



Deconvolution strategies for efficient parametric variance estimation in stochastic media transport problems

Gianluca Geraci and Aaron J. Olson

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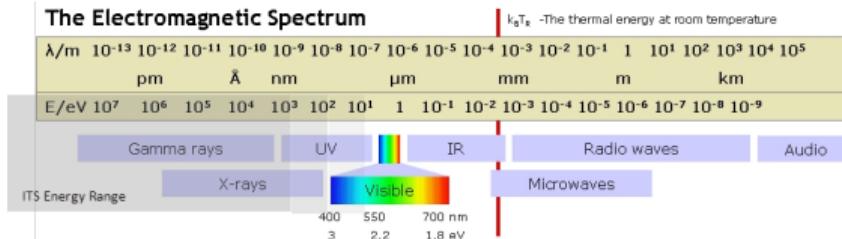
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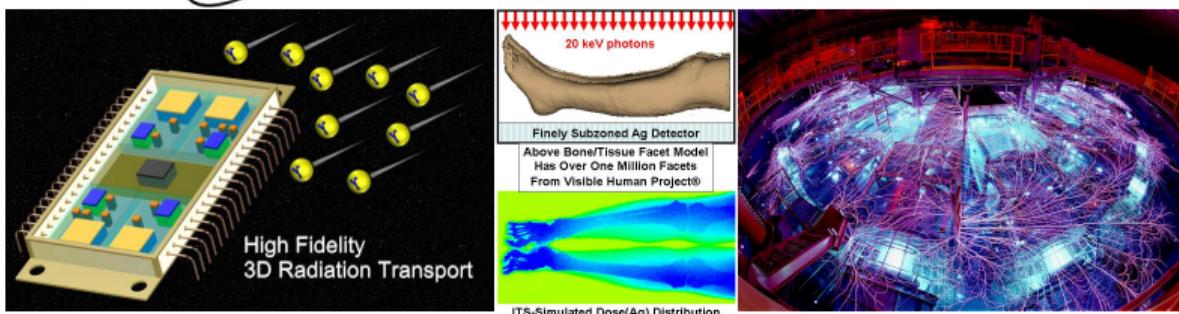
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UNCERTAINTY QUANTIFICATION FOR RADIATION TRANSPORT

CONTEXT AND CHALLENGES



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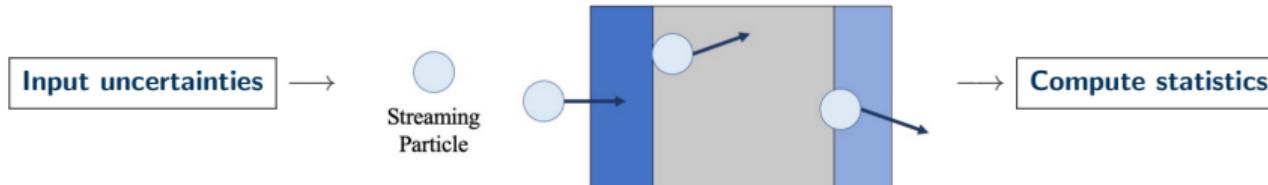
Figures courtesy of Brian Franke and Shawn Pautz

High-fidelity radiation transport modeling and simulation with HPC

- **Severe** simulations **budget constraints**
- **Significant UQ parameter dimensionality** driven by model complexity

