

24 **Abstract**

25 The lack of farmers' willingness to grow perennial bioenergy crops (PBCs) presents
26 a critical barrier to the emergence of cellulosic biofuel production. The willingness relies
27 on a complex network of economic, environmental, and social drivers, among which the
28 influence of social factors (e.g., the influence of neighborhood, community, and
29 communication) is less understood. This study addresses this knowledge gap via a survey
30 analysis of midwestern farmers. The survey data are analyzed through ordinary least
31 square regression and structural equation model, which together investigate the individual
32 and interactive impacts of multiple factors on farmers' decisions to adopt PBCs. Based on
33 a farm-scale analysis, six statistically significant predictors of farmer willingness to grow
34 PBCs are identified: perception of PBCs' environment benefits, education level,
35 willingness to take risks, familiarity with PBCs, portion of peers already growing PBCs,
36 and support of biorefineries locating in the local community. Among these, the latter
37 three predictors are social support variables. It is found that familiarity with the crops is
38 the most significant predictor of willingness; familiarity is also an important intermediate
39 variable that mediates the influence of many other predictors. In addition, peer adoption
40 can both directly and indirectly affect willingness *via* its influence on familiarity. These
41 findings suggest that it is a pressing need to improve farmers' knowledge of PBCs to
42 promote the adoption of such crops.

43

44 **Keywords:**

45 Cellulosic bioeconomy; Farm-scale; Social support; Structural equation model; Survey;
46 Willingness.

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49 **1. Introduction**

50 Bioenergy is a crucial field that addresses the world's progress toward renewable
51 energy, economic development, and environmental sustainability [1–3]. While existing
52 biofuels are mostly produced from food crops such as corn and soybean using hydrolysis,

53 fermentation, or transesterification, many believe cellulosic biofuels produced from
54 perennial bioenergy crops (PBCs) (such as Miscanthus, *Miscanthus x giganteus*, and
55 switchgrass, *Panicum virgatum*) using biochemical or thermochemical technologies are
56 more promising given the high yield of the PBCs and their various environmental and
57 ecological benefits [4,5]. In addition, PBCs can be grown on marginal lands that are not
58 suitable for food crop production, hence ameliorating the conflict between food and
59 energy for corn-based and soybean-based biofuels [6,7]. Despite the benefits highlighted
60 above, the PBC development in the United States (U.S.) is still limited to small-scale
61 trials mainly for research [8]. As a result, the production of cellulosic biofuel (the primary
62 use of PBCs) in 2020 only met 5.9% of the target of 10.5 billion gallons, the mandate of
63 the Renewable Fuel Standard (RFS) [9]. Therefore, it is of great importance to improve
64 understanding of potential drivers, as well as barriers of PBC adoption, which could be
65 used to design more effective promotion policies for PBCs.

66 The adoption of PBCs relies on decisions at the farm scale. Economic factors,
67 including the availability of the biomass market, expected return of PBCs, contracts with
68 biorefineries, farmland ownership, etc., are identified as the most important drivers of
69 PBCs adoption [10–14]. Fewell et al. [14] surveyed Kansas farmers for their switchgrass
70 adoption decisions under contract through a choice experiment approach. Their results
71 suggested that uncertainties in PBC production contribute to farmers' lack of willingness
72 to adopt such crops, and the profitability of PBCs must be compatible with other food
73 crops to encourage farmer adoption. Through a similar approach, Khanna et al. [15]
74 found that farmers' willingness to adopt PBCs is negatively correlated with their discount
75 rate and the energy crops' upfront establishment costs, which are the additional crop-
76 specific investment for PBCs. Other non-economic factors, mostly behavior-related
77 factors such as farmer's age, education level, attitude toward the environment, knowledge
78 of PBCs, etc., also play important roles in PBC adoption [11,16–19]. In addition, large-
79 scale adoption of PBCs also requires additional infrastructure for transportation, water,
80 and access to biorefineries [1,20,21].

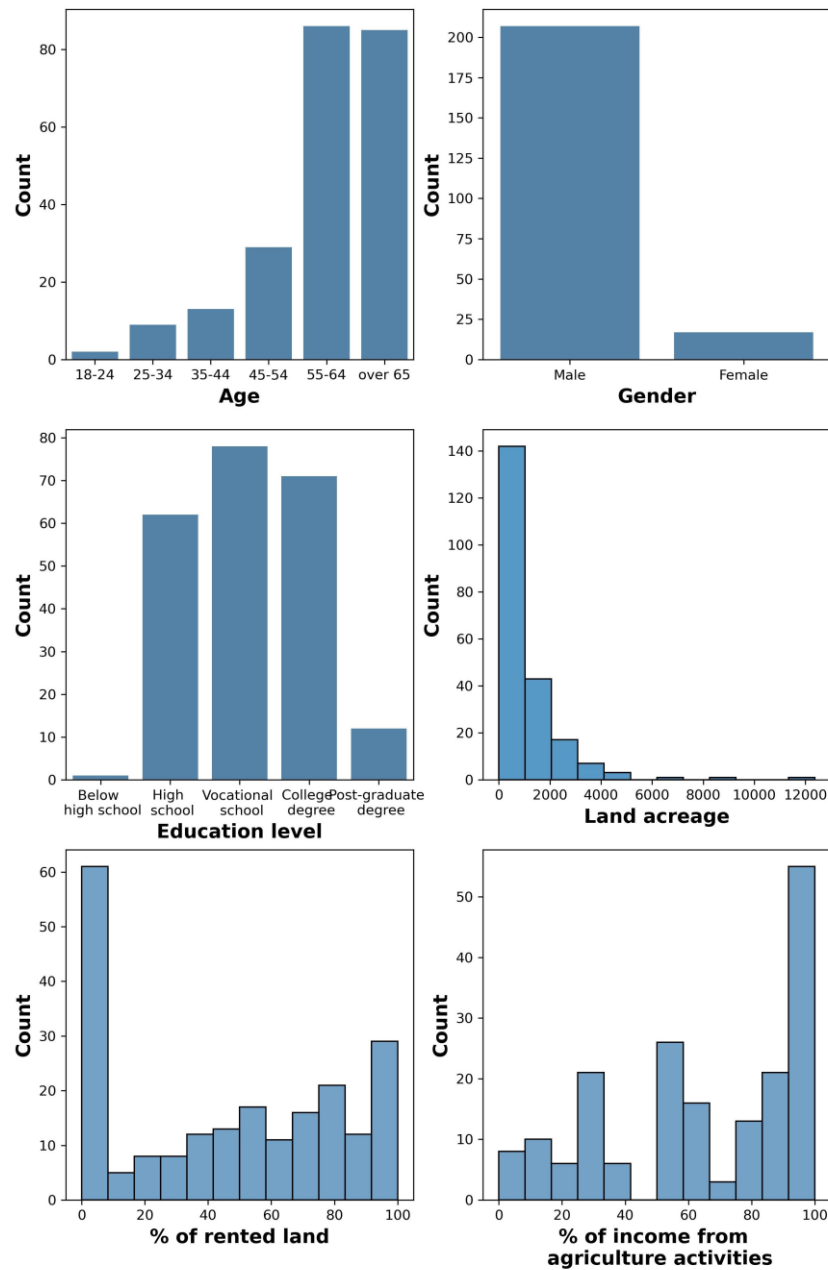
81 Despite many studies on PBCs adoption, few investigated the influence of social
82 factors, such as the information and support from their neighbors, local communities, and
83 the society. The factors could be explained as subjective norms following the theory of

84 planned behavior [24], and their impacts have been proven essential for the adoption of
85 novel land use options [22,23]. In general, farmers tend to mimic their neighbors'
86 practices as a result of peer pressure and information diffusion [25–27]; they make land
87 use decisions under specific community and social contexts such as mass media
88 propaganda, access to urban areas, and the proximity of private enterprises and public
89 services [23,28]. Yet, there still needs additional evidences for the social support factor
90 impact as mentioned above, specifically on PBCs adoption. Our study aims to add the
91 evidence by analyzing data from a recent land use survey.

