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Parameter estimation for system submodels with limited or missing data using a data-free inference procedure

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Introduction

Models for physical systems of practical interest often are often characterized by hierarchies of models, e.g. continuum approximations

Such approaches inevitably require closure models for sub-continuum processes, e.g. constitutive relations

This may involve the evaluation of detailed models at the sub-continuum level, from which statistics may be extracted, or more commonly carefully designed experiments are performed to constrain these models with (typically noisy) measurement data

A robust treatment of uncertainty in such systems should involve analysis these experimental data sources, perhaps computing a joint probability density function (JPDF) to describe the model parameters



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Model parameter estimation in the absence of data

Given noisy evidential data, \mathbf{D}_0 , (e.g. from an experiment) statistical methodologies (e.g. Bayesian inference) can be employed to estimate joint posterior densities on model parameters, λ :

$$p(\lambda|\mathbf{D}_0) \propto p(\mathbf{D}_0|\lambda)p(\lambda) \quad (1)$$

However, raw experimental data are usually unreported, and for many legacy experiments the underlying data are missing or lost forever

What is reported are limited (low-dimensional) summary statistics, $S(\mathbf{D}_0)$, of the data, typically in the form of nominal fitting model parameters, interval uncertainty

Also perhaps nominal information about the experiment and fitting context (fitting model form, experiment initial and boundary conditions, etc.)

This limited form of reporting often omits important information about the parametric uncertainty e.g. correlations, hindering uncertainty quantification efforts



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Data-free inference (DFI)

Can we estimate the desired JPDF given the available information?

If the reported statistics of the data, $S(\mathbf{D}_0)$, are **sufficient** for the parameters λ , then computing $p(\lambda|S(\mathbf{D}_0))$ is equivalent to computing the desired $p(\lambda|\mathbf{D}_0)$.

The reported statistics $S(\mathbf{D}_0)$, are a unique mapping of the missing data. Estimating the inverse mapping would give access to plausible data sets, \mathbf{D} , consistent with these statistics, of which \mathbf{D}_0 is a member.

In a Bayesian framework, we can construct an inference to explore a space of data and associated parameters that satisfy the statistics, interpreted as constraints, with distribution $\Rightarrow p(\lambda, \mathbf{D}|S(\mathbf{D}_0))$

Applying Bayes' Law directly:

$$p(\lambda, \mathbf{D}|S(\mathbf{D}_0)) \propto p(S(\mathbf{D}_0)|\lambda, \mathbf{D}) p(\lambda|\mathbf{D}) p(\mathbf{D}) \quad (2)$$

$p(\lambda|\mathbf{D})$: can be computed if the fitting model is known

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Data-free inference: Two-step algorithm and Approximate Bayesian Computation (ABC)

Using a two-step approach, first constructing an inference on data:

$$p(\mathbf{D}|S(\mathbf{D}_0)) \propto p(S(\mathbf{D}_0)|\mathbf{D})p(\mathbf{D}) \quad (3)$$

$p(S(\mathbf{D}_0)|\mathbf{D})$: Approximate Bayesian Computation (ABC) approaches can be invoked to construct an approximate likelihood that enforces consistency of proposed data, \mathbf{D} , by comparing statistics of λ through the associated parameter posterior $p(\lambda|\mathbf{D})$ (which we can compute) with the reported statistics.

This requires constructing an inference on parameters:

$$p(\lambda|\mathbf{D}) \propto p(\mathbf{D}|\lambda)p(\lambda) \quad (4)$$

$p(S(\mathbf{D}_0)|\mathbf{D})$ can then be designed to enforce consistency (approximately) e.g.:

$$p(S(\mathbf{D}_0)|\mathbf{D}) \approx \exp\left(-\delta\left(\frac{S(\mathbf{D}_0) - S(\mathbf{D})}{S(\mathbf{D}_0)}\right)^2\right) \quad (5)$$

a Gaussian penalty kernel.



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Data-free inference: Maximum Entropy (MaxEnt)

With this nested inference construction we can generate samples from $p(\lambda, \mathbf{D} | S(\mathbf{D}_0))$, with \mathbf{D} from the outer inference on data and λ from the inner inference on parameters

Computing $p(\lambda | S(\mathbf{D}_0))$ is then achieved by marginalizing over data-space:

$$p(\lambda | S(\mathbf{D}_0)) = \int_{\mathbf{D}} p(\lambda, \mathbf{D}' | S(\mathbf{D}_0)) d\mathbf{D}' = \int_{\mathbf{D}} p(\lambda | \mathbf{D}') p(\mathbf{D}' | S(\mathbf{D}_0)) d\mathbf{D}' \quad (6)$$

corresponding to linear pooling of consistent parameter posteriors.

