

Polynomial Chaos

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Uncertainty Quantification

Topic 1: UQ Methods / PDF assignment

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Acknowledgements

Cosmin Safta Khachik Sargsyan Kenny Chowdhary Habib Najm Prashant Rai	Sandia National Laboratories Livermore, CA
Omar Knio Justin Winokur	Duke University, Raleigh, NC
Roger Ghanem	University of Southern California Los Angeles, CA
Olivier Le Maître	LIMSI-CNRS, Orsay, France

As well as Katy Jarvis, Sarah de Bord, Sarah Castorena, Majid Latif, Katherine Johnston, and many others ...

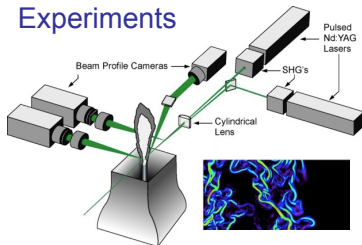
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Aspects of Uncertainty Quantification

Experimental data is used to calibrate models

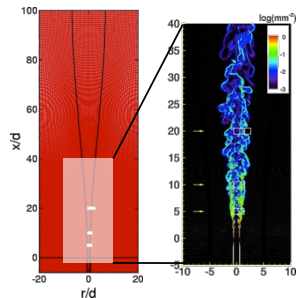
Experiments



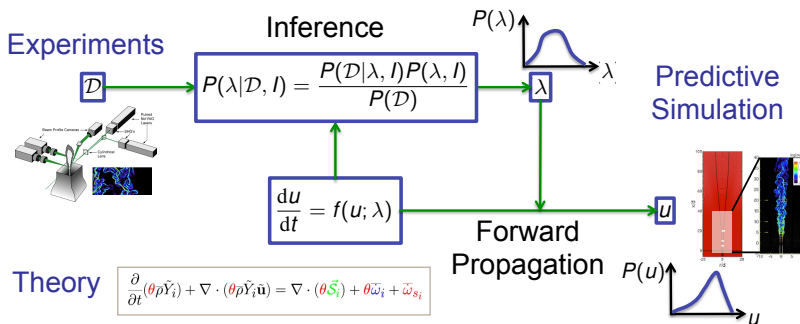
$$\frac{\partial}{\partial t}(\theta \tilde{p} \tilde{Y}_i) + \nabla \cdot (\theta \tilde{p} \tilde{Y}_i \tilde{\mathbf{u}}) = \nabla \cdot (\theta \tilde{\mathbf{S}}_i) + \theta \tilde{\omega}_i + \tilde{\omega}_{s_i}$$

Theory

Predictive Simulation



UQ methods extract information from all sources to enable predictive simulation



- UQ not just about propagating uncertainties
- The term UQ covers a wide range of methods

UQ assesses confidence in model predictions and allows resource allocation for fidelity improvements

- Parameter inference
 - Determine parameters from data
 - Characterize uncertainties in inferred parameters
- Propagate input uncertainties through computational models
 - Account for uncertainty from all sources
 - Resolve coupling between sources
- Analysis
 - Sensitivity analysis
 - Attributions
- Model calibration, validation, selection, averaging

Types of Uncertainty

- Epistemic uncertainty
 - Variable has one particular value, but it is not known
 - Reducible: by taking more measurements, we can get to know the value of the variable better
 - Examples
 - The mass of the planet Neptune
 - Ocean temperature at a particular point and time
- Aleatory uncertainty
 - Intrinsic or inherent uncertainty: variable is random; different value each time it is observed
 - Irreducible: taking more measurements will not reduce uncertainty in the value of the variable
 - Examples:
 - Collision interactions in molecular systems
 - Sampling noise

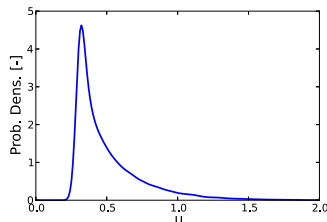
Epistemic or aleatory can depend on point of view

- Epistemic versus aleatory sometimes depends on scale level
 - In fully resolved turbulent simulations, velocity field near a wall is deterministic
 - In mesoscale channel flow, near wall velocity field is turbulent forcing term (aleatory)
 - In macroscale channel flow, near wall velocity field provides friction term (epistemic)
- These lectures follow the Bayesian view: probability represents the degree of belief in the value of a variable

Polynomial Chaos Expansions represent random variables

$$u = \sum_{k=0}^P u_k \psi_k(\xi)$$

- u : Random Variable (RV) represented with 1D PCE
- u_k : PC coefficients (deterministic)
- ψ_k : 1D Hermite polynomial of order k
- ξ : Gaussian RV

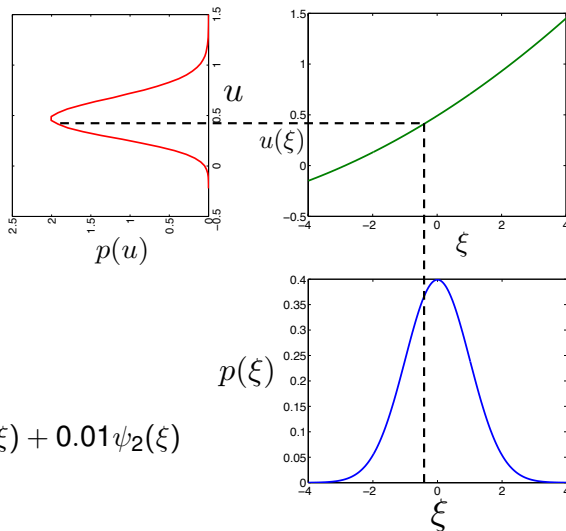


$$u = 0.5 + 0.2\psi_1(\xi) + 0.1\psi_2(\xi)$$

Expansion in terms of functions of random variables multiplied with deterministic coefficients

- Set of deterministic PC coefficients fully describes RV
- Separates randomness from deterministic dimensions

PCEs are a functional map from standard RVs to the represented RV



$$u = 0.5 + 0.2\psi_1(\xi) + 0.01\psi_2(\xi)$$

One-Dimensional Hermite Polynomials

$$\psi_0(\xi) = 1$$

$$\psi_k(\xi) = (-1)^k e^{\xi^2/2} \frac{d^k}{d\xi^k} e^{-\xi^2/2}, \quad k = 1, 2, \dots$$

$$\psi_1(\xi) = \xi, \quad \psi_2(\xi) = \xi^2 - 1, \quad \psi_3(\xi) = \xi^3 - 3\xi, \quad \dots$$

The Hermite polynomials form an orthogonal basis over $[-\infty, \infty]$ with respect to the inner product

$$\langle \psi_i \psi_j \rangle \equiv \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} \psi_i(\xi) \psi_j(\xi) w(\xi) d\xi = \delta_{ij} \langle \psi_i^2 \rangle$$

where $w(\xi) = e^{-\xi^2/2}$ is the weight function.