92 Yang et al. [11] reported a survey on Midwestern farmers' PBC adoption decisions.
93 The survey is designed with a balance of economic, behavioral, and social support factors
94 for land use decisions, including farmers' familiarity with PBCs, their neighbors' land
95 use decisions, and community support. Their analysis revealed that economic factors
96 such as the expected return of PBCs and marginal land availability are the most relevant
97 drivers. The analysis also found social related factors such as information sources and
98 concern on social related barriers for growing PBC as potential indirect predictors.
99 However, as mentioned in Yang et al. [11], such analysis for social factors is exploratory
100 and requires further validation based on more rigorous statistical analyses. Further, the
101 analysis in Yang et al. [11] focused on adoption decisions under choice experiment
102 scenarios, and social factors are not explicitly represented in the scenario design. Thus, it
103 is not clear how social factors may affect farmer willingness to adopt PBCs. Following
104 Yang et al. [11] and using the same survey, the current study specifically addresses the
105 role of social support on the willingness to grow PBCs in general, without the influence
106 of any choice experiment scenario. The collected data are analyzed with ordinary least
107 square regression to investigate how PBC adoption decisions are influenced by those
108 economic, behavioral, and social support factors. Following that, this study further
109 applies a structural equation model to investigate how those factors interactively
110 influence PBCs adoption willingness via a hierarchical structure. Findings from this study
111 would be beneficial for the understanding and promotion of PBC adoption at the farm
112 scale.

113

114 **2. Survey and data preprocessing**



115
116

Figure 1. Histograms of respondents' socio-economic variables.

117 The survey, conducted in late 2019 and early 2020, was sent to 2,500 Midwestern
 118 farmers identified from the Farm Market ID database [29]. To ensure an appropriate
 119 coverage of farmers with small and large farm sizes, the 2,500 farmers were sampled
 120 from the Farm Market ID database using a stratified random sampling technique [30]
 121 based on farm size. We collected 242 survey responses from the 2,500 farmers, resulting

122 in a response rate of 9.7%. 17 of the 242 responses contain missing values and were
123 excluded in the analysis of Yang et al. [11]. This study fills the missing values using a
124 KNN model trained with the original sample data and thus uses the full set of 242
125 samples for the analysis.

126 Figure 1 shows histograms respondents' socio-economic variables. The respondents
127 are mostly male, aged farmers (85.5% male and 70.7% over 55), and most of them have a
128 high school/vocational school/college degree. Over 75% of respondents work on farms
129 which are smaller than 2000 acres, and very few on large farms over 10,000 acres. About
130 60% of the respondents own all the farmland they operate; about 19% of them depend
131 heavily on rented land (with >80% of their land rented). About 68% of the respondents
132 have over 50% of their income from agricultural activities, while the rest are less
133 dependent on agricultural incomes.

134 The survey covers questions on farmers' willingness to adopt three PBCs –
135 Miscanthus, switchgrass, and energy sorghum – in five priority levels from “most
136 unwilling” to “most willing.” It should be noted that the willingness defined with our
137 survey questionnaire is different from the adoption decision preference as defined in
138 Yang et al. [11]. Instead, it can be understood as a proxy of inclination or likelihood to
139 adopt PBCs. We designed three sets of companion questions for a comprehensive
140 analysis considering multiple factors: i) economics related (e.g., farm size and
141 participation in crop insurance programs), ii) behavior related (e.g., age and education
142 level), and iii) social support related (e.g., neighbors' land use decisions and participation
143 in government conservation practices). A farmer's willingness to adopt the three PBCs is
144 transformed into numeric values ranging from 1 (most unwilling) to 5 (most willing), and
145 these are averaged to calculate the farmer's average willingness to grow PBCs (*Will_ave*).
146 The survey also includes a choice experiment which asked farmers to choose the acreage
147 of land to be converted into PBCs under specific economic scenarios. As discussed in
148 section 1, answers to the choice experiment are not included in the current analysis.
149 Responses to other survey questions are transformed into explanatory variables in Table
150 S1 in the Supporting Information (SI) to model *Will_ave*.

151 Several variables in Table S1 are composite variables derived from responses to more
152 than one survey question. *Env_Pct* is calculated as a farmer's average agreement with the
153 importance of two objectives in land management: environmental protection and land
154 preservation. *Land_Mang* is calculated as a farmer's average agreement on the
155 importance of two environmental issues on their land management decisions: climate
156 change and fertilizer and pesticide overuse. *High_Profit* is calculated as a farmer's
157 average agreement with two statements that adopting PBCs can: increase revenue and
158 increase profit reliability under poor weather conditions. *Bene_Env* is calculated as
159 farmer's average agreement with three statements relating to the environmental benefits
160 of PBCs: local water quality improvement, soil quality improvement, and national energy
161 security improvement. *Drawback* is calculated as farmer's average agreement with three
162 concerns of growing PBCs: competition with other land uses, lack of market, and lack of
163 transportation infrastructures.

164 Table S1 includes six sets of social support variables. *Govp* is included with the
165 assumption that farmers who participated in government conservation programs would be
166 more willing to try new conservation practices such as PBCs. *Fam_ave* was usually
167 included in similar innovation land use diffusion studies as an important intermediate
168 variable to explain farmers' adoption decisions [22,31]. An increase in *Fam_ave* can be
169 considered a result of peer communication, community promotion, and public
170 propaganda. These types of neighborhood communication have shown positive impacts
171 in the adoption of certain farming techniques or crops [32]. *Sup_bio* represents the
172 availability of biomass end-users (i.e., biorefinery) for farmers and is expected to have a
173 positive impact on PBCs adoption. *Com_sup* represents the community's support for
174 growing PBCs in the community, which could be considered as a part of social
175 judgement of PBCs adoption. *Peer* can have multiple impacts on PBCs adoption: i) the
176 increase in peers adopting PBCs would increase a farmer's exposure to information and
177 knowledge of PBCs; ii) as PBCs become more popular, there is peer pressure for farmers
178 to adopt this environmentally friendly land use. Finally, *Info_use* could be used to
179 investigate if farmers' adoption of PBCs could be affected by any source of information
180 they receive.

181

182 3. Methods

183 As the target variable *Will_ave* (farmers' average willingness to adopt PBCs) is a
184 continuous ranging from 1 to 5, we can use a mix of empirical regression models, namely
185 Ordinary Least Square regression (OLS) and Structural Equation Model (SEM), to
186 investigate the predictors of the variable, as the two models are widely used for treating
187 continuous variables. Through a preliminary analysis (Figure S2 of the SI), farmers'
188 familiarity with PBCs (*Fam_ave*) appears to be the most important predictor of *Will_ave*.
189 We use OLS to identify the key drivers for willingness and familiarity, respectively. We
190 then combine the two OLS models into one SEM to show how various factors directly
191 influence adoption willingness or indirectly affect the target variable via the influence of
192 familiarity (here referred to as the mediation effect). In practice, the causal-relationships
193 of *Will_ave* and its influencing factors may show a hierarchical structure [11], but here
194 we limit the mediation variable in the SEM to be only *Fam_ave*, given the SEM
195 complexity constraint posed by the data sample size (242). As a complement to the main
196 SEM, for other important predictors of *Will_ave*, we develop separate OLS models to
197 identify their driving factors and compare the results crossing the multiple OLS models,
198 especially on the key predictors in the OLS models.

199 During the analysis, a special focus is given to the variables in the social support
200 category. The hypothesis is that social support variables such as *Fam_ave* and
201 community support (*Com_sup*) are key parts in the model of PBC adoption willingness.

202 3.1 Ordinary Least Square Regression

203 The OLS model calculates a farmer *i*'s willingness to adopt PBCs as:

$$204 \quad Will_{ave_i} = X_i^T \beta + \varepsilon_i \quad (1)$$

205 where β is a vector of regression coefficients, X refers to the explanatory variables, and
206 ε is the error term. While the variables in Table 1 could all potentially affect farmers'
207 willingness to adopt PBCs, we add two constraints on the predictors of the willingness
208 OLS model to make the model concise: i) the correlation coefficient between the
209 predictor and *Will_ave* should be larger than 0.03 to make sure the predictor is influential

210 [33,34], and ii) the correlation coefficient between two predictors should be less than 0.3
211 to avoid collinearity [35].

212 Several key predictors (*Risk*, *Fam_ave*, and *Peer*) show clear linear relationships with
213 *Will_ave* (Figures S1-S3 in the SI), which supports the choice of the OLS model. Among
214 them, *Fam_ave* appears to be the most important predictor of *Will_ave* (given its
215 correlation coefficient with *Will_ave* equals 0.282). In fact, familiarity has been identified
216 as an important intermediate variable that mediates the influence of many other more
217 fundamental drivers of willingness in similar land use studies [22,31]. Hence, we build a
218 separate OLS model for familiarity (*Fam_ave*) to identify the drivers that could indirectly
219 influence farmers' willingness to adopt PBCs through the mediation of *Fam_ave*.

