



Model Discrepancy Sensitivities to Enable Decision-Making

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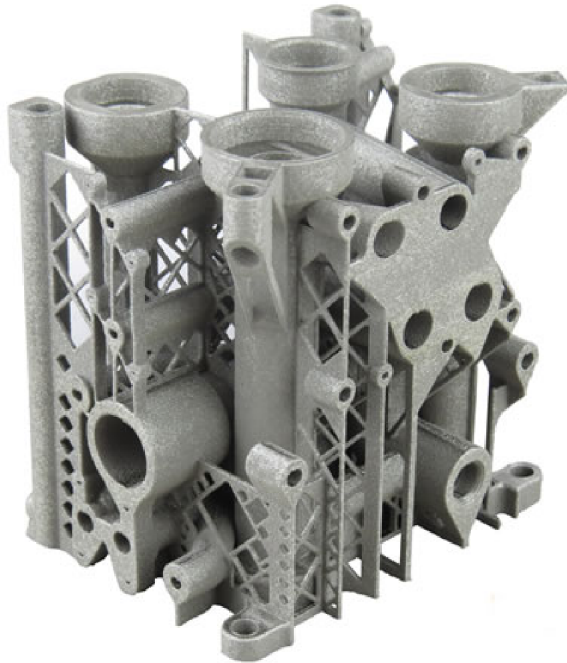
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All models are wrong...



...but some are useful (George Box 76)

- Computational models support decision making when:
 - ✓ the models are computationally efficient enough
 - ✓ uncertainty may be accounted for and propagated through the analysis
- Many models of complex systems do not meet these criteria
- Data science supports the development of computationally efficient models, but may introduce additional errors and uncertainty.

$$\min_{z \in \mathcal{Z}} J(\tilde{S}(z), z)$$

- J is the objective
- z is a design, control, or inversion parameter
- $\tilde{S}(z)$ is an approximate model

Our goals are:

- Use the limited high-fidelity evaluations to improve the solution
- Characterize uncertainty in the optimal solution due to $S - \tilde{S}$



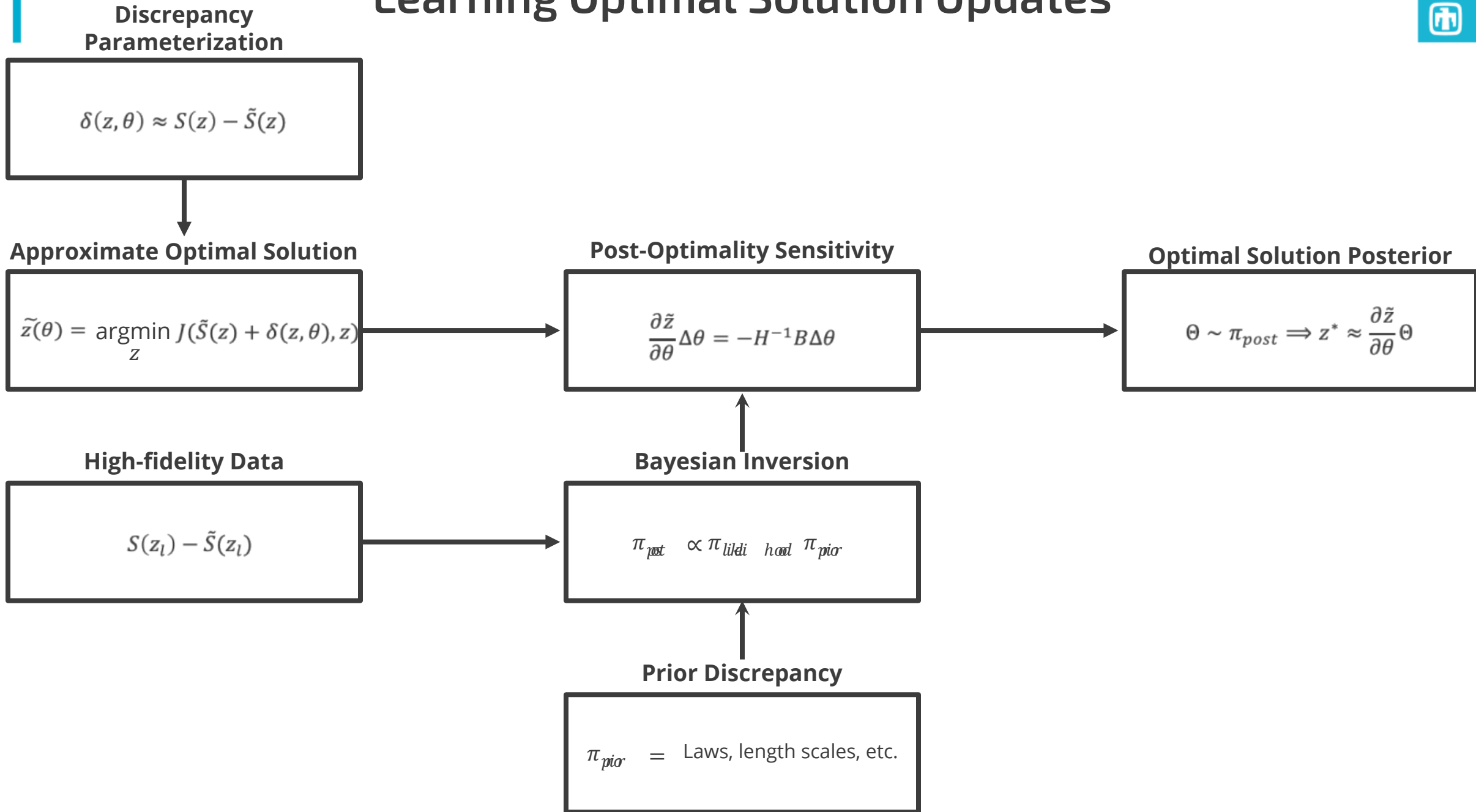
Approximate Optimal Solution

$$\tilde{z}(\theta) = \operatorname{argmin}_z J(\tilde{S}(z), z)$$

High-fidelity Data

$$S(z_l) - \tilde{S}(z_l)$$

Learning Optimal Solution Updates





$$\min_z \frac{1}{2} \int_0^1 (\tilde{S}(z) - T(x))^2 dx + \frac{\beta}{2} \int_0^1 z \mathcal{E} z$$