Note that $\frac{e^{-\xi^2/2}}{\sqrt{2\pi}}$ is the density of a standard normal random variable

Multidimensional Hermite Polynomials

The multidimensional Hermite polynomial $\Psi_i(\xi_1, \dots, \xi_n)$ is a tensor product of the 1D Hermite polynomials, with a suitable multi-index $\alpha^i = (\alpha_1^i, \alpha_2^i, \dots, \alpha_n^i)$,

$$\Psi_i(\xi_1, \dots, \xi_n) = \prod_{k=1}^n \psi_{\alpha_k^i}(\xi_k)$$

For example, 2D Hermite polynomials:

i	p	Ψ_i	α^i
0	0	1	(0,0)
1	1	ξ_1	(1,0)
2	1	ξ_2	(0,1)
3	2	$\xi_1^2 - 1$	(2,0)
4	2	$\xi_1 \xi_2$	(1,1)
5	2	$\xi_2^2 - 1$	(0,2)
...	...	...	...

Multidimensional Polynomial Chaos Expansion

$$u = \sum_{k=0}^P u_k \Psi_k(\xi_1, \dots, \xi_n)$$

- u : Random Variable (RV) represented with multi-D PCE
- u_k : PC coefficients (deterministic)
- Ψ_k : Multi-D Hermite polynomials up to order p
- ξ_i : Gaussian RV
- n : Dimensionality of stochastic space
- $P + 1$: Number of PC terms: $P + 1 = \frac{(n+p)!}{n!p!}$

The number of dimensions represents the number of independent inputs, degrees of freedom for u

- E.g. one stochastic dimension per uncertain model parameter
- Contributions from each uncertain input can be identified

Obtaining PC coefficients for arbitrary random variables is not trivial

- This is a hard problem in general
- Random Variable can be specified in a variety of ways, but often incomplete
 - Probability Density Function (PDF)
 - Samples
 - Expert opinion (*e.g.* “somewhere between 2 and 4”)
- Particular case of a random variable specified by a PDF is generally tractable

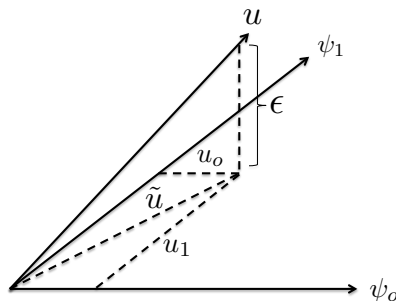
The PC basis functions are orthogonal with respect to the probability measure of the associated RVs

$$\begin{aligned}\langle \Psi_i \Psi_j \rangle &\equiv \int \dots \int \Psi_i(\xi) \Psi_j(\xi) g(\xi_1) g(\xi_2) \dots g(\xi_n) d\xi_1 d\xi_2 \dots d\xi_n \\ &= \prod_{k=1}^n \left\langle \psi_{\alpha_k^i}(\xi_k) \psi_{\alpha_k^j}(\xi_k) \right\rangle = \delta_{ij} \langle \Psi_i^2 \rangle\end{aligned}$$

where,

$$g(\xi) = \frac{e^{-\xi^2/2}}{\sqrt{2\pi}}$$

Orthogonality enables a Galerkin projection to determine the PC coefficients



$$\begin{aligned}
 u &\approx \sum_{k=0}^P u_k \psi_k & \Rightarrow & \langle \psi_i u \rangle = \sum_{k=0}^P u_k \langle \psi_i \psi_k \rangle = u_i \langle \psi_i^2 \rangle \\
 & & \Rightarrow & u_i = \frac{\langle u \psi_i \rangle}{\langle \psi_i^2 \rangle}
 \end{aligned}$$

Generalized Polynomial Chaos

PC Type	Domain	Density $w(\xi)$	Polynomial	Free parameters
Gauss-Hermite	$(-\infty, +\infty)$	$\frac{1}{\sqrt{2\pi}} e^{-\frac{\xi^2}{2}}$	Hermite	none
Legendre-Uniform	$[-1, 1]$	$\frac{1}{2}$	Legendre	none
Gamma-Laguerre	$[0, +\infty)$	$\frac{\xi^\alpha e^{-\xi}}{\Gamma(\alpha+1)}$	Laguerre	$\alpha > -1$
Beta-Jacobi	$[-1, 1]$	$\frac{(1+\xi)^\alpha (1-\xi)^\beta}{2^{\alpha+\beta+1} B(\alpha+1, \beta+1)}$	Jacobi	$\alpha > -1, \beta > -1$

Inner product: $\langle \psi_i \psi_j \rangle \equiv \int_a^b \psi_i(\xi) \psi_j(\xi) w(\xi) d\xi = \delta_{ij} \langle \psi_i^2 \rangle$

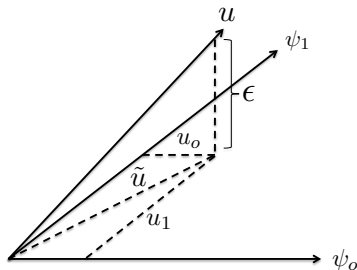
- Wiener-Askey scheme provides a hierarchy of possible continuous PC bases [Xiu and Karniadakis, 2002]
 - Legendre-Uniform is special case of Beta-Jacobi
- Input parameter domain often dictates the most convenient choice of PC
- Polynomials can also be tailored to be orthogonal w.r.t. chosen, arbitrary density

How do I know my PCE is converged?

- Approximation error in PCE is topic of a lot of research
- Rules of thumb:
 - Try to use a PC basis in terms of a random variable with a distribution similar to the uncertainties you expect to see
 - Higher order PC coefficients should decay
 - Increase order until results no longer change
 - Not always fail-proof ...