220 **3.2 Structural Equation Model**

221 To further reveal how multiple drivers directly influence farmer willingness to adopt
222 PBCs and indirectly influence willingness through the mediation of farmer familiarity
223 with these crops, we use a structural equation model (SEM) to link the two OLS models
224 through the familiarity variable *Fam_ave*.

225 SEM is a multivariate statistical analysis tool for analyzing structural relationships
226 [36–38]. SEM combines factor analysis and multiple regression analysis to analyze the
227 direct and indirect relationships in a causal-relationship structure that involve one or more
228 latent variables [37]. Given its advantage in analyzing complex multiple variable models
229 and revealing relationships in hidden structures, SEM has been widely applied in social
230 science studies [39–41].

231 **4. Results**

232 **4.1 The Willingness OLS Model**

233 The results of the OLS model for farmer willingness to adopt PBCs (hereafter referred
234 to as the willingness OLS model) are shown in Table 1. The modeling results show a
235 good agreement with the farmer responses, where about 84.8% of the variation in
236 farmers' average willingness to grow the three surveyed PBCs could be explained by the
237 model (with an adjusted R^2 value of 0.85). The model also has an F-statistic of 114.3 and

238 a p-value < 0.0001, which all suggest that the model contains helpful information for
 239 predicting the target variable *Will_ave*.

240

241 Table 1. Coefficients of the OLS model for farmer willingness to adopt PBCs.

	ESTIMATE OF COEFFICIENTS	T-VALUE	P-VALUE
<i>Old</i>	0.198	0.409	0.683
<i>Male</i>	0.128	0.270	0.788
<i>Edu</i>	-0.040	-0.539	0.590
<i>IL</i>	0.120	0.812	0.418
<i>Govp</i>	-0.138	-0.630	0.529
<i>Crop_ins</i>	-0.026	-0.131	0.896
<i>Risk</i>	0.152**	2.042	0.042
<i>Fam_ave</i>	0.405***	4.146	1E-04
<i>Sup_bio</i>	0.117*	1.915	0.057
<i>Peer</i>	0.177*	1.94	0.054
Multiple R ²	0.855	Adjusted R ²	0.848
F-Statistic	114.3	P-value of F- stat	2.20E-16

242 Note: *** $P < 0.01$, ** $P < 0.05$, * $P < 0.1$; coefficients in bold are statistically significant.

243 Altogether, ten variables are selected as the predictors for the willingness OLS model,
 244 among which three are economics variables (*IL*, *Risk* and *Crop_Ins*), three are behavior
 245 variables (*Old*, *Male*, and *Edu*), and four are social support variables (*Govp*, *Fam_ave*,
 246 *Sup_Bio*, and *Peer*) (Table 1). Four of the predictors show statistical significance, and
 247 three of them are in the social support category. Among them, farmer’s familiarity with
 248 PBCs (*Fam_ave*) turns out to be the most important predictor of adoption willingness
 249 with a p-value < 0.0001. With one unit increase in farmers’ familiarity with PBCs, the
 250 average willingness to grow PBCs will increase by approximately 8.9% (Table 1). The
 251 impact of *Fam_ave* shows that farmers’ willingness to adopt a novel land use option
 252 highly depends on their familiarity with the option; this phenomenon is consistent with
 253 what was found by Ariti et al. [31] and Lee et al. [22].

254 **4.2 The Familiarity OLS Model**

255 The willingness OLS model shows how various factors affect farmers’ willingness to
 256 adopt PBCs, among which farmers’ familiarity with the crops is most influential. In the
 257 following, we further explore potential predictors of familiarity via the familiarity OLS
 258 model. Results from the familiarity OLS model provide hints on the potential factors that
 259 could indirectly influence willingness via their impacts on familiarity. As shown in Table
 260 2, the familiarity OLS model also shows a good agreement with the farmer responses,
 261 where about 95.7% of the variation in farmers’ average familiarity to the three surveyed
 262 PBCs could be explained by the model (adjusted R² value equals 0.957). The model has
 263 an F-statistic of 527.9 and a p-value < 0.0001, which suggests a stronger regression
 264 performance than the willingness OLS model.

265 Table 2. Coefficients of the OLS model for farmer familiarity to PBCs.

	ESTIMATE OF COEFFICIENTS	T-VALUE	P- VALUE
<i>Env_pct</i>	0.007	0.155	0.877
<i>Land_mang</i>	0.027	0.463	0.644
<i>Peer</i>	0.310 ^{***}	5.042	1.74E-07
<i>High_profit</i>	0.359 ^{***}	4.120	5.41E-05
<i>Bene_env</i>	0.482 ^{***}	6.331	1.4E-09
<i>Drawback</i>	0.101	1.283	0.201
<i>Knowledge</i>	-0.057	-1.020	0.309
<i>Edu</i>	-0.076	-1.484	0.139
Multiple R ²	0.9567	Adjusted R ²	0.9549
F-Statistic	527.9	P-value of F-stat	<2.20E- 16

266 Note: ^{***}*P*<0.01, ^{**}*P*<0.05, ^{*}*P*<0.1; coefficients in bold are statistically significant.

267 Altogether eight variables are selected as the predictors for the familiarity OLS model,
 268 among which one is an economic variable (*High_Profit*), six are behavior variables
 269 (*Env_Pct*, *Land_Mang*, *Bene_Env*, *Drawback*, *Knowledge*, and *Edu*), and one is social
 270 support variable (*Peer*). Three of the predictors show statistical significance, and one of

271 them (*Peer*) is in the social factor category. Among the statistically significant variables,
 272 *Peer* appears in both the willingness OLS model and the familiarity OLS model,
 273 suggesting that it may influence willingness directly or indirectly via the influence of
 274 *Fam_ave*. *High_profit* and *Bene_env* only appear in the familiarity OLS model,
 275 suggesting their impact on farmer willingness may be indirect via the mediation of
 276 *Fam_ave*. It should be noted that the potential indirect influences of *Peer*, *High_profit*,
 277 and *Bene_env*, as identified in Table 2, are based on theoretical reasoning, which requires
 278 further validation, as detailed the following SEM model in section 4.3.

279 4.3 The Structural Equation Model

280 To further validate the direct and indirect influences of the factors identified in
 281 sections 4.1-4.2, a structural equation model (SEM) is constructed, using the same factors
 282 defined in sections 4.1-4.2. The SEM assumes a hierarchical structure where the impacts
 283 of various indirect influencing factors are realized via the mediation of *Fam_ave*. More
 284 specifically, it is assumed that *Fam_ave* directly influences willingness, while other
 285 factors either directly influence willingness or indirectly influence willingness via their
 286 influence on *Fam_ave*.

287 Table 3. Statistical report of the structural equation model.

Model Statistic	Value	Criteria	Source
Comparative Fit Index (CFI)	0.884	> .95	[42,43]
Tucker-Lewis Index (TLI)	0.723	> .95	[43]
Root Mean Square Error of Approximation (RMSEA)	0.045	< .06	[43]
Root Mean Square Residual (SRMR)	0.022	< .08	[42]

288 A critical step in the SEM development is the fitness test. Since relying on one fitness
 289 index might lead to inappropriate model design, many suggest that a combination of
 290 fitness indices should be used for evaluating SEM models [42,44]. In this study, the
 291 model fitness is evaluated via four widely adopted indices [42]: the root mean square
 292 error of approximation (RMSEA), the root mean square residual (SRMR), the
 293 comparative fit index (CFI), and the Tucker-Lewis index (TLI). The former two indices
 294 belong to absolute fit indices for SEMs, which assess how well an a priori SEM model

295 reproduces the sample data. The latter two are incremental fit indices, which measure the
 296 proportionate improvement in fit by comparing a target model with a baseline model
 297 assuming uncorrelated observed variables. A good performing model should have high
 298 CFI and TLI values, and low RMSEA and SRMR values. The recommended ranges of
 299 good performance are listed in Table 3. The CFI of our SEM is close to 0.9 (though not
 300 over 0.95), and the TLI is also relatively high. The RMSEA and SRMR are smaller than
 301 0.05, indicating that the model's fitting performance is ideal. Considering the relatively
 302 small sample size of the survey data, it is reasonable to conclude that the performance of
 303 the SEM is statistically valid.

