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Bayesian Nonlocal Operator Regression (BNOR): Towards the Characterization of Uncertainty in Heterogeneous Materials

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Introduction

- Consider the **modeling of heterogeneous materials**.
- Material properties (microstructure, interfacial conditions, environments, etc.) cause variability in material response.
- Non-trivial to provide quantitative characterization for each sample.

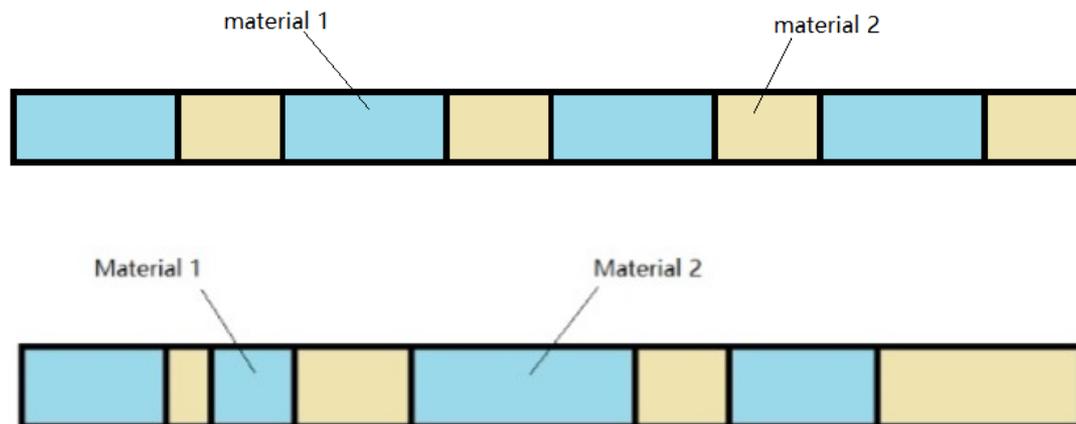


Figure: One-dimensional bar with periodic microstructure and disordered microstructure.

Why use nonlocal model?

Local model:

- Characterized by **differential** operators,
- Interactions happen at **contact**,
- Regularity requirements on the solution.

- Local Poisson's equation

$$-\Delta \mathbf{u}(\mathbf{x}) = \mathbf{f}(\mathbf{x})$$

Nonlocal model:

- Characterized by **integral** operators,
- Long-range interactions of size δ (**horizon**),
- No regularity requirements on the solution.

- Nonlocal Poisson's equation [3]

$$-2 \int_{B(\mathbf{x}, \delta)} K(\mathbf{x}, \mathbf{y}) (\mathbf{u}(\mathbf{y}) - \mathbf{u}(\mathbf{x})) d\mathbf{y} = \mathbf{f}(\mathbf{x})$$

Introduction

- Objective: Develop a Bayesian framework to characterize the uncertainty in using a nonlocal model to describe material response.
- Approach: **Bayesian Nonlocal Operator Regression (BNOR) technique**. To be specific, an **MCMC** Bayesian inference method, to identify the probability distribution of the nonlocal constitutive law that embeds the material properties.
- Application: Wave propagation through a heterogeneous bar, with disordered microstructure layers.

High-Fidelity Data

- Define a set of materials parameterized by the disorder parameter $\mathcal{D} \in [0,1]$ such that each layer material 1 or 2 has size $w \sim \mathcal{U}((1 - \mathcal{D})w_i, (1 + \mathcal{D})w_i)$, where $i = 1,2$ and $w_1 = (1 - \phi)\lambda$, $w_2 = \phi\lambda$, where λ is the mean period of the microstructure. In our experiments we set $L=0.2$ (the bar length), $E_1=1$, $E_2=0.25$ (the Young's Moduli), $\rho=1$ (the density), and $\Omega = (-b, b)$ (the spatial domain representing the bar).