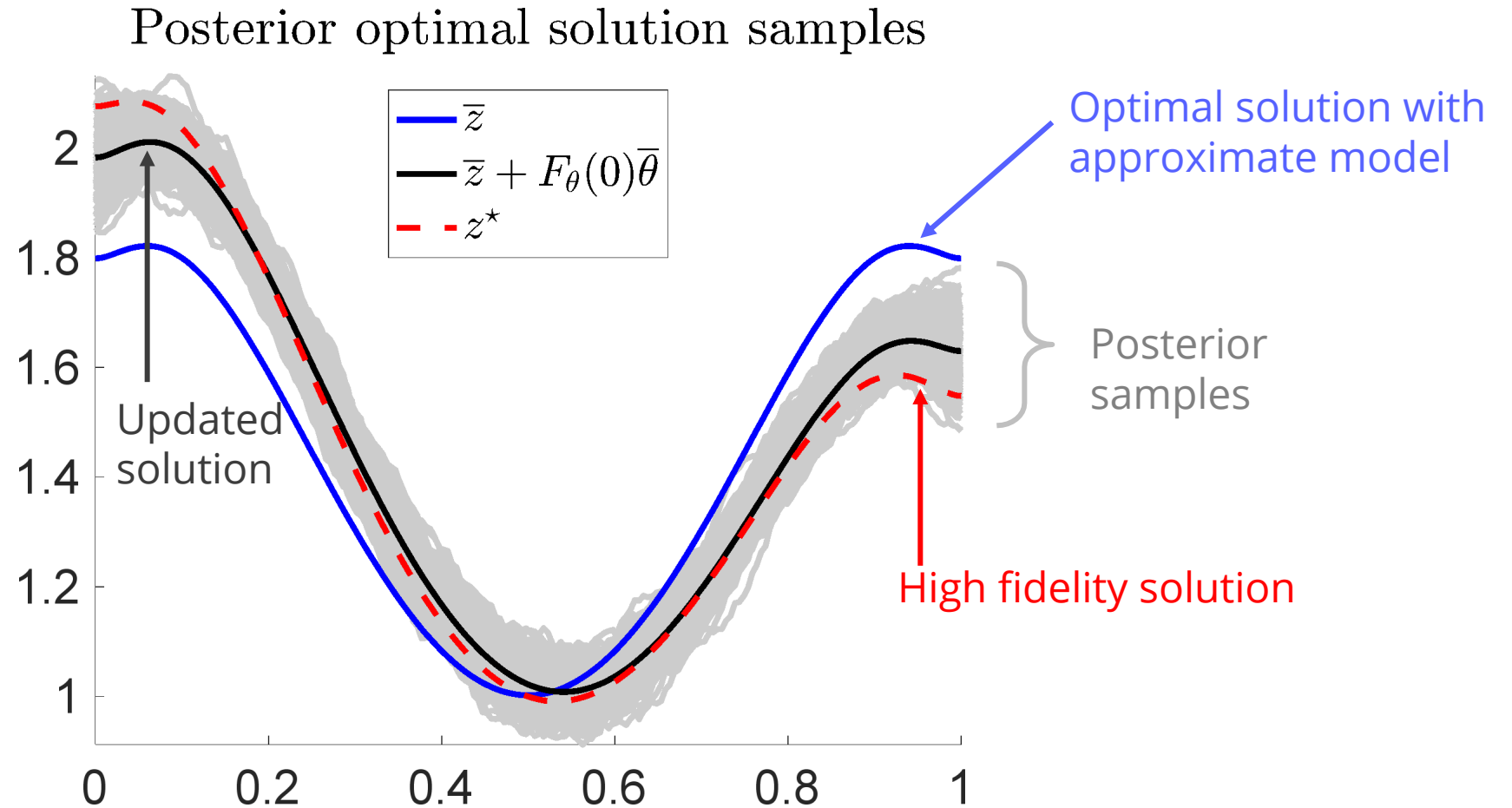
where $\tilde{S}(z)$ is the solution operator for

$$\begin{aligned} -\kappa u'' &= z && \text{on } (0, 1) \\ \kappa u' &= hu && \text{on } \{0, 1\} \end{aligned}$$

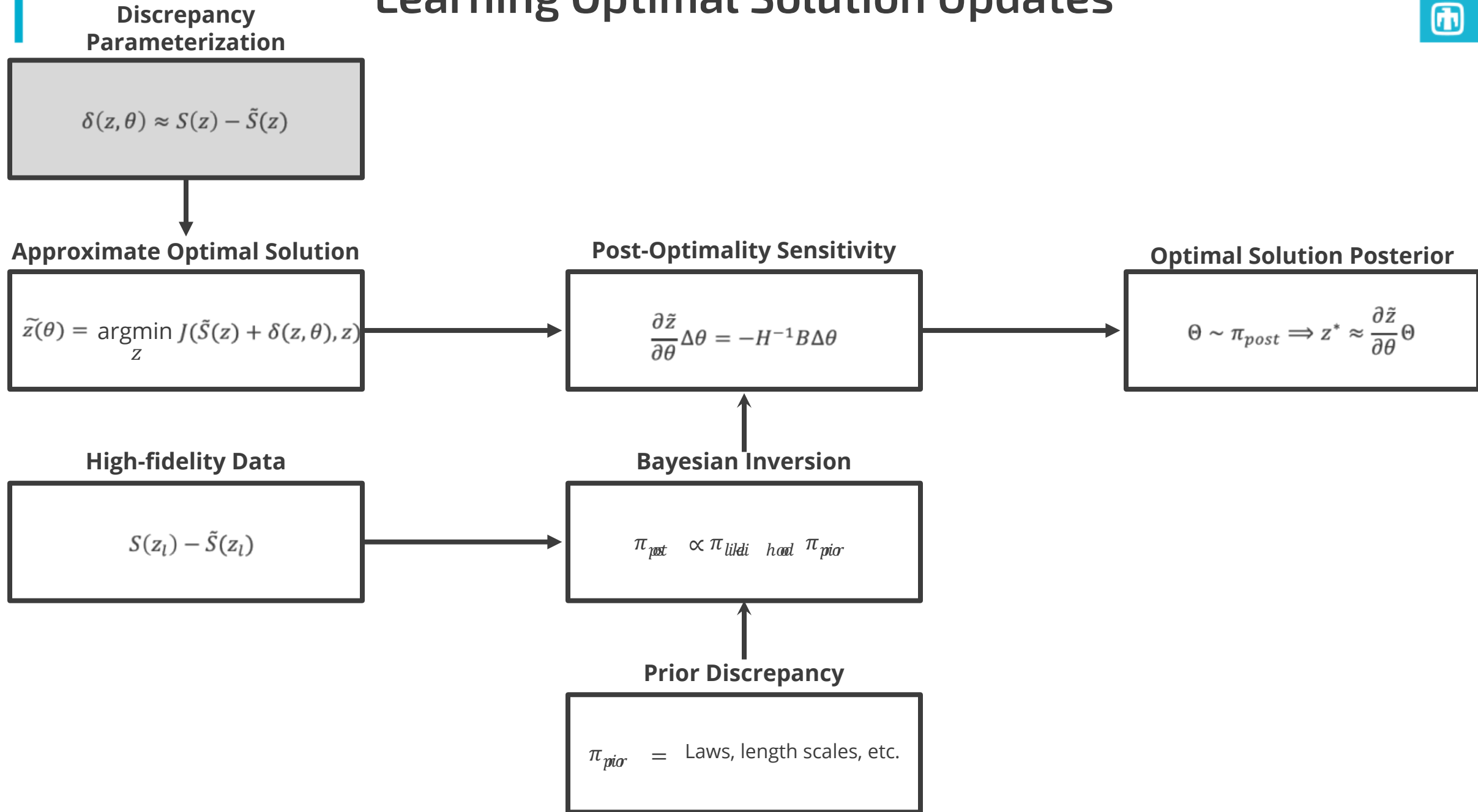
The high-fidelity model S solves

$$\begin{aligned} -\kappa u'' + vu' &= z && \text{on } (0, 1) \\ \kappa u' &= hu && \text{on } \{0, 1\} \end{aligned}$$

Given the high-fidelity solution $S(z)$ for 2 different source terms, improve and characterize uncertainty in the low-fidelity optimal source.