From an entropy perspective of Bayesian updating, the posterior $p(\lambda | \mathbf{D}_0)$ maximizes entropy relative to the prior.

For situations where the posterior conditioning variable is described by a distribution, the relative-entropy maximizing distribution corresponds to linear pooling of posteriors according to that distribution. This is precisely the case with the marginalization in Eq. 6, making $p(\lambda | S(\mathbf{D}_0))$ the maximum entropy (MaxEnt) density given the available information



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Example - line fit to noisy data

Consider the case of noisy data, fit with a simple line model assuming zero-mean additive *iid* Gaussian noise. The noisy data model is:

$$y(x) = mx + c + \sigma\varepsilon \quad (7)$$

where ε is a standard normally distributed random variable, σ is the standard deviation of the noise, and the model parameters are $\lambda = \{m, c\}$

This stochastic data model implies a data likelihood:

$$p(\mathbf{D}|\lambda) = \frac{1}{\sqrt{2\pi}\sigma} \prod_i \exp\left(-\frac{(y(x_i) - [mx_i + c])^2}{2\sigma^2}\right) \quad (8)$$

which enables construction of an inference on λ given data

Since the true data \mathbf{D}_0 , or the posterior $p(\lambda|\mathbf{D}_0)$, are typically unreported, the DFI construction employs the available information to approximate them



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$$S(\mathbf{D}_0) = \{\mu_m, \mu_c, \sigma_m, \sigma_c\} \quad (9)$$

The two-step DFI construction, outer (data) inference:

$$p(\mathbf{D}|S(\mathbf{D}_0)) = \frac{p(S(\mathbf{D}_0)|\mathbf{D})p(\mathbf{D})}{p(S(\mathbf{D}_0))} \quad (10)$$

where $p(S(\mathbf{D}_0)|\mathbf{D})$ is constructed using ABC:

$$p(S(\mathbf{D}_0)|\mathbf{D}) = \prod_k \exp\left(-\delta_k \left(\frac{S_k(\mathbf{D}_0) - S_k(\mathbf{D})}{S_k(\mathbf{D}_0)}\right)^2\right) \quad (11)$$

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$$p(\lambda|\mathbf{D}) = \frac{p(\mathbf{D}|\lambda)p(\lambda)}{p(\mathbf{D})} \quad (12)$$

Both outer and inner inferences performed numerically using Markov Chain Monte Carlo methods to compute samples $\{\mathbf{D}, \lambda\}$



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Define the data space, number of data points (e.g. defined in experiment)

Initialize the data in a region of high likelihood

Typically run the outer data inference to discover 1000s of data sets consistent with the statistics

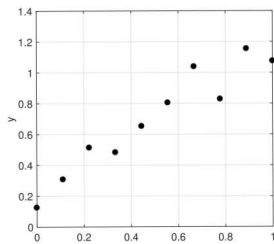
Linear pooling achieved by averaging over all consistent parameter posteriors

Is the statistics are sufficient, each consistent posterior should be identical, but some discrepancy from numerical approximations

Each individual consistent posterior constructed from consistent parameter samples using Kernel Density Estimation (KDE)

Result is the full approximation of the missing posterior, $p(\lambda | S(\mathbf{D}_0))$

sample noisy data set (\mathbf{D}) from outer inference



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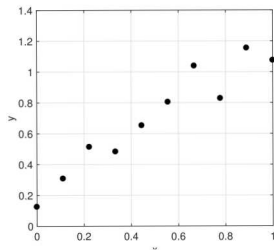
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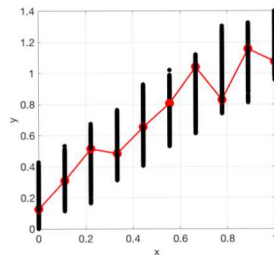
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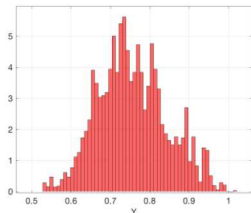
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histogram of data samples for $\mathbf{D}(x = 5\Delta x)$