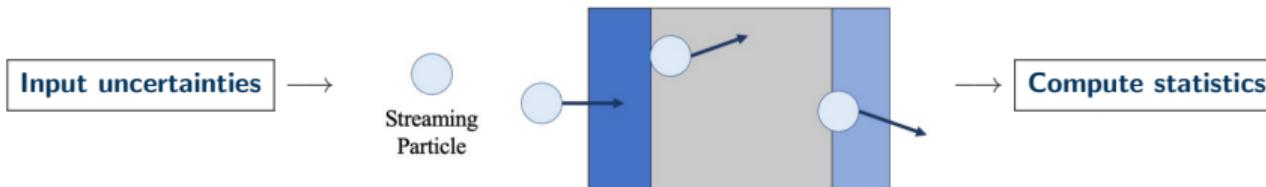
UNCERTAINTY QUANTIFICATION FOR MONTE CARLO RADIATION TRANSPORT

THE CHALLENGE OF UNDER-RESOLVED COMPUTATIONS

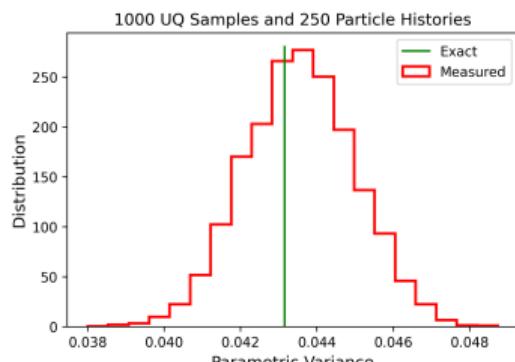
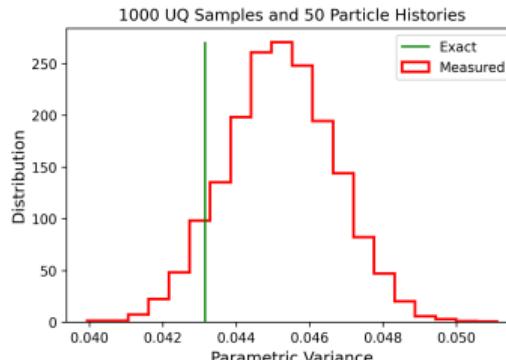
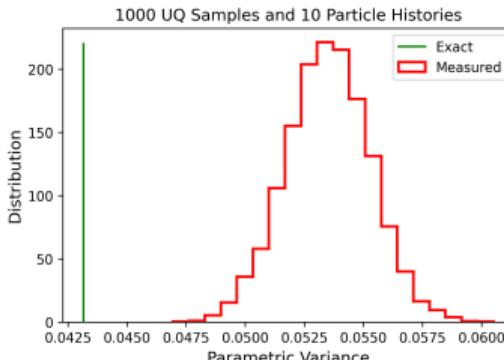


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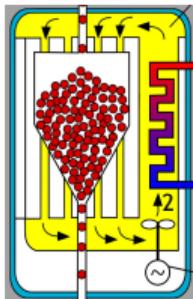
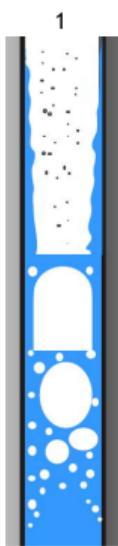


How much should your MC RT be resolved?

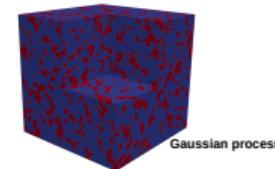
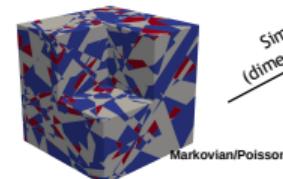
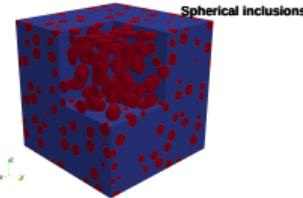


STOCHASTIC MEDIA IN RADIATION TRANSPORT

FROM REAL-WORLD APPLICATIONS TO NUMERICS



REAL WORLD



3D COMPUTATIONAL MODELS



1D Markovian geometry - Material type forms continuous Markov process

Simplify (dimensionality)

Simplify (closed-form solutions)



1D "Cell" geometry - Cell boundaries uniform, but material fill random

Real-world examples

- Two-phase flow in **Boiling Water Reactor** nuclear power coolant (1)
- **Pebble distribution** in Pebble-Bed nuclear power reactors (2)
- Distribution of **TRISO fuel** particles in Pebble-Bed pebble (3)
- Raleigh-Taylor **instabilities** in Inertial Confinement Fusion reactors
- **Accident scenarios** in various nuclear power reactor cores



Contributions of this talk:

- How can we obtain accurate parametric variance estimations given a **limited** computational budget for **particle histories**?
- How can we extend this to account for **limited stochastic media realizations**?

