

# Enabling and interpreting hyper-differential sensitivity analysis for Bayesian inverse problems

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RISE Group Meeting

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- 1 Bayesian inverse problems
- 2 Hyper-differential sensitivity analysis (HDSA)
- 3 The Bayesian interpretation of HDSA
- 4 Enabling HDSA for ill-posed problems
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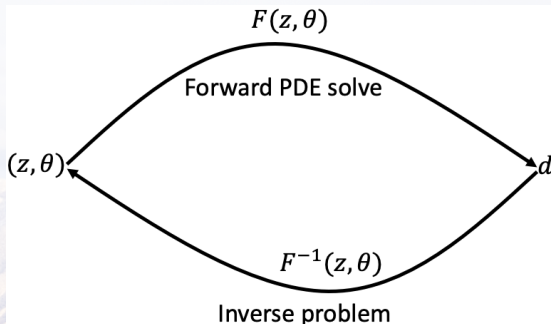
# Inverse problems

Find parameters  $(z, \theta)$  such that

$$F(z, \theta) = \mathcal{Q}(u(z, \theta)) \approx \mathbf{d}$$

where

- $\mathbf{d}$  are sparse and noisy observations of a state variable  $u(z, \theta)$ .
- $u(z, \theta)$  is the solution of a PDE and  $\mathcal{Q}$  is the observation operator.



# The joint Bayesian formulation

- Assume a prior distribution for  $(\mathbf{z}, \boldsymbol{\theta}) \sim \mathcal{N}((\mathbf{z}_{\text{prior}}, \boldsymbol{\theta}_{\text{prior}}), \Gamma_{\text{prior}})$ .
- Assume that  $\mathbf{d} = F(\mathbf{z}^*, \boldsymbol{\theta}^*) + \epsilon$  where  $\epsilon \sim \mathcal{N}(0, \Gamma_{\text{noise}})$ .
- The joint posterior probability density function (PDF) is

$$\pi_{\text{post}}(\mathbf{z}, \boldsymbol{\theta}) \propto \pi_{\text{like}}(\mathbf{d}|\mathbf{z}, \boldsymbol{\theta})\pi_{\text{prior}}(\mathbf{z}, \boldsymbol{\theta})$$

where

- $\pi_{\text{like}}$  is the likelihood function,
- $\pi_{\text{prior}}$  is the prior PDF.

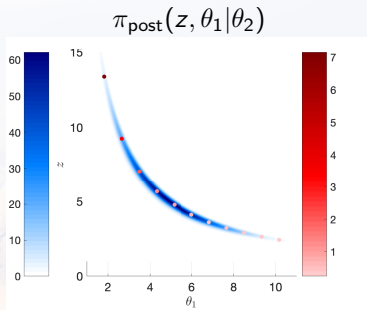
Analyzing properties of  $\pi_{\text{post}}(\mathbf{z}, \boldsymbol{\theta})$  provide a wealth of information, but may be computationally intractable.

# The conditional Bayesian formulation

- The posterior probability density function (PDF) for  $\mathbf{z}$  given  $\boldsymbol{\theta} = \boldsymbol{\theta}_{\text{prior}}$  is

$$\pi_{\text{post}}(\mathbf{z}|\boldsymbol{\theta}_{\text{prior}}) \propto \pi_{\text{like}}(\mathbf{d}|\mathbf{z}, \boldsymbol{\theta}_{\text{prior}})\pi_{\text{prior}}(\mathbf{z}, \boldsymbol{\theta}_{\text{prior}}).$$

- Fixing  $\boldsymbol{\theta} = \boldsymbol{\theta}_{\text{prior}}$  to its prior mean simplifies the analysis.
- How does  $\boldsymbol{\theta}_{\text{prior}}$  influence the posterior distribution for  $\mathbf{z}$ ?



