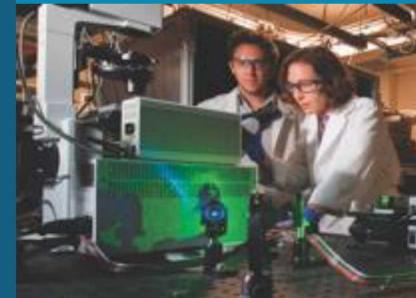




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TENSOR-BASED INFERENTIAL PROCEDURES FOR VIDEO DATA



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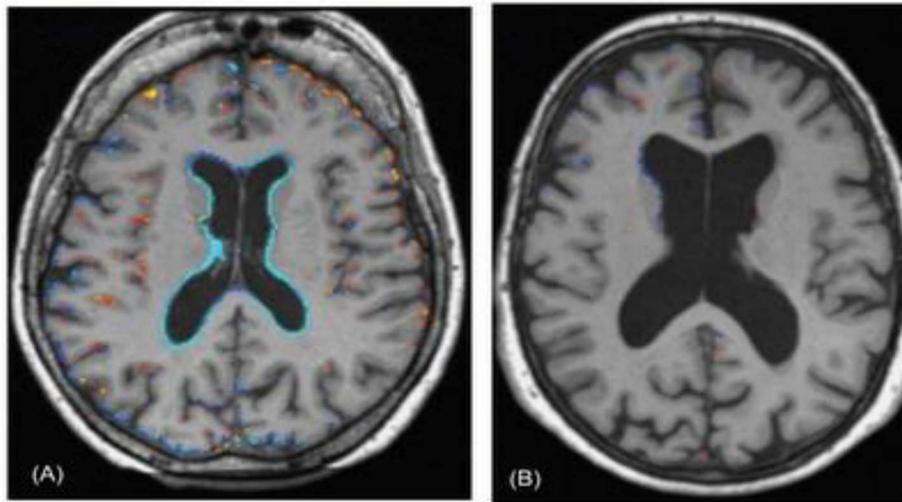


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APPLICATION: ANALYSIS OF POPULATION OF IMAGES



- Suppose we want to test a treatment for Alzheimer's disease.
- Data: Brain images of two groups of Alzheimer's patients with multiple images per patient taken over time
- Two groups: treatment and control



Two questions:

- How do we analyze these high-dimensional images ($120 \times \sim 277,000$) collectively?
- How do we determine if the treatment is effective?

APPLICATION TO TRAIN VIDEO DATA



BNSF train car



Single-stack train car



Double-stack train car

Two questions:

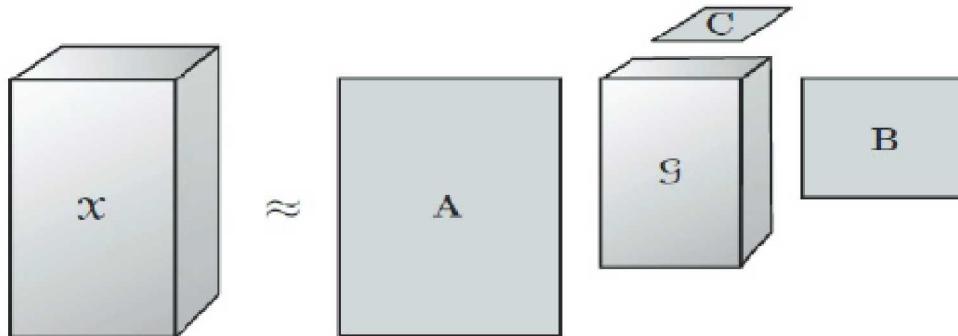
- How do we reduce the dimensions of these video frames?
- How do we detect significant differences between and within video segments?

OUTLINE



1. Three-way Tucker Decomposition and Tensor Preliminaries
2. Inferential Procedures
 - General Framework
 - Likelihood-Ratio Test
 - Score Test
 - Regression Based Inference
 - Simulations
 - Application to Train Video Data
3. Future Work

Tucker decomposition of a three-way array:



$$\mathfrak{X} \approx \mathcal{G} \times_1 \mathbf{A} \times_2 \mathbf{B} \times_3 \mathbf{C} = \sum_{p=1}^P \sum_{q=1}^Q \sum_{r=1}^R g_{pqr} \mathbf{a}_p \circ \mathbf{b}_q \circ \mathbf{c}_r = [[\mathcal{G}; \mathbf{A}, \mathbf{B}, \mathbf{C}]]$$

$$x_{ijk} \approx \sum_{p=1}^P \sum_{q=1}^Q \sum_{r=1}^R g_{pqr} a_{ip} b_{jq} c_{kr} \quad \text{for } i = 1, \dots, I, j = 1, \dots, J, k = 1, \dots, K$$

Question: Can we develop hypothesis testing procedures for this framework, where **C** measures the temporal correlation between frames?

MODE- n UNFOLDINGS



Tensors can be unfolded into various matrices that contain all of the elements of the tensor.