Propagation of Uncertain Inputs Represented with PCEs

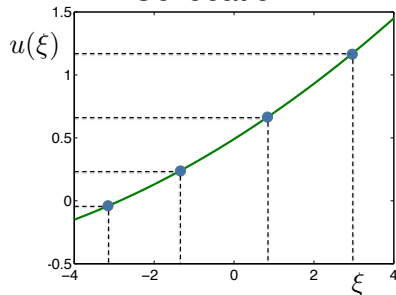
Galerkin Projection



$$u_k = \frac{\langle u \psi_k \rangle}{\langle \psi_k^2 \rangle}, \quad k = 0, \dots, P$$

Residual orthogonal to space covered by basis functions

Collocation



Match PCE to random variable at chosen sample points: interpolation or regression

Galerkin projection methods are either intrusive or non-intrusive

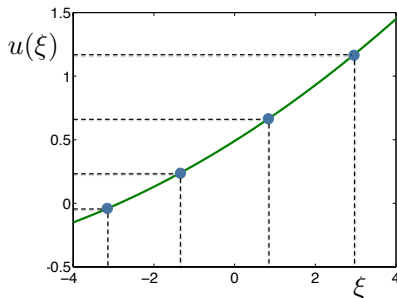
- Use same projection but in different ways

$$u_k = \frac{\langle u \psi_k \rangle}{\langle \psi_k^2 \rangle}, \quad k = 0, \dots, P$$

- Intrusive methods apply Galerkin projection to governing equations
 - Results in set of equations for the PC coefficients
 - Requires redesign of computer code
 - PCEs for all uncertain variables in system
- Non-intrusive approaches apply Galerkin projection to outputs of interest
 - Sampling to evaluate projection operator
 - Can use existing code as black box
 - Only computes PCEs for quantities of interest

Collocation approaches are non-intrusive and minimize errors at sample points

$$\sum_{k=0}^P u_k \psi_k(\xi_i) = u(\xi_i)$$
$$i = 1, \dots, N_c$$



- Use functional representation point of view
- Can use interpolation, e.g. Lagrange interpolants
- Or use regression approaches: $P + 1$ degrees of freedom to fit N_c points
- Can position points where most accuracy desired

Intrusive Galerkin projection reformulates original equations

- Assume $v = f(u; a, \lambda)$, with
 - a deterministic parameter(s)
 - λ uncertain parameter(s)
 - u, v variables of interest (deterministic or uncertain)
- Represent uncertain variables with PCEs

$$\lambda = \sum_{k=0}^P \lambda_k \psi_k(\xi), \quad v = \sum_{k=0}^P v_k \psi_k(\xi)$$

- Apply Galerkin projection to get PC coefficients of v

$$v_k = \frac{\langle v \psi_k \rangle}{\langle \psi_k^2 \rangle} = \frac{\langle f(u; a, \lambda) \psi_k \rangle}{\langle \psi_k^2 \rangle}, \quad k = 0, \dots, P$$

- Results in larger, but deterministic set of equations

Projection of product: $v = \gamma \lambda$

- Assume $v = \gamma \lambda$, with

$$\gamma = \sum_{k=0}^P \gamma_k \psi_k, \lambda = \sum_{k=0}^P \lambda_k \psi_k, \text{ and } v = \sum_{k=0}^P v_k \psi_k$$

- Galerkin projection

$$v_k = \frac{\langle v \psi_k \rangle}{\langle \psi_k^2 \rangle} = \frac{\langle (\gamma \lambda) \psi_k \rangle}{\langle \psi_k^2 \rangle}, \quad k = 0, \dots, P$$

$$\begin{aligned} \langle (\gamma \lambda) \psi_k \rangle &= \left\langle \left(\sum_{i=0}^P \gamma_i \psi_i \sum_{j=0}^P \lambda_j \psi_j \right) \psi_k \right\rangle \\ &= \left\langle \sum_{i=0}^P \sum_{j=0}^P \gamma_i \lambda_j \psi_i \psi_j \psi_k \right\rangle \\ &= \sum_{i=0}^P \sum_{j=0}^P \gamma_i \lambda_j \langle \psi_i \psi_j \psi_k \rangle \end{aligned}$$

Projection of product: $v = \gamma \lambda$

$$\Rightarrow v_k = \sum_{i=0}^P \sum_{j=0}^P \gamma_i \lambda_j \frac{\langle \psi_i \psi_j \psi_k \rangle}{\langle \psi_k^2 \rangle} = \sum_{i=0}^P \sum_{j=0}^P \gamma_i \lambda_j C_{ijk}$$

- $C_{ijk} = \frac{\langle \psi_i \psi_j \psi_k \rangle}{\langle \psi_k^2 \rangle}$
- The C_{ijk} tensor can be computed up front for any given PC order and dimension and stored for use whenever two RVs are multiplied
- This tensor is sparse, *i.e.* many of its elements are zero

Non-intrusive Galerkin Projection

$$u_k = \frac{\langle u \psi_k \rangle}{\langle \psi_k^2 \rangle} = \frac{1}{\langle \psi_k^2 \rangle} \int u \psi_k(\xi) w(\xi) d\xi, \quad k = 0, \dots, P$$

Evaluate projection integrals numerically

- Pick samples of uncertain parameters, e.g. $b(\xi) = \sum b_k \psi_k(\xi)$ by sampling ξ
- Run deterministic forward model for each of the sampled input parameter values $b^i = b(\xi^i)$
- Integration using random sampling or quadrature methods

Reconstruct uncertain model output

$$u(x, t; \theta) = \sum_{k=0}^P u_k(x, t) \psi_k(\xi(\theta))$$

Random Sampling Approaches for Galerkin Projection

- Evaluate integral through sampling

$$\int u \psi_k(\xi) w(\xi) d\xi = \frac{1}{N_s} \sum_{i=1}^{N_s} u(\xi^i) \psi_k(\xi^i)$$

- Samples are drawn according to the distribution of ξ
 - Monte-Carlo (MC)
 - Latin-Hypercube-Sampling (LHS)
- Pros:
 - Can be easily made fault tolerant
 - Sometimes random samples is all we have
- Cons: slow convergence, but less dependent on number of stochastic dimensions

Quadrature Approaches for Galerkin Projection

- Use numerical quadrature rules to evaluate integrals

$$\int u \psi_k(\xi) w(\xi) d\xi = \sum_{i=1}^{N_q} q^i u(\xi^i) \psi_k(\xi^i)$$

- The $N_q \xi^i$ are quadrature points, with corresponding weights q^i
- Choice of quadrature points important for accuracy
 - Also referred to as *deterministic sampling* approach
- Pros:
 - Can use existing codes as black box to evaluate $u(\xi^i)$
 - Embarrassingly parallel
- Cons: Tensor product rule for d dimensions requires N_q^d samples

Gauss quadrature rules are very efficient

$$\int u \Psi_k(\xi) w(\xi) d\xi = \sum_{i=1}^{N_q} q^i u(\xi^i) \Psi_k(\xi^i)$$

- In one dimension, N_q quadrature points can integrate polynomial of order $2N_q - 1$ exactly
- Gauss-Hermite and Gauss-Legendre quadrature tailored to specific choices of the weight function $w(\xi)$

Further Reading

An (incomplete) list of references about the material covered

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- D.S. Sivia, "Data Analysis: A Bayesian Tutorial," Oxford Science, 1996.

