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Optimal Experimental Design and Uncertainty Propagation in Physical Systems

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About me

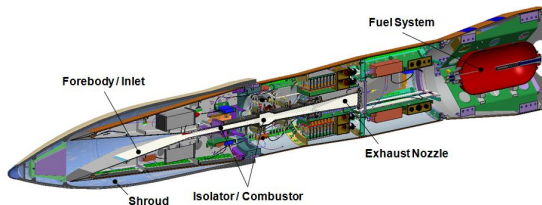
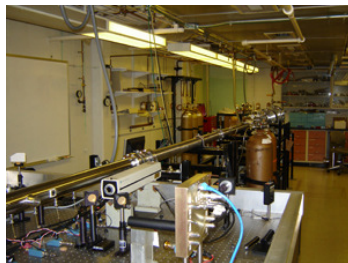
Postdoc at Sandia National Labs, California

- Ph.D. in Computational Science and Engineering, MIT, 2015
- S.M. in Aerospace Engineering, MIT, 2010
- B.A.Sc. in Engineering Science (Aerospace), Toronto, 2008

Research focus: uncertainty quantification



Uncertainty in engineering and science applications



Uncertainty is everywhere

“Uncertainty is everywhere and you cannot escape from it.”

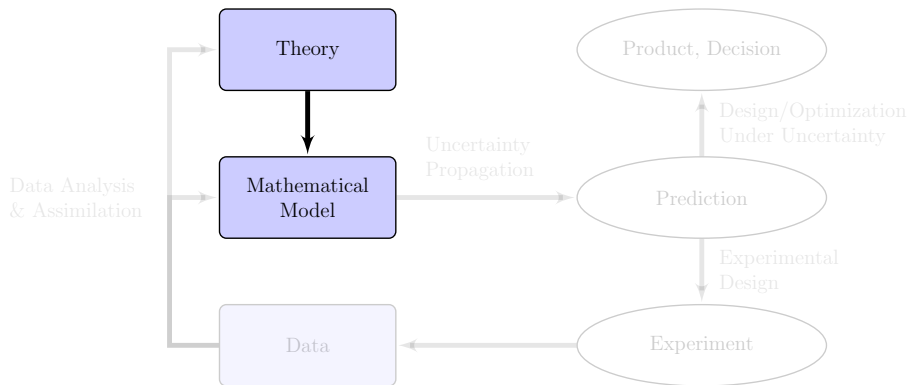
— Dennis Lindley

Uncertainty quantification (UQ) provides a systematic approach to:

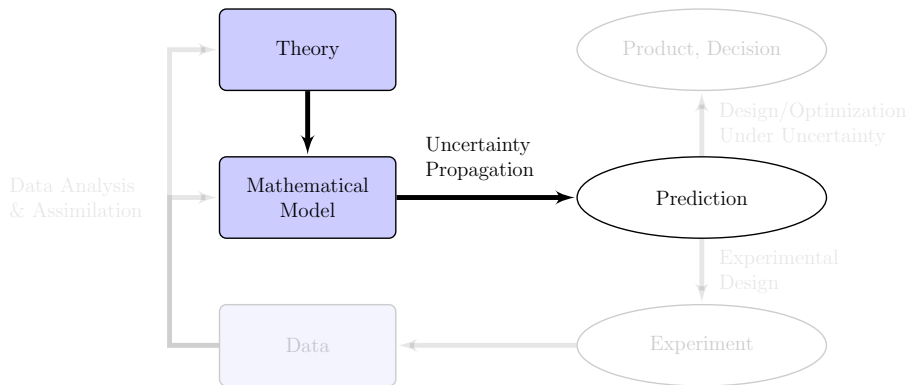
- characterize
- incorporate
- propagate
- reduce

... uncertainty for computational and real-world applications

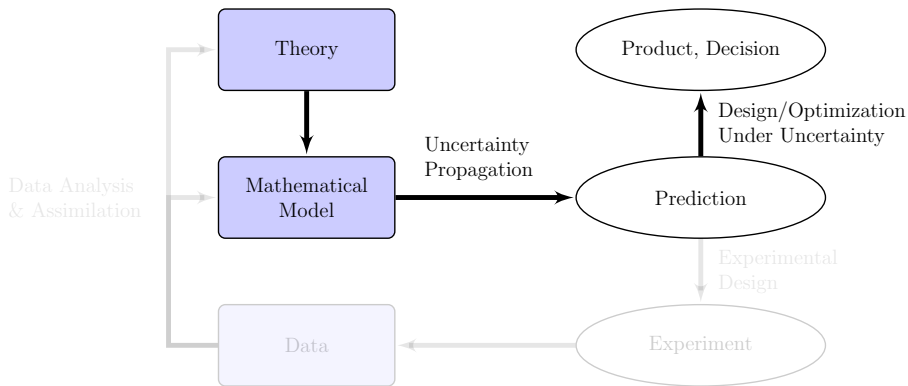
UQ in engineering and science: big picture



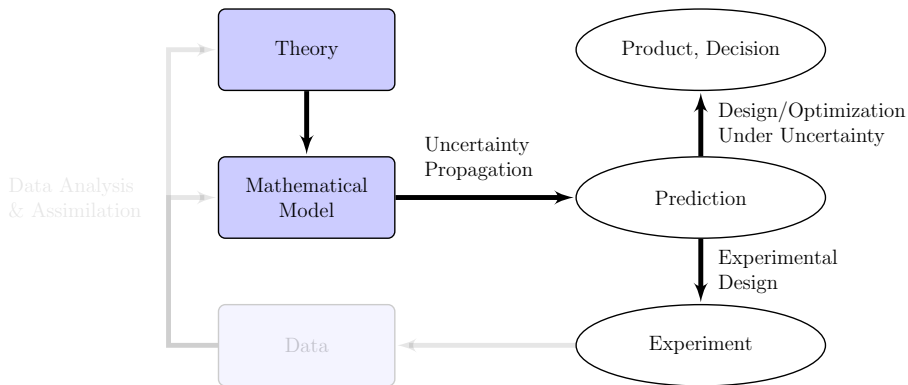
UQ in engineering and science: big picture



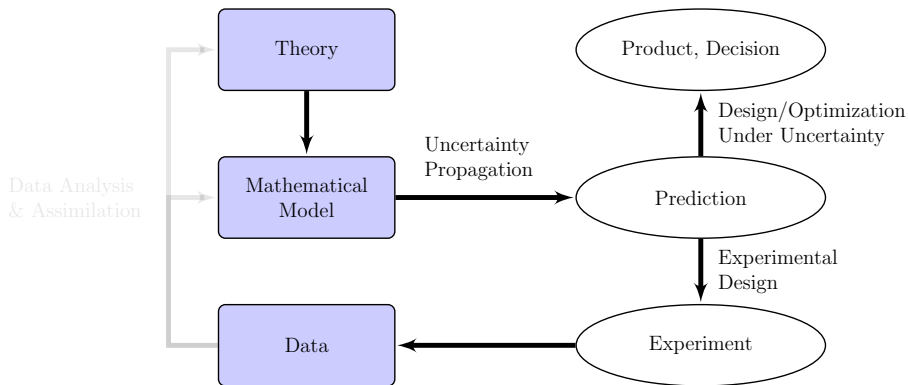
UQ in engineering and science: big picture