304 Table 4. Coefficients of the SEM for farmer familiarity to PBCs.

	ESTIMATE OF COEFFICIENTS	STANDARD DEVIATION	Z- VALUE	P- VALUE
<i>Env_pct</i>	0.019	0.045	0.432	0.666
<i>Land_mang</i>	0.000 ^a	0.056	-0.009	0.993
<i>Peer</i>	0.271 ^{***}	0.056	4.866	0.000 ^a
<i>High_profit</i>	0.100	0.111	0.904	0.366
<i>Bene_env</i>	0.257 ^{***}	0.097	2.665	0.008
<i>Drawback</i>	0.005	0.080	0.060	0.952
<i>Knowledge</i>	-0.058	0.054	-1.086	0.278
<i>Edu</i>	-0.109 ^{***}	0.050	-2.177	0.029

305 Note: *** $P < 0.01$, ** $P < 0.05$, * $P < 0.1$; coefficients in bold are statistically significant.

306 ^athe p-value is small and is rounded to 0.000 for consistency.

307

308 Tables 4 and 5 show the coefficients of the final SEM model. To avoid correlation
 309 among independent variables, age and sex are excluded from the main models. A
 310 comparison between the SEM and the two OLS models suggests no significant difference
 311 between the statistically significant variables. In the SEM model, *Peer* and *Bene_env*
 312 remain influential to familiarity, and the four most important predictors of willingness
 313 remain to be *Fam_ave*, *Sup_bio*, *Peer*, and *Risk*.

314 A visualization of the SEM (Figure 2) shows the central role of familiarity (*Fam_ave*)
 315 in predicting willingness (*Will_ave*). *Fam_ave* has the largest coefficient value in the

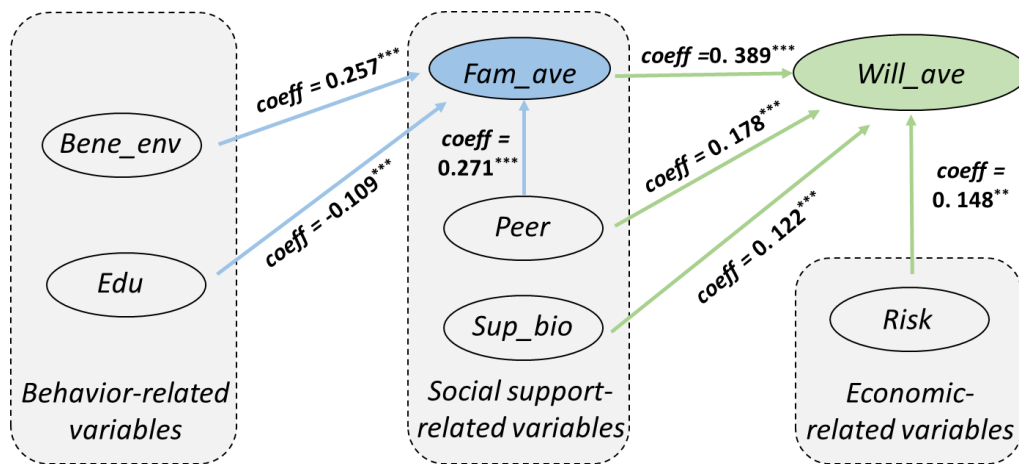
316 SEM (which indicates the most important influence), and it mediates the indirect
 317 influence of many variables (*Bene_env*, *Edu*, and *Peer*) on *Will_ave*. Meanwhile,
 318 economic variables (*Risk*) directly influence willingness without mediation, and behavior
 319 variables (*Bene_env* and *Edu*) only indirectly influence willingness via the mediation of
 320 familiarity. Among the social support variables, *Peer* both directly and indirectly
 321 influences willingness; *Sup_bio* has only a direct impact. Clearly, peer information (*Peer*)
 322 appears to be another key predictor of willingness, a conclusion supported by a variety of
 323 agricultural land use publications [11,22].

324 Table 5. Coefficients of the SEM for farmer willingness to adopt PBCs.

	ESTIMATE OF COEFFICIENTS	STANDARD DEVIATION	Z- VALUE	P- VALUE
<i>Edu</i>	-0.044	0.077	-0.568	0.570
<i>IL</i>	0.109	0.142	0.769	0.442
<i>Govp</i>	-0.145	0.213	-0.678	0.497
<i>Crop_ins</i>	-0.008	0.189	-0.042	0.967
<i>Risk</i>	0.148 ^{**}	0.076	1.941	0.052
<i>Fam_ave</i>	0.389 ^{***}	0.101	3.870	0.000 ^a
<i>Sup_bio</i>	0.122 ^{***}	0.060	2.050	0.040
<i>Peer</i>	0.178 ^{***}	0.089	1.997	0.046

325 Note: ^{***} $P < 0.01$, ^{**} $P < 0.05$, ^{*} $P < 0.1$; coefficients in bold are statistically significant.

326 ^athe p-value is small and is rounded to 0.000 for consistency.



327

328 Figure 2. Visualization of the SEM; only statistically significant variables are shown in the

329 figure (^{***} $P < 0.01$, ^{**} $P < 0.05$, ^{*} $P < 0.1$).

330 **5. Discussion**

331 **5.1 Impacts of social support variables**

332 Among the six statistically significant predictors of farmer willingness to grow PBCs
333 (*Will_ave*), three of them (farmer familiarity with the crops, *Fam_ave*; portion of peers
334 already adopting the crops, *Peer*; and farmer support for biorefineries locating in his/her
335 community, *Sup_bio*) are social support variables; familiarity appears to be the most
336 important one. The results clearly demonstrate the importance of social support in farmer
337 decision-making regarding PBCs.

338 The SEM model in Figure 2 suggests an interesting result that the factors affecting
339 farmer willingness to grow PBCs may be indirect via the mediation of familiarity. The
340 model structure follows an intuitive belief that farmers can only be willing to grow PBCs
341 after they have enough knowledge about the crops. In a relevant study about PBCs,
342 Signorini et al. [45] argued that “as agricultural producer learns more about switchgrass
343 and its conveniences, they are likely to understand that the crop itself carries multiple
344 attributes they show a preference for.” Similar results are also drawn from Fewell et al.
345 [14] and Skevas et al. [46]. Building upon the findings of the studies listed above, this
346 study further demonstrates the intermediate role of familiarity in mediating the impacts of
347 some factors that indirectly affect willingness. Our conclusion is echoed in a study that
348 investigated the adoption of another novel land use option. Lee et al. [22] used a causal
349 mediation analysis and identified a similar hierarchical structure that the awareness of
350 cover crop options mediates the influence of multiple background characteristics on
351 farmer willingness to adopt such crops in Iowa. Our results, together with those of Lee et
352 al. [22], suggest the central role of familiarity in promoting novel land use options, as
353 discussed in section 4.3.

354 The model structure in Figure 2 also partially reflects the conclusions in our previous
355 study, which estimates the causal relationship structure of farmers’ PBC adoption
356 decisions via Bayesian network analysis in a choice experiment [11]. In that study,
357 familiarity was identified as one of the most influential variables in the adoption decision,
358 and it mediates the impact of peer adoption, which has been identified in Figure 2.
359 Familiarity mediates the impacts of two other variables shown in Figure 2, *Edu* and

360 *Bene_env*, while such relationships are not identified in Yang et al. [11]. The causal
361 relationship structure in Yang et al. [11] was identified based on both optimal search and
362 manual adjustment, while the model structure in Figure 2 is believed to be more
363 appropriate given the statistical rigor of the SEM. Figure 2 suggests that behavior factors
364 such as *Edu* and *Bene_env* affect farmer willingness to grow PBCs but only via their
365 influence on farmer's familiarity with the crops. Among them, the impact of *Bene_env* is
366 positive, suggesting that a farmer's awareness of PBCs' environmental benefits is an
367 important part of their knowledge on the crops. Unexpectedly, the impact of education is
368 negative, while most previous studies found education level positively affects willingness
369 and/or familiarity [19,47]. Possible explanation may be related to observed naivety with
370 less educated/informed farmers [48]. The naivety can drive the farmers to believe they
371 know enough about PBCs. The result thus suggests a necessity to provide a more
372 objective measure of familiarity, possibly via a set of quizzes about PBCs to survey
373 respondents. Meanwhile, the result implies that more outreach programs on PBCs are
374 needed to improve farmers' knowledge about such crops.