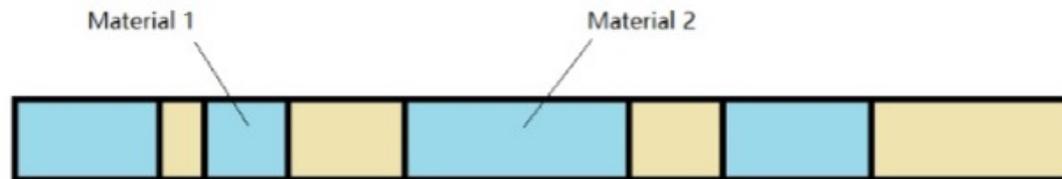


Figure: One-dimensional bar with disordered microstructure.

High-Fidelity Data

- Both training and validation datasets are generated via high-fidelity (HF) simulations of the propagation of stress wave through a one-dimensional heterogeneous bar. The HF model is a classical wave equation: find $u(x, t)$ such that, for $(x, t) \in \Omega \times [0, T]$,

$$\frac{\partial^2 u_{\text{HF}}}{\partial t^2} - \mathcal{L}_{\text{HF}}[u_{\text{HF}}](x, t) = f(x, t)$$

with force loading term $f(x, t)$, provided some initial conditions at $t = 0$ and boundary conditions on $\partial\Omega$.

- The HF-model is numerically solved using a [Direct Numerical Solver \(DNS\)](#) in [2], which guarantees that the wave velocity is computed exactly.

High-Fidelity Data

- Three types of data are generated where type 1 and type 2 are used for training, and type 3 is used for validation. In what follows, v presents the velocity.

- Type 1: Oscillating source. We set $b=50$, $T=2$, $v(x,0)=u(x,0)=0$,

$$f(x, t) = e^{-\left(\frac{2x}{5kL}\right)^2} e^{-\left(\frac{t-t_0}{t_p}\right)} \cos^2\left(\frac{2\pi x}{kL}\right), k = 1, 2, \dots, 20, t_0 = t_p = 0.8.$$

- Type 2: Plane wave with ramp. For $b=50$, $T=2$, $f(x,t)=0$ and $u(x,0)=0$,

$$\text{we prescribe } v(-b, t) = \begin{cases} \cos(\omega t) \sin^2\left(\frac{\pi t}{30}\right), & \text{if } t \leq 15 \\ \cos(\omega t), & \text{if } t > 15 \end{cases}, \text{ for } \omega =$$

0.35, 0.7, ..., 3.85.

- Type 3: Wave packet. $b=133.3$, $T=100$, $f(x,t)=0$ and $v(-b, t) = \sin(\omega t) e^{\left\{-\left(\frac{t}{5}-3\right)^2\right\}}$, for $\omega = 1.0, 2.0, 3.0, 4.0$.

Nonlocal Operator Regression (NOR)

- Following the method in [1], we proposed a nonlocal model to approximate the HF-model at large scales:

$$\frac{\partial^2 u_{NL}}{\partial t^2} - \mathcal{L}_{\mathcal{K}}[u_{NL}](x, t) = f(x, t)$$

where

$$\mathcal{L}_{\mathcal{K}}[u](x, t) = \int_{\bar{\Omega}} K(|x - y|) (u_{NL}(y, t) - u_{NL}(x, t)) dy$$

- We represent the kernel K as a linear combination of Bernstein basis polynomials:

$$K(|x - y|) = \sum_{m=0}^M \frac{C_m}{\delta^{d+2}} B_{m,M} \left(\left| \frac{x - y}{\delta} \right| \right)$$

where

$$B_{m,M}(x) = \binom{M}{m} x^m (1 - x)^{M-m}$$

for $0 \leq x \leq 1$ and $C_m \in \mathbb{R}$.