The best way to understand how this works is by considering an example. Let the frontal slides of a tensor $\mathcal{X} \in \mathbb{R}^{3 \times 4 \times 2}$ be

$$\mathbf{X}_1 = \begin{bmatrix} 1 & 4 & 7 & 10 \\ 2 & 5 & 8 & 11 \\ 3 & 6 & 9 & 12 \end{bmatrix}, \mathbf{X}_2 = \begin{bmatrix} 13 & 16 & 19 & 22 \\ 14 & 17 & 20 & 23 \\ 15 & 18 & 21 & 24 \end{bmatrix}.$$

Then the three mode- n unfoldings are

$$\mathbf{X}_{(1)} = \begin{bmatrix} 1 & 4 & 7 & 10 & 13 & 16 & 19 & 22 \\ 2 & 5 & 8 & 11 & 14 & 17 & 20 & 23 \\ 3 & 6 & 9 & 12 & 15 & 18 & 21 & 24 \end{bmatrix},$$

$$\mathbf{X}_{(2)} = \begin{bmatrix} 1 & 2 & 3 & 13 & 14 & 15 \\ 4 & 5 & 6 & 16 & 17 & 18 \\ 7 & 8 & 9 & 19 & 20 & 21 \\ 10 & 11 & 12 & 22 & 23 & 24 \end{bmatrix},$$

$$\mathbf{X}_{(3)} = \begin{bmatrix} 1 & 2 & 3 & 4 & 5 & \dots & 9 & 10 & 11 & 12 \\ 13 & 14 & 15 & 16 & 17 & \dots & 21 & 22 & 23 & 24 \end{bmatrix}.$$

TENSOR NORMAL DISTRIBUTION

Suppose Y_1, \dots, Y_r are dependent images that make up the *slices* of a tensor of order 3 that follows a tensor normal distribution with the following parameters:

$$\mathfrak{X} = Y_1, \dots, Y_r, \mathcal{G} = V_1, \dots, V_r, \mathbf{A}, \mathbf{B}, \mathbf{C} = \Omega,$$

where $\mathfrak{X} = Y_1, \dots, Y_r$ is of size $T \times F \times r$, $\mathcal{G} = V_1, \dots, V_r$ is of size $t \times f \times r$, $\mathbf{A} = P$ is of size $T \times t$, $\mathbf{B} = D$ is of size $f \times F$, and $\mathbf{C} = \Omega$ is of size $r \times r$.

Additionally, there will be the row and covariance matrices of Y_i , Σ and Ψ , respectively.

We assume that \mathfrak{X} follows a *tensor normal distribution*, written as
 $\mathfrak{X} \sim N_{T,F,r}(\mathfrak{B} \times \mathbf{A}, \mathbf{C}, \mathbf{D}, \Sigma, \Psi, \Omega)$.

The probability density function is

$$f_{\mathfrak{X}}(\mathfrak{X}) = (2\pi)^{-\frac{TFn}{2}} |\Sigma|^{-\frac{-Fn}{2}} |\Psi|^{-\frac{-Tn}{2}} |\Omega|^{-\frac{-TF}{2}} \exp\left\{-\frac{1}{2}(\mathfrak{X} - \mathbf{ABC}) \times_{1\dots 3} (\circ_{j=1}^3 U_j^{-1}) \times_{1\dots 3} (\mathfrak{X} - \mathbf{ABC})\right\},$$

where \circ denotes the outer product and \times denotes the tensor product.

A useful property of the tensor normal distribution:

$$\text{vec}(\mathfrak{B} \times \{\mathbf{A}, \mathbf{C}, \mathbf{D}\}) = (\mathbf{D} \otimes \mathbf{C} \otimes \mathbf{A}) \text{vec}(\mathfrak{B}) \quad (1)$$

$$\sim N_{TFr}((\mathbf{D} \otimes \mathbf{C} \otimes \mathbf{A}) \text{vec}(\mathfrak{B}), \Omega \otimes \Psi \otimes \Sigma) \quad (2)$$

ONE-SAMPLE PROBLEM

Suppose we have n i.i.d. tensors, denoted as \mathfrak{X}_i , $i = 1, \dots, n$, that follow a tensor normal distribution. For a one-sample hypothesis testing problem, we assume that \mathbf{A} , \mathbf{C} , and \mathbf{D} are fixed and computed.

Assumptions:

- One population:

$$\mathfrak{X}_i = \mathfrak{B} \times \mathbf{A}, \mathbf{C}, \mathbf{D} + \mathfrak{E}_i, i = 1, \dots, n$$

$$\mathfrak{X} \sim N_{T,F,r}(\mathfrak{B} \times \mathbf{A}, \mathbf{C}, \mathbf{D}, \Sigma, \Psi, \Omega)$$

$$\begin{aligned} \Rightarrow \text{vec}(\mathfrak{B} \times \mathbf{A}, \mathbf{C}, \mathbf{D}) &\sim N_{TFr}((\mathbf{D} \otimes \mathbf{C} \otimes \mathbf{A})\text{vec}(\mathfrak{B}), \Omega \otimes \Psi \otimes \Sigma) \\ &= (\mathbf{D} \otimes \mathbf{C} \otimes \mathbf{A})\text{vec}(\mathfrak{B}) \end{aligned}$$

- Row covariance matrix Σ is $T \times T$, fixed, p.d., but unknown
- Column covariance matrix Ψ is $F \times F$, fixed, p.d., but unknown
- Slice covariance matrix Ω is $r \times r$, fixed, p.d., but unknown

We are testing

$$H_0 : \mathcal{B} = \mathcal{B}_0$$

$$H_a : \mathcal{B} \neq \mathcal{B}_0$$

MAXIMUM-LIKELIHOOD ESTIMATION



We assume that \mathbf{A} and \mathbf{B} are calculated and fixed. We assume the covariances matrices Σ , Ψ , and Ω are unknown.