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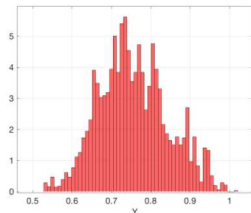
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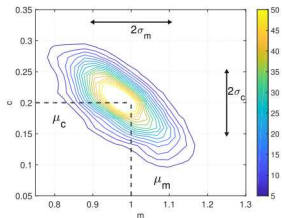
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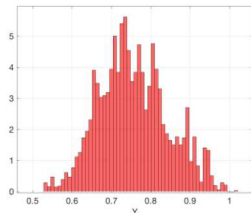
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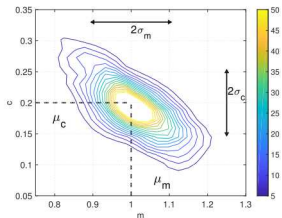
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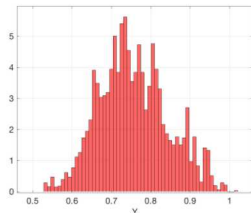
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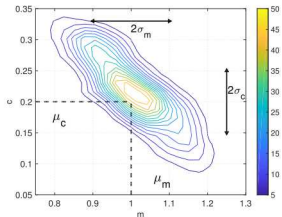
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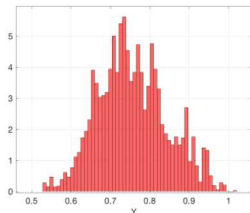
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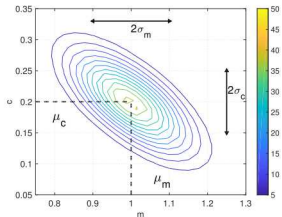
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Data-free inference: Model Replacement, Data Fusion

$$p(\lambda|S) = \int_{\mathbf{D}} p(\lambda, \mathbf{D}'|S) d\mathbf{D}' = \int_{\mathbf{D}} p(\lambda|\mathbf{D}') p(\mathbf{D}'|S) d\mathbf{D}' \quad (13)$$

MaxEnt parameter posterior corresponding to linear pooling of consistent parameter posteriors.

Model replacement: linearly pooling over a different posterior

$$p(\theta|S) = \int_{\mathbf{D}} p(\theta, \mathbf{D}'|S) d\mathbf{D}' = \int_{\mathbf{D}} p(\theta|\mathbf{D}') p(\mathbf{D}'|S) d\mathbf{D}' \quad (14)$$

Why do this? Desire a different fitting context than that used by the experimentalists

Data fusion: pooling over a posterior computed using multiple consistent data sources, e.g. from different reported experiments

$$\begin{aligned} p(\theta|S_1, S_2) &= \int_{\mathbf{D}_1, \mathbf{D}_2} p(\theta, \mathbf{D}'_1, \mathbf{D}'_2|S_1, S_2) d\mathbf{D}'_1 d\mathbf{D}'_2 \\ &= \int_{\mathbf{D}_1, \mathbf{D}_2} p(\theta|\mathbf{D}'_1, \mathbf{D}'_2) p(\mathbf{D}'_1|S_1) p(\mathbf{D}'_2|S_2) d\mathbf{D}'_1 d\mathbf{D}'_2 \end{aligned} \quad (15)$$



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Hierarchical models

Often we are interested in analyzing data at different conditions (e.g. temperature) to construct hierarchical models that span these conditions.

Chemical kinetics experiments: noisy exponential decay data model at different temperatures

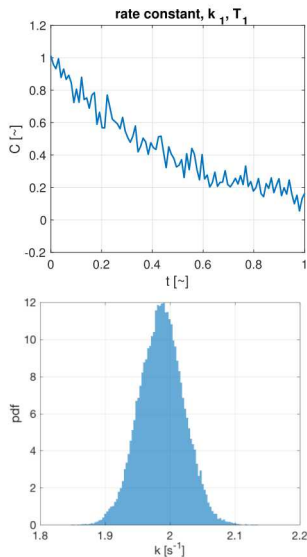
$$c(t) = \exp(-k(T)t) + \sigma\varepsilon \quad (16)$$

Hierarchical (Arrhenius) model couples the exponential decay constants across temperatures

$$k(T) = A \exp\left(-\frac{E}{T}\right) \quad (17)$$

DFI approach:

Infer consistent data for the reported statistics of $k(T)$



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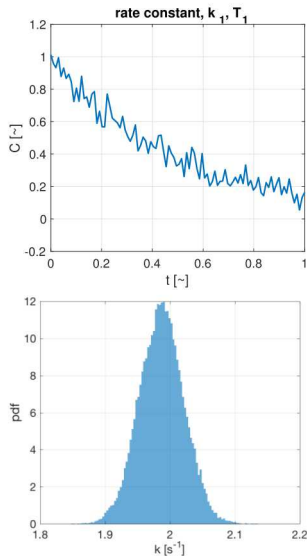
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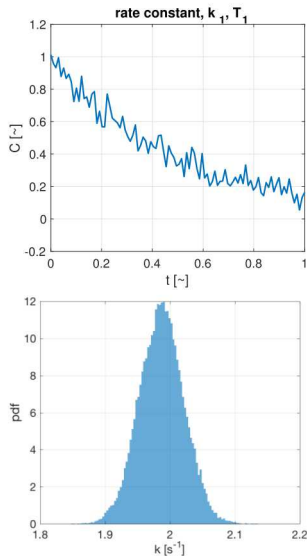
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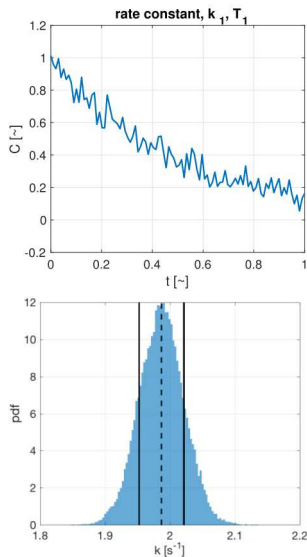
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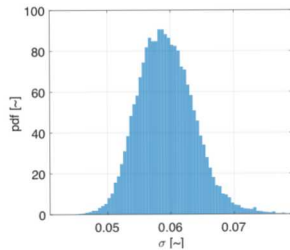
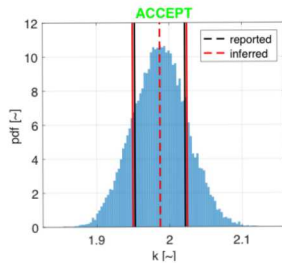
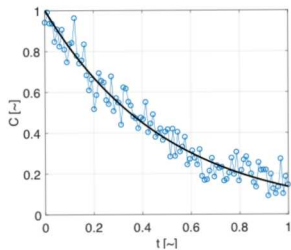
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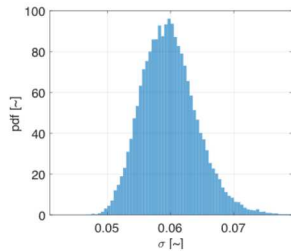
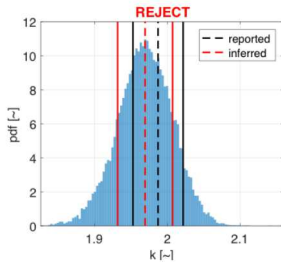
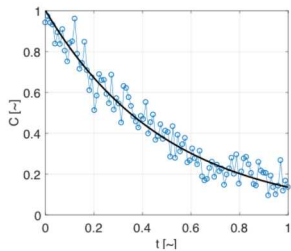
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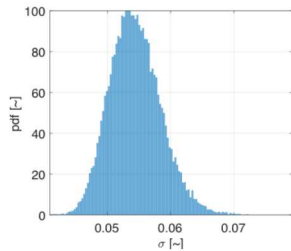
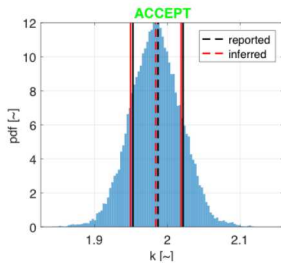
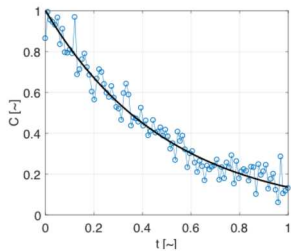
Hierarchical models



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Hierarchical models

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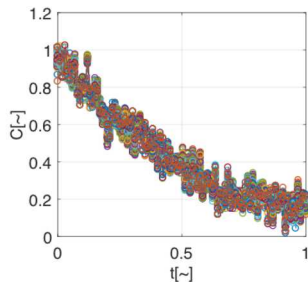
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But really want to calibrate the hierarchical model across temperatures

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Infer data for this second experiment, given statistics on $k(T_2)$

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Use data fusion and model replacement to learn the posterior on the parameters of the Arrhenius (hierarchical) model, $p(\log A, E | S_1, S_2)$