375 Peer adoption (*Peer*) is another important variable as it affects both familiarity and
376 willingness. Yang et al. [11] also identified peer adoption as an important factor in farmer
377 adoption willingness, but its impact is only indirect via the mediation of familiarity.
378 Results regarding the impact of *Peer* in both studies reflect the spatial spillover effect of
379 PBCs adoption willingness, as identified by Skevas et al. [47], Swinton et al. [49], and
380 Jiang et al. [19], who suggested that peer communication and peer pressure regarding
381 PBCs could be the major driver. However, the direct impact of *Peer* on willingness
382 identified in Figure 2 demonstrates that noticing the adoption of neighbors can already
383 increase a farmer's willingness to adopt PBCs, even if they still know little about such
384 crops. Therefore, the impact from early adopters in a community may be quicker and
385 stronger than expected, as the farmers in the neighborhood may decide to grow PBCs
386 before they have the same level of knowledge and/or experiences as the early adopters
387 have. It should be noted that the inclusion of *Peer* in the regression models may raise
388 concern regarding the issue of endogeneity; that is to say, the error term might be
389 correlated with the predictor (*Peer*) if the predictor is affected by the regressing target
390 (*Will_ave*) [50,51]. However, our model may not be subject to the endogeneity issue,

391 because the respondents in the survey do not know each other and their adoption would
392 not affect the peer adoption ratio of other respondents. Meanwhile, our survey reflects the
393 value of *Peer* and *Will_ave* at only one specific point of time, and *Will_ave* would only
394 affect *Peer* at future time points.

395 Farmers' support for biorefineries located in their community (*Sup_bio*) is another key
396 behavior factor identified in Figure 2. A high value of *Sub_bio* provides a sign that
397 farmers are considering future markets of PBCs, which further indicates a higher
398 possibility that they are willing to adopt such crops. A similar result was also identified in
399 previous studies. For example, Embaye et al. [52] noted that the availability of nearby
400 crushing facilities is important for PBCs adoption because well-developed nearby
401 facilities could potentially contribute to a better market for the crops.

402 Two behavior variables (i.e., perception of the environmental benefits of PBCs,
403 *Bene_env*; education level, *Edu*) in Figure 2 indirectly influence the adoption of PBCs
404 and the impact of one economic variable (willingness to take risk, *Risk*) in Figure 2 is
405 direct. The impacts of the three variables are consistent with the finding of other studies
406 [11,45], except for the impact of *Edu*. As explained above, the negative impact of *Edu*
407 may be related to the observation that less educated/informed farmers may falsely believe
408 they know enough about PBCs [48], and/or they can be convinced with simple or quick
409 knowledge (e.g., following neighbors as discussed above) without deep thoughts or
410 concerns. The two behavior variables are closely connected to the social support factors
411 although they are not included in the social support category of this study. This relates to
412 the mechanism that impacts of behavior factors are realized. As shown in Lee et al. [22]
413 and Wang et al. [53], behavior factors such as education level, age, and gender can have
414 impacts on farmers' land use decisions, but only via the mediation of intermediate
415 variables (many of which relate to the social support variables) such as awareness, social
416 status, and economic capability.

417 **5.2 Limitations**

418 In addition to familiarity (*Fam_ave*), farmer's support for biorefineries (*Sup_bio*) and
419 farmers' willingness to take risks (*Risk*) appear to be statistically significant predictors of
420 their willingness to adopt PBCs (*Will_ave*) in the OLS and SEM models (Tables 1 and 5),

421 and they may also mediate the impacts of other indirect variables. However, the
422 developed SEM in Section 4.3 only explores the factors that can indirectly influence the
423 willingness due to the mediation of familiarity, given the relatively small sample size
424 (242).

425 Meanwhile, additional survey samples would be needed to verify the difference in
426 model structures identified in Figure 2 and in Yang et al. [11]. We suggest that future
427 research should include more survey samples to further verify the assumptions developed
428 in this study. In addition, a multi-year survey effort would further improve the validity of
429 the model developed in this study, as well as identifying potential changes in behavior
430 resulting from market or policy changes [1,60,61].

431 **5.3 Implications**

432 Dedicated bioenergy crops such as PBCs growing on marginal lands are believed to be
433 a possible solution of the grand global challenge of food, energy, and sustainability [3].
434 While the first generation biofuel produced from food crops (e.g., corn and soybean)
435 faces the ‘food versus fuel’ dilemma, such conflicts can be alleviated for advanced
436 biofuels produced from PBCs as such crops can be mostly produced on marginal lands
437 (which amounts to over 55.1 million hectares in the U.S. [7]). Meanwhile, PBCs bring
438 multiple environmental benefits including but not limited to erosion prevention, soil
439 nutrient recycling, soil fertility improvement, and pollination; furthermore, it is believed
440 that these environmental benefits can improve the crop production on surrounding
441 agriculture lands [62]. Further, it is believed that PBCs will mitigate carbon emissions
442 from soil [63]. As the entire world is cooperatively trying to achieve the objective of
443 carbon neutralization by 2050 [64], this study could contribute to a more targeted design
444 of promotion policies for bioenergy crop expansion and bioenergy development while
445 supporting other carbon emission reduction measures. Some implications for PBCs
446 policy design are discussed below.

447 The SEM result in Figure 2 suggests the important role of social support in promoting
448 PBCs. In particular, the peer adoption ratio could directly increase farmer willingness to
449 grow PBCs, and it could indirectly increase willingness by increasing farmer familiarity
450 with the crops. Therefore, a potential policy to promote PBCs could be pilot projects that

451 (financially) support farmers to demonstrate the environmental and economic advantages
452 of PBCs. As simulated in Yang et al. [61] via an agent-based model, it is hypothesized
453 that a successful pilot project would encourage early adoption among an agricultural
454 community, which could further increase farmers' knowledge about the crops and their
455 willingness to adopt. Alternatively, pilot projects for bioenergy and bioproduct facilities,
456 e.g., cellulosic biorefineries or small-scale bio-facilities, could foster an early market for
457 biomass and hence encourage early adoptions of PBCs. Such a strategy echoes a
458 stakeholder's suggestion to "get the supply chain rolling" in a Midwestern bioeconomy
459 stakeholder focus group discussion [61]. The core idea of the stakeholder's suggestion is
460 to establish an early market for the bioeconomy supply chain (including biomass
461 producers such as farmers), and the supply chain would mature as participants learn
462 during the processes [61]. The establishment of pilot projects for bioenergy and
463 bioproduct facilities would also affect farmer willingness to adopt PBCs via the influence
464 of *Sup_bio*, if the pilot project could provide a good price for biomass and win the
465 support from local farmers.

466 Another potential means to improve PBCs adoption is to increase farmers' familiarity
467 with the crops. As validated by multiple past researchers [65–67], advertising and
468 outreach programs focusing on the characteristics, economic, and environmental payoffs
469 of PBCs are prioritized in promotion campaigns. Potential beneficial courses of action are
470 related to another variable *Bene_env*, which looks at whether the farmers are aware of the
471 environmental benefits of growing PBCs. In other words, and as also mentioned by past
472 research [14,68,69], the current mission is to alert farmers of the unique characteristics of
473 PBCs. Some of those characteristics are especially preferred by them, including their
474 potential contribution to the environment and other benefits of the nation.

475 Alternatively, policymakers targeting PBCs promotion could focus on the variable
476 *Risk*. If farmers can manage their risk and be 20% less risk-averse, they tend to be ~15%
477 more willing to grow PBCs on average (Figure 2). Therefore, policymakers could
478 manage to reduce the risk that are involved in economic loss, climate threats, and lack of
479 local market [1,14]. Agricultural subsidies and government insurance projects could also
480 be potential actions to take [70–72].