Nzabanita et al (2015) derive the maximum likelihood estimates for all of the parameters of the tensor normal distribution when the mean has the structure $\mathcal{M} = \mathcal{B} \times \{\mathbf{A}, \mathbf{C}, \mathbf{D}\}$.

These methods *unfold* the tensor into matrices. The model

$$\mathcal{X} \sim N_{T,F,r}(\mathcal{B} \times \{\mathbf{A}, \mathbf{C}, \mathbf{D}\}, \Sigma, \Psi, \Omega) \quad (3)$$

in matrix form using three different modes as

$$\mathbf{X}_{(1)} \sim N_{T,Fr}(\mathbf{A}\mathbf{B}_{(1)}(\mathbf{D} \otimes \mathbf{C})', \Sigma, \Omega \otimes \Psi) \quad (4)$$

$$\mathbf{X}_{(2)} \sim N_{F,Tr}(\mathbf{C}\mathbf{B}_{(2)}(\mathbf{D} \otimes \mathbf{A})', \Psi, \Omega \otimes \Sigma) \quad (5)$$

$$\mathbf{X}_{(3)} \sim N_{r,TF}(\mathbf{D}\mathbf{B}_{(3)}(\mathbf{C} \otimes \mathbf{A})', \Omega, \Psi \otimes \Sigma). \quad (6)$$

The likelihoods for each of the three modes are equivalent. MLEs for all of these parameters $(\mathbf{B}_{(1)}, \mathbf{B}_{(2)}, \mathbf{B}_{(3)}, \Sigma, \Psi, \Omega)$ are derived.

MATRIX NORMAL DISTRIBUTION



A random matrix X of dimensions $T \times F$ that follows the matrix normal distribution has the pdf

$$p(X|M, \Sigma, \Omega) = \frac{\exp(-\frac{1}{2}\text{tr}[\Sigma^{-1}(X - M)\Omega^{-1}(X - M)'])}{(2\pi)^{TF/2}|\Omega|^{T/2}|\Sigma|^{F/2}},$$

where M is the $T \times F$ mean matrix, Σ is the $T \times T$ row covariance matrix, and Ω is the $F \times F$ column covariance matrix.

The matrix normal distribution is related to the multivariate normal distribution.

$$X \sim MN_{T \times F}(M, \Sigma, \Omega)$$

if and only if

$$\text{vec}(X) \sim N_{TF}(\text{vec}(M), \Omega \otimes \Sigma).$$

ONE-SAMPLE PROBLEM: LIKELIHOOD-RATIO TEST

The likelihood using the mode-1 model with

$$\mathbf{X}_{(1)} \sim N_{T, Fr}(\mathbf{AB}_{(1)}(\mathbf{D} \otimes \mathbf{C})', \Sigma, \Omega \otimes \Psi)$$

is

$$L(B_{(1)} | \mathbf{A}, \mathbf{C}, \mathbf{D}, x_{(1),1}, \dots, x_{(1),n}) = \frac{\exp(-\frac{1}{2} \sum_{i=1}^n \text{tr}\{(\Omega \otimes \Psi)^{-1} [\mathbf{X}_{(1),i} - \mathbf{AB}_{(1)}(\mathbf{D} \otimes \mathbf{C})']' \Sigma^{-1} [(\mathbf{D} \otimes \mathbf{C})(\mathbf{D} \otimes \mathbf{C})']\})}{(2\pi)^{TFm/2} |\Omega \otimes \Psi|^{nT/2} |\Sigma|^{TFr/2}}.$$

We want to compute the likelihood-ratio test statistic

$$\begin{aligned} \Lambda &= \frac{L_{H_0}}{L_{H_a}} \\ &= \left(\frac{\hat{\Omega}_A \otimes \hat{\Psi}_A}{\hat{\Omega}_0 \otimes \hat{\Psi}_0} \right)^{\frac{nT}{2}} \left(\frac{\hat{\Sigma}_A}{\hat{\Sigma}_0} \right)^{\frac{TFr}{2}}. \end{aligned}$$

ONE-SAMPLE PROBLEM: MAXIMUM-LIKELIHOOD ESTIMATION



$$\begin{aligned}\hat{\mathbf{B}}_{(1)} &= \frac{1}{N} (\mathbf{A}' \Sigma^{-1} \mathbf{A})^{-1} \sum_{i=1}^n \mathbf{A}' \Sigma^{-1} \mathbf{X}_{(1),i} (\Omega \otimes \Psi)^{-1} (\mathbf{D} \otimes \mathbf{C}) [(\mathbf{D} \otimes \mathbf{C})' (\Omega \otimes \Psi) (\mathbf{D} \otimes \mathbf{C})]^{-1} \\ &= (\mathbf{A}' \Sigma^{-1} \mathbf{A})^{-1} \mathbf{A}' \Sigma^{-1} \bar{\mathbf{X}}_{(1)} (\Omega \otimes \Psi)^{-1} (\mathbf{D} \otimes \mathbf{C}) [(\mathbf{D} \otimes \mathbf{C})' (\Omega \otimes \Psi) (\mathbf{D} \otimes \mathbf{C})]^{-1} \\ \hat{\Sigma} &= \frac{\sum_{i=1}^n (\mathbf{X}_{(1),i} - \mathbf{A}\mathbf{B}_{(1)}(\mathbf{D} \otimes \mathbf{C})')' (\hat{\Omega} \otimes \hat{\Psi})^{-1} (\mathbf{X}_{(1),i} - \mathbf{A}\mathbf{B}_{(1)}(\mathbf{D} \otimes \mathbf{C})')}{nFr} \\ \hat{\Psi} &= \frac{\sum_{i=1}^n (\mathbf{X}_{(2),i} - \mathbf{C}\mathbf{B}_{(2)}(\mathbf{D} \otimes \mathbf{A})')' (\hat{\Omega} \otimes \hat{\Sigma})^{-1} (\mathbf{X}_{(2),i} - \mathbf{C}\mathbf{B}_{(2)}(\mathbf{D} \otimes \mathbf{A})')}{nTr} \\ \hat{\Omega} &= \frac{\sum_{i=1}^n (\mathbf{X}_{(3),i} - \mathbf{D}\mathbf{B}_{(3)}(\mathbf{C} \otimes \mathbf{A})')' (\hat{\Psi} \otimes \hat{\Sigma})^{-1} (\mathbf{X}_{(3),i} - \mathbf{D}\mathbf{B}_{(3)}(\mathbf{C} \otimes \mathbf{A})')}{nTF}\end{aligned}$$