# Leveraging PDE-constrained optimization

- The maximum a posteriori probability (MAP) point(s) for  $\pi_{\text{post}}(\mathbf{z}|\boldsymbol{\theta}_{\text{prior}})$  are local minima of

$$\min_{\mathbf{z} \in \mathbb{R}^m} J(\mathbf{z}; \boldsymbol{\theta}_{\text{prior}}) := M(\mathbf{z}, \boldsymbol{\theta}_{\text{prior}}) + R(\mathbf{z}, \boldsymbol{\theta}_{\text{prior}})$$

where  $M(\mathbf{z}, \boldsymbol{\theta}_{\text{prior}})$  and  $R(\mathbf{z}, \boldsymbol{\theta}_{\text{prior}})$  are the negative log likelihood and prior.

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## Benefit of PDE-constrained optimization

- Leverages computationally scalable, matrix free, and parallel algorithms.

## Limitation of PDE-constrained optimization

- Only provides a limited characterization of the posterior distribution.



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This work focuses using hyper-differential sensitivity analysis (HDSA) to analyze the influence of  $\boldsymbol{\theta}_{\text{prior}}$  on the MAP point for  $\mathbf{z}|\boldsymbol{\theta}_{\text{prior}}$ .

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# Post optimality sensitivity analysis

$$\min_{\mathbf{z} \in \mathbb{R}^m} J(\mathbf{z}; \boldsymbol{\theta}_{\text{prior}}) := M(\mathbf{z}, \boldsymbol{\theta}_{\text{prior}}) + R(\mathbf{z}, \boldsymbol{\theta}_{\text{prior}})$$

- Let  $\mathbf{z}^*$  denote a local minimum when  $\boldsymbol{\theta} = \boldsymbol{\theta}_{\text{prior}}$  is fixed,

$$\nabla_{\mathbf{z}} J(\mathbf{z}^*, \boldsymbol{\theta}_{\text{prior}}) = 0 \quad \text{and} \quad \nabla_{\mathbf{z}, \mathbf{z}} J(\mathbf{z}^*, \boldsymbol{\theta}_{\text{prior}}) \succ 0$$

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- The implicit function theorem gives

$$\mathcal{G} : \mathcal{N}(\boldsymbol{\theta}_{\text{prior}}) \rightarrow \mathcal{N}(\mathbf{z}^*),$$

defined on neighborhoods of  $\boldsymbol{\theta}_{\text{prior}}$  and  $\mathbf{z}^*$ , such that

$$\nabla_{\mathbf{z}} J(\mathcal{G}(\boldsymbol{\theta}), \boldsymbol{\theta}) = 0 \quad \forall \boldsymbol{\theta} \in \mathcal{N}(\boldsymbol{\theta}_{\text{prior}})$$

# Post optimality sensitivity analysis

$$\mathcal{G} : \mathcal{N}(\boldsymbol{\theta}_{\text{prior}}) \rightarrow \mathcal{N}(\mathbf{z}^*)$$

- $\mathcal{G}$  associates parameters  $\boldsymbol{\theta}$  with the corresponding MAP points for  $\mathbf{z}$  given  $\boldsymbol{\theta}$
- Further,  $\mathcal{G}$  is differentiable at  $\boldsymbol{\theta}_{\text{prior}}$  and its Jacobian is given by

$$\mathcal{G}'(\boldsymbol{\theta}_{\text{prior}}) = -\mathcal{H}^{-1}\mathcal{B} \in \mathbb{R}^{m \times n}$$

- $\mathcal{H} = \nabla_{\mathbf{z}, \mathbf{z}} J$  is the Hessian of  $J$  and,
- $\mathcal{B} = \nabla_{\mathbf{z}, \boldsymbol{\theta}} J$  is the Jacobian of  $\nabla_{\mathbf{z}} J$  with respect to  $\boldsymbol{\theta}$

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$\mathcal{G}'(\boldsymbol{\theta}_{\text{prior}})$  is like a Newton step to update the MAP point given a perturbation of  $\boldsymbol{\theta}$ .

# Hyper-differential sensitivity analysis

Compute properties of

$$\mathcal{G}'(\boldsymbol{\theta}_{\text{prior}}) = -\mathcal{H}^{-1}\mathcal{B} \in \mathbb{R}^{m \times n},$$

a large dense matrix which is only accessible through matrix vector products.