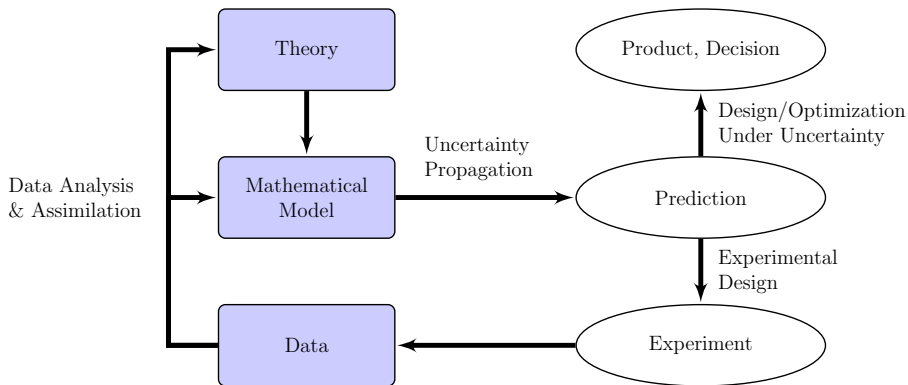
UQ in engineering and science: big picture



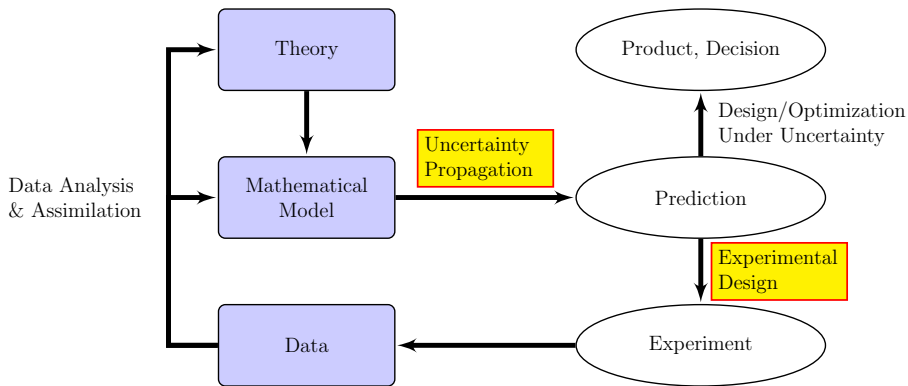
UQ in engineering and science: big picture



UQ in engineering and science: big picture



UQ in engineering and science: big picture



Outline

- 1 Sequential Optimal Experimental Design
 - Formulation
 - Numerical Methods
 - Results
- 2 Uncertainty Propagation in Scramjet Computations
- 3 Summary and Future Work

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Some experiments are more useful than others

Experiments are:

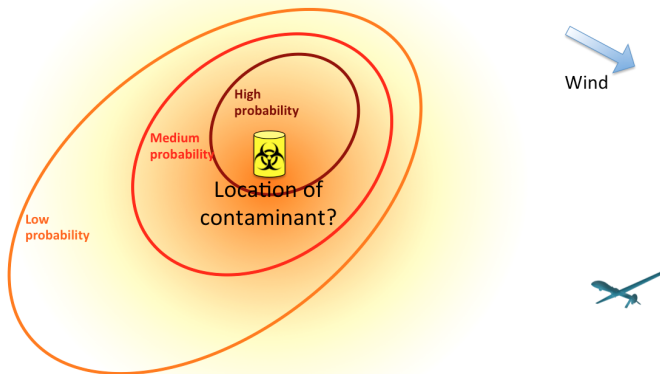
- Expensive
- Time-consuming
- Delicate to perform

Experimental design helps address:

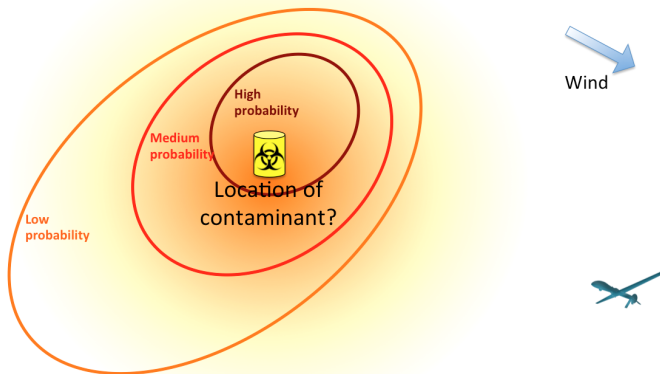
- Under what **conditions** to perform the experiment?
- **What** to measure?
- **Where** to measure?
- **When** to measure?

Today's talk: design of **multiple** experiments

Planning measurements: batch (non-sequential) design



Planning measurements: sequential design



Sequential experimental design is less developed

Batch experimental design:

- Linear: Fisher information matrix (e.g., A -, D -optimal)
 - Nonlinear: advances beyond linearization and Gaussianization
 - Information-based experimental design [Lindley 56]
-

Greedy (myopic) design:

- Repeated application of batch design [Solonen 12, Drovandi 14, Kim 14]
- *Not optimal*

Dynamic programming:

- Fully optimal description (has 1. feedback, 2. forward looking)
- Thus far limited to discrete variables, special problem and solution structures [Carlin, Bradley 98, Brockwell 03, Berry 02]
- Only simple objectives

Contribution and scope

Contribution:

Develop a mathematical framework and numerical tools to find optimal sequential experimental designs in a computationally feasible manner

Scope:

- **Finite** number of experiments
- **Nonlinear** and physically realistic models
- **Continuous** parameter, design, and data spaces of multiple dimensions
- **Bayesian** treatment of uncertainty
- **Non-Gaussian** distributions
- **Information measure** objective (design for parameter inference)

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Core components of general sequential design formulation

Experiment: $k = 0, \dots, N - 1$, total N experiments; $N < \infty$

State: $x_k = [x_{k,b}, x_{k,p}]$ all information needed for optimal future designs

- *Belief state:* $x_{k,b}$ current state of uncertainty
- *Physical state:* $x_{k,p}$ deterministic design-relevant variables

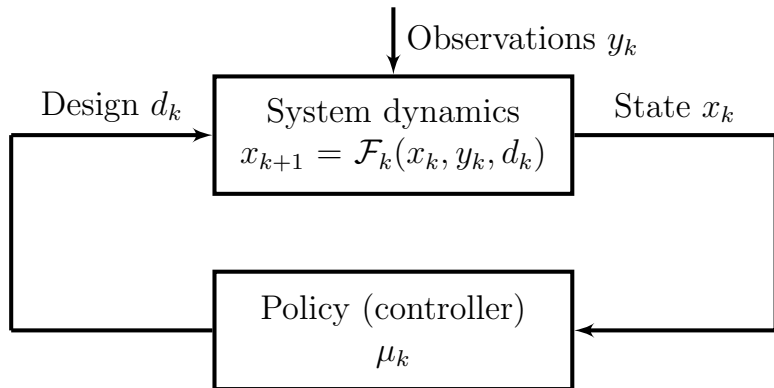
Design: $d_k = \mu_k(x_k)$

seek good *policy* $\pi \equiv \{\mu_0, \mu_1, \dots, \mu_{N-1}\}$

Observations: y_k distributed according to likelihood $f(y_k|\theta, d_k)$
(e.g., $y_k = G(\theta, d_k) + \epsilon$, with ϵ Gaussian)