An iterative algorithm is used to estimate $\hat{\mathbf{B}}_{(1)}$, $\hat{\Sigma}$, $\hat{\Psi}$, and $\hat{\Omega}$.



ASYMPTOTIC DISTRIBUTION OF $-2 \log \Lambda$

Because we are testing

$$H_0 : \mathcal{B} = \mathcal{B}_0$$

$$H_a : \mathcal{B} \neq \mathcal{B}_0$$

we have a simple null hypothesis. Therefore, by Wilks' Theorem, as $n \rightarrow \infty$,

$$-2 \log \Lambda \sim \chi^2_{tfr}.$$

ONE-SAMPLE PROBLEM: FORMULATION AS REGRESSION PROBLEM



Recall

$$\hat{\mathbf{B}}_{(1)} = (\mathbf{A}'\Sigma^{-1}\mathbf{A})^{-1} \sum_{i=1}^n \mathbf{A}'\sigma^{-1}\bar{\mathbf{X}}_{(1)}(\Omega \otimes \Psi)^{-1}(\mathbf{D} \otimes \mathbf{C})[(\mathbf{D} \otimes \mathbf{C})'(\Omega \otimes \Psi)(\mathbf{D} \otimes \mathbf{C})]^{-1}$$

Then we can formulate the following regression problem:

$$\underbrace{\text{vec}(\bar{\mathbf{X}}_{(1)})}_Y = \underbrace{(\mathbf{D} \otimes \mathbf{C} \otimes \mathbf{A})}_{X} \underbrace{\text{vec}(\mathbf{B}_{(1)})}_{\beta} + \underbrace{\text{vec}(\mathbf{E})}_{\epsilon}$$

If the errors are not homoscedastic ($\text{vec}(\mathbf{E}) \sim N(0, \frac{1}{n}\Omega \otimes \Psi \otimes \Sigma)$, where none of Σ , Ψ , and Ω are equal to $\sigma^2 I$), then we take the Cholesky decomposition to get a matrix C such that

$$C'C = \left(\frac{1}{n}\Omega \otimes \Psi \otimes \Sigma\right)^{-1} = n\Omega^{-1} \otimes \Psi^{-1} \otimes \Sigma^{-1}$$

and process like we would for generalized least squares.

$$C\text{vec}(\bar{\mathbf{X}}_{(1)}) = C(\mathbf{D} \otimes \mathbf{C} \otimes \mathbf{A})\text{vec}(\mathbf{B}_{(1)}) + C\text{vec}(\mathbf{E})$$

Test statistic: $F \sim F_{tfr, TFr - tfr}$

ONE-SAMPLE PROBLEM: SCORE TEST

Under H_0 , $\mathcal{X}_i = \mathcal{B}_0 \times \{\mathbf{A}, \mathbf{C}, \mathbf{D}\} + \mathcal{E}$

$$L(\mathbf{B}_{(1)} | \mathbf{A}, \mathbf{C}, \mathbf{D}, x_{(1),1}, \dots, x_{(1),n})$$

$$= \frac{\exp(-\frac{1}{2} \sum_{i=1}^n \text{tr}\{(\Omega \otimes \Psi)^{-1} [\mathbf{X}_{(1),i} - \mathbf{A}\mathbf{B}_{(1)}(\mathbf{D} \otimes \mathbf{C})']' \Sigma^{-1} [\mathbf{X}_{(1),i} - \mathbf{A}\mathbf{B}_{(1)}(\mathbf{D} \otimes \mathbf{C})']\})}{(2\pi)^{TFrn/2} |\Omega \otimes \Psi|^{nT/2} |\Sigma|^{TFr/2}}$$

$$l(\mathbf{B}_{(1)} | \mathbf{A}, \mathbf{C}, \mathbf{D}, x_{(1),1}, \dots, x_{(1),n}) = -\frac{1}{2} \sum_{i=1}^n [\{\text{vec}(\mathbf{X}_{(1),i})' - [(\mathbf{D} \otimes \mathbf{C} \otimes \mathbf{A})\text{vec}(\mathbf{B}_{(1)})]'\} \times \\ (\Omega^{-1} \otimes \Psi^{-1} \otimes \Sigma^{-1}) \times \{\text{vec}(\mathbf{X}_{(1),i}') - (\mathbf{A} \otimes \mathbf{D} \otimes \mathbf{C})\text{vec}(\mathbf{B}_{(1)}')\}]$$