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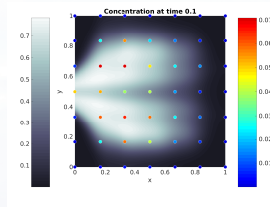
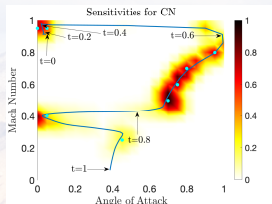
- Compute matrix-vector products with  $\mathcal{H}$  and  $\mathcal{B}$  using
  - adjoint-based derivative computations,
  - matrix free linear algebra.
- Approximation  $\mathcal{G}'(\boldsymbol{\theta}_{\text{prior}})$  via a randomized generalized SVD
  - efficiently expose low rank structure,
  - embarrassingly parallel,
  - appropriate inner products to facilitate the interpretation.



# Hyper-differential sensitivity analysis: Previous work

Has been used in the context of:

- Parameter uncertainty in PDE-constrained control
- Nuisance parameter uncertainty in deterministic inverse problems
- Data sensitivity to augment optimal experimental design
- Feedback controller robustness
- Model form error in PDE-constrained optimization



Has been applied in:

- Aerospace vehicle trajectory planning
- Thermal fluid system control
- Subsurface source inversion
- Ice sheet bedrock topography inversion

# Pulling it all together

- State-of-the-art tools enables efficient analysis for the sensitivity of  $\mathbf{z}$ 's MAP point with respect to perturbations of  $\theta_{\text{prior}}$ .
- Scalability in the parameter dimension is achieved through adjoint calculations and low rank approximation.

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**There are two questions we must address to facilitate this analysis:**

1. What is the Bayesian interpretation of  $\mathcal{G}'(\boldsymbol{\theta}_{\text{prior}})$ ?
2. Can I still compute/interpret  $\mathcal{G}'(\boldsymbol{\theta}_{\text{prior}})$  for ill-posed problems where:
  - An ill-conditioned Hessian introduces theoretical and computational challenges,
  - The optimizer may struggle to solve the MAP point estimation problem to optimality (satisfaction of the first order optimality condition)?

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# The linear case

## Theorem

*If  $F(\mathbf{z}, \boldsymbol{\theta}) = A\mathbf{z} + B\boldsymbol{\theta}$  then the posterior is Gaussian with covariance*

$$\Sigma_{post} = \begin{pmatrix} \Sigma_{z,z} & \Sigma_{z,\theta} \\ \Sigma_{\theta,z} & \Sigma_{\theta,\theta} \end{pmatrix}$$

*and the post-optimality sensitivity is given by*

$$\mathcal{G}'(\boldsymbol{\theta}_{prior}) = \Sigma_{z,\theta} \Sigma_{\theta,\theta}^{-1}.$$

- The post-optimality sensitivity is a correlation between  $\mathbf{z}$  and  $\boldsymbol{\theta}$ .

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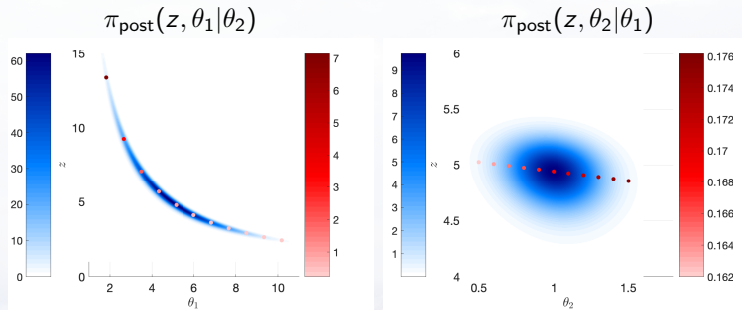
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- The post-optimality sensitivity is a correlation between  $\mathbf{z}$  and  $\boldsymbol{\theta}$ .
- Connection between **optimization/analysis** and **Bayesian statistics**.
- Local correlation for nonlinear inverse problems (Laplace approximation).