Polynomial Chaos Based Uncertainty Quantification

Topic 2: Sensitivity Analysis

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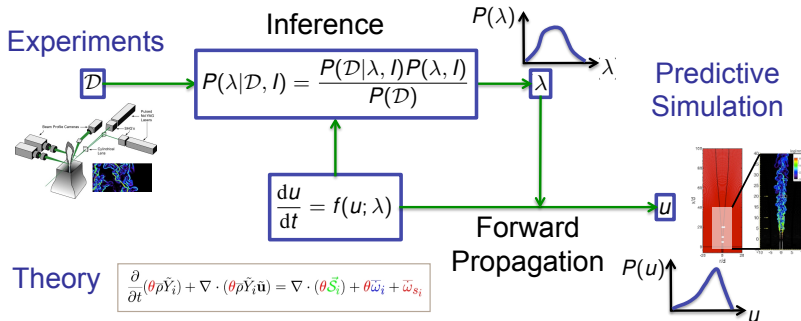
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Sensitivity analysis gives insight into key sources of uncertainty



- Obtaining global sensitivity analysis from PCEs
 - Identify dominant sources of uncertainty
 - Attribution

PC Postprocessing: global sensitivity information is readily obtained from PCE

$$g(\xi_1, \dots, \xi_d) = \sum_{k=0}^P c_k \psi_k(\xi)$$

- Global sensitivity analysis $\equiv$ Variance decomposition
- Total variance

$$\text{Var}[g(\xi)] = \sum_{k>0} c_k^2 \|\psi_k\|^2$$

PC Postprocessing: Main Effect and Joint Sensitivity Indices

- Main effect sensitivity indices

$$S_i = \frac{\text{Var}[\mathbb{E}(g(\xi|\xi_i))]}{\text{Var}[g(\xi)]} = \frac{\sum_{k \in \mathbb{I}_i} c_k^2 \|\psi_k\|^2}{\sum_{k > 0} c_k^2 \|\psi_k\|^2}$$

- $\mathbb{I}_i$ is the set of bases with only ξ_i involved
- S_i is the uncertainty contribution that is due to i -th parameter only
- Joint sensitivity indices

$$S_{ij} = \frac{\text{Var}[\mathbb{E}(g(\xi|\xi_i, \xi_j))]}{\text{Var}[g(\xi)]} - S_i - S_j = \frac{\sum_{k \in \mathbb{I}_{ij}} c_k^2 \|\psi_k\|^2}{\sum_{k > 0} c_k^2 \|\psi_k\|^2}$$

- $\mathbb{I}_{ij}$ is the set of bases with only ξ_i and ξ_j involved
- S_{ij} is the uncertainty contribution that is due to (i, j) parameter pair

PC Postprocessing: Total Effect Sensitivity Indices

- Total effect sensitivity indices

$$T_i = 1 - \frac{\text{Var}[\mathbb{E}(g(\xi|\xi_{-i}))]}{\text{Var}[g(\xi)]} = \frac{\sum_{k \in \mathbb{I}_i^T} c_k^2 ||\psi_k||^2}{\sum_{k > 0} c_k^2 ||\psi_k||^2}$$

- The notation ξ_{-i} indicates terms that do not have ξ_i in them
- $\mathbb{I}_i^T$ is the set of bases with ξ_i involved, including all its interactions
- The sum of all T_i is usually > 1

Sensitivity indices are directly computable from PC

$$g(\xi) = \sum_{k=0}^P c_k \Psi_k(\xi)$$

Consider dimensionality $d = 3$, total order $p = 2$,
number of PC terms $P + 1 = (d + p)! / (d! p!) = 10$.

$$g(\xi_1, \xi_2, \xi_3) = c_0 + c_1 \psi_1(\xi_1) + c_2 \psi_1(\xi_2) + c_3 \psi_1(\xi_3) + \\ + c_4 \psi_2(\xi_1) + c_5 \psi_1(\xi_1) \psi_1(\xi_2) + c_6 \psi_1(\xi_1) \psi_1(\xi_3) + c_7 \psi_2(\xi_2) + c_8 \psi_1(\xi_2) \psi_1(\xi_3) + c_9 \psi_2(\xi_3)$$

Variance contributions

$$\text{Var}(g) = 0 + c_1^2 \langle \psi_1^2 \rangle + c_2^2 \langle \psi_1^2 \rangle + c_3^2 \langle \psi_1^2 \rangle + \\ + c_4^2 \langle \psi_2^2 \rangle + c_5^2 \langle \psi_1^2 \rangle \langle \psi_1^2 \rangle + c_6^2 \langle \psi_1^2 \rangle \langle \psi_1^2 \rangle + c_7^2 \langle \psi_2^2 \rangle + c_8^2 \langle \psi_1^2 \rangle \langle \psi_1^2 \rangle + c_9^2 \langle \psi_2^2 \rangle$$

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Variance contributions

$$\text{Var}(g) = 0 + c_1^2 \langle \psi_1^2 \rangle + c_2^2 \langle \psi_1^2 \rangle + c_3^2 \langle \psi_1^2 \rangle + \\ + c_4^2 \langle \psi_2^2 \rangle + c_5^2 \langle \psi_1^2 \rangle \langle \psi_1^2 \rangle + c_6^2 \langle \psi_1^2 \rangle \langle \psi_1^2 \rangle + c_7^2 \langle \psi_2^2 \rangle + c_8^2 \langle \psi_1^2 \rangle \langle \psi_1^2 \rangle + c_9^2 \langle \psi_2^2 \rangle$$

Main effect sensitivities ξ_1 ξ_2 ξ_3

Sensitivity indices are directly computable from PC

$$g(\xi) = \sum_{k=0}^P c_k \Psi_k(\xi)$$

Consider dimensionality $d = 3$, total order $p = 2$,
number of PC terms $P + 1 = (d + p)!/(d!p!) = 10$.

