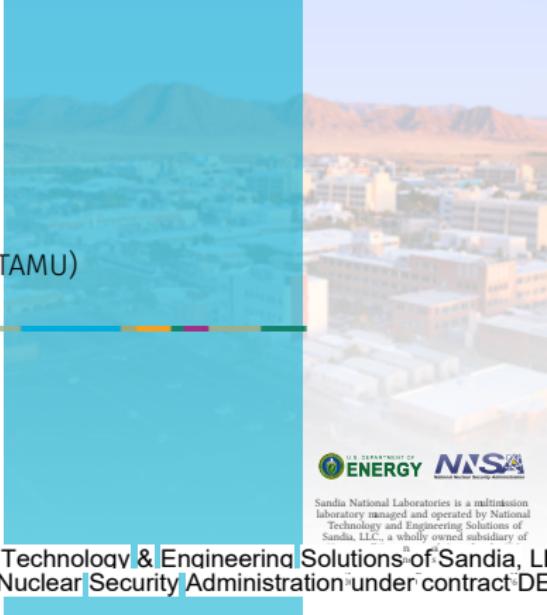
Sandia
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Towards dynamic quantile function models for anomaly detection

PRESENTED BY

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Outline



- 1 Introduction: Problem Definition, Motivating Applications, and Prior Work
- 2 Our Approach: General Framework, Constraints, and Current Model
- 3 Conclusions

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Problem Definition



We observe univariate $X_1 \sim F_1, X_2 \sim F_2, \dots$ in a stream

$$F_1, F_2, \dots$$

vary through time. At each time $t \geq 1$, we want to estimate the

$$0 < q_1 < q_2 < \dots < q_K < 1$$

quantiles of F_t

Motivating Applications



Motivating Applications

Detecting Malicious Activity in a Stream of Computer Network Data

- Multivariate points in the stream of network data
 X_1, X_2, \dots are converted via feature engineering to a discriminative 1d stream
 X_1, X_2, \dots as in (Barata, 2021)
- Raises in the .9 quantile without changes in the .85 quantile could indicate a small group of machines with a common characteristic have become infected



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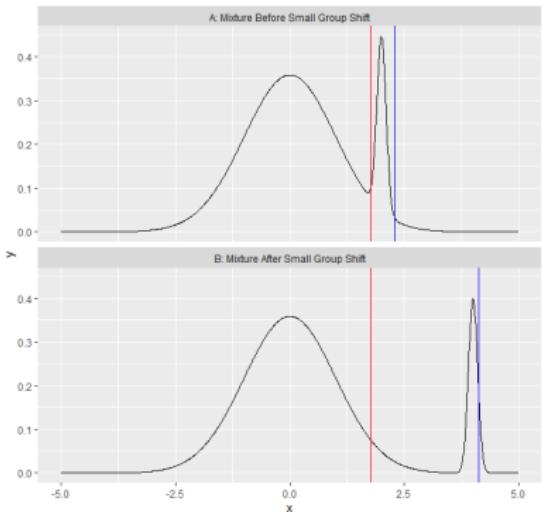


Figure: Red line is .86 quantile. Blue line is .99 quantile

Motivating Applications





Monitoring for Oil Price Shocks over Days

- Changes in the .9 quantile of world oil prices without changes in the .99 quantile indicates that oil producing nations that had expensive prices became more expensive but the countries with the most costly oil still had the most costly oil
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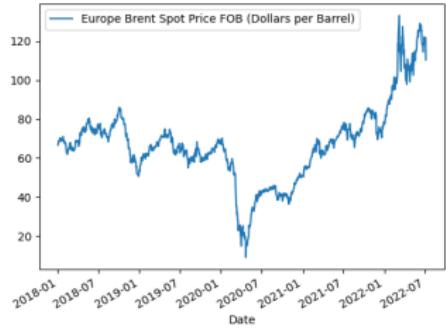


Figure: Time Series of BRENT Crude oil Dollars Per Barrel



Prior Work

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- Extension to multiple quantiles via rules that constrain the order on the estimates
- Consistency in iid setting is proved: If $F_1 = F_2 = \dots = F$, then in probability if $Q_q = F^{-1}(q) > 0$, and $Q_q(0) > 0$

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- In the non-dynamic setting $F = F_1 = F_2 = \dots$, frequentist approaches have worked out asymptotically valid interval estimation for a wide range of distributions Barata (2021), Barata (2021), Barata (2021)
 - These methods are inherently unable to adapt to change because they weight each member of the sample equally