System dynamics: $x_{k+1} = \mathcal{F}_k(x_k, y_k, d_k)$ state evolution

Sequential design exhibits a closed-loop behavior



The sOED problem: find optimal policy that maximizes the expected total reward

Stage reward: $g_k(x_k, y_k, d_k)$

Terminal reward: $g_N(x_N)$

The sequential optimal experimental design (sOED) problem:

Find π^* where

$$\pi^* = \operatorname{argmax}_{\pi = \{\mu_0, \dots, \mu_{N-1}\}} \mathbb{E}_{y_0, \dots, y_{N-1} | \pi} \left[\sum_{k=0}^{N-1} g_k(x_k, y_k, \mu_k(x_k)) + g_N(x_N) \right]$$

$$\begin{aligned} \text{s.t.} \quad & x_{k+1} = \mathcal{F}_k(x_k, y_k, d_k), \forall k \\ & \mu_k(x_k) \in \mathcal{D}_k, \forall x_k, k \end{aligned}$$

Difficult to solve directly, involves optimization of a functional

The sOED problem in dynamic programming (DP) form

Re-express using Bellman's Principle of Optimality [Bellman 53]

Dynamic programming form (Bellman equations): (e.g., [Bertsekas 05])

$$J_k(x_k) = \max_{d_k \in \mathcal{D}_k} \mathbb{E}_{y_k | x_k, d_k} [g_k(x_k, d_k, y_k) + J_{k+1}(\mathcal{F}_k(x_k, d_k, y_k))]$$

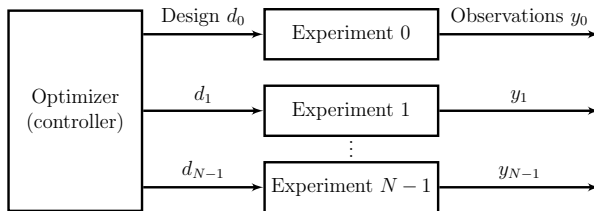
$$J_N(x_N) = g_N(x_N)$$

$k = 0, \dots, N - 1$; $J_k(x_k)$ are value functions

- A set of smaller tail subproblems
- Optimal policy functions implicitly in argmax: $d_k^* = \mu_k^*(x_k)$
- “Curse of dimensionality”: exponential scenario growth from recursion
- Large body of approximate methods: *approximate dynamic programming* (e.g., [Bertsekas 96, Kaelbling 96, Sutton 98, Powell 11])

Batch (non-sequential) design is a special case of the sOED problem, and thus suboptimal

- Has no feedback
- Designs all experiments concurrently as a batch
- Finds optimal *designs* (vectors) rather than a policy



$$\{d_0^*, \dots, d_{N-1}^*\} = \underset{d_0, \dots, d_{N-1}}{\operatorname{argmax}} \mathbb{E}_{y_0, \dots, y_{N-1} | d_{0:N-1}} \left[\sum_{k=0}^{N-1} g_k(x_k, y_k, d_k) + g_N(x_N) \right]$$

Greedy (myopic) design is a special case of the sOED problem (DP form), and thus suboptimal

- Uses feedback
- Considers the *next experiment* only
- Has no future effects

$$J_k(x_k) = \max_{d_k \in \mathcal{D}_k} \mathbb{E}_{y_k | x_k, d_k} \left[g_k(x_k, y_k, d_k) + \cancel{J_{k+1}(\mathcal{F}_k(x_k, d_k, y_k))} \right]$$

$$J_N(x_N) = g_N(x_N)$$

subject to $x_{k+1} = \mathcal{F}_k(x_k, y_k, d_k)$

Sequential Bayesian inference

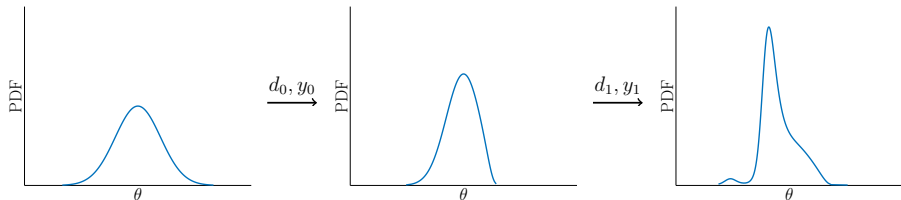
For the k -th experiment:

$$\underbrace{f(\theta|y_k, d_k, I_k)}_{\text{posterior}} = \frac{\underbrace{f(y_k|\theta, d_k, I_k)}_{\text{likelihood}} \underbrace{f(\theta|I_k)}_{\text{prior}}}{\underbrace{f(y_k|d_k, I_k)}_{\text{evidence}}}$$

θ — parameters to infer

I_k — information from previous experiments, $I_k \equiv \{d_0, y_0, \dots, d_{k-1}, y_{k-1}\}$

Conceptually: belief state is posterior random variable $x_{k,b} = \theta|I_k$



Information gain objective for parameter inference

We choose to use total information gain at end of all experiments (Kullback-Leibler (KL) divergence from final posterior to prior)

$g_k(x_k, d_k, y_k)$ = reflects experimental cost

$$g_N(x_N) = D_{KL}(f(x_{N,b}) || f(x_{0,b})) = \int_{\mathcal{H}} f(x_{N,b}) \ln \left[\frac{f(x_{N,b})}{f(x_{0,b})} \right] d\theta$$

Corresponding system dynamics:

- *Belief state*: Bayes' Theorem

$$f(x_{k+1,b}) = \frac{f(y_k | \theta, d_k, l_k) f(x_{k,b})}{f(y_k | d_k, l_k)}$$

- *Physical state*: physical process

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Represent a policy using one-step lookahead form

One-step lookahead policy representation: (e.g., [Bertsekas 05])

$$\mu_k(x_k) = \operatorname{argmax}_{d_k \in \mathcal{D}_k} \mathbb{E}_{y_k | x_k, d_k} \left[g_k(x_k, y_k, d_k) + \tilde{J}_{k+1}(\mathcal{F}_k(x_k, y_k, d_k)) \right]$$

Approximate value functions using **linear architecture**:

$$\tilde{J}_k(x_k) = r_k^\top \phi_k(x_k)$$

ϕ_k features (selected from heuristics), r_k weights

Approximate value iteration (backward induction with regression):

$$\tilde{J}_k(x_k) = \mathcal{P} \left\{ \max_{d_k \in \mathcal{D}_k} \mathbb{E}_{y_k | x_k, d_k} [g_k(x_k, d_k, y_k) + \tilde{J}_{k+1}(\mathcal{F}_k(x_k, d_k, y_k))] \right\}$$

Start with $\tilde{J}_N(x_N) \equiv g_N(x_N)$, and proceed backwards $k = N - 1, \dots, 1$

\mathcal{P} : regression operator, samples from **exploration** and **exploitation**

Belief state representation

Conceptually: belief state is posterior random variable

How to numerically represent it ...