$$g(\xi_1, \xi_2, \xi_3) = c_0 + c_1\psi_1(\xi_1) + c_2\psi_1(\xi_2) + c_3\psi_1(\xi_3) + \\ + c_4\psi_2(\xi_1) + c_5\psi_1(\xi_1)\psi_1(\xi_2) + c_6\psi_1(\xi_1)\psi_1(\xi_3) + c_7\psi_2(\xi_2) + c_8\psi_1(\xi_2)\psi_1(\xi_3) + c_9\psi_2(\xi_3)$$

Variance contributions

$$\text{Var}(g) = 0 + c_1^2 \langle \psi_1^2 \rangle + c_2^2 \langle \psi_1^2 \rangle + c_3^2 \langle \psi_1^2 \rangle + \\ + c_4^2 \langle \psi_2^2 \rangle + c_5^2 \langle \psi_1^2 \rangle \langle \psi_1^2 \rangle + c_6^2 \langle \psi_1^2 \rangle \langle \psi_1^2 \rangle + c_7^2 \langle \psi_2^2 \rangle + c_8^2 \langle \psi_1^2 \rangle \langle \psi_1^2 \rangle + c_9^2 \langle \psi_2^2 \rangle$$

Main effect sensitivities ξ_1 ξ_2 ξ_3

Sensitivity indices are directly computable from PC

$$g(\xi) = \sum_{k=0}^P c_k \Psi_k(\xi)$$

Consider dimensionality $d = 3$, total order $p = 2$,
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Total sensitivities $\xi_1 \quad \xi_2 \quad \xi_3$

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Joint sensitivities (ξ_1, ξ_2) (ξ_1, ξ_3) (ξ_2, ξ_3)

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Joint sensitivities (ξ_1, ξ_2) (ξ_1, ξ_3) (ξ_2, ξ_3)

PC Postprocessing: Sampling-Based Approaches

$$g(\xi_1, \dots, \xi_d) = \sum_{k=0}^P c_k \psi_k(\xi)$$

- In some cases, need to resort to Monte-Carlo estimation, e.g.
 - Piecewise-PC with irregular subdomains
 - Output transformations, e.g. build PC for $\log g(\xi)$, but inquire sensitivity with respect to $g(\xi)$
- A brute-force sampling of $\text{Var}[\mathbb{E}(g(\xi|\xi_i))]$ is extremely inefficient.

PC Postprocessing: Sampling-Based Approaches

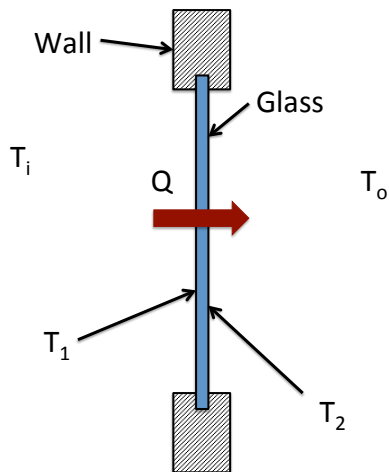
- Tricks are available, given a single set of sampled input [\[Saltelli, 2002\]](#). E.g., use

$$\mathbb{E}[g(\boldsymbol{\xi}|\xi_i)^2] = \mathbb{E}[g(\boldsymbol{\xi}|\xi_i)g(\boldsymbol{\xi}'|\xi_i)] = \frac{1}{N-1} \sum_{r=1}^N g(\boldsymbol{\xi}^{(r)})g(\tilde{\boldsymbol{\xi}}^{(r)}),$$

where $\tilde{\boldsymbol{\xi}}$ is $\boldsymbol{\xi}'$ with i -th element replaced by ξ_i .

- Similar formulae available for joint sensitivity indices.
- Con: as all Monte-Carlo algorithms, converges slowly.
- Pro: sampling is cheap.

Heat Transfer through a Window



$$h_i(T_i - T_1) = k_w \frac{(T_1 - T_2)}{d_w}$$

$$k_w \frac{(T_1 - T_2)}{d_w} = h_o(T_2 - T_o)$$

6 Uncertain, Gaussian parameters

$$T_i = 293\text{K}, \sigma = 0.5\%$$

$$T_o = 273\text{K}, \sigma = 0.5\%$$

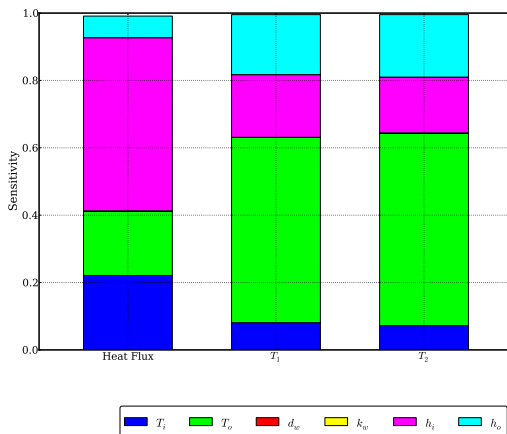
$$d_w = 0.01\text{m}, \sigma = 1\%$$

$$k_w = 1\text{W/mK}, \sigma = 5\%$$

$$h_i = 2\text{W/m}^2\text{K}, \sigma = 15\%$$

$$h_o = 6\text{W/m}^2\text{K}, \sigma = 15\%$$

Outputs are most sensitive to ambient temperatures and convective heat transfer coefficients



- Main effect sensitivities
 - Sum to 1 only if coupling terms do not matter
- k_w has minimal contribution due to its low uncertainty

Further Reading

An (incomplete) list of references about the material covered

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- M. Rosenblatt, "Remarks on a Multivariate Transformation," *Ann. Math. Statist.*, 23:3, pp. 470–472, 1952.

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Polynomial Chaos Based Uncertainty Quantification

Topic 3: Parameter Dependencies / Interactions

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Reservoirs UQ Workshop – Aug 1-3, 2017



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As well as Katy Jarvis, Sarah de Bord, Sarah Castorena, Majid Latif, Katherine Johnston, and many others ...

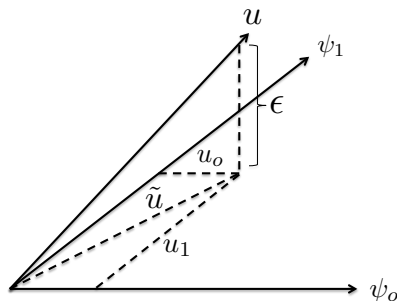
This material is based upon work supported by the U.S. Department of Energy, Office of Science, Office of Advanced Scientific Computing Research (ASCR), Scientific Discovery through Advanced Computing (SciDAC) and Applied Mathematics Research (AMR) programs.