Learning Optimal Solution Updates



Model Discrepancy Representation



- General form for a (discretized) operator

$$\sum_{i=1}^m f_i(z) \phi_i$$

- Since post-optimality analysis only depends on the mixed (z, θ) derivative, assume f_i 's are linear, Reisz representation yields

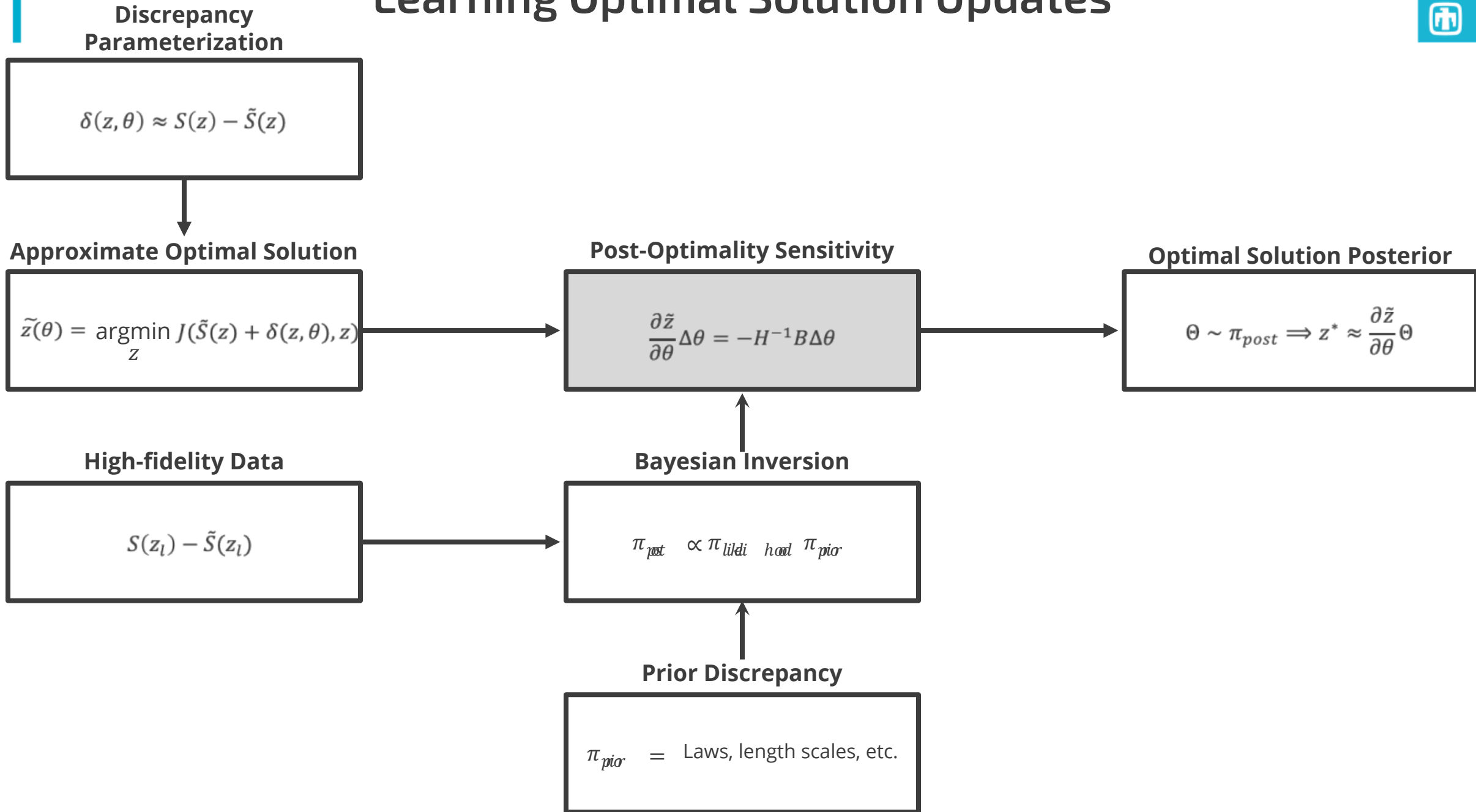
$$\sum_{i=1}^m \left(\theta_{i,0} + \sum_{j=1}^n \theta_{i,j}(z, \psi_j) z \right) \phi_i$$

↓

$$\delta(\mathbf{z}, \theta) = (\mathbf{I}_m \quad \mathbf{I}_m \otimes \mathbf{z}^T \mathbf{M}_z) \theta$$

- Discretized $\delta : \mathbb{R}^n \times \mathbb{R}^p \rightarrow \mathbb{R}^m$ is parameterized by $\theta \in \mathbb{R}^p$
- $p = m(n + 1)$ so the dimension of θ may be $\mathcal{O}(\text{mesh size}^2)$
- Evaluate $\delta(z, \theta)$ efficiently using Kronecker product
- $(M_z)_{i,j} = (\psi_i, \psi_j)_Z$ - mass matrix that defines the inner product on Z_h

Learning Optimal Solution Updates





$$\min_{\mathbf{z}} \mathbf{J}(\tilde{\mathbf{S}}(\mathbf{z}) + \boldsymbol{\delta}(\mathbf{z}, \theta), \mathbf{z}) \quad (1)$$

- $\tilde{\mathbf{z}}^*$ solves (1) when $\boldsymbol{\delta}(\mathbf{z}, \theta_0) = \mathbf{0}$, the problem solved in practice
- Under mild assumptions, applying the Implicit Function Theorem to

$$\nabla J(\tilde{\mathbf{z}}^*, \theta_0) = \mathbf{0}$$

gives

$$\mathcal{F} : \mathcal{N}(\theta_0) \rightarrow \mathcal{N}(\tilde{\mathbf{z}}^*)$$

such that $\mathcal{F}(\theta_0)$ solves (1) when $\theta = \theta_0$ and

$$\mathcal{F}'_{\theta}(\theta_0) = -\mathcal{H}^{-1}\mathcal{B}$$

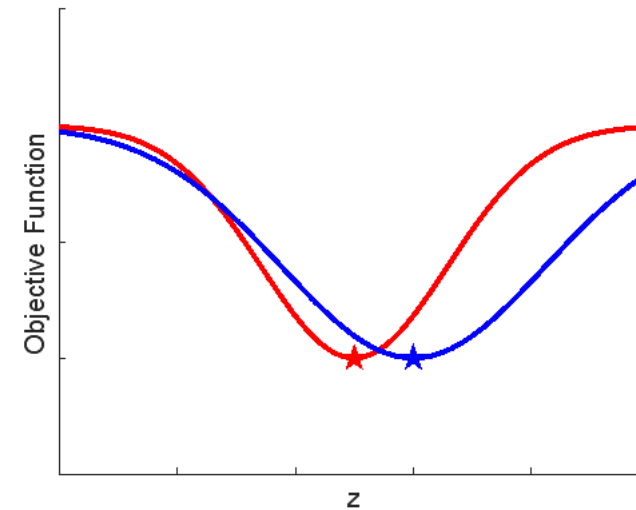
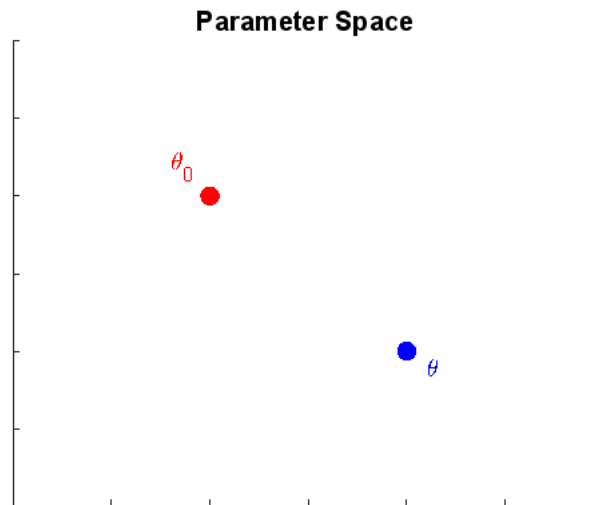
is the sensitivity of the optimal solution with respect to model discrepancy