# An illustrative example

$$F(z, \theta_1, \theta_2) = e^{\frac{1}{10}z\theta_1} + \theta_2$$



- High sensitivity (left panel) corresponds to stronger correlations
- Approximately Gaussian distribution (right panel) has nearly constant sensitivity (local correlations = global correlations for Gaussians)

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$$\mathcal{G}'(\boldsymbol{\theta}_{\text{prior}}) = -\mathcal{H}^{-1}\mathcal{B}$$

- For ill-posed inverse problems,  $\mathcal{H}$  may be ill-conditioned.
- Can yield high sensitivity as a result of lacking information/data.
- Analyzing  $\mathcal{G}'(\boldsymbol{\theta}_{\text{prior}})$  will be numerically troublesome and the resulting sensitivities may be dominated by what the data does not tell you.

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**Proposed Approach:** Compute sensitivities

$$\mathcal{P}\mathcal{G}'(\boldsymbol{\theta}_{\text{prior}}) = -\mathcal{P}\mathcal{H}^{-1}\mathcal{B}$$

where  $\mathcal{P}$  projects onto the likelihood informed subspace <sup>2</sup>.

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# Likelihood informed subspaces (LIS)

$$\mathcal{H}_M \mathbf{v}_j = \lambda_j \mathcal{H}_R \mathbf{v}_j$$

$$\mathcal{H} = \mathcal{H}_M + \mathcal{H}_R$$

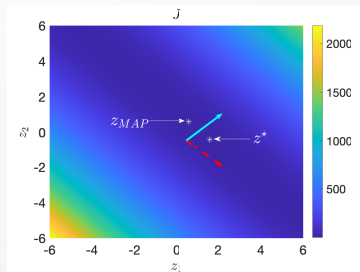
The eigenvalue

$$\lambda_j = \frac{\mathbf{v}_j^T \mathcal{H}_M \mathbf{v}_j}{\mathbf{v}_j^T \mathcal{H}_R \mathbf{v}_j}$$

measures the ratio of the **likelihood** and **prior** in the direction of  $\mathbf{v}_j$ .

$$\mathcal{P}\mathcal{G}'(\boldsymbol{\theta}_{\text{prior}}) = -\mathcal{P}\mathcal{H}^{-1}\mathcal{B}$$

Project the sensitivities onto the span of  $\{\mathbf{v}_1, \mathbf{v}_2, \dots, \mathbf{v}_r\}$ , the subspace which is most informed by the **likelihood**.



## Theorem

*The LIS sensitivity is given by*

$$S(\bar{\theta}) = \|\mathcal{P}\mathcal{H}^{-1}\mathcal{B}\bar{\theta}\|_{W_Z} = \sqrt{\sum_{k=1}^r \sum_{j=1}^r \left( \frac{\mathbf{v}_j^T \mathcal{B}\bar{\theta}}{1 + \lambda_j} \right) \left( \frac{\mathbf{v}_k^T \mathcal{B}\bar{\theta}}{1 + \lambda_k} \right) \mathbf{v}_k^T W_Z \mathbf{v}_j}.$$

- Need the leading eigenpairs

$$\mathcal{H}_M \mathbf{v}_j = \lambda_j \mathcal{H}_R \mathbf{v}_j$$

rather than  $\mathcal{H}^{-1}$ .

- Sensitivities inherit the likelihood to prior ratio interpretation.
- $W_Z$  measures the inner products in the original function space.

# Failure to satisfy the first order optimality condition?

It may not be practical to solve the MAP point estimation problem

$$\min_{\mathbf{z} \in \mathbb{R}^m} J(\mathbf{z}; \boldsymbol{\theta}_{\text{prior}}) := M(\mathbf{z}, \boldsymbol{\theta}_{\text{prior}}) + R(\mathbf{z}, \boldsymbol{\theta}_{\text{prior}})$$

to optimality if ill-conditioning yields slow convergence.

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**Question:** HDSA assume satisfaction of the optimality criteria. What can I do when converging the optimization is impractical/unnecessary?