- Dynamic credible interval construction for a single quantile is also well studied Barata (2021)
- This work suffers from a stochastic ordering violation
 - The model can be fit to 2 quantiles separately but if estimating quantiles q_1 and q_2 where $q_1 < q_2$ it is possible there will be a $c \in \mathbb{R}$ s.t $P(Q_{1t} > c | X_1, \dots, X_t) > P(Q_{2t} > c | X_1, \dots, X_t)$
 - This type of contradiction is likely to occur when estimating extreme low quantiles

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General Framework: HMM

We will approach the dynamic many quantile estimation problem using a hidden markov model structure as well. Recall we seek to estimate the $0 < q_1 < \dots < q_K < 1$ quantiles of F_t for each t . We propose the following model

Initial Distribution

$$\mathbf{Q}_0 = (Q_{10}, \dots, Q_{K0}) \sim G_0$$

where G_0 is some to be determined parametric family of distributions on $\{(x_1, \dots, x_K) | -\infty < x_1 < x_2 < \dots < x_K < \infty\}$. And for $t \geq 1$

Transition Distribution

$$Q_t | \mathbf{Q}_{t-1} \sim G(\mathbf{q}_{t-1}, \mathbf{V}_t)$$

where G is some to be determined parametric family of distributions on $\{(x_1, \dots, x_K) | -\infty < x_1 < x_2 < \dots < x_K < \infty\}$ parameterized by the underlying quantile vector at time $t - 1$ and a vector of parameters dictating the variance of transition at time



Emission Distribution

$$X_t | \mathbf{Q}_t \sim \text{Hist}_{q_1, \dots, q_K}(\mathbf{Q}_t, s_t, e_t)$$

where $\text{Hist}_{q_1, \dots, q_K}(\mathbf{Q}_t, s_t, e_t)$ is the distribution coming from the family

$$\begin{aligned}
 \mathcal{F}_{(q_1, \dots, q_K)} &= \{ \text{Hist}_{q_1, \dots, q_K}(\mathbf{Q}, s, e) = p(x|\mathbf{Q}, s, e) \\
 &= q_1(s * \exp(-(q_1 - x)s) \mathbb{I}(x < q_1) + \\
 &\quad \sum_{i=2}^{K-1} (q_i - q_{i-1}) \mathbb{I}(q_i \leq x < q_{i+1}) \frac{1}{q_i - q_{i-1}} \\
 &\quad + (1 - q_K) \mathbb{I}(x \geq q_K) (e * \exp(-(x - q_{M-1})e)) \mid \infty < q_1 < \dots < q_K, s > 0, e > 0 \}
 \end{aligned} \tag{1}$$

Constraints



The following 3 constraints allow this work to build upon what has already been done

- Order preservation
- Consistent estimation
- Adaptable



Constraint

Order Preserving Estimation: For each time $t \geq 0$ the marginal filter distributions should be stochastically ordered. In other words $p(Q_{1t}|X_1 = x_1, \dots, X_{t-1} = x_{t-1}) \preceq \dots \preceq p(Q_{Kt}|X_1 = x_1, \dots, X_{t-1} = x_{t-1})$



Constraint

Consistent in iid Circumstances For every F , if $F_1 = F_2 = \dots = F$, i.e if the stream is truly an iid stream possessing true quantiles $Q_1 = F^{-1}(q_1), \dots, Q_K = F^{-1}(q_k)$, then for $1 \leq k \leq K$ and every $\epsilon > 0$

$$\lim_{t \|\mathbf{v}_t\| \rightarrow \infty, \|\mathbf{v}_t\| \rightarrow 0} p(|Q_{kt} - Q_k| > \epsilon | X_1 = x_1, \dots, X_{t-1} = x_{t-1}) = 0$$

Constraints: Adaptable

Constraint

Fast Adaptation In practice we require \mathbf{V}_t to not tend to zero so that adaptation to distributional change is possible. At least, we must empirically verify fast adaption to change for non-varying \mathbf{V} . Ideally we can specify a deterministic algorithm for dynamically varying \mathbf{V}_t that maintains consistency but allows adaptation as in Barata (2021).