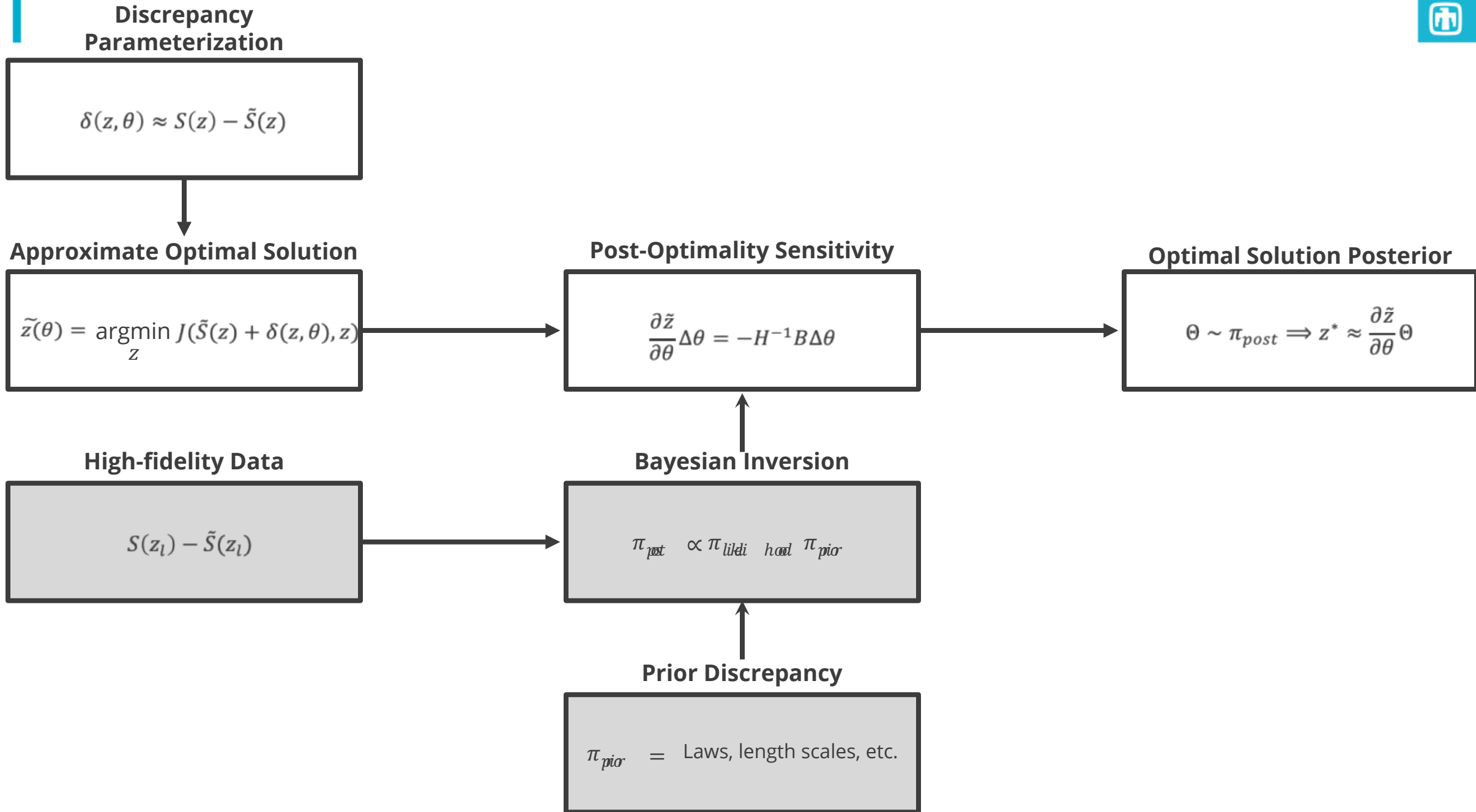
Post-optimality Sensitivities



$$\mathcal{F}'_{\theta}(\theta_0) = -\mathcal{H}^{-1}\mathcal{B}$$

- \mathcal{H} is the Hessian of the objective function with respect to z
- \mathcal{B} is the mixed second derivative of the objective with respect to z and θ
- Acts like a Newton step to update the optimal solution after a perturbation of the model discrepancy





Bayesian Inverse Problem - Prior Discrepancy



- Measure size of δ :

$$\|\delta(z, \theta)\|_L^2 = \theta^T \begin{pmatrix} \mathbf{L} & \mathbf{L} \otimes \mathbf{z}^T \mathbf{M}_z \\ \mathbf{L} \otimes \mathbf{M}_z \mathbf{z} & \mathbf{L} \otimes \mathbf{M}_z \mathbf{z} \mathbf{z}^T \mathbf{M}_z \end{pmatrix} \theta$$

- Marginalize out \mathbf{z} :

$$\mathbb{E}_{\mathbf{z}}[\|\delta(z, \theta)\|_L^2] = \theta^T \mathbf{M}_{\theta} \theta$$

where

$$\mathbf{M}_{\theta} = \begin{pmatrix} \mathbf{L} & \mathbf{L} \otimes \bar{\mathbf{z}}^T \mathbf{M}_z \\ \mathbf{L} \otimes \mathbf{M}_z \bar{\mathbf{z}} & \mathbf{L} \otimes \mathbf{E} \end{pmatrix}$$

- \mathbf{L} encodes known physics of the discrepancy - in our case a Laplacian like operator and \mathbf{L}^{-1} represents the prior covariance
- $\mathbf{\Gamma}$ is a covariance matrix on the control space \mathcal{Z}
- Hence \mathbf{M}_{θ} defines an inner product for θ to measure the size of the model discrepancy $\delta(\mathbf{z}, \theta)$ according to our prior knowledge imposed in \mathbf{L} and $\mathbf{\Gamma}$

Bayesian Inverse Problem - notation



- for notational simplicity, we define

$$\mathbf{A}_\ell = (\mathbf{I}_m \quad \mathbf{I}_m \otimes \mathbf{z}_\ell^T \mathbf{M}_z) \in \mathbb{R}^{m \times p}, \quad \ell = 1, 2, \dots, N,$$

so that $\delta(\mathbf{z}_\ell, \theta) = \mathbf{A}_\ell \theta$, and the concatenation of these matrices

$$\mathbf{A} = \begin{pmatrix} \mathbf{A}_1 \\ \mathbf{A}_2 \\ \vdots \\ \mathbf{A}_N \end{pmatrix} \in \mathbb{R}^{mN \times p}$$

so that $\mathbf{A}\theta \in \mathbb{R}^{mN}$ corresponds to evaluating $\delta(\mathbf{z}, \theta)$ for the inputs \mathbf{z}_ℓ .