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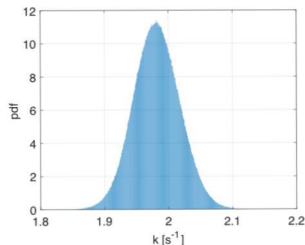
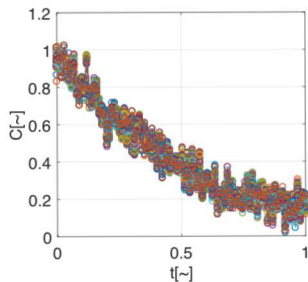
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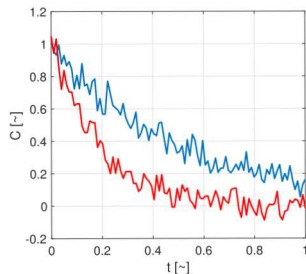
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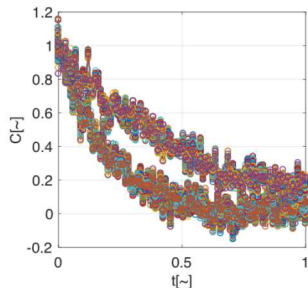
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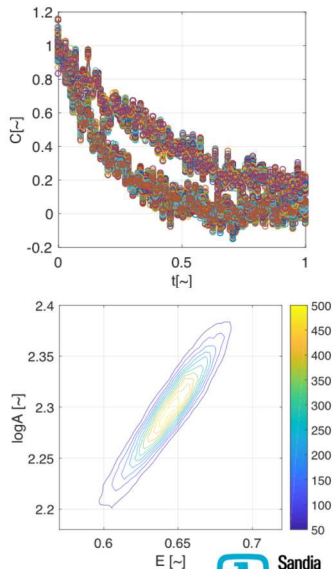
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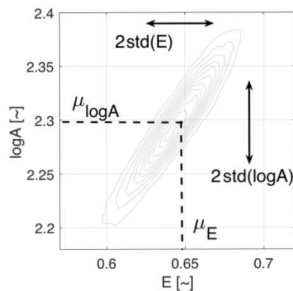


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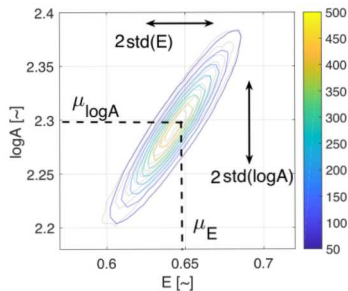


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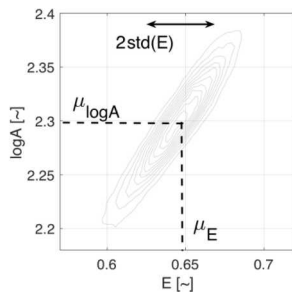
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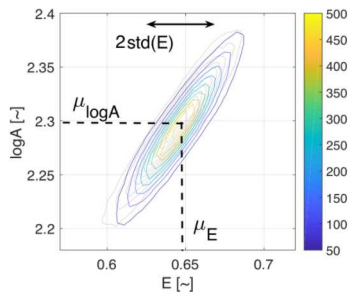
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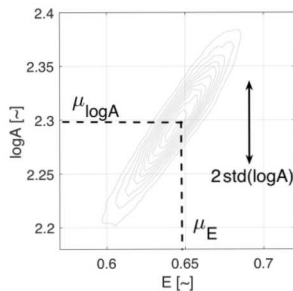
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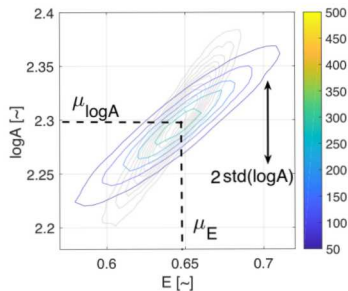
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Real experiments - combustion chemical kinetics

Can extend this hierarchical learning approach to real experiments, e.g. H_2-O_2 chemical kinetics

chemical ODE model: assembly of parameterized reactions, $k_i(T) = A_i \exp\left(-\frac{E_i}{T}\right)$



Pirraglia et al., J. Phys. Chem. (1989)