# Failure to satisfy the first order optimality condition?

**Idea:** Compute sensitivities of a nearby problem which is solved to optimality.



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- Assume that  $\mathbf{z}^*$  is an approximation of the MAP point but

$$\nabla_{\mathbf{z}} J(\mathbf{z}^*, \boldsymbol{\theta}_{\text{prior}}) \neq 0.$$

- Find a minimum norm perturbation  $\tilde{R}$  so that  $\nabla_{\mathbf{z}} J(\mathbf{z}^*; \boldsymbol{\theta}_{\text{prior}}) + \nabla_{\mathbf{z}} \tilde{R}(\mathbf{z}^*) = 0$ ,

$$\begin{aligned} \min_{\tilde{R} \in Q} \|\tilde{R}\|_{L^1(\mu)} \\ \text{s.t. } \nabla_{\mathbf{z}} \tilde{R}(\mathbf{z}^*) = -\nabla_{\mathbf{z}} J(\mathbf{z}^*; \boldsymbol{\theta}_{\text{prior}}) \end{aligned}$$

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where

- $Q = \{\tilde{R} : \mathbb{R}^m \rightarrow \mathbb{R} \mid \tilde{R} \geq 0, \tilde{R} \text{ is quadratic, } \tilde{R} \text{ is convex}\}$
- $L^1(\mu)$  is defined by a Gaussian measure  $\mu$  with mean  $\mathbf{z}^*$  and covariance  $\alpha^2 I$
- $\alpha$  is a user defined length scale parameter (will revisit it later)

# First order a posteriori update

Can solve

$$\begin{aligned} \min_{\tilde{R} \in Q} \|\tilde{R}\|_{L^1(\mu)} \\ \text{s.t. } \nabla_z \tilde{R}(\mathbf{z}^*) = -\nabla_z J(\mathbf{z}^*; \theta_{\text{prior}}) \end{aligned}$$

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in closed form (with judicious algebraic manipulations) to find the update

$$\tilde{R}(\mathbf{z}) = \frac{\alpha}{2} \|\mathbf{g}\|_2 - (\mathbf{z} - \mathbf{z}^*)^T \mathbf{g} + \frac{1}{2} (\mathbf{z} - \mathbf{z}^*)^T \frac{1}{\alpha \|\mathbf{g}\|_2} \mathbf{g} \mathbf{g}^T (\mathbf{z} - \mathbf{z}^*),$$

where  $\mathbf{g} = \nabla_z J(\mathbf{z}^*; \boldsymbol{\theta}_{\text{prior}})$ .

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where  $\mathbf{g} = \nabla_z J(\mathbf{z}^*; \boldsymbol{\theta}_{\text{prior}})$ .

- $\tilde{R}$  is a “nice” function.
- The length scale parameter  $\alpha$  dictates the curvature.
- Computational cost is negligible since a closed form expression is available.

# The perturbed MAP point problem

$$\min_{\mathbf{z} \in \mathbb{R}^m} J(\mathbf{z}; \boldsymbol{\theta}_{\text{prior}}) + \tilde{R}(\mathbf{z})$$

- $\mathbf{z}^*$  satisfies the first order optimality condition.
- Post-optimality sensitivities are well defined for this perturbed problem.

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Some important questions:

- What is the Bayesian interpretation of  $\tilde{R}$ ?
- How does the perturbation influence the sensitivities?
- How do the sensitivities depend on  $\alpha$  (the length scale parameter)?



# A perturbed Gaussian prior

- What is the Bayesian interpretation of  $\tilde{R}$ ?
- How does the perturbation influence the sensitivities?
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## Theorem

*The perturbed inverse problem has a Gaussian prior with mean*

$$\tilde{\mathbf{z}}_{prior} = \mathbf{z}_{prior} + \frac{\alpha - (\mathbf{z}^* - \mathbf{z}_{prior})^T \mathbf{s}}{\alpha - \mathbf{v}^T \mathbf{s}} \mathbf{v}$$

*and covariance*

$$\tilde{\Gamma}_{prior} = \Gamma_{prior} - \frac{1}{\|\mathbf{g}\|_2} \frac{1}{\alpha - \mathbf{v}^T \mathbf{s}} \mathbf{v} \mathbf{v}^T$$

*where*

$$\mathbf{g} = \nabla_{\mathbf{z}} J(\mathbf{z}^*; \boldsymbol{\theta}_{prior}), \quad \mathbf{s} = -\frac{\mathbf{g}}{\|\mathbf{g}\|_2}, \quad \text{and} \quad \mathbf{v} = \Gamma_{prior} \mathbf{g}.$$

*The perturbation shifts the mean and reduces uncertainty in the direction  $\mathbf{v}$ .*

# Difference in sensitivities

- What is the Bayesian interpretation of  $\tilde{R}$ ?
- How does the perturbation influence the sensitivities?
- How do the sensitivities depend on  $\alpha$  (the length scale parameter)?