Nonlocal Operator Regression (NOR)

- From now on, denote K_C and $u_{NL,C}$ as the nonlocal kernel and nonlocal solution corresponding to a specific set of kernel parameters $C := \{C_m\}_{m=0}^M$.
- Using central-difference scheme in time

$$u_{NL,C}(x_i, t_{n+1}) = 2u_{NL,C}(x_i, t_n) - u_{NL,C}(x_i, t_n - 1) + dt^2 \left(\mathcal{L}_{\mathcal{K}_C, h}[u_{NL,C}](x_i, t_n) + f(x_i, t_n) \right)$$

- Learning procedure: minimize the following cost function with regularization

$$\sum_{s=1}^S \frac{\|u_{NL,C}^s - u_{DNS}^s\|_{l_2(\Omega \times [0, t])}^2}{\|u_{NL,C}^s\|_{l_2(\Omega \times [0, t])}^2} + \lambda \|C\|_{l_2}^2$$

- $u_{NL,C}$ satisfies the central difference equation and K_C satisfies physical-based constraints[1].



- Note1: NOR is a purely deterministic approach. It cannot quantify the uncertainty in using a nonlocal model to describe material response.
- Note2: We can use this preliminary result as prior knowledge in Bayesian inference.

Bayesian Inference: likelihood

- In this work, we take ground truth u_{HF} as the numerical solution u_{DNS} generated from the DNS solver.
- Denote the nonlocal solution corresponding to a specific set of kernel parameters $\{C_m\}$ as $u_{\{C_m\}}(x, t)$, model the error between the nonlocal solution and the ground truth as a Gaussian variable

$$u_{DNS}^s(x, t) = u_{NL,C}^s(x, t) + \epsilon(x, t), \epsilon \sim N(0, \sigma_s^2)$$

where $\sigma_s = \sigma \left\| u_{NL,C}^s \right\|_{l_2(\Omega \times [0,t])}^2$, σ is a constant independent of s .

- The negative log-likelihood reads

$$\sum_{s=1}^S \left(\frac{\left\| u_{NL,C}^s - u_{DNS}^s \right\|_{l_2(\Omega \times [0,T])}^2}{\left\| u_{NL,C}^s \right\|_{l_2(\Omega \times [0,t])}^2} + N \log(\sigma \left\| u_{NL,C}^s \right\|_{l_2(\Omega \times [0,t])}^2) \right)$$

Bayesian Inference: prior

- Assume $C = \{C_m\}$ has a Gaussian prior, i.e. $C_m, m = 0, \dots, M$ are independent Gaussian variables with $C_m \sim N\left(C_{0,m}, \frac{\sigma_0^2}{\lambda}\right)$. Here $\{C_{0,m}\}$ is the set of parameters of the kernel learnt from the deterministic NOR, and σ_0 is the standard deviation calculated using the current nonlocal solution associated with $\{C_m\}$.

- The negative log-prior reads

$$\lambda \frac{\|C - C_0\|^2}{2\sigma_0^2}$$

- Combining the negative log-likelihood and the negative log-prior, we have the negative log posterior

$$\sum_{s=1}^S \left(\frac{\|u_{NL,C}^s - u_{DNS}^s\|_{l_2(\Omega \times [0,T])}^2}{\|u_{NL,C}^s\|_{l_2(\Omega \times [0,t])}^2} + N \log(\sigma \|u_{NL,C}^s\|_{l_2(\Omega \times [0,t])}^2) \right) + \lambda \frac{\|C - C_0\|_{l_2}^2}{2\sigma_0^2}$$