Bonus point

- Designed a verification test case, *i.e.* closed form solution for necessary statistics



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Mathematical Framework



What are the sources of variability/uncertainty?

- **Uncertain parameters**, e.g. cross sections: ξ
- **MC RT (internal) randomness**: η
- **Material arrangements/realizations**: ω

What is the mathematical counterpart of a MC RT simulation?

- Particle histories can be interpreted as elementary events: $f = f(\xi, \omega, \eta)$
- MC RT QoI can be interpreted as an average of f over the histories with fixed UQ parameters and material arrangements

$$Q(\xi, \omega) = \mathbb{E}_\eta [f(\xi, \omega, \eta)] \stackrel{MC\ RT}{\approx} \frac{1}{N_\eta} \sum_{j=1}^{N_\eta} f(\xi, \omega, \eta^{(j)}) \stackrel{\text{def}}{=} \tilde{Q}_{N_\eta}(\xi, \omega)$$

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Problem formulation

We focus on statistics over **material arrangements**

$$\mathbb{P}_{\mathbb{E}}(\xi) \stackrel{\text{def}}{=} \mathbb{E}_{\omega} [Q(\xi, \omega)],$$

while the UQ analysis aims at evaluating the parametric variance, i.e. $\text{Var}_{\xi} [\mathbb{P}_{\mathbb{E}}(\xi)]$

Brute force approach

$$\text{Var}_{\xi} [\mathbb{P}_{\mathbb{E}}(\xi)] \approx \text{Var}_{\xi} \left[\tilde{\mathbb{P}}_{N_{\omega}}^{\mathbb{E}}(\xi) \right],$$

where

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UNDERSTANDING THE VARIANCE CONTRIBUTIONS FROM THE LAW-OF-TOTAL VARIANCE TO DECONVOLUTION

Definitions
$$\begin{cases} \tilde{\mathbb{P}}_{N_\omega}^{\mathbb{E}}(\xi) \stackrel{\text{def}}{=} \frac{1}{N_\omega} \sum_{k=1}^{N_\omega} \tilde{Q}_{N_\eta}(\xi, \omega^{(k)}), \\ \tilde{Q}_{N_\eta}(\xi, \omega) \stackrel{\text{def}}{=} \frac{1}{N_\eta} \sum_{j=1}^{N_\eta} f(\xi, \omega, \eta^{(j)}) \end{cases}$$

Let's start by 'decomposing' the **variance contribution** (law-of-total-variance)

$$\text{Var} \left[\tilde{\mathbb{P}}_{N_\omega}^{\mathbb{E}} \right] = \text{Var}_\xi \left[\mathbb{E}_{\eta, \omega} \left[\tilde{\mathbb{P}}_{N_\omega}^{\mathbb{E}} \right] \right] + \mathbb{E}_\xi \left[\text{Var}_{\eta, \omega} \left[\tilde{\mathbb{P}}_{N_\omega}^{\mathbb{E}} \right] \right]$$

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- Law-of-total-variance can be applied recursively

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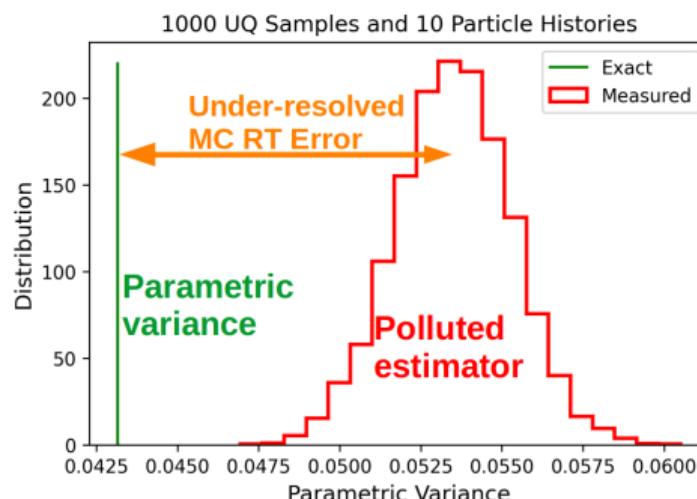
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VARIANCE DECONVOLUTION