Sandia National Laboratories is a multimission laboratory managed and operated by National Technology and Engineering Solutions of Sandia, LLC., a wholly owned subsidiary of Honeywell International, Inc., for the U.S. Department of Energy's National Nuclear Security Administration under contract DE-NA-0003525.

Obtaining PC coefficients for arbitrary random variables is not trivial

- Characterizing PCEs for uncertain inputs is a really difficult problem
- Inputs specified in a variety of ways, and often incomplete
 - Probability density function
 - Samples, *e.g.* from inverse problem solution
 - Expert opinion (*e.g.* “about 3.5”)
- Particular case of a random variable specified by a PDF is generally tractable

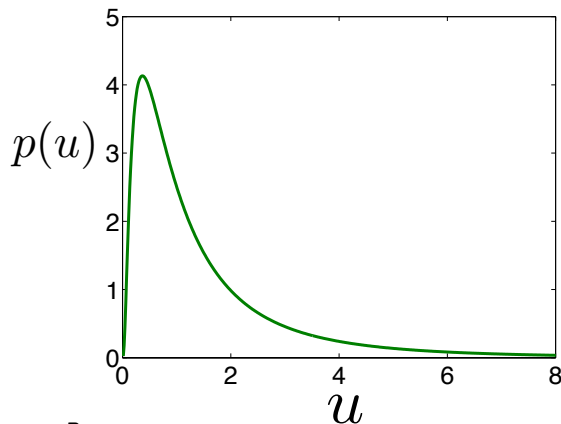
Orthogonality enables a Galerkin projection to determine the PC coefficients



$$u \approx \sum_{k=0}^P u_k \psi_k \quad \Rightarrow \quad \langle \psi_i u \rangle = \sum_{k=0}^P u_k \langle \psi_i \psi_k \rangle = u_i \langle \psi_i^2 \rangle$$

$$\Rightarrow \quad u_i = \frac{\langle u \psi_i \rangle}{\langle \psi_i^2 \rangle}$$

Galerkin projection requires functional relationship between random variable and germ of PCE



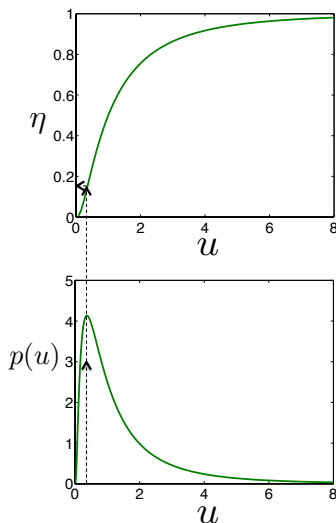
$$u = \sum_{k=0}^P u_k \Psi_k(\xi) \quad \Rightarrow \quad u_i = \frac{\langle u \Psi_i(\xi) \rangle}{\langle \Psi_i^2 \rangle}$$

Cumulative Distribution Function (CDF) maps arbitrary random variable to a uniform random variable

- Consider u with PDF $p(u)$
- CDF of u is given by

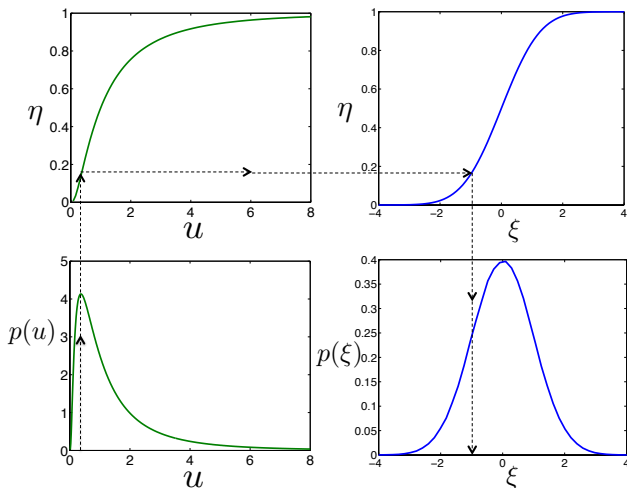
$$F(u) = \int_{-\infty}^u p(s) ds$$

- $F(u)$ maps u to η , uniform on $[0, 1]$



Inverse CDF mapping enables Galerkin Projection

- $\eta = F(u)$
- $\eta = \Phi(\xi)$
maps uniform η to normal RV ξ
- $u = F^{-1}(\Phi(\xi))$



$$\langle u \Psi_i(\xi) \rangle = \int \underbrace{F^{-1}(\Phi(\xi))}_u \Psi_i(\xi) w(\xi) d\xi$$

Constructing an n D PCE for a RV with a Given PDF

- Given RV $z \in \mathbb{R}$ with PDF: $g(z)$, define:

$$z = \sum_{i=0}^P z_i \Psi_i(\xi_1, \xi_2, \dots, \xi_n), \quad P + 1 = \frac{(n + p)!}{n!p!}$$

- No general procedure
- Construct PCE as model choice to represent what is known about RV
- Can choose $\{n, p\}$ and the mode strengths by ensuring accurate capture of
 - the PDF $g(z)$
 - select moments of z
 - some observable of interest $\phi(z)$

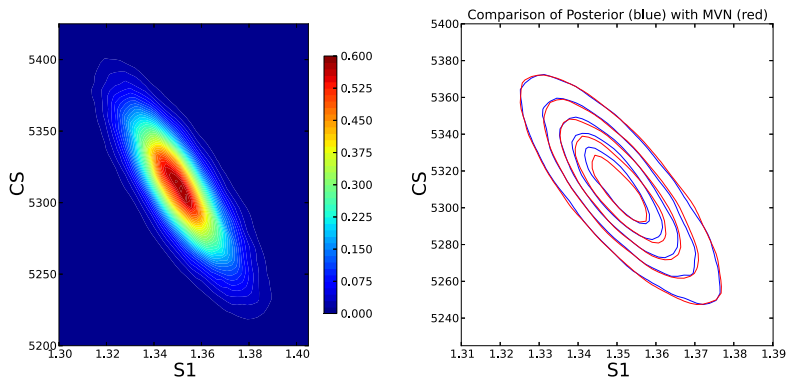
Multivariate Normal Approximation (MVN)