## Theorem

*The quantities*

$$S(\bar{\theta}) = \|\mathcal{P}\mathcal{H}^{-1}\mathcal{B}\bar{\theta}\|_{w_z} \quad \text{and} \quad \tilde{S}(\bar{\theta}) = \|\mathcal{P}\tilde{\mathcal{H}}^{-1}\tilde{\mathcal{B}}\bar{\theta}\|_{w_z}$$

*satisfy*

$$\frac{|\tilde{S}(\bar{\theta}) - S(\bar{\theta})|}{\|\mathcal{H}^{-1}\mathcal{B}\bar{\theta}\|_{w_z}} \leq \frac{\|\mathcal{P}\mathbf{n}\|_{w_z}}{\mathbf{s}^T \mathbf{n} + \alpha},$$

*where*

$$\mathbf{g} = \nabla_z J(\mathbf{z}^*; \theta_{\text{prior}}), \quad \mathbf{s} = -\frac{\mathbf{g}}{\|\mathbf{g}\|_2}, \quad \text{and} \quad \mathbf{n} = -\mathcal{H}^{-1}\mathbf{g}.$$

# Robustness with respect to $\alpha$

- What is the Bayesian interpretation of  $\tilde{R}$ ?
- How does the perturbation influence the sensitivities?
- How do the sensitivities depend on  $\alpha$  (the length scale parameter)?

## Theorem

Letting  $\tilde{S}_\alpha(\bar{\theta})$  be the sensitivity as a function of  $\alpha$ ,

$$\frac{|\tilde{S}_{\alpha+\alpha\beta}(\bar{\theta}) - \tilde{S}_\alpha(\bar{\theta})|}{\|\mathcal{H}^{-1}\mathcal{B}\mathbf{e}_i\|_{W_Z}} < |\beta| \cdot \frac{\|\mathcal{P}\mathbf{n}\|_{W_Z}}{\mathbf{s}^T \mathbf{n} + \alpha(1 + \beta)} \quad \text{for } -1 < \beta < 1$$

where

$$\mathbf{g} = \nabla_{\mathbf{z}} J(\mathbf{z}^*; \boldsymbol{\theta}_{\text{prior}}), \quad \mathbf{s} = -\frac{\mathbf{g}}{\|\mathbf{g}\|_2}, \quad \text{and} \quad \mathbf{n} = -\mathcal{H}^{-1} \mathbf{g}.$$

# First order a posteriori update

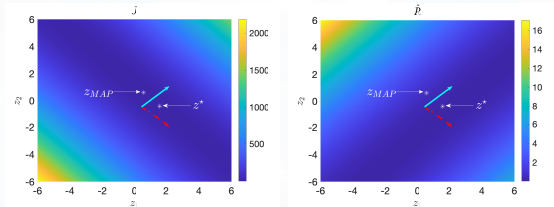
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# First order a posteriori update

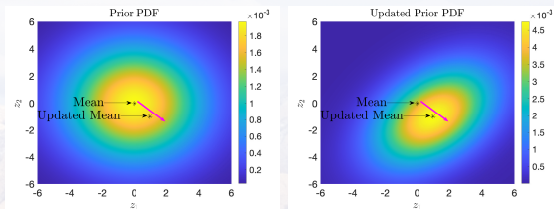
- What is the Bayesian interpretation of  $\tilde{R}$ ?
- How does the perturbation influence the sensitivities?
- How do the sensitivities depend on  $\alpha$  (the length scale parameter)?
- $\tilde{R}$  shifts the prior mean and reduces the variance in the direction  $\mathbf{v} = \Gamma_{\text{prior}} \mathbf{g}$ .
- The change in the sensitivity indices and their robustness with respect to  $\alpha$  are bounded by quantities proportional to  $\|\mathcal{P}\mathbf{n}\|$ .