NOTES ON PRACTICAL IMPLEMENTATION



Variance deconvolution recipe:

1 Measure the **polluted variance**

$$\mathbb{Var} \left[\tilde{\mathbb{P}}_{N_\omega}^{\mathbb{E}} \right] \rightarrow \tilde{S}^2 = \frac{1}{N_\xi - 1} \sum_{i=1}^{N_\xi} \left(\tilde{\mathbb{P}}_{N_\omega}^{\mathbb{E}}(\xi^{(i)}) - \frac{1}{N_\xi} \sum_{q=1}^{N_\xi} \tilde{\mathbb{P}}_{N_\omega}^{\mathbb{E}}(\xi^{(q)}) \right)^2$$

2 Measure the **under-resolved statistics**

$$\begin{aligned} & \mathbb{E}_\xi \left[\frac{\mathbb{Var}_\omega [Q(\xi, \omega)]}{N_\omega} \right] + \mathbb{E}_\xi \left[\frac{\mathbb{E}_\omega \left[\sigma_\eta^2(\xi, \omega) \right]}{N_\omega N_\eta} \right] \\ & \rightarrow \Delta S^2 = \frac{1}{N_\omega N_\xi} \sum_{i=1}^{N_\xi} \left[\frac{1}{N_\omega} \sum_{k=1}^{N_\omega} \left(\tilde{Q}_{N_\eta}(\xi^{(i)}, \omega^{(k)}) - \frac{1}{N_\omega} \sum_{q=1}^{N_\omega} \tilde{Q}_{N_\eta}(\xi^{(i)}, \omega^{(q)}) \right)^2 \right] \end{aligned}$$

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$$\mathbb{Var}_\xi [\mathbb{P}_{\mathbb{E}}] \rightarrow S^2 = \tilde{S}^2 - \Delta S^2$$

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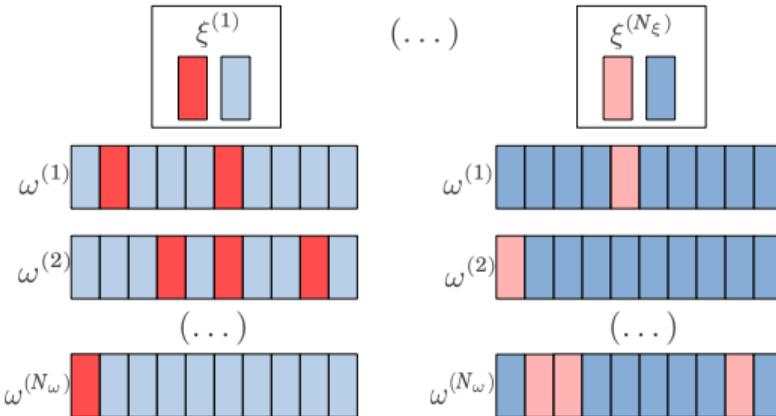
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Numerical Results



Material	$\Sigma_{t,m}^0 [\text{cm}^{-1}]$	$\Sigma_{t,m}^\Delta [\text{cm}^{-1}]$	p_m
A	1.0	0.95	0.05
B	0.4	0.25	0.95

TABLE: Uncertain cross sections properties

- 1D slab, neutral particle, **absorption-only** mono-energetic steady state radiation transport
- Normally incident beam with unitary magnitude
- Random cross sections** ($m = A, B$): $\Sigma_{t,m}(\xi_m) = \Sigma_{t,m}^0 + \Sigma_{t,m}^\Delta \xi_m$, where $\xi_A, \xi_B \sim \mathcal{U}(-1, 1)$
- Total number of **material sections**: $N_{tot} = 10$ (with $\Delta x = 0.15 \text{ cm}$)
- The **QoI** is the **transmittance**