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```

ELEMENTS
O H AR N
END
SPECIES
H2 H 0 O2 OH H2O HO2 H2O2
AR N2
END
! THERMO
! Insert GRI-Mech thermodynamics here or use in default file
!END
REACTIONS
H+O2<=>O+OH 3.558E+15 -0.410 16600.00
H2+OH<=>H2O+H 2.160E+08 1.510 3430.00
H2O2(+M)<=>2OH(+M) 2.950E+14 0.000 48400.00
LOW / 1.200E+17 0.000 45500.0/
O+H2<=>H+OH 5.000E+04 2.670 6290.00
O+H2O<=>2OH 2.970E+06 2.020 13400.00
H2+H<=>2H+M 4.500E+19 -1.400 104300.00
H2/2.5/ H2O/12.0/
H2+AR<=>2H+AR 5.840E+10 -1.100 104300.00
2O+M<=>O2+M 6.160E+15 -0.500 0.00
H2/2.5/ H2O/12.0/
2O+AR<=>O2+AR 4.710E+13 0.000 -1790.00
O+H+M<=>OH+M 4.710E+13 0.000 -1790.00
H2/2.5/ H2O/12.0/
H+OH+H<=>H2O+H 3.800E+22 -2.000 0.00
H2/2.5/ H2O/12.0/ AR/0.38/
H+O2(+M)<=>HO2(+M) 1.400E+12 0.600 0.00
LOW / 9.040E+19 -1.500 490.00/
TROE / 0.50 1.0E+10 1.0E-10 /
HO2+H<=>H2+O2 1.660E+13 0.000 820.00
HO2+H<=>2OH 7.000E+13 0.000 300.00
HO2+O<=>OH+O2 3.250E+13 0.000 0.00
HO2+OH<=>H2O+O2 2.890E+13 0.000 -500.00
2HO2<=>O2+H2O2 4.200E+14 0.000 11900.00
DUPLICATE
2HO2<=>O2+H2O2 1.300E+11 0.000 -1603.00
DUPLICATE
H2O2+H<=>H2O+OH 2.410E+13 0.000 3970.00
H2O2+H<=>H2+HO2 4.820E+13 0.000 7950.00
H2O2+O<=>OH+HO2 9.550E+06 2.000 3970.00
H2O2+OH<=>H2O+HO2 1.000E+12 0.000 0.00
DUPLICATE
H2O2+OH<=>H2O+HO2 5.800E+14 0.000 9560.00
DUPLICATE
END
    