- Many distributions are unimodal and somewhat shaped like Gaussians
- MultiVariate Normal approximations capture mean and correlation structure of the random variables
- Easy to extract from a set of samples
 - In 1D: just compute mean and standard deviation:
 $u = u_0 + u_1 \xi$
 - Multi-D: Cholesky factorization of covariance

$$C = LL^T$$
$$u = L\xi$$

```
# Compute mean parameter values
par_mean = numpy.mean(samples,axis=0)
# Compute the covariance
par_cov = numpy.cov(samples,rowvar=0)
# Compute the Cholesky Decomposition
chol_lower = numpy.linalg.cholesky(par_cov)
```

MVN Approximation of Distribution from Samples



$$\begin{aligned} S_1 &= 1.351 + 0.01367\xi_1 \\ CS &= 5310 - 26.25\xi_1 + 20.26\xi_2 \end{aligned}$$

Rosenblatt Transformation for Multi-D RVs

- Assume samples of multi-D RVs are (e.g. from MCMC sampling of posterior parameter distribution)
- Rosenblatt transformation maps any (not necessarily independent) set of random variables $(\lambda_1, \dots, \lambda_d)$ to uniform i.i.d.'s $\{\eta_i\}_{i=1}^d$ (Rosenblatt, 1952).

$$\eta_1 = F_1(\lambda_1)$$

$$\eta_2 = F_{2|1}(\lambda_2|\lambda_1)$$

$$\vdots$$

$$\eta_d = F_{d|d-1,\dots,1}(\lambda_d|\lambda_{d-1},\dots,\lambda_1)$$

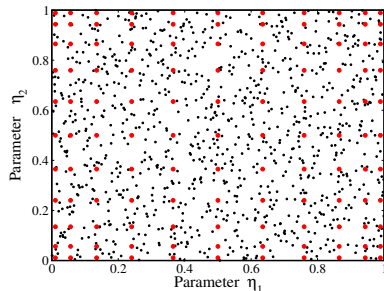
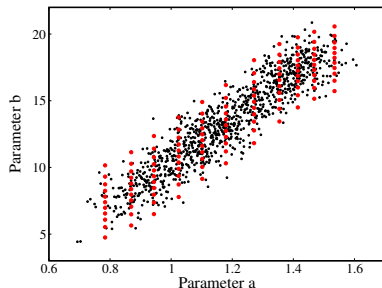
- Rosenblatt transformation is a multi-D generalization of 1D CDF mapping.
- Conditional CDFs are harder to evaluate in high dimensions

Rosenblatt transformation of a given sample set requires KDE

- Given samples $\{\lambda^{(n)}\}_{n=1}^N$ of the random variable $\lambda = (\lambda_1, \dots, \lambda_d)$
- Kernel Density Estimation (KDE) is needed to compute conditional CDFs

$$\begin{aligned}
 \eta_k &= F_{k|k-1, \dots, 1}(\lambda_k | \lambda_{k-1}, \dots, \lambda_1) = \\
 &= \int_{-\infty}^{\lambda_k} p_{k|k-1, \dots, 1}(\lambda'_k | \lambda_{k-1}, \dots, \lambda_1) d\lambda'_k = \int_{-\infty}^{\lambda_k} \frac{p_{k, \dots, 1}(\lambda'_k, \lambda_{k-1}, \dots, \lambda_1)}{p_{k-1, \dots, 1}(\lambda_{k-1}, \dots, \lambda_1)} d\lambda'_k \\
 &\approx \frac{1}{h\sqrt{2\pi}} \int_{-\infty}^{\lambda_k} \frac{\sum_{n=1}^N \exp\left(-\frac{(\lambda_1 - \lambda_1^{(n)})^2 + \dots + (\lambda'_k - \lambda_k^{(n)})^2}{2h^2}\right)}{\sum_{n=1}^N \exp\left(-\frac{(\lambda_1 - \lambda_1^{(n)})^2 + \dots + (\lambda_{k-1} - \lambda_{k-1}^{(n)})^2}{2h^2}\right)} d\lambda'_k \\
 &= \int_{-\infty}^{\lambda_k} \frac{\sum_{n=1}^N \exp\left(-\frac{(\lambda_1 - \lambda_1^{(n)})^2 + \dots + (\lambda_{k-1} - \lambda_{k-1}^{(n)})^2}{2h^2}\right) \times \frac{1}{h\sqrt{2\pi}} \exp\left(-\frac{(\lambda'_k - \lambda_k^{(n)})^2}{2h^2}\right)}{\sum_{n=1}^N \exp\left(-\frac{(\lambda_1 - \lambda_1^{(n)})^2 + \dots + (\lambda_{k-1} - \lambda_{k-1}^{(n)})^2}{2h^2}\right)} d\lambda'_k \\
 &= \frac{\sum_{n=1}^N \exp\left(-\frac{(\lambda_1 - \lambda_1^{(n)})^2 + \dots + (\lambda_{k-1} - \lambda_{k-1}^{(n)})^2}{2h^2}\right) \times \Phi\left(\frac{\lambda_k - \lambda_k^{(n)}}{h}\right)}{\sum_{n=1}^N \exp\left(-\frac{(\lambda_1 - \lambda_1^{(n)})^2 + \dots + (\lambda_{k-1} - \lambda_{k-1}^{(n)})^2}{2h^2}\right)},
 \end{aligned}$$

Rosenblatt transformation enables Galerkin projection



$(a, b) = R^{-1}(\xi_1, \xi_2)$ ensures a well-defined quadrature integration

$$a = \sum_{k=0}^P a_k \psi_k(\xi)$$

$$a_k \propto \int \underbrace{R_a^{-1}(\xi)}_a \psi_k(\xi) w(\xi) d\xi$$

$$b = \sum_{k=0}^P b_k \psi_k(\xi)$$

$$b_k \propto \int \underbrace{R_b^{-1}(\xi)}_b \psi_k(\xi) w(\xi) d\xi$$

Further Reading

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Compressed Sensing

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Bayesian Inference

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Polynomial Chaos Based Uncertainty Quantification

Topic 4: Surrogate Models / Meta-models

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Reservoirs UQ Workshop – Aug 1-3, 2017



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Surrogate Models

- Approaches:
 - Polynomial Chaos Expansions
 - Gaussian Process Models
 - Low-rank Approximations
- Challenges:
 - High-dimensionality of input space
 - Computational cost of forward model

Mitigation Approaches

- High-Dimensionality:
 - Dimensionality reduction (sensitivity analysis)
 - Adaptive Sparse Quadrature approaches
 - Compressed Sensing
 - Low-rank approximations
- Computational Cost:
 - Multifidelity - Multilevel methods (see DAKOTA team)