Specific Model Under Analysis



This model maintains the $(1/K, 2/K, \dots, (K-1)/K)$ quantiles in a transformed space that allows for multivariate normal transitions
 Let $\mathbf{a}_t = (a_{1t}, a_{2t}, \dots, a_{(K-1)t}, \alpha_t, \beta_t)$ for $t \geq 1$. Also define for $t \geq 1$.

$$Q_{1t}(\mathbf{a}_t) = a_{1t}$$

and for $2 \leq j \leq K-1$

$$Q_{jt}(\mathbf{a}_t) = a_{1t} + \sum_{t=2}^j \exp(a_{jt})$$

and

$$s(\mathbf{a}_t) = \exp(\alpha_t)$$

and

$$e(\mathbf{a}_t) = \exp(\beta_t)$$

In the “a” space, the first dimension is the $1/K$ quantile, the second dimension is the log of the difference between the $2/K$ and $1/K$ quantile, the third dimension is the log of the difference between $3/K$ and $2/K$ quantile, and so on.

Transition Distributions: Specification of G

$\mathbf{a}_t | \mathbf{a}_{t-1} \sim MVN(\mathbf{a}_{t-1}, \Sigma_t)$ where Σ_t is $K+1 \times K+1$ diagonal and

$$Var(\mathbf{a}_{1t} | \mathbf{a}_{1(t-1)}) = \sigma_L^2(t)$$

and for $2 \leq j \leq K-1$

$$Var(\mathbf{a}_{jt} | \mathbf{a}_{j(t-1)}) = \zeta(t, K)$$

and

$$Var(\alpha_t | \alpha_{t-1}) = Var(\beta_t | \beta_{t-1}) = \sigma_B^2(t)$$

Note in this model that $\mathbf{V}_t := (\sigma_L^2(t), \zeta(t, K), \sigma_B^2(t))$

Initial Distribution: Specification of G_0

$\mathbf{a}_1 \sim MVN([0, c, c, \dots, c, d, d]^T, \Sigma_0)$. Σ_0 is diagonal; the first entry is $\sigma_L^2(0)$. The last two entries are $\sigma_B^2(0)$. The other diagonal entries are $\zeta(0, K)$

Simulations



Simulations

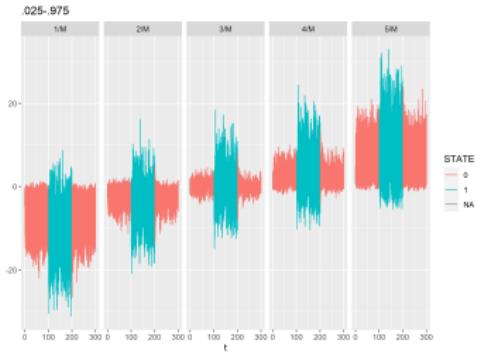


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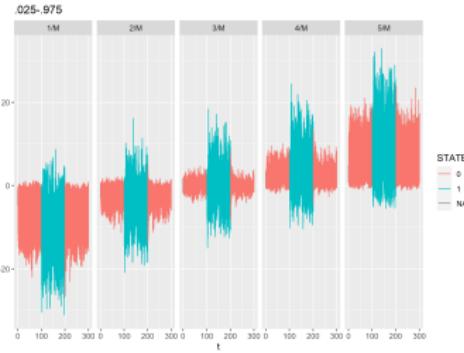


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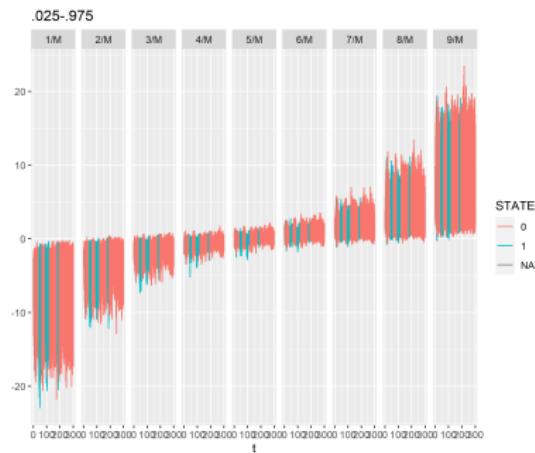


Figure: Quantiles over a t-distribution.

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Conclusions



- We have built a flexible framework for modeling multiple quantiles for streaming data.
- Simulated streaming data using an example model class shows promise in using this method for detecting anomalies.
- Regulatory conditions are needed to ensure the estimated quantile function is proper along the stream.
- Establishing posterior consistency for the dynamic setting requires focus on the iid setting which is currently in progress.
- Once established, will focus on credible intervals of the quantiles can be used to do anomaly detection with risk quantification.

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