481 **6. Conclusion**

482 This study investigates the knowledge gap in the impacts of social factors (e.g.,
483 neighborhood knowledge, community communication, and shared practices) on farmers'
484 willingness to grow PBCs. As revealed through a survey analysis of midwestern farmers
485 and a structural model that together investigate the economic, behavioral, and social
486 drivers on farmer decision-making, this study highlights the importance of social
487 variables, such as familiarity with PBCs (*Fam_ave*), peer adoption ratio (*Peer*), and
488 support of biorefineries (*Sup_bio*) on farmer willingness to grow PBCs. Combining two
489 Ordinary Least Square (OLS) regression models that address farmers' familiarity and
490 willingness with PBCs, a Structural Equation Model (SEM) provides insight into the
491 variables' direct and indirect drivers. From both OLS models, six statistically significant
492 predictors of farmer willingness to grow PBCs are identified (farmer agreement with the
493 environmental benefits of PBCs, *Bene_env*; education level, *Edu*; farmer's average
494 familiarity of PBCs, *Fam_ave*; portion of peers already adopting PBCs, *Peer*; support of
495 biorefinery building in the local community, *Sup_bio*; and willingness to take risks, *Risk*),
496 and three of them (*Fam_ave*, *Peer*, and *Sup_bio*) are social support variables. Overall,
497 *Fam_ave* appears to be the most significant predictor of farmer willingness to adopt
498 PBCs, and the impacts of many other predictors are only realized via their influence on
499 *Fam_ave*. Other variables such as *Peer* affect both familiarity and willingness. Our
500 findings suggest the potential for improving PBCs policies through social programs such
501 as pilot projects for farmers and bioenergy facilities. The pilot projects could establish an
502 early market for PBCs, which could further attract early adopters of PBCs and increase
503 other farmers' willingness of adoption via the influence of *Peer* and *Fam_ave*.
504 Meanwhile, the cellulosic bioeconomy could be further enhanced by increasing farmers'
505 familiarity with PBCs and minimizing farmers' risk perceptions.

506

507 **Supporting Information**

508 Supporting information include three tables (Tables S1-S3) and three figures (Figures S1-
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- 522 [1] C. Leibensperger, P. Yang, Q. Zhao, S. Wei, X. Cai, The Synergy Between Stakeholders
523 for Cellulosic Biofuel Development—Perspectives, Opportunities, and Barriers, *Renew.*
524 *Sustain. Energy Rev.* 137 (2020), 110613. <https://doi.org/10.1016/j.rser.2020.110613>.
- 525 [2] W. V Reid, M.K. Ali, C.B. Field, The future of bioenergy, *Glob. Chang. Biol.* 26 (1)
526 (2020) 274–286. <https://doi.org/10.1111/gcb.14883>.
- 527 [3] G.P. Robertson, S.K. Hamilton, B.L. Barham, B.E. Dale, R.C. Izaurralde, R.D. Jackson,
528 D.A. Landis, S.M. Swinton, K.D. Thelen, J.M. Tiedje, Cellulosic biofuel contributions to
529 a sustainable energy future: Choices and outcomes, *Science* 356 (6345) (2017), eaal2324.
530 <https://doi.org/10.1126/science.aal2324>.
- 531 [4] Z. Qin, J.B. Dunn, H. Kwon, S. Mueller, M.M. Wander, Soil carbon sequestration and
532 land use change associated with biofuel production: empirical evidence, *Glob. Chang.*
533 *Biol.* 8 (1) (2016) 66–80. <https://doi.org/https://doi.org/10.1111/gcbb.12237>.
- 534 [5] B.P. Werling, T.L. Dickson, R. Isaacs, H. Gaines, C. Gratton, K.L. Gross, H. Liere, C.M.
535 Malmstrom, T.D. Meehan, L. Ruan, B.A. Robertson, G.P. Robertson, T.M. Schmidt, A.C.
536 Schrottenboer, T.K. Teal, J.K. Wilson, D.A. Landis, Perennial grasslands enhance
537 biodiversity and multiple ecosystem services in bioenergy landscapes, *Proc. Natl. Acad.*
538 *Sci.* 111 (4) (2014) 1652–1657. <https://doi.org/10.1073/pnas.1309492111>.
- 539 [6] S. Ahmed, T. Warne, E. Smith, H. Goemann, G. Linse, M. Greenwood, J. Kedziora, M.
540 Sapp, D. Kraner, K. Roemer, J.H. Haggerty, M. Jarchow, D. Swanson, B. Poulter, P.C.
541 Stoy, Systematic review on effects of bioenergy from edible versus inedible feedstocks on
542 food security, *NPJ Sci. Food* 5 (1) (2021), 9. <https://doi.org/10.1038/s41538-021-00091-6>.
- 543 [7] P. Yang, Q. Zhao, X. Cai, Machine learning based estimation of land productivity in the
544 Contiguous US using biophysical predictors, *Environ. Res. Lett.* 15 (7) (2020), 074013.
545 <https://doi.org/10.1088/1748-9326/ab865f>.
- 546 [8] CAEP Explorer Course 2018, Eric Rund Farm Visit. [Accessed: March 13, 2020].
547 Available from: <https://publish.illinois.edu/caep-2018/2018/07/09/rund-farm/>
- 548 [9] United States Environmental Protection Agency, Final Renewable Fuel Standards for
549 2020, and the Biomass-Based Diesel Volume for 2021; 2020 [Accessed: September 24,
550 2020] Available from: [https://www.epa.gov/renewable-fuel-standard-program/final-](https://www.epa.gov/renewable-fuel-standard-program/final-renewable-fuel-standards-2020-and-biomass-based-diesel-volume)
551 [renewable-fuel-standards-2020-and-biomass-based-diesel-volume](https://www.epa.gov/renewable-fuel-standard-program/final-renewable-fuel-standards-2020-and-biomass-based-diesel-volume)
- 552 [10] P. Burli, P. Lal, B. Wolde, S. Jose, S. Bardhan, Factors affecting willingness to cultivate
553 switchgrass: Evidence from a farmer survey in Missouri, *Energy Econ.* 80 (2019) 20–29.
554 <https://doi.org/10.1016/j.eneco.2018.12.009>.
- 555 [11] P. Yang, X. Cai, C. Leibensperger, M. Khanna, Adoption of perennial energy crops in the
556 US Midwest: Causal and heterogeneous determinants, *Biomass Bioenergy* 155 (2021),
557 106275. <https://doi.org/10.1016/j.biombioe.2021.106275>.
- 558 [12] A.M. Rossi, C.C. Hinrichs, Hope and skepticism: Farmer and local community views on
559 the socio-economic benefits of agricultural bioenergy, *Biomass Bioenergy.* 35 (4) (2011)
560 1418–1428. <https://doi.org/https://doi.org/10.1016/j.biombioe.2010.08.036>.
- 561 [13] T. Skevas, N.J. Hayden, S.M. Swinton, F. Lupi, Landowner willingness to supply
562 marginal land for bioenergy production, *Land Use Policy* 50 (2016) 507–517.
563 <https://doi.org/10.1016/j.landusepol.2015.09.027>.