- for general non-Gaussian continuous random variables
- in a finite-dimensional manner
- to easily perform Bayesian inference repeatedly in evaluating

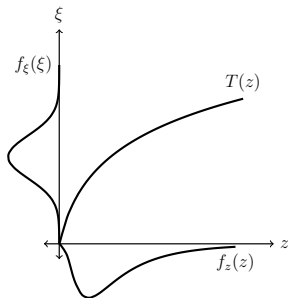
$$\tilde{J}_k(x_k) = \mathcal{P} \left\{ \max_{d_k \in \mathcal{D}_k} \mathbb{E}_{y_k | x_k, d_k} \left[g_k(x_k, d_k, y_k) + \tilde{J}_{k+1}(\mathcal{F}_k(x_k, d_k, y_k)) \right] \right\}$$

Traditional approaches:

- Gaussian approximation and model linearization
- Gridding or functional approximation of its PDF or CDF
- Non-parametrics (with particle filter, MCMC)

Often expensive and some do not scale well to multiple dimensions. We seek an approach that can quickly perform **many** Bayesian inferences.

Transport map transforms *distributions* (e.g., [Villani 08])



- $\xi \sim$ reference distribution, $z \sim$ target distribution
- Equivalence in distribution $\xi \stackrel{i.d.}{=} T(z)$
- Knothe-Rosenblatt (KR) maps: defined by conditional distributions, is triangular and monotone, exists and is unique [Rosenblatt 52, Carlier 10]
- Easy to construct from samples: convex optimization problem
- Target *joint* distribution for fast approx Bayesian inference

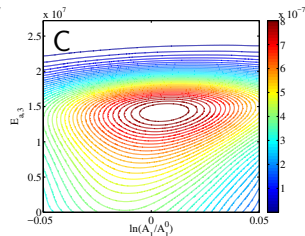
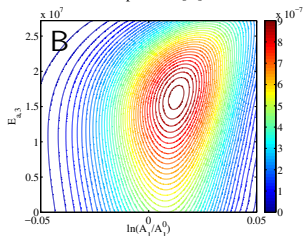
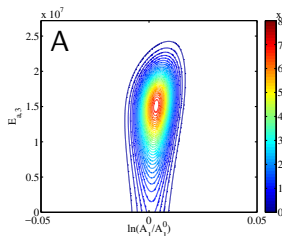
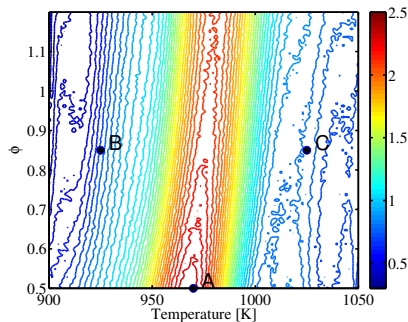
Final algorithm for sOED

- ① **Set parameters**
- ② **Initial exploration**
- ③ **Make joint map**
- ④ **Iterate to refine . . .**
 - (a) **Exploration**
 - (b) **Exploitation**
 - (c) **Approximate value iteration**
- ⑤ **Extract final policy parameterization**

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Shock tube experiments for combustion kinetics



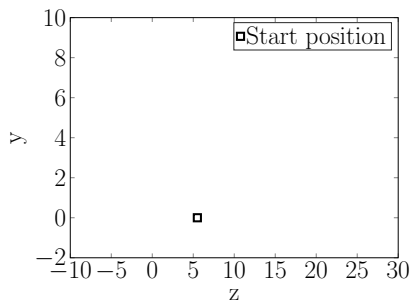
1D source inversion problem: problem settings

$$y_k = \frac{s}{\sqrt{2\pi}2\sqrt{(0.3 + Dt)}} \exp\left(-\frac{\|\theta + d_w(t) - z_{k+1}\|^2}{2(4)(0.3 + Dt)}\right) + \epsilon_k$$

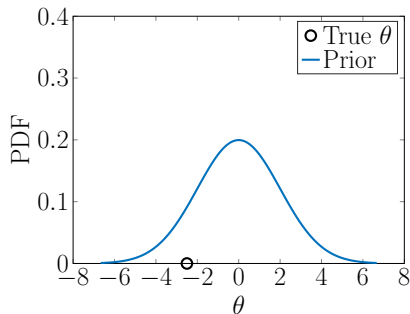
- 2 experiments
- $\theta \sim \mathcal{N}(0, 2^2)$ starting location: 5.5
- Strong wind blows to the right after first experiment
- Quadratic movement penalty

1D source inversion problem: state and plume evolution

An example scenario:



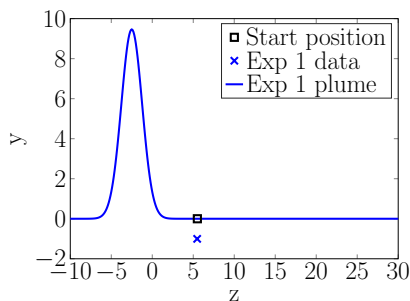
physical state and plume



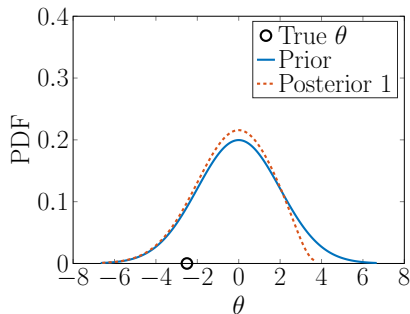
belief state

1D source inversion problem: state and plume evolution

An example scenario:



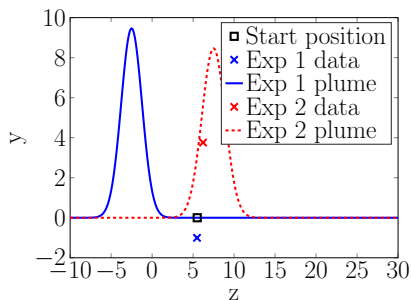
physical state and plume



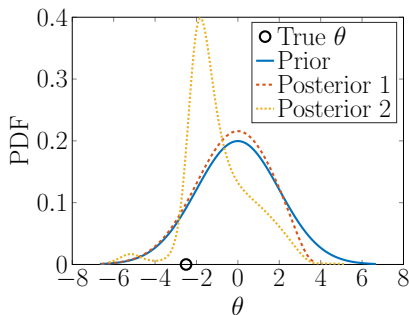
belief state

1D source inversion problem: state and plume evolution

An example scenario:



physical state and plume



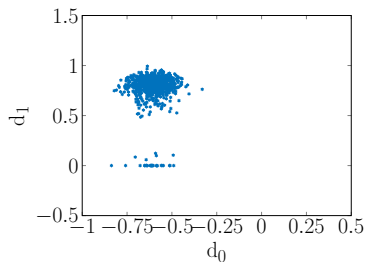
belief state

1D source inversion problem: case 1

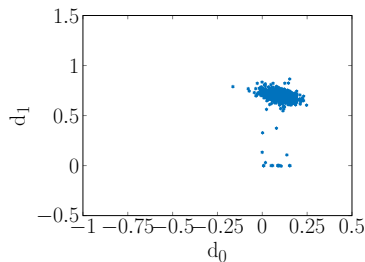
advantages of sOED over greedy design

Greedy design does not account for future wind conditions

Expected reward: greedy (0.07), sOED (0.15)



greedy design



sOED

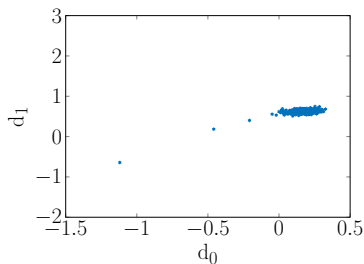
1D source inversion problem: case 2

advantages of sOED over batch design

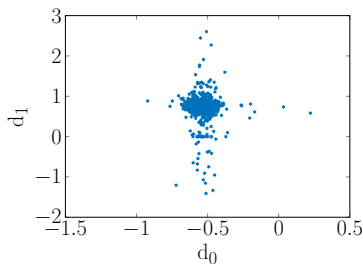
A more precise instrument available only if prior variance < 3

Batch design does not have feedback

Expected reward: batch (0.15), sOED (0.26)



batch design

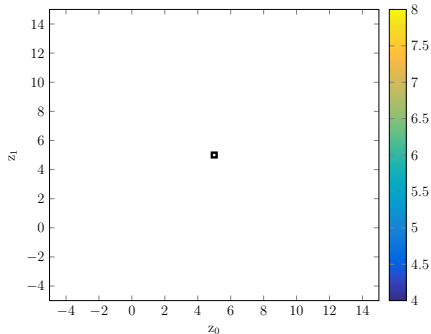


sOED

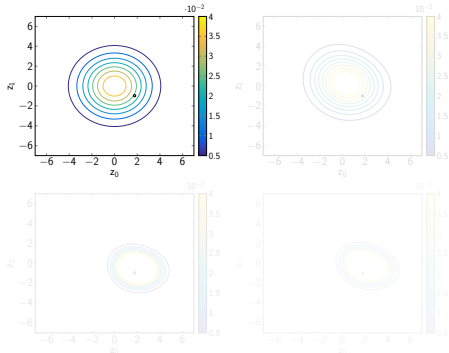
2D source inversion problem: trajectory example #1

An example scenario:

physical state and plume



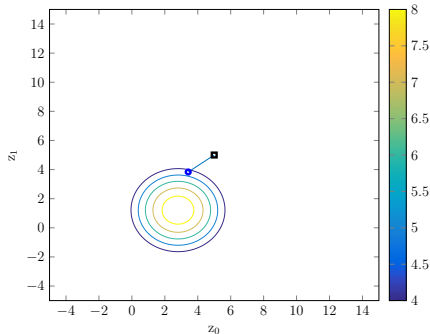
belief states



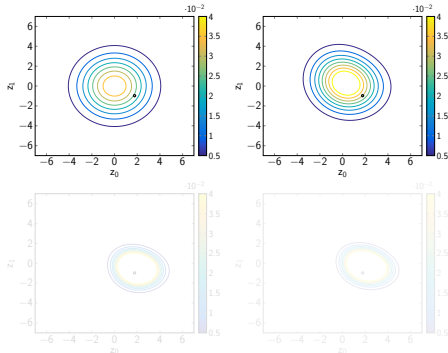
2D source inversion problem: trajectory example #1

An example scenario:

physical state and plume



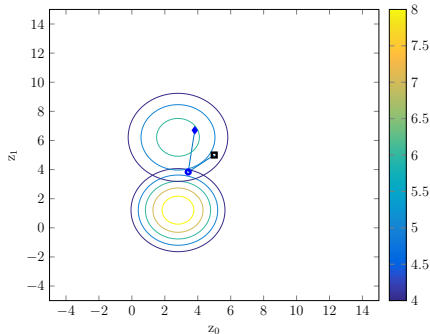
belief states



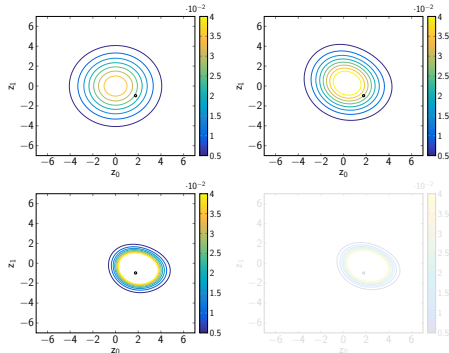
2D source inversion problem: trajectory example #1

An example scenario:

physical state and plume



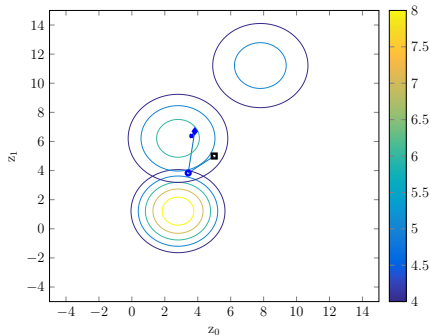
belief states



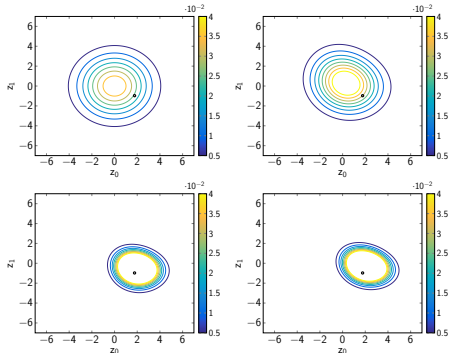
2D source inversion problem: trajectory example #1

An example scenario:

physical state and plume



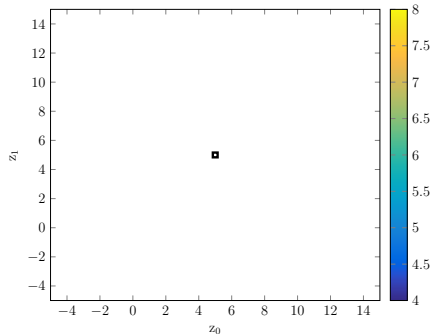
belief states



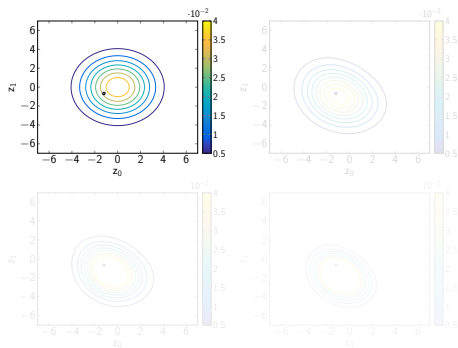
2D source inversion problem: trajectory example #2