$$- \frac{nTFr}{2} \log(2\pi) - \frac{nT}{2} \log |\Omega \otimes \Psi| - \frac{nFr}{2} \log |\Sigma|$$

$$U(\mathbf{B}_{(1)}) = \frac{\partial l}{\partial \mathbf{B}_{(1)}} = (\mathbf{D}' \Omega^{-1} \otimes \mathbf{C}' \Psi^{-1} \otimes \mathbf{A}' \Sigma^{-1}) \sum_{i=1}^n [\text{vec}(\mathbf{X}_{(1),i}) - (\mathbf{D} \otimes \mathbf{C} \otimes \mathbf{A})\text{vec}(\mathbf{B}_{(1)})]$$

$$\frac{\partial^2 l}{\partial \mathbf{B}_{(1)}^2} = -n(\mathbf{D}' \Omega^{-1} \mathbf{D} \otimes \mathbf{C}' \Psi^{-1} \mathbf{C} \otimes \mathbf{A}' \Sigma^{-1} \mathbf{A})$$

$$I(\mathbf{B}_{(1)}) = -E[-n(\mathbf{D}' \Omega^{-1} \mathbf{D} \otimes \mathbf{C}' \Psi^{-1} \mathbf{C} \otimes \mathbf{A}' \Sigma^{-1} \mathbf{A})] = n(\mathbf{D}' \Omega^{-1} \mathbf{D} \otimes \mathbf{C}' \Psi^{-1} \mathbf{C} \otimes \mathbf{A}' \Sigma^{-1} \mathbf{A})$$

ONE-SAMPLE PROBLEM: SCORE TEST STATISTIC



$$\begin{aligned} & U(\mathbf{B}_{(1),0})' I(\mathbf{B}_{(1),0})^{-1} U(\mathbf{B}_{(1),0}) \\ &= \left\{ \sum_{i=1}^n [\text{vec}(\mathbf{X}_{(1),i}) - (\mathbf{D} \otimes \mathbf{C} \otimes \mathbf{A}) \text{vec}(\mathbf{B}_{(1),0})] \right\}' \times \\ & \frac{1}{n} (\mathbf{D}' \Omega^{-1} (\mathbf{D}' \Omega^{-1} \mathbf{D})^{-1} \mathbf{D}' \Omega^{-1} \otimes \mathbf{C}' \Psi^{-1} (\mathbf{C}' \Psi^{-1} \mathbf{C})^{-1} \mathbf{C}' \Psi^{-1} \otimes \mathbf{A}' \Sigma^{-1} (\mathbf{A}' \Sigma^{-1} \mathbf{A})^{-1} \mathbf{A}' \Sigma^{-1} \\ & \left\{ \sum_{i=1}^n [\text{vec}(\mathbf{X}_{(1),i}) - (\mathbf{D} \otimes \mathbf{C} \otimes \mathbf{A}) \text{vec}(\mathbf{B}_{(1),0})] \right\} \\ & \sim \chi^2_{tfr} \end{aligned}$$

ONE-SAMPLE TESTS: SIMULATIONS

Model:

$$\mathcal{X} = \mathcal{B} \times \{\mathbf{A}, \mathbf{C}, \mathbf{D}\} + \mathcal{E}$$

Simulated Data (under $H_0 : \mathcal{B} = B_0, \dots, B_0$):

$$\mathcal{X} \sim N_{T,F,r}(\mathcal{B} \times \{\mathbf{A}, \mathbf{C}, \mathbf{D}\}, \Sigma, \Psi, \Omega)$$

where

\mathcal{X} is a $10 \times 10 \times 3$ tensor ($T, F = 10, r = 3$),

\mathbf{A} is a 10×4 arbitrary, orthogonal matrix ($t = 4$),

B_0 is a $4 \times 2 \times 3$ tensor, with each slice consisting of the 4×2 matrix B_0

B_0 is a 4×2 matrix consisting of independent $N(0, 10^2)$ observations,

\mathbf{C} is a 2×10 arbitrary, orthogonal matrix ($f = 2$),

\mathbf{D} is a 3×3 arbitrary, orthogonal matrix,

\mathcal{E} is a $10 \times 10 \times 3$ tensor with $N_{T,F,r}(0, \Sigma, \Psi, \Omega)$ distribution,

Σ is a 10×10 arbitrary symmetric, positive-definite covariance matrix,

Ψ is a 10×10 arbitrary symmetric, positive-definite covariance matrix,

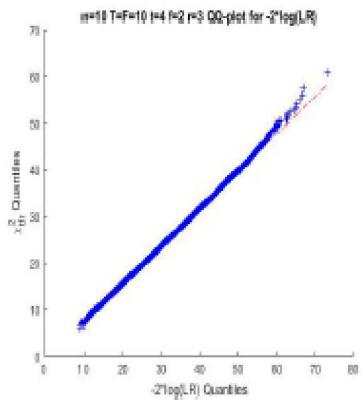
Ω is a 3×3 arbitrary symmetric, positive-definite covariance matrix.

We run 10,000 simulations under $H_0 : \mathcal{B} = B_0, \dots, B_0$ using MATLAB.