```



Real experiments - combustion chemical kinetics

Can extend this hierarchical learning approach to real experiments, e.g. H_2-O_2 chemical kinetics

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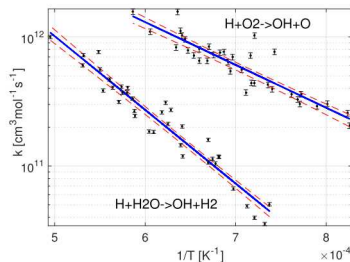


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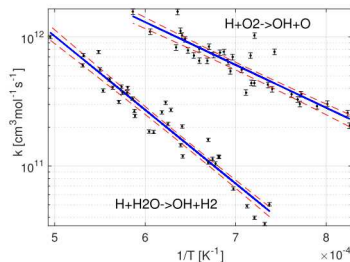


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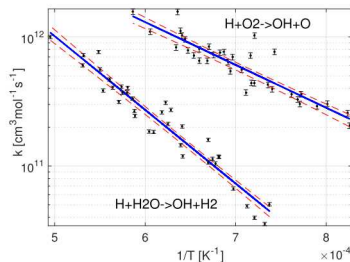


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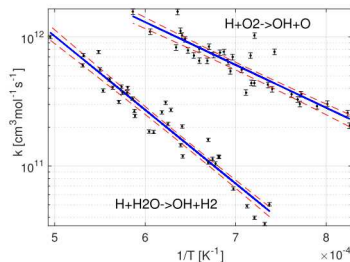


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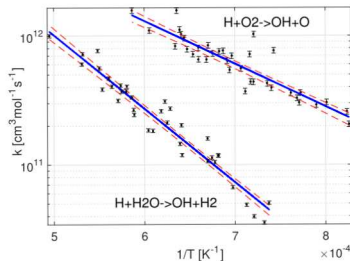
Real experiments - combustion chemical kinetics

Reaction 1:

Perform DFI procedure using reported $k(T)$ information at each temperature (35 in total)
Use data fusion and model replacement to learn the Arrhenius parameters of reaction 1 across all the data

Reaction 2:

repeat the procedure to learn the Arrhenius parameters of reaction 2



Real experiments - combustion chemical kinetics

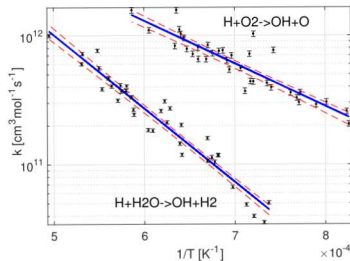
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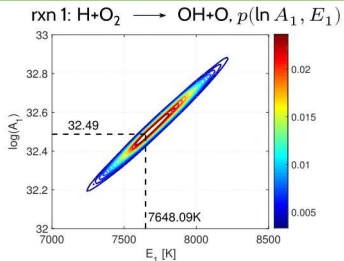
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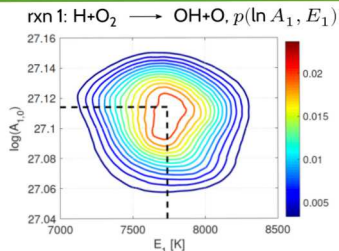
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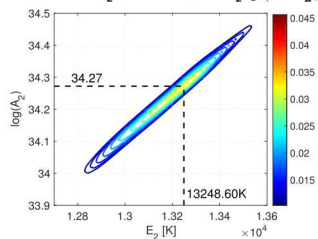
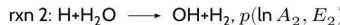
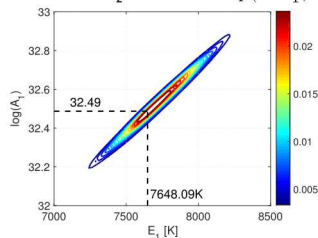
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Multiple reactions - combined learning

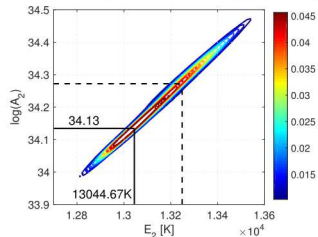
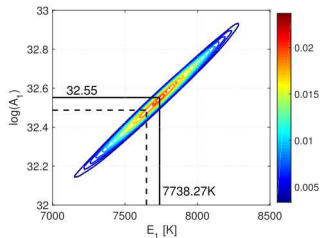
Learning all parameters across both reactions simultaneously

Data from rxn 1 experiment is informative on both sets of parameters
Marginal PDFs augmented compared to separate learning

Reveal subtle correlations between parameters across reactions

Cross-reaction correlation arises from experiment design, activity of 'nuisance' reactions
Analysis reveals temperature range where reactions are most coupled
Opportunity to learn parameters of both reactions from a single experiment

Can be extended to experiments for all reactions in the chemical model, to construct the full JPDP



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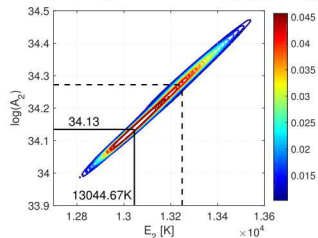
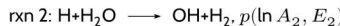
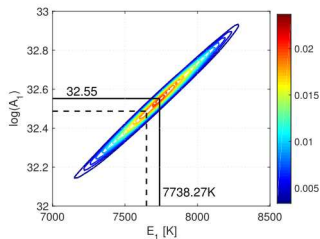
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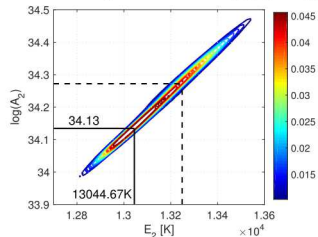
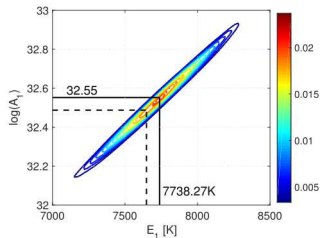
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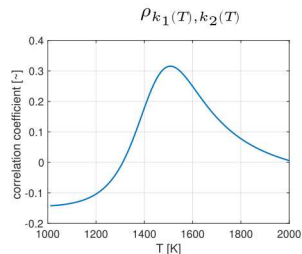
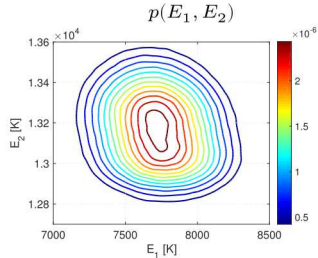
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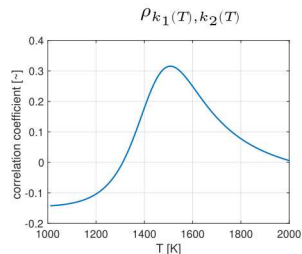
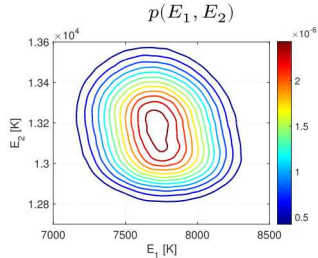