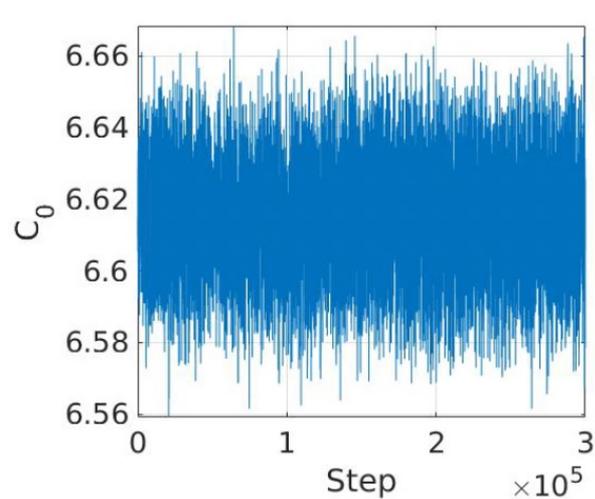
MCMC: algorithm and evaluation

- We adopt the **adaptive MCMC** algorithm proposed in [4]. In practice, we couple **PyUQTK** [5] with our nonlocal model to apply the MCMC algorithm.
- Inspect the **trace plot visually** and evaluate the **acceptance rate** (the percentage of the accepted proposal states) of MCMC.
- **Effective sample size (ESS)** is a guidance on sub-sampling, also a criterion which could tell us how long an MCMC chain we need. In this work, the ESS is computed following the multivariate extension defined in [6].
- Plot the **probability density function (PDF)** based on equally spaced samples of the chain, where the number of the subsamples is determined by the ESS.

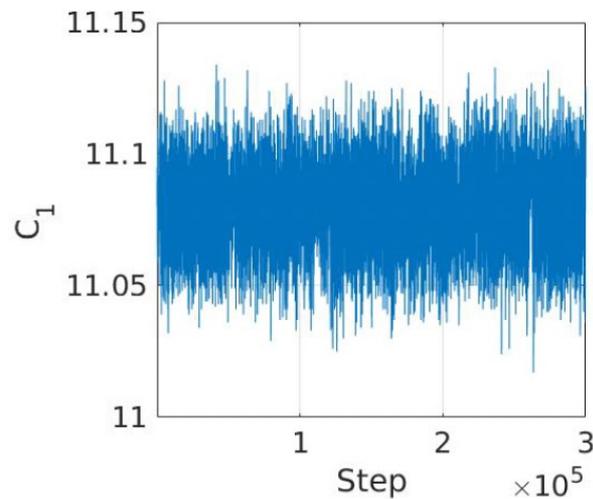
- [4] Haario, H., Saksman, E., & Tamminen, J. (1999). Adaptive proposal distribution for random walk Metropolis algorithm. *Computational statistics*, 14(3), 375-395.
- [5] Sargsyan, K., Saffa, C., Boll, L., Johnston, K., Khalil, M., Chowdhary, K., ... & Debusschere, B. (2022). UQTK Version 3.1. 2 User Manual (No. SAND2022-0377). Sandia National Lab.(SNL-NM), Albuquerque, NM (United States).
- [6] Vats, D., Flegal, J. M., & Jones, G. L. (2019). Multivariate output analysis for Markov chain Monte Carlo. *Biometrika*, 106(2), 321-337.