- 564 [14] J.E. Fewell, J.S. Bergtold, J.R. Williams, Farmers' willingness to contract switchgrass as a
565 cellulosic bioenergy crop in Kansas, *Energy Econ.* 55 (2016) 292–302.
566 <https://doi.org/10.1016/j.eneco.2016.01.015>.
- 567 [15] M. Khanna, J. Louviere, X. Yang, Motivations to grow energy crops: the role of crop and
568 contract attributes, *Agric. Econ.* 48 (3) (2017) 263–277.
569 <https://doi.org/10.1111/agec.12332>.
- 570 [16] W.M. Eaton, M. Burnham, K. Running, C.C. Hinrichs, T. Selfa, Symbolic meanings,
571 landowner support, and dedicated bioenergy crops in the rural northeastern United States,
572 *Energy Res. Soc. Sci.* 52 (2019) 247–257.
573 <https://doi.org/https://doi.org/10.1016/j.erss.2019.02.005>.
- 574 [17] R. Helliwell, Where did the marginal land go? Farmers perspectives on marginal land and
575 its implications for adoption of dedicated energy crops, *Energy Policy* 117 (2018) 166–
576 172. <https://doi.org/10.1016/j.enpol.2018.03.011>.
- 577 [18] P. Yang, X. Cai, M. Khanna, Farmers' heterogeneous perceptions of marginal land for
578 biofuel crops in US Midwestern states considering biophysical and socioeconomic factors,
579 *Glob. Change Biol. Bioenergy* 13 (5) (2021) 849–861.
580 <https://doi.org/10.1111/gcbb.12821>.
- 581 [19] W. Jiang, K.Y. Zipp, M.H. Langholtz, M.G. Jacobson, Modeling spatial dependence and
582 economic hotspots in landowners' willingness to supply bioenergy crops in the
583 northeastern United States, *Glob. Change Biol. Bioenergy* 11 (9) (2019) 1086–1097.
584 <https://doi.org/https://doi.org/10.1111/gcbb.12617>.
- 585 [20] T.L. Ng, X. Cai, Y. Ouyang, Some implications of biofuel development for engineering
586 infrastructures in the United States, *Biofuels Bioprod. Biorefin.* 5 (5) (2011) 581–592.
587 <https://doi.org/10.1002/bbb.309>.
- 588 [21] M. Housh, M.A. Yaeger, X. Cai, G.F. McIsaac, M. Khanna, M. Sivapalan, Y. Ouyang, I.
589 Al-Qadi, A.K. Jain, Managing Multiple Mandates: A System of Systems Model to
590 Analyze Strategies for Producing Cellulosic Ethanol and Reducing Riverine Nitrate Loads
591 in the Upper Mississippi River Basin, *Environ. Sci. Technol.* 49 (19) (2015) 11932–11940.
592 <https://doi.org/10.1021/acs.est.5b02712>.
- 593 [22] D. Lee, J.G. Arbuckle, Z. Zhu, L. Nowatzke, Conditional Causal Mediation Analysis of
594 Factors Associated With Cover Crop Adoption in Iowa, USA, *Water Resour. Res.* 54 (11)
595 (2018) 9566–9584. <https://doi.org/10.1029/2017WR022385>.
- 596 [23] E. Celio, T. Koellner, A. Grêt-Regamey, Modeling land use decisions with Bayesian
597 networks: Spatially explicit analysis of driving forces on land use change, *Environ. Model.*
598 *Softw.* 52 (2014) 222–233. <https://doi.org/10.1016/j.envsoft.2013.10.014>.
- 599 [24] I. Ajzen, The theory of planned behavior, *Organ. Behav. Hum. Decis. Process.* 50 (2)
600 (1991) 179–211. [https://doi.org/10.1016/0749-5978\(91\)90020-T](https://doi.org/10.1016/0749-5978(91)90020-T).
- 601 [25] X. Rueda, M.A. Velez, L. Moros, L. Rodriguez, Beyond proximate and distal causes of
602 land-use change: Linking individual motivations to deforestation in rural contexts, *Ecol.*
603 *Soc.* 24 (1) (2019), 4. <https://doi.org/10.5751/ES-10617-240104>.
- 604 [26] T. Tahmasebi, E. Karami, M. Keshavarz, Agricultural land use change under climate
605 variability and change: Drivers and impacts, *J. Arid Environ.* 180 (2020), 104202.
606 <https://doi.org/10.1016/j.jaridenv.2020.104202>.

- 607 [27] H. Zhang, Y. Vorobeychik, Empirically grounded agent-based models of innovation
608 diffusion: a critical review, *Artif. Intell. Rev.* 52 (1) (2019) 707–741.
609 <https://doi.org/10.1007/s10462-017-9577-z>.
- 610 [28] G.P. Shivakoti, W.G. Axinn, P. Bhandari, N.B. Chhetri, The impact of community context
611 on land use in an agricultural society, *Popul. Environ.* 20 (1999) 191–213.
- 612 [29] Farm Market ID, 2020 [Accessed: October 3, 2020] Available from:
613 <https://www.farmmarketid.com/>
- 614 [30] H. Aoyama, A study of stratified random sampling, *Ann. Inst. Stat. Math* 6 (1) (1954) 1–
615 36.
- 616 [31] A.T. Ariti, J. van Vliet, P.H. Verburg, Farmers’ participation in the development of land
617 use policies for the Central Rift Valley of Ethiopia, *Land Use Policy* 71 (2018) 129–137.
618 <https://doi.org/10.1016/j.landusepol.2017.11.051>.
- 619 [32] M.T. Konrad, G. Levin, M. Termansen, Landowners’ motivation for adopting perennial
620 energy crops: drivers, barriers and neighbourhood effects, *Eur. Rev. Agric. Econ.* 45 (5)
621 (2018) 809–829. <https://doi.org/https://doi.org/10.1093/erae/jby015>.
- 622 [33] A.G. Asuero, A. Sayago, A.G. Gonzalez, The correlation coefficient: An overview, *Crit.*
623 *Rev. Anal. Chem.* 36 (1) (2006) 41–59. <https://doi.org/10.1080/10408340500526766>.
- 624 [34] M.M. Mukaka, A guide to appropriate use of correlation coefficient in medical research,
625 *Malawi Med. J.* 24 (3) (2012) 69–71.
- 626 [35] G.W. Stewart, Collinearity and least squares regression, *Stat. Sci.* 2 (1) (1987) 68–84.
627 <https://doi.org/10.1214/ss/1177013439>.
- 628 [36] M.E. Civelek, *Essentials of Structural Equation Modeling*, Zea Books, 1st ed., 2018.
- 629 [37] B.M. Byrne, *Structural equation modeling with EQS: Basic concepts, applications, and*
630 *programming*, Routledge, 2nd ed., 2013.
- 631 [38] J.F. Hair, *Multivariate data analysis*, Prentice Hall, 7th ed., 2009.
- 632 [39] H.Y. Durak, M. Saritepeci, Analysis of the relation between computational thinking skills
633 and various variables with the structural equation model, *Comput. Educ.* 116 (2018) 191–
634 202. <https://doi.org/10.1016/j.compedu.2017.09.004>.
- 635 [40] J. Chen, J. Huang, X. Huang, S. Sun, Y. Hao, H. Wu, How does new environmental law
636 affect public environmental protection activities in China? Evidence from structural
637 equation model analysis on legal cognition, *Sci. Total Environ.* 714 (2020), 136558.
638 <https://doi.org/10.1016/j.scitotenv.2020.136558>.
- 639 [41] Y.K. Dwivedi, E. Ismagilova, P. Sarker, A. Jeyaraj, Y. Jadir, L. Hughes, A meta-analytic
640 structural equation model for understanding social commerce adoption, *Inf. Syst. Front.*
641 (2021) 1–17. <https://doi.org/10.1007/s10796-021-10172-2>.
- 642 [42] L. Hu, P.M. Bentler, Cutoff criteria for fit indexes in covariance structure analysis:
643 Conventional criteria versus new alternatives, *Struct. Equ. Modeling* 6 (1) (1999) 1–55.
644 <https://doi.org/10.1080/10705519909540118>.
- 645 [43] Y. Xia, Y. Yang, RMSEA, CFI, and TLI in structural equation modeling with ordered
646 categorical data: The story they tell depends on the estimation methods, *Behav. Res.*
647 *Methods* 51 (2019) 409–428. <https://doi.org/10.3758/s13428-018-1055-2>.
- 648 [44] H.W. Marsh, K.-T. Hau, Assessing goodness of fit: Is parsimony always desirable?, *J.*
649 *Exp. Educ.* 64 (4) (1996) 364–390. <https://doi.org/10.1080/00220973.1996.10806604>.