**Take away message:** If the optimizer has converged in the likelihood informed subspace, then HDSA is a robust, interpretable, scalable, and efficient (RISE) way to assess correlations in the joint Bayesian posterior distribution.

# An illustrative example



**Figure:** The solid cyan arrow indicates the likelihood informed subspace and the broken red arrow indicates the uninformed subspace.



**Figure:** The magenta arrow indicates the direction of  $\mathbf{v} = \Gamma_{\text{prior}} \mathbf{g}$ .

# The two questions

1. What is the Bayesian interpretation of  $\mathcal{G}'(\theta_{\text{prior}})$ ?

The local correlation between  $\mathbf{z}$  and  $\theta$  in the Bayesian posterior.

2. Can I still compute/interpret  $\mathcal{G}'(\theta_{\text{prior}})$  for ill-posed problems where:

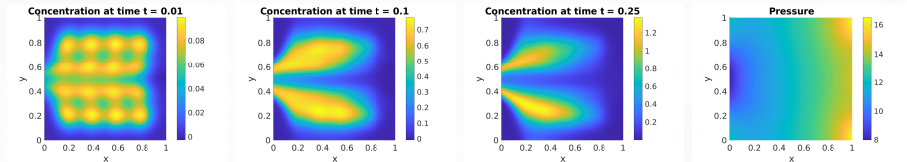
- An ill-conditioned Hessian introduces theoretical and computational challenges, Project on likelihood informed subspaces.
- The optimizer may struggle to solve the MAP point estimation problem to optimality (satisfaction of the first optimality condition)?

A posteriori update.

- 1 Bayesian inverse problems
- 2 Hyper-differential sensitivity analysis (HDSA)
- 3 The Bayesian interpretation of HDSA
- 4 Enabling HDSA for ill-posed problems
- 5 Numerical Results



# Subsurface permeability inversion



$$-\nabla \cdot (e^{\kappa} \nabla p) = 0$$

in  $\Omega$

$$c_t - \nabla \cdot (\epsilon(\theta) \nabla c) + \nabla \cdot (-e^{\kappa} \nabla p c) = g(\theta)$$

in  $[0, T] \times \Omega$

$$p = p_1(\theta)$$

on  $\Gamma_1$

$$p = p_2(\theta)$$

on  $\Gamma_3$

$$e^{\kappa} \nabla p \cdot n = 0$$

on  $\Gamma_0 \cup \Gamma_2$

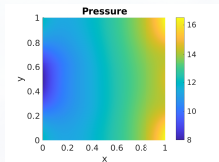
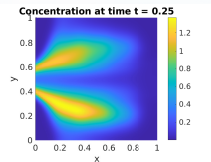
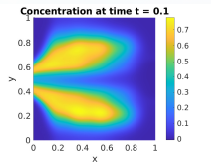
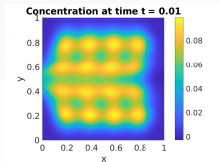
$$\nabla c \cdot n = 0$$

on  $[0, T] \times \{\Gamma_0 \cup \Gamma_1 \cup \Gamma_2 \cup \Gamma_3\}$

$$c(0, x) = 0$$

in  $\Omega$

# Subsurface permeability inversion



$$-\nabla \cdot (e^{\kappa} \nabla p) = 0$$

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on  $\Gamma_0 \cup \Gamma_2$

$$\nabla c \cdot n = 0$$

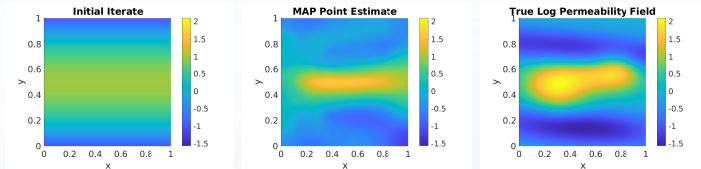
on  $[0, T] \times \{\Gamma_0 \cup \Gamma_1 \cup \Gamma_2 \cup \Gamma_3\}$

$$c(0, x) = 0$$

in  $\Omega$

# Subsurface permeability inversion

$$\min_{\kappa} \sum_{i=1}^{n_c} w_c(\mathcal{Q}_c^i c(\kappa) - d_c^i)^2 + \sum_{i=1}^{n_p} w_p(\mathcal{Q}_p^i p(\kappa) - d_p^i)^2 + \gamma_1 \|\nabla \kappa\|^2 + \gamma_2 \|\kappa\|^2$$