Another example scenario:

physical state and plume



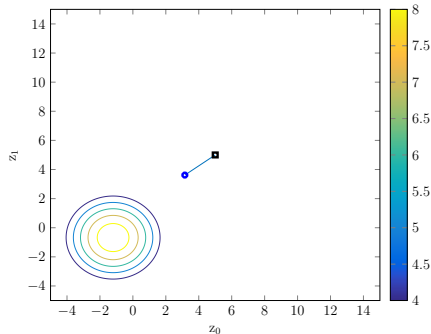
belief states



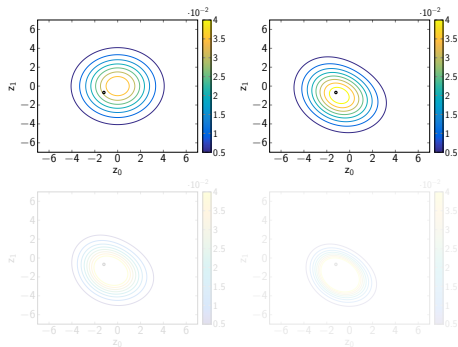
2D source inversion problem: trajectory example #2

Another example scenario:

physical state and plume



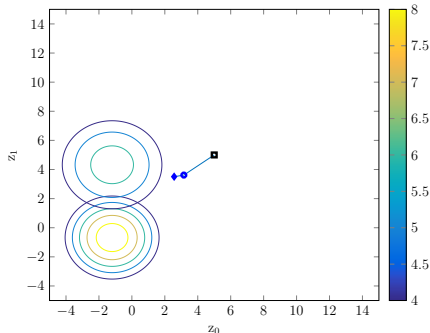
belief states



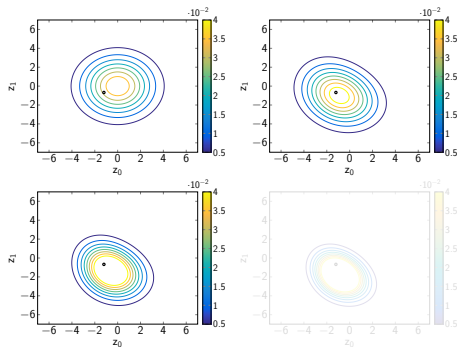
2D source inversion problem: trajectory example #2

Another example scenario:

physical state and plume



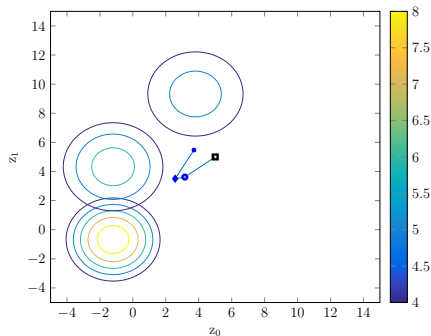
belief states



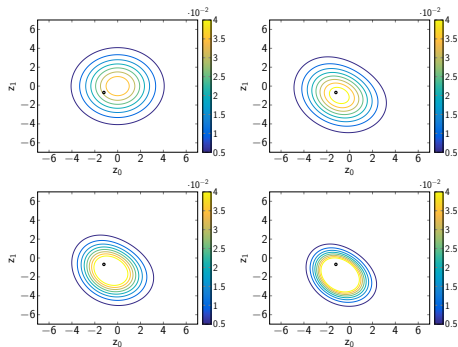
2D source inversion problem: trajectory example #2

Another example scenario:

physical state and plume



belief states



Summary

- Formulated the sequential optimal experimental design (sOED) problem rigorously (has 1. feedback, 2. forward looking)
- Developed new numerical methods to solve the sOED problem in a computationally-feasible manner, using
 - approximate dynamic programming
 - transport maps
- Demonstrated computational effectiveness on realistic applications

References:

1. Huan & Marzouk, "Simulation-based optimal Bayesian experimental design for nonlinear systems," *Journal of Computational Physics*, 232(1):288-317, 2013.
2. Huan & Marzouk, "Gradient-Based Stochastic Optimization Methods in Bayesian Experimental Design," *International Journal of Uncertainty Quantification*, 4(6):479-510, 2014.
3. Huan & Marzouk, "Sequential Bayesian Optimal Experimental Design via Approximate Dynamic Programming," *arXiv: 1604.08320*, 2016.

Outline

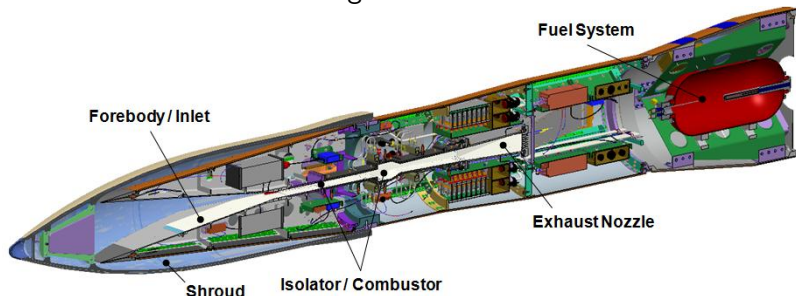
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HIFiRE-II Scramjet

Development of scramjet¹ engine involves

- flow simulations
- uncertainty quantification (UQ)
- design optimization

We focus on the HIFiRE-II² configuration:



¹supersonic combustor ramjet

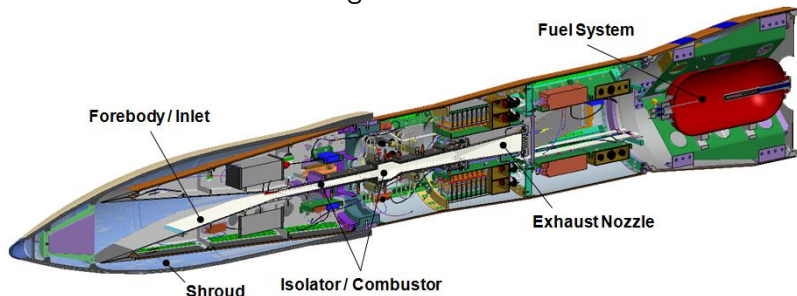
²Hypersonic International Flight Research and Experimentation-II

HIFiRE-II Scramjet

Development of scramjet¹ engine involves

- flow simulations
- **uncertainty quantification (UQ)**
- design optimization

We focus on the HIFiRE-II² configuration:

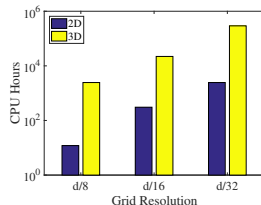
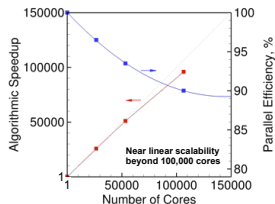
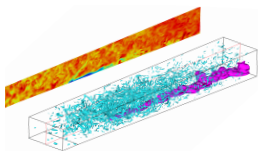


¹supersonic combustor ramjet

²Hypersonic International Flight Research and Experimentation-II

Direct UQ would be intractable

RAPTOR: LES solver by Oefelein *et al.* at Sandia [Oefelein 06]



Highly-scalable but still **very expensive**

Major challenges: for uncertainty quantification

- Many uncertain parameters (high stochastic dimension)
- Expensive simulations

⇒ direct exploration of parameter space **intractable!**

Global sensitivity analysis: Sobol indices

Global sensitivity analysis (GSA) [Saltelli 04, Saltelli 08]

- For a given quantify of interest (Qol) ...
- Variance of Qol decomposed into contributions from each parameter
- Sobol indices rank parameters by their contributions [Sobol 03]

$$\text{Total effect} \quad S_{T_i} = \frac{\mathbb{E}_{\lambda \sim i} [\text{Var}_{\lambda_i} (f(\lambda) | \lambda_i)]}{\text{Var}(f(\lambda))}$$

S_{T_i} small \Rightarrow low impact parameter \Rightarrow fix value (i.e. dim. eliminated)

How to compute?