- 650 [45] G. Signorini, D.L. Ortega, R.B. Ross, H.C. Peterson, Heterogeneity in farmers'
651 willingness to produce bioenergy crops in the Midwest USA, *Agric. Resour. Econ. Rev.*
652 50 (2) (2021) 367-393. <https://doi.org/10.1017/age.2021.8>.
- 653 [46] T. Skevas, S.M. Swinton, N.J. Hayden, What type of landowner would supply marginal
654 land for energy crops?, *Biomass Bioenergy* 67 (2014) 252–259.
655 <https://doi.org/10.1016/j.biombioe.2014.05.011>.
- 656 [47] T. Skevas, I. Skevas, S.M. Swinton, Does spatial dependence affect the intention to make
657 land available for bioenergy crops?, *J. Agric. Econ.* 69 (2) (2018) 393–412.
658 <https://doi.org/10.1111/1477-9552.12233>.
- 659 [48] C.M. Mattia, S.T. Lovell, A. Davis, Identifying barriers and motivators for adoption of
660 multifunctional perennial cropping systems by landowners in the Upper Sangamon River
661 Watershed, Illinois, *Agrofor. Syst.* 92 (2018) 1155–1169. [https://doi.org/10.1007/s10457-](https://doi.org/10.1007/s10457-016-0053-6)
662 [016-0053-6](https://doi.org/10.1007/s10457-016-0053-6).
- 663 [49] S.M. Swinton, S. Tanner, B.L. Barham, D.F. Mooney, T. Skevas, How willing are
664 landowners to supply land for bioenergy crops in the Northern Great Lakes Region?,
665 *Glob. Change Biol. Bioenergy* 9 (2) (2017) 414–428. <https://doi.org/10.1111/gcbb.12336>.
- 666 [50] G.J. Duncan, K.A. Magnuson, J. Ludwig, The endogeneity problem in developmental
667 studies, *Res. Hum. Dev.* 1 (1-2) (2004) 59–80.
668 <https://doi.org/10.1080/15427609.2004.9683330>.
- 669 [51] W.R.Q. Anton, G. Deltas, M. Khanna, Incentives for environmental self-regulation and
670 implications for environmental performance, *J. Environ. Econ. Manage.* 48 (1) (2004)
671 632–654. <https://doi.org/10.1016/j.jeem.2003.06.003>.
- 672 [52] W.T. Embaye, J.S. Bergtold, D. Archer, C. Flora, G.C. Andrango, M. Odening, J. Buysse,
673 Examining farmers' willingness to grow and allocate land for oilseed crops for biofuel
674 production, *Energy Econ.* 71 (2018) 311–320.
675 <https://doi.org/10.1016/j.eneco.2018.03.005>.
- 676 [53] Y. Wang, Q. Zhang, R. Bilsborrow, S. Tao, X. Chen, K. Sullivan-Wiley, Q. Huang, J. Li,
677 C. Song, Effects of payments for ecosystem services programs in China on rural
678 household labor allocation and land use: Identifying complex pathways, *Land Use Policy*
679 99 (2020), 105024. <https://doi.org/10.1016/j.landusepol.2020.105024>.
- 680 [54] S. Kim, B.E. Dale, All biomass is local: The cost, volume produced, and global warming
681 impact of cellulosic biofuels depend strongly on logistics and local conditions, *Biofuels*
682 *Bioprod. Biorefin.* 9 (4) (2015) 422–434. <https://doi.org/10.1002/bbb.1554>.
- 683 [55] K. Krah, D.R. Petrolia, A. Williams, K.H. Coble, A. Harri, R.M. Rejesus, Producer
684 preferences for contracts on a risky bioenergy crop, *Appl. Econ. Perspect Policy* 40 (2)
685 (2018) 240–258. <https://doi.org/10.1093/aep/pxp034>.
- 686 [56] J.S. Bergtold, A. Shanoyan, J.E. Fewell, J.R. Williams, Annual bioenergy crops for
687 biofuels production: Farmers' contractual preferences for producing sweet sorghum,
688 *Energy* 119 (2017) 724-731. <https://doi.org/10.1016/j.energy.2016.11.032>.
- 689 [57] T. Nielsen, A. Keil, M. Zeller, Assessing farmers' risk preferences and their determinants
690 in a marginal upland area of Vietnam: a comparison of multiple elicitation techniques,
691 *Agric. Econ.* 44 (3) (2013) 255–273. <https://doi.org/10.1111/agec.12009>.
- 692 [58] J. Jianjun, G. Yiwei, W. Xiaomin, P.K. Nam, Farmers' risk preferences and their climate
693 change adaptation strategies in the Yongqiao District, China, *Land Use Policy* 47 (2015)
694 365–372. <https://doi.org/10.1016/j.landusepol.2015.04.028>.

- 695 [59] U. Pröbstl-Haider, N.M. Mostegl, J. Kelemen-Finan, W. Haider, H. Formayer, J.
696 Kantelhardt, T. Moser, M. Kapfer, R. Trenholm, Farmers' Preferences for Future
697 Agricultural Land Use Under the Consideration of Climate Change, *Environ. Manage.* 58
698 (2016) 446–464. <https://doi.org/10.1007/s00267-016-0720-4>.
- 699 [60] S.S. White, T. Selfa, Shifting Lands: Exploring Kansas Farmer Decision-Making in an Era
700 of Climate Change and Biofuels Production, *Environ. Manage.* 51 (2013) 379–391.
701 <https://doi.org/10.1007/s00267-012-9991-6>.
- 702 [61] P. Yang, X. Cai, X. Hu, Q. Zhao, Y. Lee, M. Khanna, Y.R. Cortés-Peña, J.S. Guest, J.
703 Kent, T.W. Hudiburg, An agent-based modeling tool supporting bioenergy and bio-
704 product community communication regarding cellulosic bioeconomy development,
705 *Renew. Sustain. Energy Rev.* 167 (2022), 112745.
706 <https://doi.org/10.1016/j.rser.2022.112745>.
- 707 [62] M. Von Cossel, B. Winkler, A. Mangold, J. Lask, M. Wagner, I. Lewandowski, B.
708 Elbersen, M. van Eupen, S. Mantel, A. Kiesel, Bridging the gap between biofuels and
709 biodiversity through monetizing environmental services of Miscanthus cultivation, *Earths*
710 *Future* 8 (10) (2020), e2020EF001478. <https://doi.org/10.1029/2020EF001478>.
- 711 [63] L.K. Tiemann, A.S. Grandy, Mechanisms of soil carbon accrual and storage in bioenergy
712 cropping systems, *Glob. Change Biol. Bioenergy* 7 (2) (2015) 161–174.
713 <https://doi.org/10.1111/gcbb.12126>.
- 714 [64] N. Höhne, M.J. Gidden, M. den Elzen, F. Hans, C. Fyson, A. Geiges, M.L. Jeffery, S.
715 Gonzales-Zuñiga, S. Mooldijk, W. Hare, Wave of net zero emission targets opens window
716 to meeting the Paris Agreement, *Nat. Clim. Change* 11 (10) (2021) 820–822.
717 <https://doi.org/10.1038/s41558-021-01142-2>.
- 718 [65] G.C. Andrango, J.S. Bergtold, D. Archer, C. Flora, Assessing extension and outreach
719 education levels for biofuel feedstock production in the Western United States, *Open*
720 *Agric.* 1 (1) (2016) 29–36. <https://doi.org/10.1515/opag-2016-0004>.
- 721 [66] S. Nepal, L.T. Tran, D.G. Hodges, Determinants of landowners' willingness to participate
722 in bioenergy crop production: a case study from Northern Kentucky, *Forests* 11 (10)
723 (2020), 1052. <https://doi.org/10.3390/f11101052>.
- 724 [67] C.A. Augustenborg, J. Finnan, L. Mcbennett, V. Connolly, U. Priegnitz, C. Müller,
725 Farmers' perspectives for the development of a bioenergy industry in Ireland, *Glob.*
726 *Change Biol. Bioenergy* 4 (5) (2012) 597–610. <https://doi.org/10.1111/j.1757-1707.2011.01151.x>.
- 728 [68] S. Ale, P. V Femeena, S. Mehan, R. Cibin, Environmental impacts of bioenergy crop
729 production and benefits of multifunctional bioenergy systems, *Bioenergy with Carbon*
730 *Capture and Storage* (2019) 195–217. <https://doi.org/10.1016/B978-0-12-816229-3.00010-7>.
- 732 [69] S.C. Davis, W.J. Parton, F.G. Dohleman, C.M. Smith, S. Del Grosso, A.D. Kent, E.H.
733 DeLucia, Comparative biogeochemical cycles of bioenergy crops reveal nitrogen-fixation
734 and low greenhouse gas emissions in a Miscanthus× giganteus agro-ecosystem,
735 *Ecosystems* 13 (2010) 144–156. <https://doi.org/10.1007/s10021-009-9306-9>.
- 736 [70] L. Traverso, M. Colangeli, M. Morese, G. Pulighe, G. Branca, Opportunities and
737 constraints for implementation of cellulosic ethanol value chains in Europe, *Biomass*
738 *Bioenergy* 141 (2020), 105692. <https://doi.org/10.1016/j.biombioe.2020.105692>.

- 739 [71] M. Banja, R. Sikkema, M. Jégard, V. Motola, J.-F. Dallemand, Biomass for energy in the
740 EU–The support framework, *Energy Policy* 131 (2019) 215–228.
741 <https://doi.org/10.1016/j.enpol.2019.04.038>.
- 742 [72] F. Ye, Z. Cai, Y. Chen, Y. Li, G. Hou, Subsidize farmers or bioenergy producer? The
743 design of a government subsidy program for a bioenergy supply chain, *Naval Res. Logist.*
744 68 (8) (2021) 1082–1097. <https://doi.org/10.1002/nav.21909>.
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747 **Figure Captions**

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749 Figure 1. Histograms of respondents' socio-economic variables.

750 Figure 2. Visualization of the SEM; only statistically significant variables are shown in
751 the figure (**P<0.01, *P<0.05, *P<0.1).

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