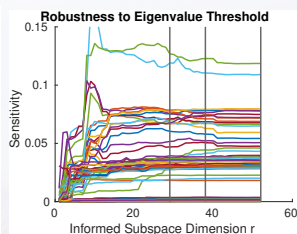
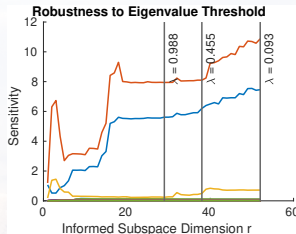
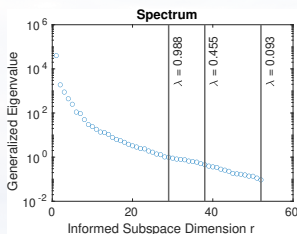


Iteration	Objective	Gradient Norm	Step Size
0	17.2	1.33	N/A
4	9.59	.697	15.6
10	3.29	.676	2.38
41	.897	.113	2.02
65	.578	.331	.115
75	.571	.102	.109
125	.529	.081	.034

Significant computational efforts gives little improvement.

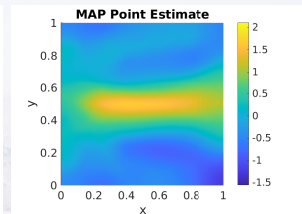
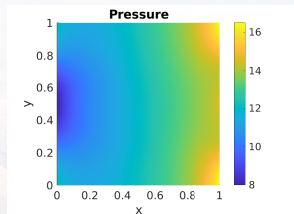
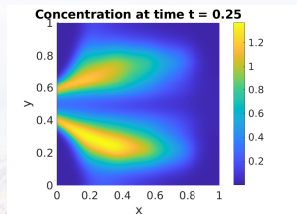
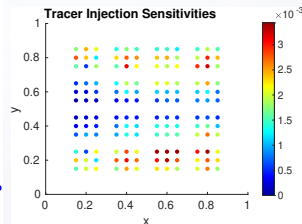
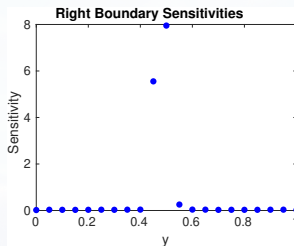
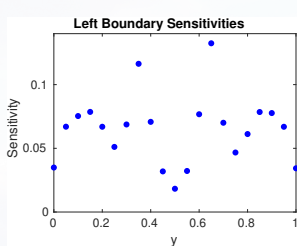
# Computing the likelihood informed subspace

$$S_i = \|\mathcal{PH}^{-1}\mathcal{B}\mathbf{e}_i\| = \sqrt{\sum_{k=1}^r \sum_{j=1}^r \left( \frac{\mathbf{v}_j^T \mathcal{B}\mathbf{e}_i}{1 + \lambda_j} \right) \left( \frac{\mathbf{v}_k^T \mathcal{B}\mathbf{e}_i}{1 + \lambda_k} \right) \mathbf{v}_k^T W_Z \mathbf{v}_j}$$



Automatically check for robustness with respect to rank truncation choice.

# Sensitivities



# Summary

- Established the Bayesian interpretation of post-optimality sensitivity analysis.
- Addressed ill-conditioning by projecting on likelihood informed subspaces.
- Theoretically justified HDSA when optimization fails to converge.
- Provided strong error bounds establishing the robustness of the analysis.
- HDSA gives a robust, interpretable, scalable, and efficient (RISE) way to assess correlations in the joint Bayesian posterior distribution.

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