Results on Disordered Material

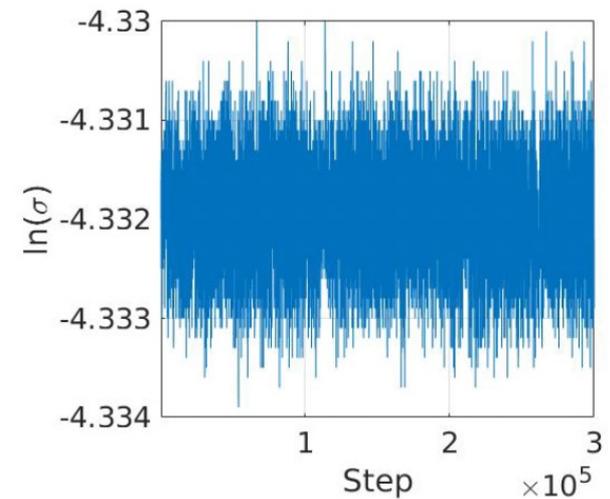
MCMC Trace Plot



(a) Parameter 1



(b) Parameter 2



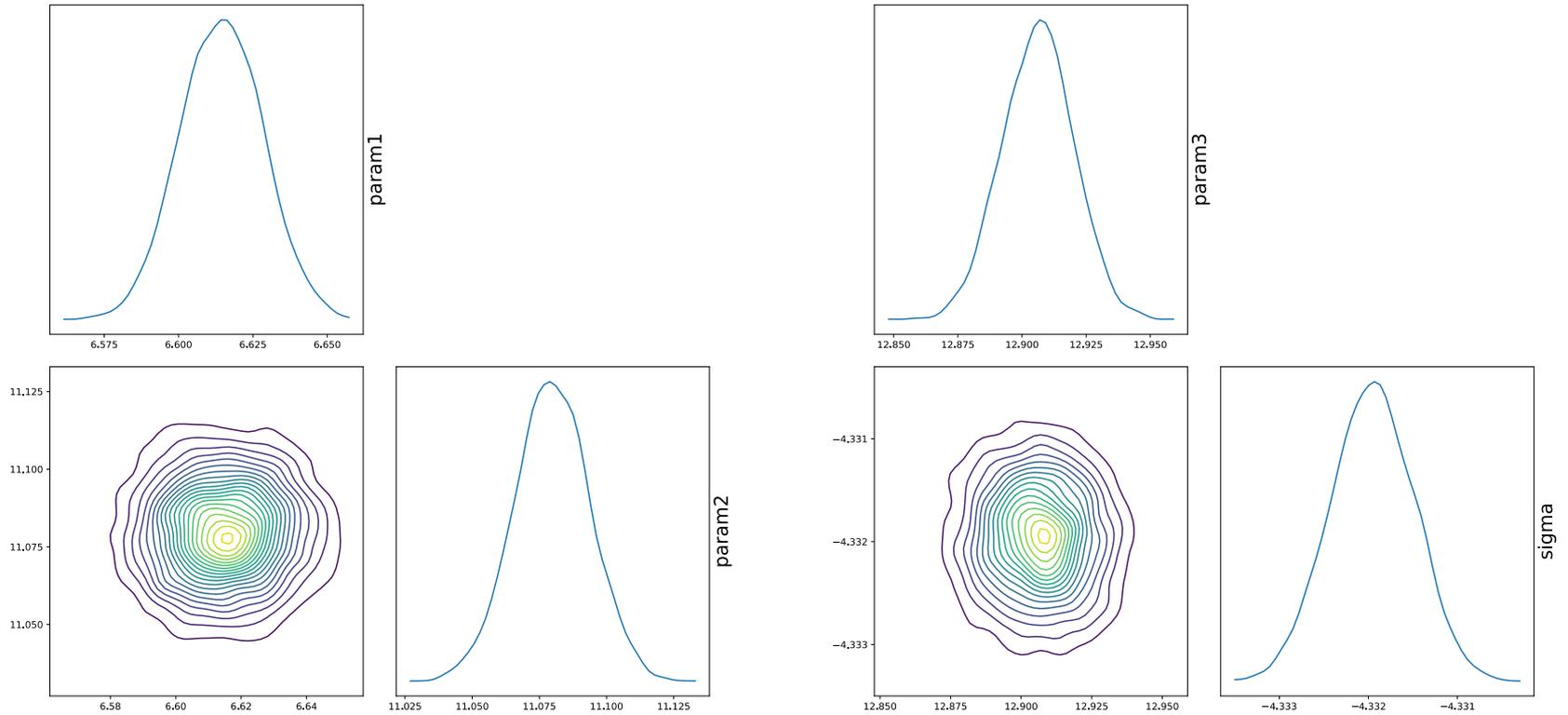
(c) $\ln(\sigma)$

- Chain length = 300,000, acceptance rate = 0.31, ESS=4212.