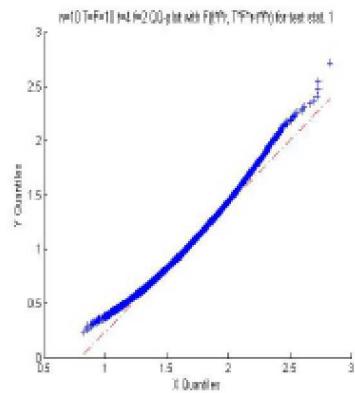
ONE-SAMPLE TESTS: SIMULATION RESULTS



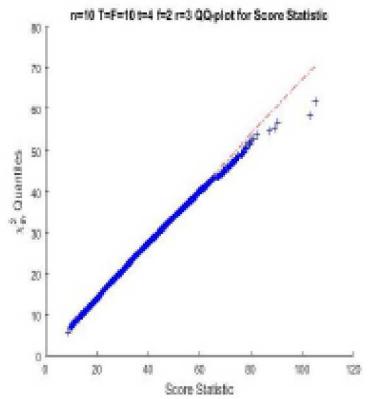
LRT ($-2 \log \Lambda \sim \chi^2_{tfr}$):



Regression Test
($F \sim F_{tfr}, TFr - tfr$):



Score Test
($U(B_0)' I(B_0)^{-1} U(B_0) \sim \chi^2_{tfr}$):



k -SAMPLE PROBLEM



Assumptions:

- Population g : $g = 1, \dots, k$ (We have k independent populations)

$$\mathfrak{X}_i^g = \mathfrak{B}_g \times \mathbf{A}, \mathbf{C}, \mathbf{D} + \mathfrak{E}_i, i = 1, \dots, N_g \left(\sum_{i=1}^g N_g = N \right)$$

$$\mathfrak{X}_i^g \sim N_{p,q,r}(\mathfrak{B}_g \times \mathbf{A}, \mathbf{C}, \mathbf{D}, \Sigma, \Psi, \Omega)$$

$$\begin{aligned} \Rightarrow \text{vec}(\mathfrak{B}_i^g \times \mathbf{A}, \mathbf{C}, \mathbf{D}) &\sim N_{prq}(\mathbf{D} \otimes \mathbf{C} \otimes \mathbf{A}) \text{vec}(\mathfrak{B}_g), \Omega \otimes \Psi \otimes \Sigma \\ &= (\mathbf{D} \otimes \mathbf{C} \otimes \mathbf{A}) \text{vec}(\mathfrak{B}_g) \end{aligned}$$

- Common \mathbf{A} , \mathbf{C} , and \mathbf{D} for all populations
- Row covariance matrix Σ is $T \times T$, fixed, p.d., but unknown
- Column covariance matrix Ψ is $F \times F$, fixed, p.d., but unknown
- Slice covariance matrix Ω is $r \times r$, fixed, p.d., but unknown

We are testing

$$H_0 : \mathcal{B}_1 = \mathcal{B}_2 = \dots = \mathcal{B}_k$$

$$H_a : \text{At least one of } \mathcal{B}_1, \dots, \mathcal{B}_k \text{ is different.}$$

k -SAMPLE PROBLEM: MAXIMUM-LIKELIHOOD ESTIMATION



Let n_1, \dots, n_k denote the total *cumulative* sample size up to and including sample k .
 $n = \sum_{i=1}^k n_i$, and $\mathbf{B}_{(1),g}$, $g = 1, \dots, k$ denote the $\mathbf{B}_{(1)}$ value corresponding to group g .

$$\begin{aligned}\hat{\mathbf{B}}_{(1),1} &= \frac{1}{n_1} (\mathbf{A}' \Sigma^{-1} \mathbf{A})^{-1} \sum_{i=1}^{n_1} \mathbf{A}' \Sigma^{-1} \mathbf{X}_{(1),i} (\Omega \otimes \Psi)^{-1} (\mathbf{D} \otimes \mathbf{C}) [(\mathbf{D} \otimes \mathbf{C})' (\Omega \otimes \Psi) (\mathbf{D} \otimes \mathbf{C})]^{-1} \\ &= (\mathbf{A}' \Sigma^{-1} \mathbf{A})^{-1} \mathbf{A}' \Sigma^{-1} \bar{\mathbf{X}}_{(1),1} (\Omega \otimes \Psi)^{-1} (\mathbf{D} \otimes \mathbf{C}) [(\mathbf{D} \otimes \mathbf{C})' (\Omega \otimes \Psi) (\mathbf{D} \otimes \mathbf{C})]^{-1} \\ \hat{\mathbf{B}}_{(1),g} &= \frac{1}{n_g - n_{g-1}} (\mathbf{A}' \Sigma^{-1} \mathbf{A})^{-1} \sum_{i=n_{g-1}+1}^{n_g} \mathbf{A}' \Sigma^{-1} \mathbf{X}_{(1),i} (\Omega \otimes \Psi)^{-1} (\mathbf{D} \otimes \mathbf{C}) \times \\ &\quad [(\mathbf{D} \otimes \mathbf{C})' (\Omega \otimes \Psi) (\mathbf{D} \otimes \mathbf{C})]^{-1} \\ &= (\mathbf{A}' \Sigma^{-1} \mathbf{A})^{-1} \mathbf{A}' \Sigma^{-1} \bar{\mathbf{X}}_{(1),g} (\Omega \otimes \Psi)^{-1} (\mathbf{D} \otimes \mathbf{C}) [(\mathbf{D} \otimes \mathbf{C})' (\Omega \otimes \Psi) (\mathbf{D} \otimes \mathbf{C})]^{-1}, \\ g &= 2, \dots, k\end{aligned}$$