- let $\mathbf{b} \in \mathbb{R}^{mN}$ be defined by stacking $\mathbf{y}_\ell = \mathbf{S}(\mathbf{z}_\ell) - \tilde{\mathbf{S}}(\mathbf{z}_\ell)$, $\ell = 1, 2, \dots, N$, into a vector so that we seek

$$\mathbf{A}\theta \approx \mathbf{b}$$

- infinite number of θ directions because the problem is underdetermined

Bayesian Inversion Problem



- Arrange data and discrepancy representation so that we seek θ such that

$$\mathbf{A}\theta \approx \mathbf{b}$$

- Given Gaussian prior and noise models, linearity of $\delta(\mathbf{z}, \theta)$ in θ , the posterior is Gaussian with a negative log probability density function

$$\frac{1}{2\alpha} (\mathbf{A}\theta - \mathbf{b})^T (\mathbf{A}\theta - \mathbf{b}) + \frac{1}{2} \theta^T \mathbf{M}_\theta \theta.$$

- α balances the dependence of prior and data misfit
- The posterior mean is

$$\bar{\theta} = \frac{1}{\alpha} \Sigma \mathbf{A}^T \mathbf{b}$$

and the posterior covariance is

$$\Sigma = \left(\mathbf{M}_\theta + \frac{1}{\alpha} \mathbf{A}^T \mathbf{A} \right)^{-1}.$$

- The goal is to sample from a Gaussian distribution which may be generated by multiplying a factor of the covariance matrix with a standard normal random vector and adding the mean
- But how do we invert the sum?

$$\Sigma = \left(\mathbf{M}_\theta + \frac{1}{\alpha} \mathbf{A}^T \mathbf{A} \right)^{-1}$$

1. Factorize \mathbf{A} to rewrite $\mathbf{M}_\theta + \frac{1}{\alpha} \mathbf{A}^T \mathbf{A}$
2. Invert $\mathbf{M}_\theta + \frac{1}{\alpha} \mathbf{A}^T \mathbf{A}$
3. Factorize Σ
4. Compute matrix-vector products for posterior samples

Bayesian Inversion Problem – 4 step process



1. Factorize \mathbf{A} to rewrite $\mathbf{M}_\theta + \frac{1}{\alpha}\mathbf{A}^T\mathbf{A}$

- compute GSVD of \mathbf{A} in the $\mathbf{M}_\theta = \mathbf{C}\mathbf{C}^T$ inner product
- exploit Kronecker structure
- factor out \mathbf{M}_θ to enable inversion: $\Sigma^{-1} = \frac{1}{\alpha}\mathbf{C}\mathbf{X}\mathbf{C}^T$
- where

$$\mathbf{X} = \alpha\mathbf{I} + \mathbf{C}^T\mathbf{\Psi}\mathbf{\Phi}^2\mathbf{\Psi}^T\mathbf{C}.$$

conveniently consists of identity, orthogonal, and diagonal matrices



2. Invert Σ

- applying Sherman-Morrison-Woodbury to \mathbf{X} yields

$$\Sigma = \mathbf{M}_{\theta}^{-1} - \Psi \mathbf{P} \Psi^T.$$

- where $\mathbf{P} \in \mathbb{R}^{mN \times mN}$ is a diagonal matrix whose entries are given by $\frac{\lambda_i}{\lambda_i + \alpha \rho_j}$, a combination of the eigenvalues from \mathbf{L} and \mathbf{G} (object from computing GSVD of \mathbf{A}), along with the noise covariance α
- enables calculation of $\bar{\theta} = \frac{1}{\alpha} \Sigma \mathbf{A}^T \mathbf{b}$



3. Factorize Σ

- factorization is needed to sample Gaussian
- compute eigendecomposition of $\mathbf{X} = \alpha \mathbf{I} + \mathbf{C}^T \Psi \Phi^2 \Psi^T \mathbf{C}$

$$\mathbf{X} = \mathbf{Q} \Upsilon \mathbf{Q}^T.$$

- \mathbf{Q} is an orthonormal matrix with collected eigenvectors

$$\begin{aligned} \Sigma &= \alpha \mathbf{C}^{-T} \mathbf{X}^{-1} \mathbf{C}^{-1} = \mathbf{T} \mathbf{T}^T \\ \mathbf{T} &= \sqrt{\alpha} \mathbf{C}^{-T} \mathbf{Q} \Upsilon^{-\frac{1}{2}}. \end{aligned}$$

- this facilitates efficient sampling

Bayesian Inversion Problem – 4 step process



4. Compute matrix-vector products for posterior samples

- computing matrix-vector products with Σ and \mathbf{T} gives a sum over eigenpairs of an operator which scales with the mesh resolution
- rewrite the sum in terms of linear solves in the direction of the incoming vectors
- this gives expressions for posterior samples which scale with the number of high-fidelity data compute posterior samples as

$$\mathbf{T}\omega = \mathbf{T} (\mathbf{I}_{n+1} \otimes \mathbf{V}^L)^T \nu = \hat{\theta} + \tilde{\theta}$$

where

$$\hat{\theta} = \sqrt{\alpha} \sum_{i=1}^N \frac{1}{\sqrt{\lambda_i}} \begin{pmatrix} s_i \hat{\mathbf{u}}_i \\ \hat{\mathbf{u}}_i \otimes \mathbf{M}_{\mathbf{z}}^{-1} \mathbf{\Gamma}^{-1} \mathbf{w}_i \end{pmatrix} \quad \text{and} \quad \tilde{\theta} = \sum_{k=1}^{n-N+1} \begin{pmatrix} \tilde{s}_k \tilde{\mathbf{u}}_k \\ \tilde{\mathbf{u}}_k \otimes \tilde{\mathbf{w}}_k \end{pmatrix}$$

with

$$\hat{\mathbf{u}}_i = (\alpha \mathbf{L} + \lambda_i \mathbf{I})^{-\frac{1}{2}} \nu_i$$

$$\tilde{\mathbf{u}}_k = \mathbf{L}^{-\frac{1}{2}} \nu_{N+k}$$

$$\tilde{s}_k = -\bar{\mathbf{z}}^T \mathbf{\Gamma}^{-\frac{1}{2}} \tilde{\mathbf{z}}_k$$

$$\tilde{\mathbf{w}}_k = \mathbf{M}_{\mathbf{z}}^{-1} \mathbf{\Gamma}^{-\frac{1}{2}} \tilde{\mathbf{z}}_k$$