Transmittance: $T(\xi, \omega) = \exp [-\tau(\xi, \omega)]$, where

Slab optical thickness: $\tau(\xi, \omega) = \Delta x (N_A(\omega) \Sigma_{t,A}(\xi_A) + N_B(\omega) \Sigma_{t,B}(\xi_B))$

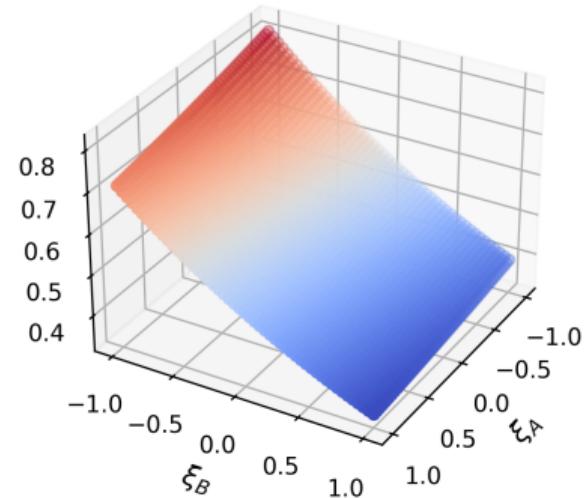
Material distribution: $N_A(\omega) \sim \mathcal{B}(N_{tot}, P_A)$, where $N_A(\omega) + N_B(\omega) = N_{tot}$

$$\mathbb{P}_E(\xi) = \mathbb{E}_\omega [T(\xi, \omega)]$$

$$= \exp \left[-N_{tot} \Delta x \Sigma_{t,B}(\xi_B) \right] \sum_{x=0}^{N_{tot}} B_\omega(x) \exp \left[-x \Delta x (\Sigma_{t,A}^\Delta \xi_A - \Sigma_{t,B}^\Delta \xi_B) \right]$$

$$B_\omega(x) \stackrel{\text{def}}{=} \frac{N_{tot}!}{x! (N_{tot} - x)!} p_A^x (1 - p_A)^{(N_{tot} - x)} \exp \left[-2x \Delta x (\Sigma_{t,A}^0 - \Sigma_{t,B}^0) \right]$$

Response $\mathbb{P}_E(\xi)$



1D TRANSPORT IN STOCHASTIC MEDIA WITH UNCERTAIN PROPERTIES

EXACT SOLUTION FOR STATISTICS – PARAMETRIC VARIANCE



$$\text{Var}_\xi [\mathbb{P}_E] = \mathbb{E}_\xi [\mathbb{P}_E(\xi)^2] - (\mathbb{E}_\xi [\mathbb{P}_E(\xi)])^2$$

where

$$\mathbb{E}_\xi [\mathbb{P}_E(\xi)] = \exp \left[-N_{tot} \Delta x \Sigma_{t,B}^0 \right] \left(\sum_{x=0}^{N_{tot}} B_\omega(x)(x) \frac{\sinh [x \Delta x \Sigma_{t,A}^\Delta]}{x \Delta x \Sigma_{t,A}^\Delta} \frac{\sinh [N_{tot} \Delta x \Sigma_{t,B}^\Delta \left(1 - \frac{x}{N_{tot}} \right) \Delta x \Sigma_{t,A}^\Delta]}{N_{tot} \Delta x \Sigma_{t,B}^\Delta \left(1 - \frac{x}{N_{tot}} \right) \Delta x \Sigma_{t,A}^\Delta} \right),$$