Bayesian Compressed Sensing

- N training data points $(\mathbf{x}_n, u_n)$ and K basis terms $\Psi_k(\cdot)$
- Projection matrix $\mathbf{P}^{N \times K}$ with $\mathbf{P}_{nk} = \Psi_k(\mathbf{x}_n)$
- Find regression weights $\mathbf{c} = (c_0, \dots, c_{K-1})$ so that

$$\mathbf{u} \approx \mathbf{P}\mathbf{c}$$

or

$$u_n \approx \sum_k c_k \Psi_k(\mathbf{x}_n)$$

- The number of polynomial basis terms grows fast; a p -th order, d -dimensional basis has a total of $K = (p + d)! / (p!d!)$ terms.
- For limited data and large basis set ($N < K$) this is a sparse signal recovery problem $\Rightarrow$ need some regularization/constraints.
- Least-squares $\operatorname{argmin}_{\mathbf{c}} \{ \|\mathbf{u} - \mathbf{P}\mathbf{c}\|_2 \}$
- The 'sparsest' $\operatorname{argmin}_{\mathbf{c}} \{ \|\mathbf{u} - \mathbf{P}\mathbf{c}\|_2 + \alpha \|\mathbf{c}\|_0 \}$
- Compressive sensing $\operatorname{argmin}_{\mathbf{c}} \{ \|\mathbf{u} - \mathbf{P}\mathbf{c}\|_2 + \alpha \|\mathbf{c}\|_1 \}$

Bayesian Compressed Sensing

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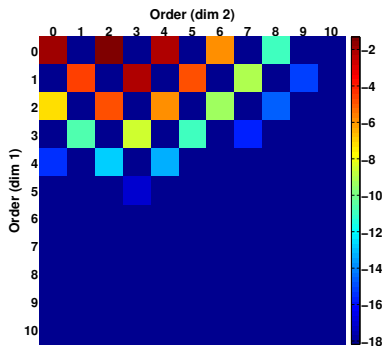
or

$$u_n \approx \sum_k c_k \psi_k(\mathbf{x}_n)$$

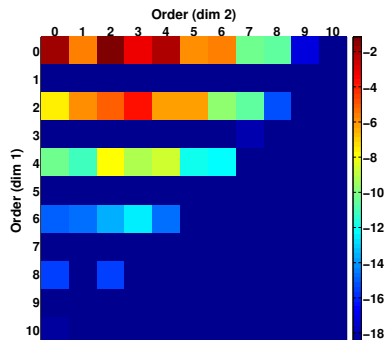
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Bayesian
Likelihood
Prior

BCS removes unnecessary basis terms

$$f(x, y) = \cos(x + 4y)$$



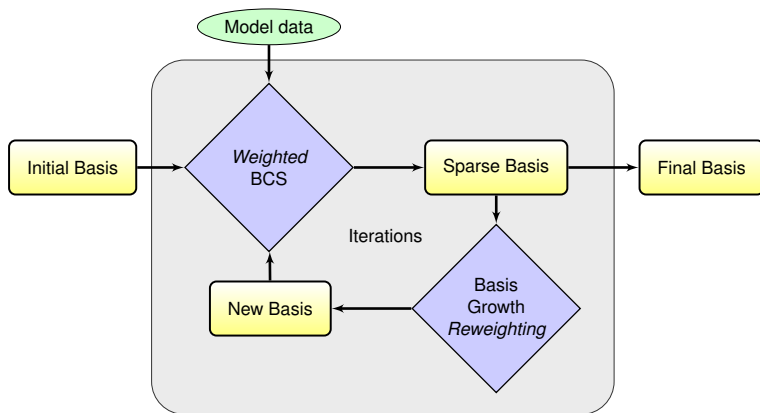
$$f(x, y) = \cos(x^2 + 4y)$$



The square (i, j) represents the (log) spectral coefficient for the basis term $\psi_i(x)\psi_j(y)$.

Iterative Bayesian Compressive Sensing (iBCS)

- *Iterative BCS*: iteratively increase the order for the relevant basis terms while maintaining the dimensionality reduction [[Sargsyan et al. 2014](#)], [[Jakeman et al. 2015](#)].
- Combine basis growth and reweighting!



Low-Rank Approximations

- Univariate function representation: $p + 1$ coefficients

$$u(\xi) = \sum_{k=0}^P u_k \psi_k(\xi)$$

- Multivariate function representation: $P + 1 = \frac{(n+p)!}{n!p!}$

$$u(\xi_1, \dots, \xi_n) = \sum_{k=0}^P u_k \prod_{i=1}^n \psi_{\alpha_i^k}(\xi_i)$$

- Low-rank approximation: $r * (p_1 + p_2 + \dots + p_n)$

$$u(\xi_1, \dots, \xi_n) = \sum_{k=1}^r w_k^{(1)}(\xi_1) \dots w_k^{(n)}(\xi_n)$$

Minimization Problem for Coefficients

- Minimization problem:

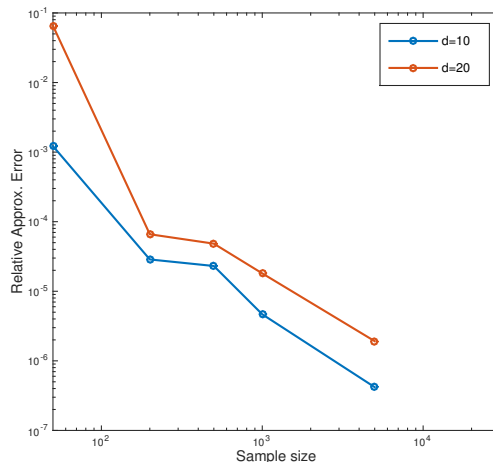
$$\min_{\tilde{u} \in \mathcal{M}} \|u(\xi) - \tilde{u}(\xi)\|^2 + \lambda \mathcal{R}(\tilde{u}(\xi))$$

where $\mathcal{M}$ is a suitable tensor subset (Canonical, Tensor Train) and $\mathcal{R}$ is a regularization function

- Selection of optimal rank r and regularization coefficient λ using cross validation
- Pros: Linear increase in parameters with dimension
- Cons: Non-linear optimization problem. Optimal approximation $\tilde{u}(\xi)$ is often not known

Low rank approximation of Genz Gaussian

- $u(\xi) = \exp\left(-\sum_{i=1}^d c^2(\xi_i - w)^2\right)$, $c = 0.1$, $w = 0.5$
- $\mathcal{M}$: Canonical rank- r , $\mathcal{R} : \ell_1$



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