Posterior Samples for Discrepancy

- Posterior samples take the form

$$\bar{\theta} + \hat{\theta} + \tilde{\theta}$$

where the mean is

$$\bar{\theta} = \frac{1}{\alpha} \sum_{\ell=1}^N \left[\left(u_{\ell} \otimes \mathbf{M}_z^{-1} \mathbf{\Gamma}^{-1} (z_{\ell} - \bar{z}) \right) - \sum_{i=1}^N b_{i,\ell} \left(\mathbf{u}_{i,\ell} \otimes \mathbf{M}_z^{-1} \mathbf{\Gamma}^{-1} \mathbf{w}_i \right) \right]$$

uncertainty in the data informed directions is

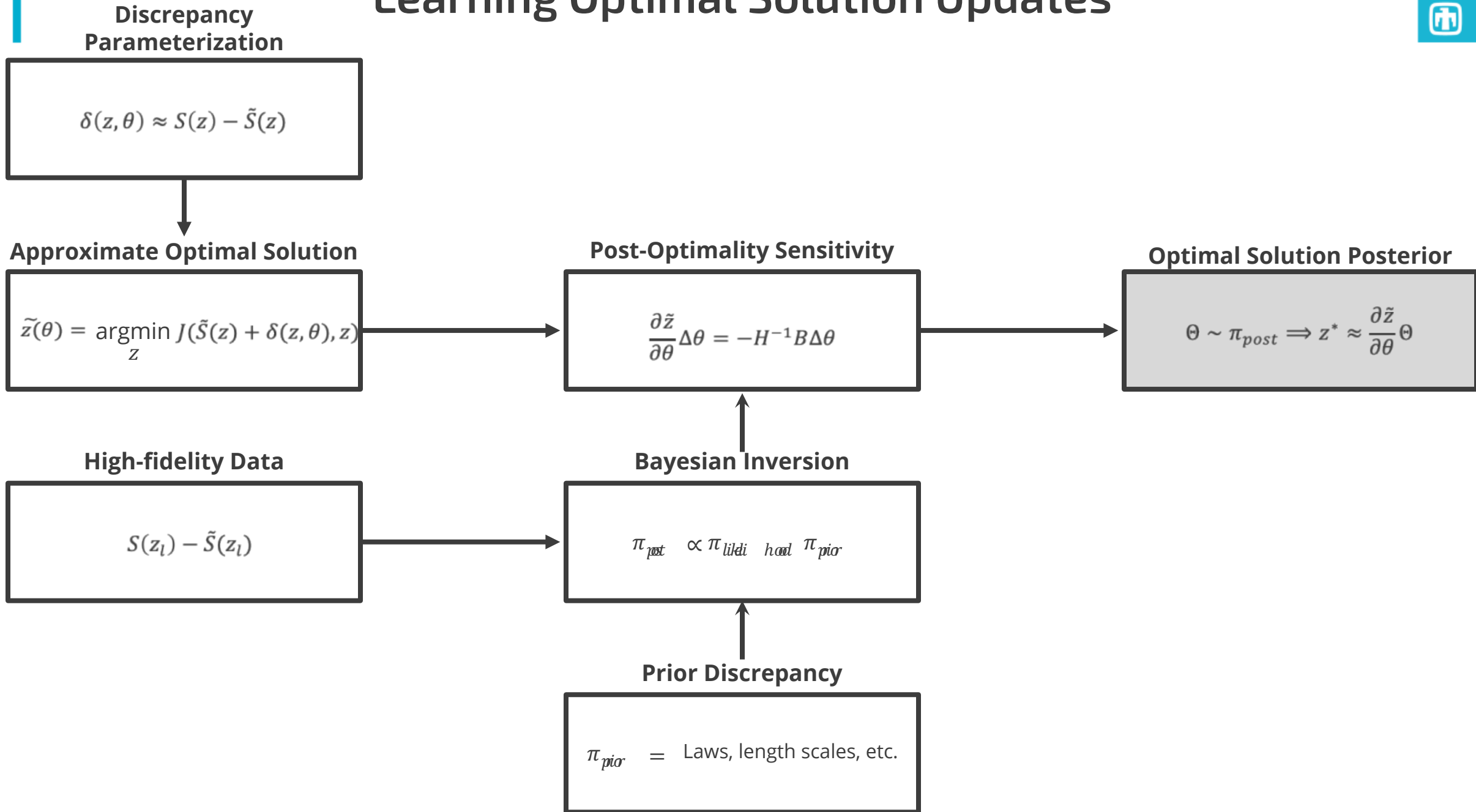
$$\hat{\theta} = \sqrt{\alpha} \sum_{i=1}^N \frac{1}{\sqrt{\lambda_i}} \left(\hat{\mathbf{u}}_i \otimes \mathbf{M}_z^{-1} \mathbf{\Gamma}^{-1} \mathbf{w}_i \right)$$

and uncertainty in the data uninformed directions is

$$\tilde{\theta} = \sum_{k=1}^{n-N+1} \left(\tilde{\mathbf{s}}_k \tilde{\mathbf{u}}_k \right)$$

Learning Optimal Solution Updates

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Sample

$$\mathcal{F}'_{\theta}(\theta_0)(\bar{\theta} + \hat{\theta} + \tilde{\theta}) = -\mathcal{H}^{-1}(\mathcal{B}\bar{\theta} + \mathcal{B}\hat{\theta} + \mathcal{B}\tilde{\theta})$$

$$\begin{aligned} \mathbf{B}\bar{\theta} &= \frac{1}{\alpha} \tilde{\mathbf{S}}_z^T \nabla_{u,u} \mathbf{J} \left[\sum_{\ell=1}^N \left(u_{\ell} - \sum_{i=1}^N b_{i,\ell} (e^T \mathbf{g}_i) \mathbf{u}_{i,\ell} \right) \right] + \frac{1}{\alpha} \sum_{\ell=1}^N (\nabla_u J u_{\ell}) \mathbf{\Gamma}^{-1} (z_{\ell} - \bar{z}) \\ &\quad - \frac{1}{\alpha} \sum_{\ell=1}^N \sum_{i=1}^N b_{i,\ell} (\nabla_u \mathbf{J} \mathbf{u}_{i,\ell}) \mathbf{\Gamma}^{-1} \mathbf{w}_i \end{aligned}$$

$$\mathbf{B}\hat{\theta} = \sqrt{\alpha} \tilde{\mathbf{S}}_z^T \nabla_{u,u} \mathbf{J} \left(\sum_{i=1}^N \frac{e^T \mathbf{g}_i}{\sqrt{\lambda_i}} \hat{u}_i \right) + \sqrt{\alpha} \sum_{i=1}^N \frac{\nabla_u \mathbf{J} \hat{u}_i}{\sqrt{\lambda_i}} \mathbf{\Gamma}^{-1} \mathbf{w}_i$$

and

$$\mathbf{B}\tilde{\theta} = \sum_{k=1}^{n-N+1} (\nabla_u \mathbf{J} \tilde{u}_k) \mathbf{\Gamma}^{-\frac{1}{2}} \tilde{z}_k.$$

A Fluid Flow Example

Optimal design of a flow controller

$$\min_z \frac{1}{2} \int_{\chi} \mathbf{v}_y(z)^2 + \frac{\beta}{2} \int_{\Omega} \|\mathbf{z}\|^2$$