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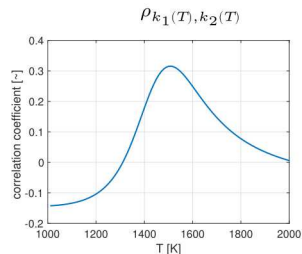
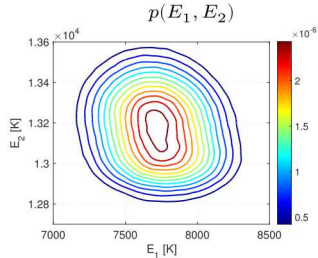
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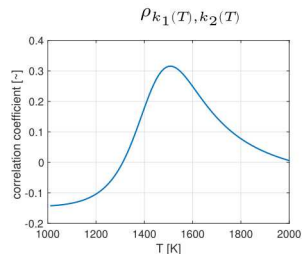
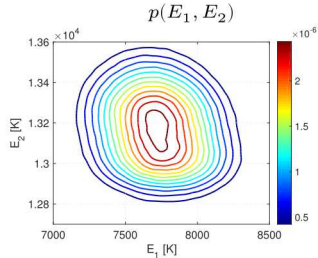
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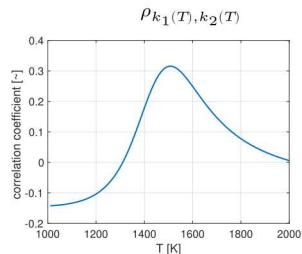
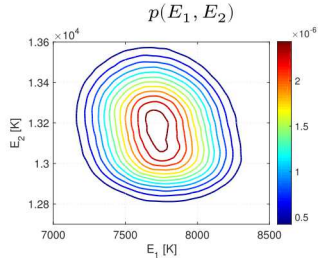
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Uncertainty propagation - predictive modeling

Drawing samples from the JPDF of all parameters, can propagate uncertainty through combustion models

Ignition delay time

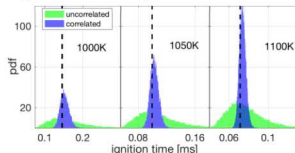
Laminar flame speed

Pressure-temperature explosion limits

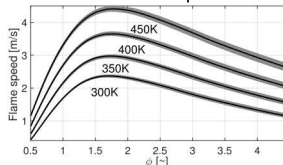
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Drastic over-prediction of uncertainty in quantities of interest (QoIs) when parameters are assumed independently distributed according to their marginal PDFs

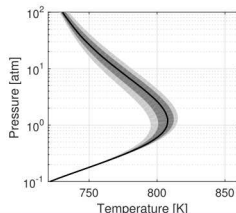
ignition time: impact of correlation



laminar flame speed



explosion limit curve



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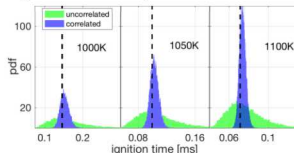
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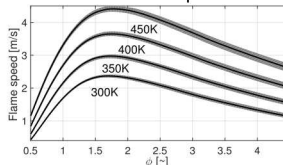
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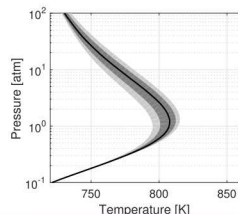
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Summary

For estimating parameter uncertainty when supporting data is unreported or missing, we have recourse to a data-free inference (DFI) procedure for constructing a MaxEnt posterior density consistent with available information

Demonstrated for simple data models, and hierarchical models with variety of reported summary statistics at different modeling levels

The MaxEnt linear pooling procedure enables combination of hypothetical data sets, model replacement for a desired fitting context

Applied to real missing experimental data sets in combustion chemical kinetics. Discover parameter uncertainty structure implied by the reported information

When missing data sets are informative on multiple parameters in the model, recover correlation of model parameters across reactions in the combined learning setting, avoiding bias in the predictive model specification

Identify missed opportunities for learning, enable model selection



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Acknowledgement

M. Khalil, X. Huan, B.J. Debuschere, K. Sargsyan, C. Safta, K. Chowdhary, R.D. Berry

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DOE Office of Basic Energy Sciences, Div. of Chem. Sci., Geosci., & Biosci.

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