- Monte Carlo estimators [Saltelli 02, Saltelli 10] still prohibitive for LES
- **Our plan:** construct affordable surrogate models via polynomial chaos expansion (PCE)

Polynomial chaos expansions

A QoI (output) random variable can be expanded as follows:

$$f(\lambda(\xi)) = \sum_{\beta \in \mathcal{J}} c_{\beta} \Psi_{\beta}(\xi)$$

- c_{β} : PCE coefficients
- ξ : reference random vector (e.g., uniform, Gaussian)
- Ψ_{β} : multivariate orthonormal polynomial (e.g., Legendre, Hermite)
- β : multi-index, reflects order of polynomial basis

Orthonormality property

⇒ extract Sobol indices analytically from coefficients (no Monte Carlo!):

$$S_{T_i} = \frac{1}{\text{Var}(f(\lambda))} \sum_{\beta \in \mathcal{J}: \beta_i > 0} c_{\beta}^2 \quad \text{where} \quad \text{Var}(f(\lambda)) = \sum_{0 \neq \beta \in \mathcal{J}} c_{\beta}^2$$

Sparse polynomial chaos expansions

Non-intrusive regression to compute expansion coefficients $Gc = f$:

$$\underbrace{\begin{bmatrix} \Psi_{\beta^1}(\xi^{(1)}) & \cdots & \Psi_{\beta^N}(\xi^{(1)}) \\ \vdots & & \vdots \\ \Psi_{\beta^1}(\xi^{(M)}) & \cdots & \Psi_{\beta^N}(\xi^{(M)}) \end{bmatrix}}_G \underbrace{\begin{bmatrix} c_{\beta^1} \\ \vdots \\ c_{\beta^N} \end{bmatrix}}_c = \underbrace{\begin{bmatrix} f(\lambda(\xi^{(1)})) \\ \vdots \\ f(\lambda(\xi^{(M)})) \end{bmatrix}}_f,$$

Challenges:

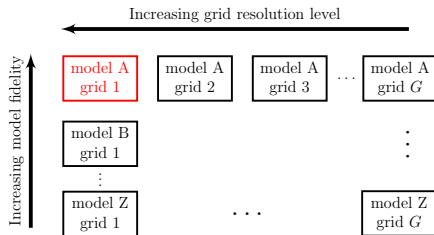
- Few LES flow solves (data), many PCE basis (e.g. total-order degree 3 in 24 dimensions: 2925 terms)
- Extremely under-determined system ($N \gg M$)

Our approach: use compressed sensing to find sparse solution (LASSO)

$$\min_c \frac{1}{2} \|c\|_1 + \tau \|Gc - f\|_2^2$$

discover and retain only basis terms with high magnitude coefficients

Multilevel and multifidelity forms



Telescopic sum:

$$f_L(\lambda) = f_0(\lambda) + \sum_{\ell=1}^L f_{\Delta_\ell}(\lambda)$$

- ℓ indicates different grid levels or fidelity of models
- Δ_ℓ indicates difference between models ℓ and $\ell - 1$

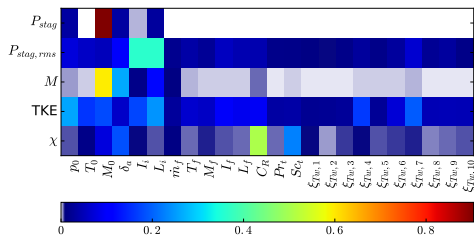
Function approximation: $f_L(\lambda) \approx \hat{f}_L(\lambda) = \hat{f}_0(\lambda) + \sum_{\ell=1}^L \hat{f}_{\Delta_\ell}(\lambda)$

Unit problem: 24 parameters

Parameter	Range	Description
Inlet boundary conditions		
p_0	[1.406, 1.554] MPa	Stagnation pressure
T_0	[1472.5, 1627.5] K	Stagnation temperature
M_0	[2.259, 2.761]	Mach number
δ_a	[2, 6] mm	Boundary layer thickness
I_i	[0, 0.05]	Turbulence intensity magnitude
L_i	[0, 8] mm	Turbulence length scale
Fuel inflow boundary conditions		
\dot{m}_f	[6.633, 8.107] $\times 10^{-3}$ kg/s	Mass flux
T_f	[285, 315] K	Static temperature
M_f	[0.95, 1.05]	Mach number
I_f	[0, 0.05]	Turbulence intensity magnitude
L_f	[0, 1] mm	Turbulence length scale
Turbulence model parameters		
C_R	[0.01, 0.06]	Modified Smagorinsky constant
Pr_t	[0.5, 1.7]	Turbulent Prandtl number
Sc_t	[0.5, 1.7]	Turbulent Schmidt number
Wall boundary conditions		
T_w	Expansion in 10 params of $\mathcal{N}(0, 1)$	Wall temperature represented via Karhunen-Loève expansion

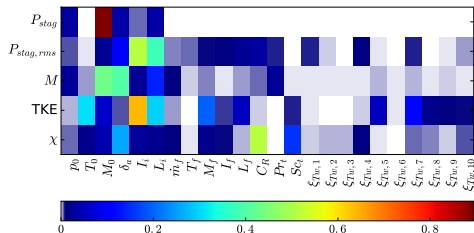
- 2D runs: 1939 (coarse grid), 79 (fine grid)
- 3D runs: 46 (coarse grid), 11 (fine grid)

Unit problem: total sensitivity



Multilevel expansion of:

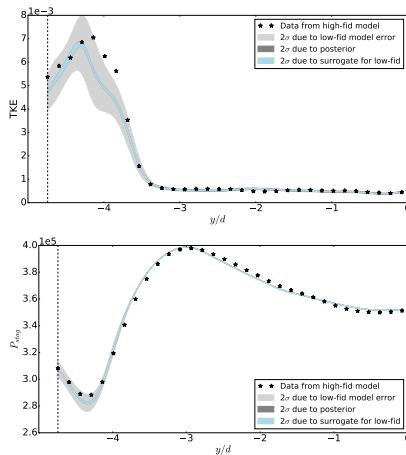
$$\hat{f}_{2D,d/16} = \hat{f}_{2D,d/8} + \hat{f}_{\Delta_{2D,d/16-2D,d/8}}$$