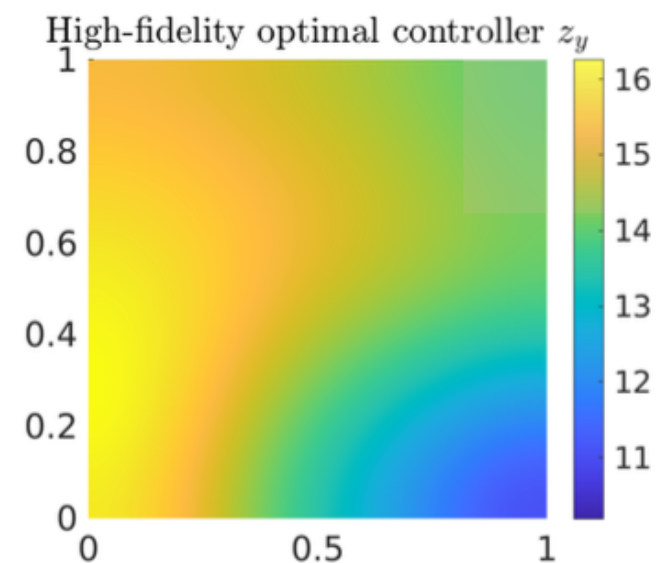
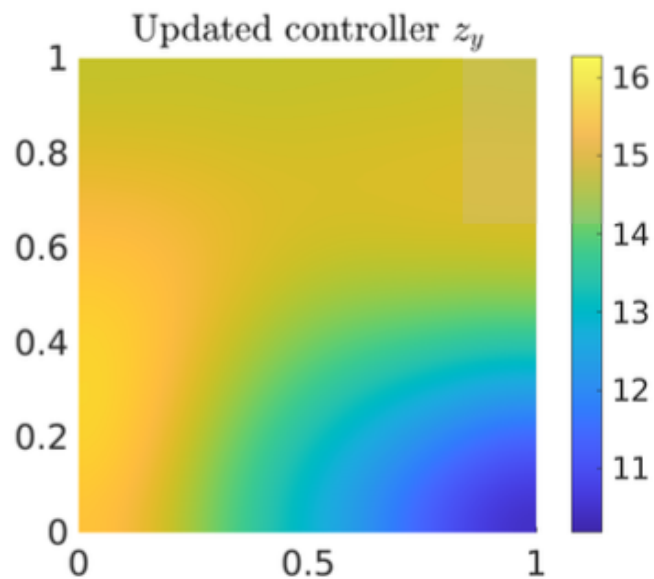
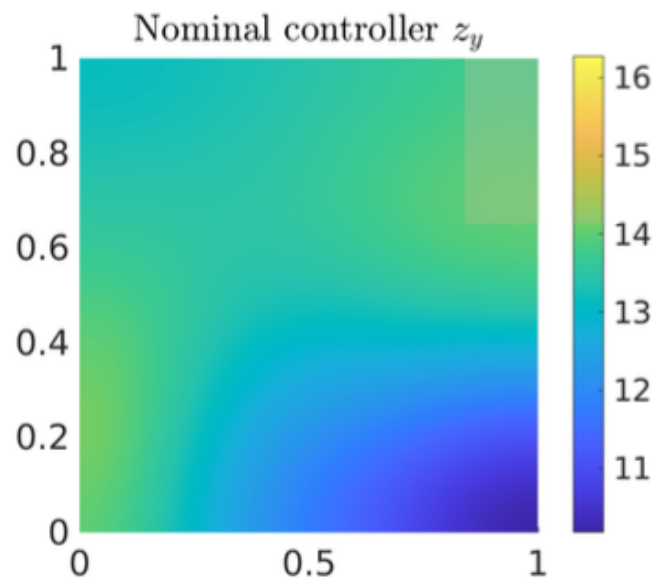
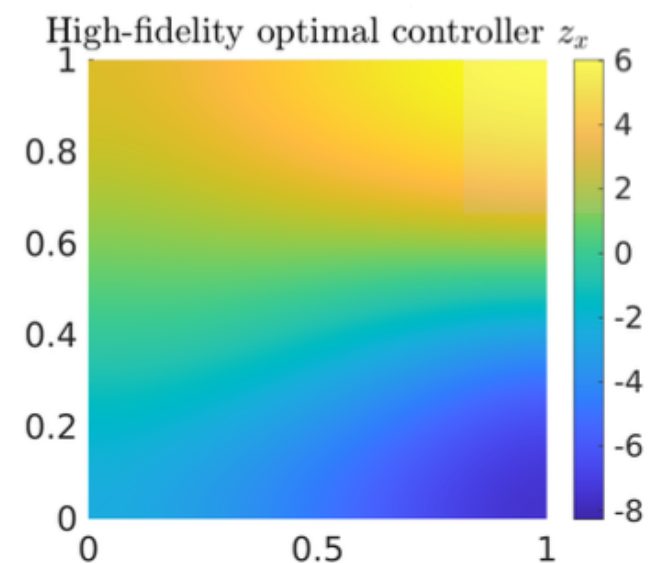
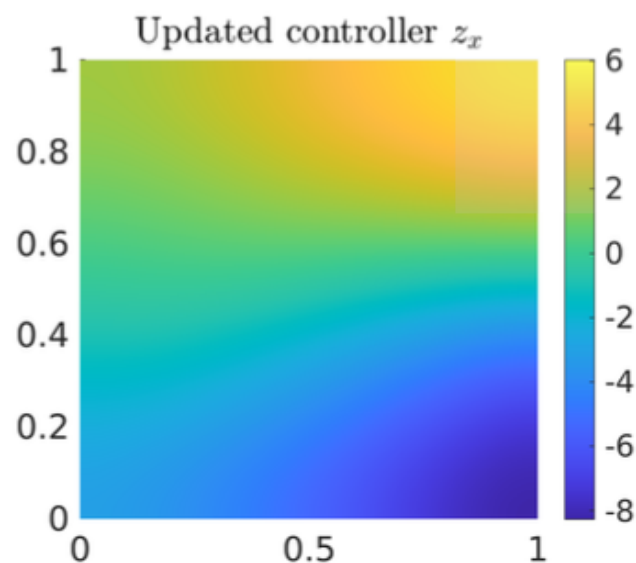
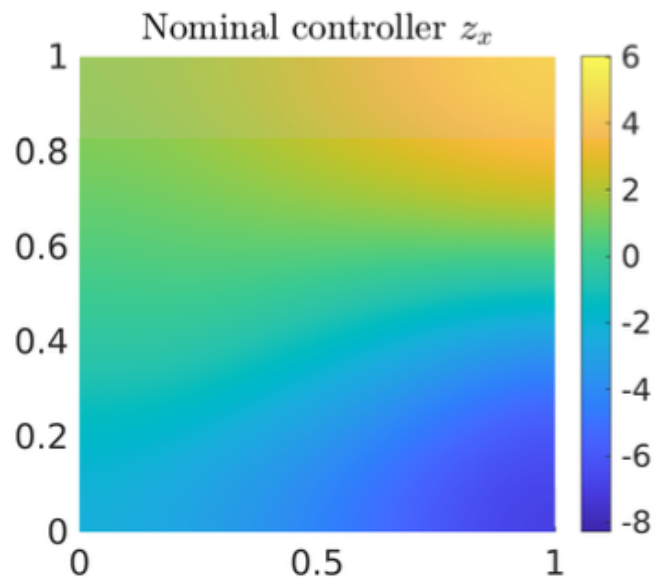
constrained by the Stokes equations

$$\begin{aligned} -\mu \nabla \mathbf{v} + \nabla p &= \mathbf{g} + \mathbf{z} && \text{on } \Omega \\ \nabla \cdot \mathbf{v} &= 0 && \text{on } \Omega \end{aligned}$$