and

$$\begin{aligned} \mathbb{E}_\xi [\mathbb{P}_E(\xi)^2] &= \exp \left[-2N_{tot} \Delta x \Sigma_{t,B}^0 \right] \left(\sum_{x=0}^{N_{tot}} B_\omega(x)^2 \frac{\sinh [2x \Delta x \Sigma_{t,A}^\Delta]}{2x \Delta x \Sigma_{t,A}^\Delta} \frac{\sinh [2N_{tot} \Delta x \Sigma_{t,B}^\Delta \left(1 - \frac{x}{N_{tot}} \right)]}{2N_{tot} \Delta x \Sigma_{t,B}^\Delta \left(1 - \frac{x}{N_{tot}} \right)} \right. \\ &\quad \left. + 2 \sum_{x=0}^{N_{tot}} \sum_{y=x+1}^{N_{tot}} B_\omega(x) B_\omega(y) \frac{\sinh [(x+y) \Delta x \Sigma_{t,A}^\Delta]}{(x+y) \Delta x \Sigma_{t,A}^\Delta} \frac{\sinh [2N_{tot} \Delta x \Sigma_{t,B}^\Delta \left(1 - \frac{x+y}{2N_{tot}} \right)]}{2N_{tot} \Delta x \Sigma_{t,B}^\Delta \left(1 - \frac{x+y}{2N_{tot}} \right)} \right) \end{aligned}$$

N_η	N_ω	\mathcal{C}	$MSE(S^2)$	$MSE(\tilde{S}^2)$	Diff. [%]
2	2	400	1.7313e-4	3.6084e-3	95.20
	4	800	4.0763e-5	9.1886e-4	95.56
	8	1600	1.3829e-5	2.3302e-4	94.06
	16	3200	6.3738e-6	6.0617e-5	89.48
	32	6400	3.6007e-6	1.7405e-5	79.31
4	2	800	5.1799e-5	9.4397e-4	94.51
	4	1600	1.4986e-5	2.4136e-4	93.79
	8	3200	6.6362e-6	6.3906e-5	89.61
	16	6400	3.7136e-6	1.7990e-5	79.35
	32	12800	2.5651e-6	5.9639e-6	56.99
8	2	1600	1.9071e-5	2.5549e-4	92.53
	4	3200	7.0521e-6	6.6686e-5	89.42
	8	6400	3.8521e-6	1.9329e-5	80.07
	16	12800	2.5911e-6	6.2441e-6	58.50
	32	25600	2.1348e-6	3.0310e-6	29.57

N_η	N_ω	\mathcal{C}	$MSE(S^2)$	$MSE(\tilde{S}^2)$	Diff. [%]
16	2	3200	8.2811e-6	7.4708e-5	88.91
	4	6400	3.9833e-6	2.0705e-5	80.76
	8	12800	2.6616e-6	6.7252e-6	60.42
	16	25600	2.1034e-6	3.0990e-6	32.12
	32	51200	1.8296e-6	2.0513e-6	10.81
32	2	6400	4.7911e-6	2.5362e-5	81.11
	4	12800	2.8784e-6	8.0892e-6	64.42
	8	25600	2.0873e-6	3.3810e-6	38.26
	16	51200	1.8619e-6	2.1960e-6	15.21
	32	102400	1.7596e-6	1.8532e-6	5.05

TABLE: MSE (based on 5,000 repetitions) for the parametric variance estimator, with (S^2) and without (\tilde{S}^2) deconvolution, $N_\xi = 100$ UQ samples and several choices of random material realizations N_ω and particle histories N_η . The total cost for each estimator is $\mathcal{C} = N_\xi \times N_\omega \times N_\eta$.

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- For fixed N_ξ and N_η , **MSE decreases** if N_ω increases