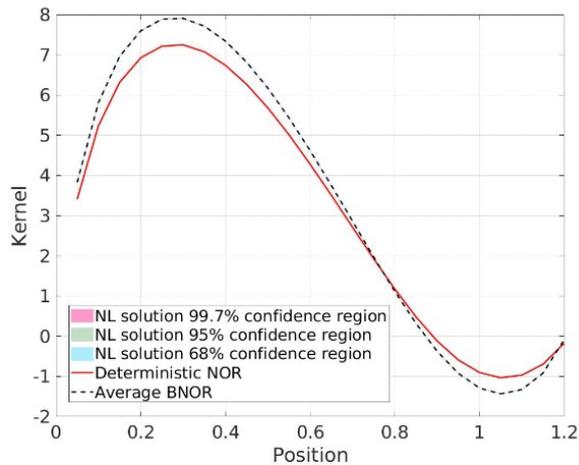
Results on Disordered Material



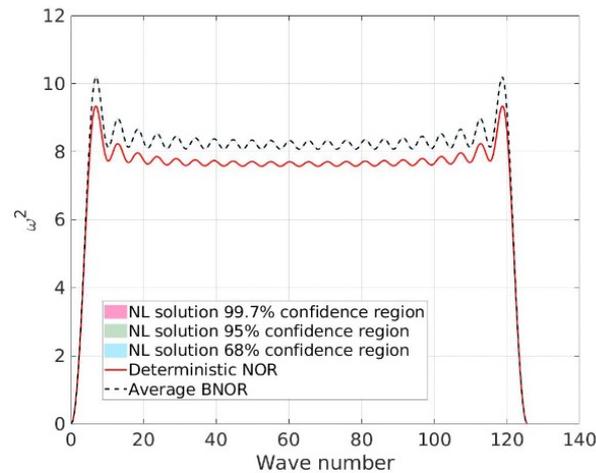
Density Plot



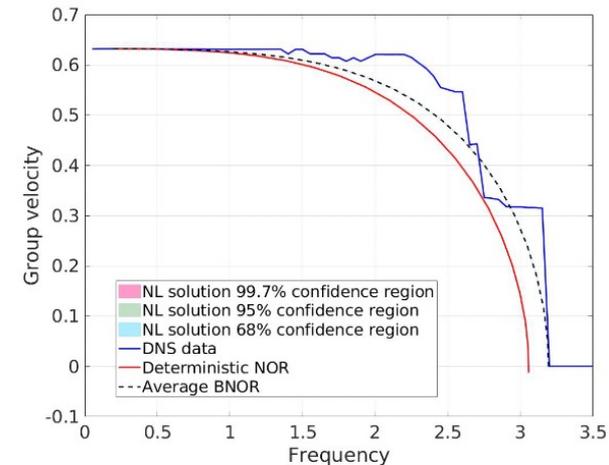
Results on Disordered Material



(a) Optimal kernel



(b) Dispersion curve



(c) Group velocity

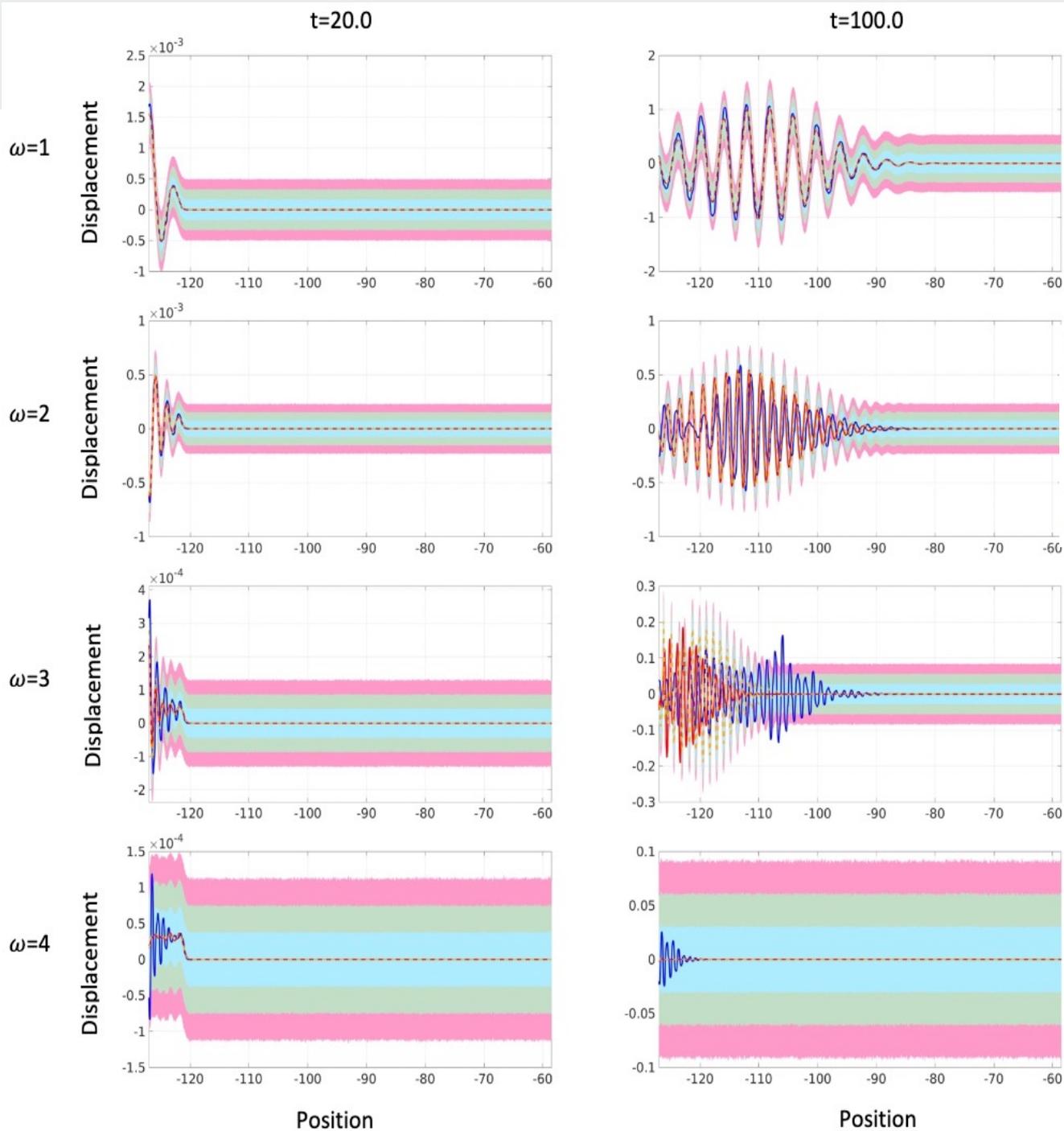
- Positive dispersion curves indicate physically stable material models
- The group velocity matches the one computed with DNS both at low frequencies and at the band stop.

Validation on wave packet through a disordered material



- Validation on the wave packet with $\omega=2.0$ at $t=20.0s$ (short time prediction) and $t=100.0s$ (long time prediction). 4 different frequencies ($\omega=1.0,2.0,3.0,4.0$) are considered. For each sample in ESS, we calculate the nonlocal kernel C and the corresponding nonlocal solution $u_{NL,C}$. Then plot the 68%-95%-99.7% confidence region for $u_{NL,C} + \epsilon$, where

$$\epsilon \sim \mathcal{N} \left(0, \sigma^2 \left\| u_{NL,C}(x, t) \right\|_{l_2([-b,b] \times [0,T])}^2 \right)$$



Conclusion



We proposed a [Bayesian nonlocal operator regression technique](#), which

- captures the nonlocal constitutive law that embeds the material properties,
- characterizes the uncertainty in using a nonlocal model for predicting wave propagation through heterogeneous materials,
- Provides models that reproduce high-fidelity data that are substantially different from the training data

Future work

- Noisy data,
- More sophisticated error model, etc.

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Thank you! Q & A