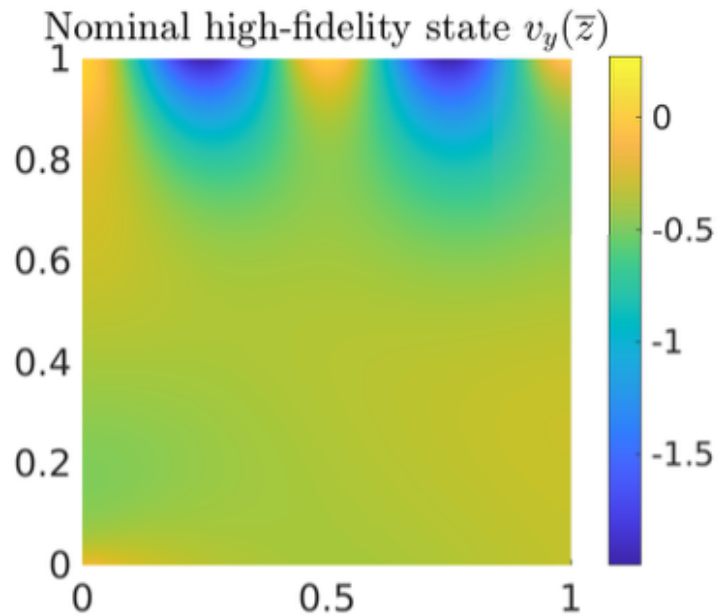
as a simplification of the Navier-Stokes equations

$$\begin{aligned} -\mu \nabla \mathbf{v} + (\mathbf{v} \cdot \nabla) \mathbf{v} + \nabla p &= \mathbf{g} + \mathbf{z} && \text{on } \Omega \\ \nabla \cdot \mathbf{v} &= 0 && \text{on } \Omega \end{aligned}$$

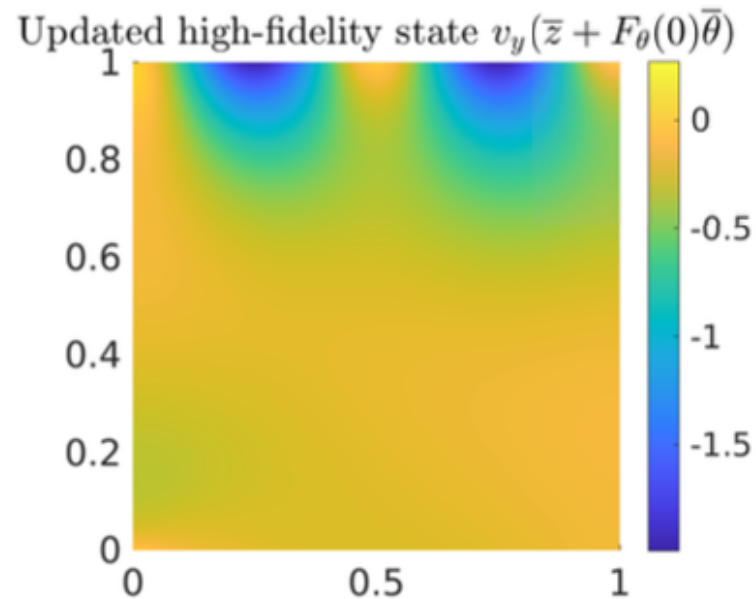
Comparison of Controllers



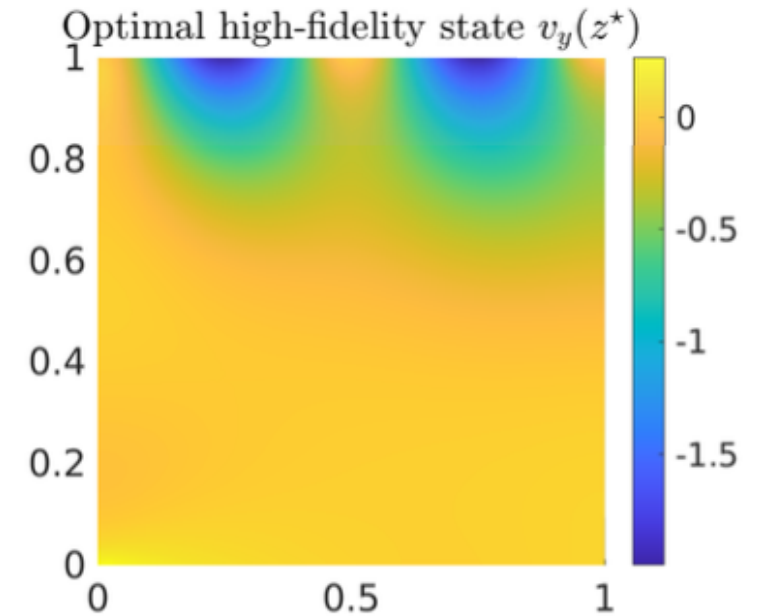
Navier-Stokes solve
with nominal control



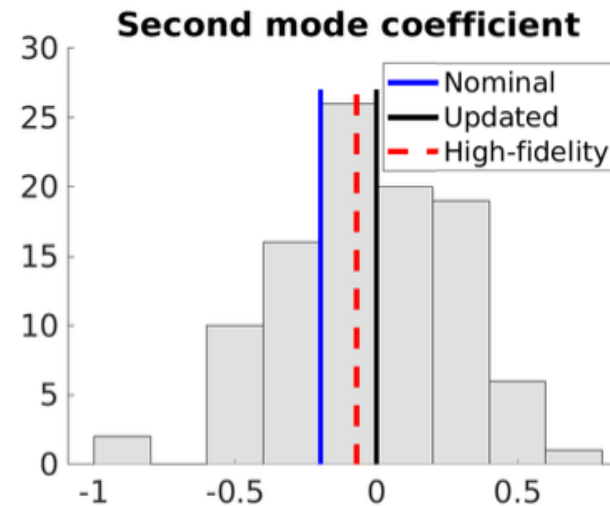
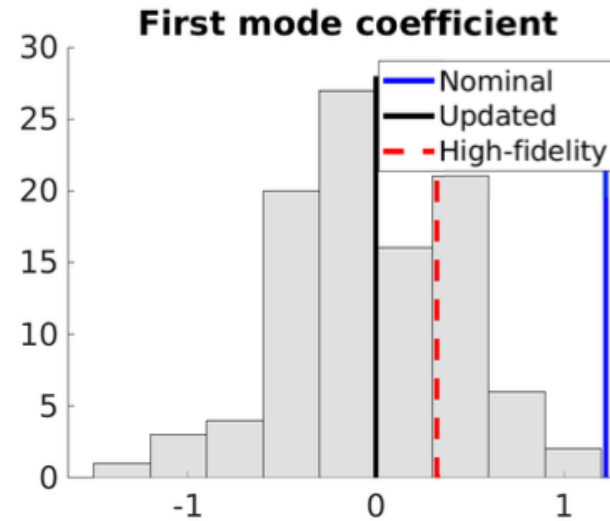
Navier-Stokes solve
with updated control



Navier-Stokes solve
with optimal control



Posterior Controller Uncertainty



- KL representation
- Histogram of posterior
- Goal is for updated to be as close as possible to high-fidelity

Conclusions

- Developed a framework to learn updates of low-fidelity optimal solutions using limited high-fidelity data
- Builds on linear approximation in post-optimality sensitivity analysis
- The discrepancy representation, inverse problem formulation, and judicious linear algebra manipulations enables closed form solution for posterior samples
- Kronecker product representation of the discrepancy facilitates computation which scales with $\dim(\mathcal{U})$ and $\dim(\mathcal{Z})$, not $\dim(\mathcal{U} \otimes \mathcal{Z})$
- Approach is non-intrusive to the high-fidelity data and hence applicable to wide range of applications

✓ Joseph Hart and Bart van Bloemen Waanders, "Hyper-Differential Sensitivity Analysis With Respect to Model Discrepancy (in preparation)