- Likelihood-ratio test (asymptotic distribution):

$$-2 \log \Lambda \sim \chi^2_{(k-1)tfr}.$$

- Regression problem framework: Test statistic $F \sim F_{ktfr, kTFr - ktfr}$
- Score test: Cannot conclude theoretically that $U(\hat{\mathbf{B}}_{(1)})' I_{\hat{\mathbf{B}}_{(1)}}^{-1} U(\hat{\mathbf{B}}_{(1)}) \sim \chi^2_{tfr}$ exactly or $U(\hat{\mathbf{B}}_{(1)})' I_{\hat{\mathbf{B}}_{(1)}}^{-1} U(\hat{\mathbf{B}}_{(1)}) \sim \chi^2_{(k-1)tfr}$ asymptotically as $n_k \rightarrow \infty$

TRAIN VIDEO DATASET

- Publicly available YouTube video at the following link:
<https://www.youtube.com/watch?v=tNT2iQZ1Wil>.
- 32-minute video consists of Amtrak, BNSF, and Metrolink trains in Santa Fe Springs, CA taken on 12/13/14.
- We take three segments of the video consisting of a BNSF train, single-stack train, and double-stack train.
- Each segment is 10 frames, and each image, X_i , $i = 1, \dots, n$, is 71×101 in size.
- Following the work of Lock et al (2011), we scale our data so that all 30 observations have the same total variability. Letting \bar{x}_i be the mean and s_i be the standard deviation of the entries of X_i , define

$$X_i^{\text{scaled}} = \frac{X_i - \bar{x}_i}{s_i}.$$

We scale all of our 30 images based on the above definition.

APPLICATION TO TRAIN VIDEO DATA: DIMENSION REDUCTION AND INFERENCE PROCEDURES



- For each video segment, we have the tensor $\mathcal{X}_i, i = 1, \dots, 40$ of size $71 \times 101 \times 10$ which incorporates all 10 frames from each video. Thus, we have three tensors.
- Using obtained values of $t = 25$ and $f = 30$ using previously established methods, we compute the Tucker decomposition on the first tensor (segment 1's images) to calculate **A**, **C**, and **D**.
- With these fixed and estimated, apply inferential procedures on dataset of 3 tensors.
 - Three types of problems: one-, two-, and three-sample problems



We wish to determine if all three video segments have the same mean, i.e. have the same mean of $\mathcal{B} \times \{\mathbf{A}, \mathbf{C}, \mathbf{D}\}$. With \mathbf{A} , \mathbf{C} , and \mathbf{D} being estimated and fixed, we want to see if they all have the same value of \mathcal{B} . To make this determination, we test the hypotheses

$$H_0 : \mathcal{B} = \mathcal{B}_0$$

$$H_0 : \mathcal{B} \neq \mathcal{B}_0,$$

where we will set \mathcal{B}_0 to be the set of images for video segment 1 (BNSF train).

APPLICATION TO TRAIN VIDEO DATA: ONE-SAMPLE TESTS

Test	Dist. of Test Statistic	Critical Value ($\alpha = 0.05$)	Test Statistic	Decision
LRT (asympt. dist.)	$-2 \log \Lambda \sim \chi^2_{tfr}$	5.7513×10^4	7.7026×10^3	Reject H_0
Score	$U(\mathbf{B}_0)' I(\mathbf{B}_0)^{-1} U(\mathbf{B}_0) \sim \chi^2_{tfr}$	1.4488×10^5	7.7026×10^3	Reject H_0
Regression	$F \sim F_{tfr, TF_r - tfr}$	1.0286	7.4041	Reject H_0

- $\mathbf{B}_0 = \mathbf{B}_{0,(1)}$
- $T = 71, F = 101, r = 10$
- $t = 25, f = 30, r = 10$
- $n = 3$

APPLICATION TO TRAIN VIDEO DATA: TWO-SAMPLE TESTS



We seek to determine if there is a significant difference in the means of the images for the video segment of the BNSF train, and the video segment of the single-stack train.

- Population 1: images of BNSF train video segment ($n_1 = 1$)
- Population 2: images of single-stack video segment ($n_2 = 2$)

With **A**, **C**, and **D** being estimated and fixed, if the mean for population 1 is $\mathcal{B}_1 \times \{\mathbf{A}, \mathbf{C}, \mathbf{D}\}$ and the mean for population 2 is $\mathcal{B}_2 \times \{\mathbf{A}, \mathbf{C}, \mathbf{D}\}$, then we want to see if $\mathcal{B}_1 = \mathcal{B}_2$. Therefore, we test the hypotheses

$$H_0 : \mathcal{B}_1 = \mathcal{B}_2 = \mathcal{B}$$

$$H_0 : \mathcal{B}_1 \neq \mathcal{B}_2.$$

APPLICATION TO TRAIN VIDEO DATA: TWO-SAMPLE TESTS



Test	Dist. of Test Statistic	Critical Value ($\alpha = 0.05$)	Test Statistic	Decision
LRT (asympt. dist.)	$-2 \log \Lambda \sim \chi^2_{tfr}$	1.1928 $\times 10^5$	7.7026 $\times 10^3$	Reject H_0
Regression	$F \sim F_{2tf, 2TF-2tf}$	8.4154	1.0202	Reject H_0

- $B_0 = B_{0,(1)}$
- $T = 71, F = 101, r = 10$
- $t = 25, f = 30, r = 10$
- $n_2 = 2, k = 2$



We seek to determine if there is a significant difference in the means of the images for:

- Population 1: BNSF train video segment ($n_1 = 1$)
- Population 2: single-stack video segment ($n_2 = 2$)
- Population 3: double-stack video segment ($n_3 = 3$)