N_η	N_ω	\mathcal{C}	$\text{MSE}(S^2)$	$\text{MSE}(\tilde{S}^2)$	Diff. [%]
2	2	400	1.7313e-4	3.6084e-3	95.20
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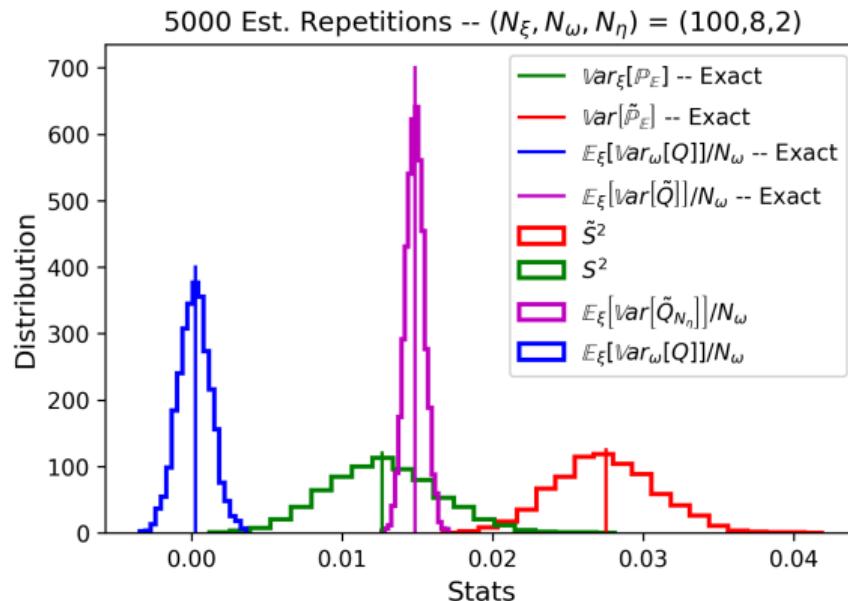
- For a given cost $\mathcal{C} = N_\xi \times N_\eta \times N_\omega$, **MSE increases** if N_η increases

N_η	N_ω	\mathcal{C}	MSE(S^2)	MSE(\tilde{S}^2)	Diff. [%]
2	2	400	1.7313e-4	3.6084e-3	95.20
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	32	6400	3.6007e-6	1.7405e-5	79.31
	64	12800	1.8003e-6	4.3730e-6	73.00
4	2	800	5.1799e-5	9.4397e-4	94.51
	4	1600	1.4986e-5	2.4136e-4	93.79
	8	3200	6.6362e-6	6.3906e-5	89.61
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	32	25600	2.1348e-6	3.0310e-6	29.57
	64	51200	1.0705e-6	1.5150e-6	25.00

N_η	N_ω	\mathcal{C}	MSE(S^2)	MSE(\tilde{S}^2)	Diff. [%]
16	2	3200	8.2811e-6	7.4708e-5	88.91
	4	6400	3.9833e-6	2.0705e-5	80.76
	8	12800	2.6616e-6	6.7252e-6	60.42
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- The difference between S^2 and S **decreases** if $\mathcal{C} = N_\xi \times N_\omega \times N_\eta$ increases



$$\text{Var} \left[\tilde{P}_{N_\omega}^E \right] = \text{Var}_\xi \left[\mathbb{P}_E \right] + \mathbb{E}_\xi \left[\frac{\text{Var}_\omega \left[Q(\xi, \omega) \right]}{N_\omega} \right] + \underbrace{\mathbb{E}_\xi \left[\frac{\mathbb{E}_\omega \left[\frac{\sigma_\eta^2(\xi, \omega)}{N_\eta} \right]}{N_\omega} \right]}_{\mathbb{E}_\xi \left[\frac{\text{Var} \left[\tilde{Q}_{N_\eta}(\xi, \omega) \right]}{N_\omega} \right]}$$

Closing remarks



Conclusions

- Uncertainty Quantification for MC RT requires a characterization of **all sources of variability**
- In the presence of stochastic media **removing the MC RT variability is unattainable**
- The use of the **law-of-total-variance** provides a rigorous approach for efficient and accurate statistics evaluation with under-resolved MC RT computations → *variance deconvolution*

Future work

- Extension to **higher-order moments** over the stochastic media realizations, *i.e.* $\mathbb{V}ar_{\omega}[T]$
- **Optimal sample allocation** strategies can be built (*in principle*) from pilot computations
- **Multifidelity** sampling approaches can be built from/for this framework¹
- Recursive use of variance deconvolution can allow for **statistics beyond variance**, *e.g.* conditional variances (Sobol' indices)

Link with other contributions at ANS

- *Numerical Investigation on the Performance of a Variance Deconvolution Estimator*,
Presented by: Kayla Clements – Tomorrow in the Uncertainty Quantification and Machine Learning session

¹G. Geraci and A. J. Olson. "Exploration of Multifidelity UQ Methods for Monte Carlo Radiation Applications in Stochastic Media". In: *SIAM Conference on Uncertainty Quantification (SIAM22)*. 2022.

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