With **A**, **C**, and **D** being estimated and fixed, if the mean for population 1 is $\mathcal{B}_1 \times \{\mathbf{A}, \mathbf{C}, \mathbf{D}\}$, the mean for population 2 is $\mathcal{B}_2 \times \{\mathbf{A}, \mathbf{C}, \mathbf{D}\}$, and the mean for population 3 is $\mathcal{B}_3 \times \{\mathbf{A}, \mathbf{C}, \mathbf{D}\}$, then we want to see if $\mathcal{B}_1 = \mathcal{B}_2 = \mathcal{B}_3$. Therefore, we test the hypotheses

$$H_0 : \mathcal{B}_1 = \mathcal{B}_2 = \mathcal{B}_3 = \mathcal{B}$$

$$H_0 : \text{At least one of } \mathcal{B}_1, \mathcal{B}_2, \mathcal{B}_3 \text{ is not equal.}$$

APPLICATION TO TRAIN VIDEO DATA: THREE-SAMPLE TESTS



Test	Dist. of Test Statistic	Critical Value ($\alpha = 0.05$)	Test Statistic	Decision
LRT (asympt. dist.)	$-2 \log \Lambda \sim \chi^2_{(k-1)tfr}$	6.5350×10^4	1.5286×10^4	Reject H_0
Regression	$F \sim F_{3tfr, 3TFr-3tfr}$	1.0165	6.5293	Reject H_0

- $\mathbf{B}_0 = \mathbf{B}_{0,(1)}$
- $T = 71, F = 101, r = 10$
- $t = 25, f = 30, r = 10$
- $n_3 = 3, k = 3$

DISCUSSION OF RESULTS

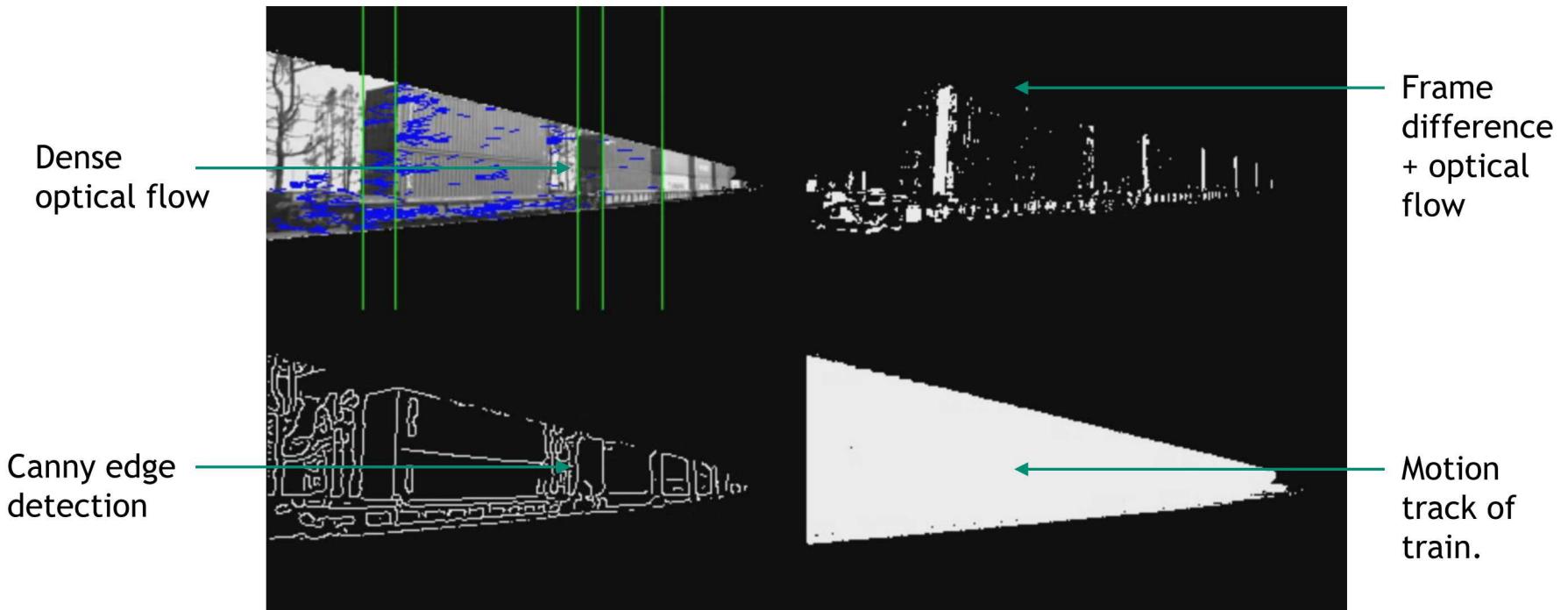


- Regression-based inference test, which has the most solid mathematical support, serves as the reference test. This test rejects H_0 for one-, two-, and three-sample problems, which is the expected result.
- Asymptotic distribution for LRT and score test (one-sample) reject H_0 .

FUTURE WORK

Integration of methods with feature detection methods in computer vision

- Deep learning
- Unsupervised machine learning





- Develop hypothesis testing procedures without unfolding tensors
- Computational issues, especially with Kronecker products of covariance matrices
- Inferential procedures for other tensor distributions, including nonparametric methods
- Goodness-of-fit tests for tensor distributions
- Using other tensor decomposition methods, such as CANDLECOMP and PARAFRAC
- Hierarchical hypothesis testing

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