Multifidelity expansion of:

$$\hat{f}_{3D,d/8} = \hat{f}_{2D,d/8} + \hat{f}_{\Delta_{3D,d/8-2D,d/8}}$$

Posterior predictives in model QoIs



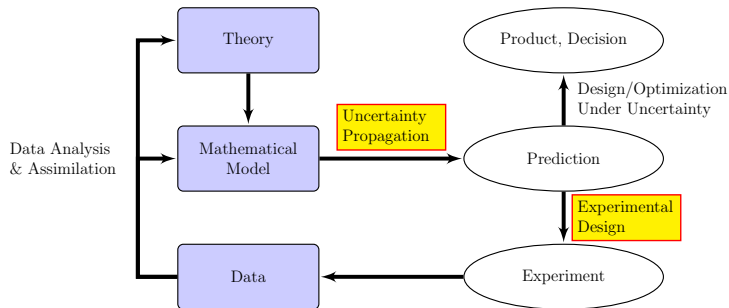
Reference:

X. Huan, C. Safta, K. Sargsyan, G. Geraci, M. S. Eldred, Z. P. Vane, G. Lacaze, J. C. Oefelein, and H. N. Najm, "Global Sensitivity Analysis and Quantification of Model Error for Large Eddy Simulation in Scramjet Design," *AIAA paper 2017-1089*, 2017.

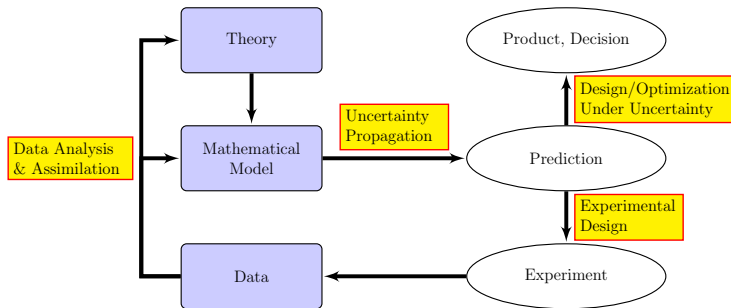
Outline

- 1 Sequential Optimal Experimental Design
 - Formulation
 - Numerical Methods
 - Results
- 2 Uncertainty Propagation in Scramjet Computations
- 3 Summary and Future Work

Summary and vision



Summary and vision



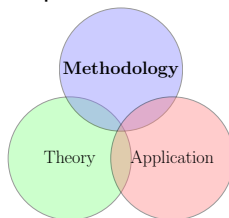
Vision: develop comprehensive UQ capability to

- ① **design experiments** for acquiring data
- ② **analyze and assimilate data** for improving model and theory
- ③ **propagate uncertainty** for making predictions
- ④ **optimize and design** in the presence of uncertainty

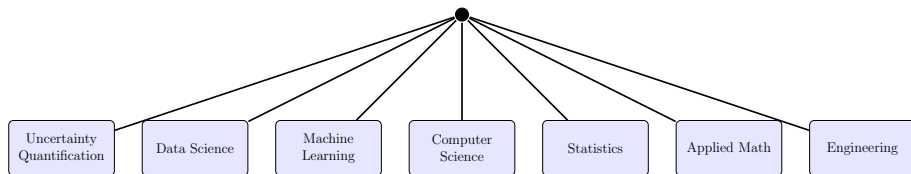
for applications with complex and realistic physics-based models

Summary and vision

Focus on **methodology** development



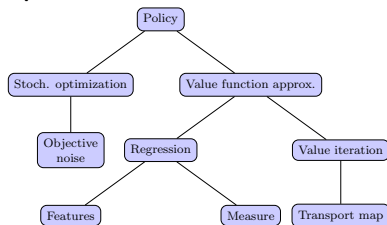
Naturally **interdisciplinary**



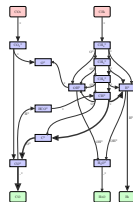
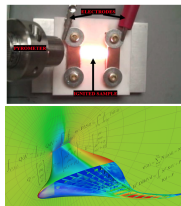
Many **collaboration** opportunities

Ideas for future work

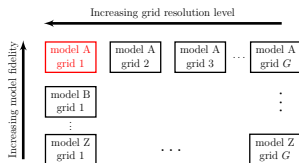
Adaptive numerical methods



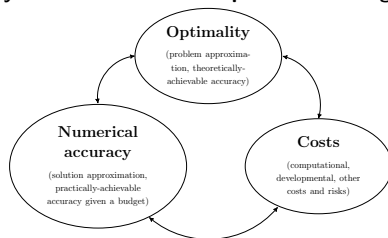
UQ for expensive, high-dimensional models



Multi-model management for experimental design and statistical inference



Hybrid frameworks for experimental design



Student education program

Strong foundational courses from multiple departments, for example:

- V&V, machine learning, data science, data mining, statistical methods in data mining, applied Bayesian statistics, statistical inference, optimization, stochastic analysis, stochastic control theory, etc.

New course ideas:

- uncertainty quantification
- statistical data analysis
- optimal experimental design / decision-making under uncertainty
- model reduction
- optimization under uncertainty
- inverse problems

Acknowledgment

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Computational Mathematics Program
- U.S. Department of Energy
Office of Science, Office of Advanced Scientific Computing Research (ASCR)
- KAUST Global Research Partnership
- Natural Sciences and Engineering Research Council of Canada (NSERC)

³Sandia National Laboratories is a multi-program laboratory managed and operated by Sandia Corporation, a wholly owned subsidiary of Lockheed Martin Corporation, for the U.S. Department of Energy's National Nuclear Security Administration under contract DE-AC04-94AL85000.

Thank You!



Supplement

Supplemental Slides

Optimal Experimental Design

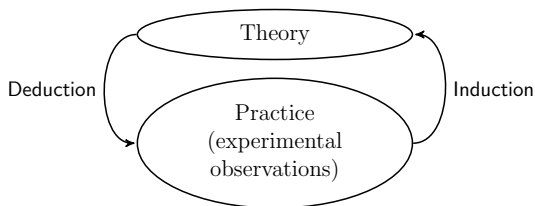
open-loop design

- classical linear design theory [Atkinson & Donev (1992)]
- Bayesian experimental design [Chaloner & Verdinelli (1995)]
- nonlinear models [Box & Lucas (1959), Lindley (1956, 1972), Sebastiani & Wynn (1997, 2000), Ryan (2003), Müller & Parmigiani (1993, 1995), Clyde *et al.* (1995), Loredó (2010)]

closed-loop design

- greedy (rolling-horizon of 1 experiment) [Gautier & Pronzato (1998), Negoescu *et al.* (2011), Solonen *et al.* (2012)]
- dynamic programming approach [Müller (2006), Brockwell & Kadane (2003), Lewis & Berry (1994), Christen & Nakamura (2003)]

Experiments are an integral part of learning



*"... science is a means whereby learning is achieved, not by mere theoretical speculation on the one hand, nor by the undirected accumulation of practical facts on the other, but rather by a motivated iteration between **theory** and **practice** ..."* [Box 76]

— George E. P. Box

Inference can be done by conditioning the joint map

Bayes' theorem: posterior is simply conditioning the joint distribution

$$f(\theta|y) = \frac{f(y|\theta)f(\theta)}{f(y)} = \frac{f(y, \theta)}{f(y)}$$

For one experiment: KR map from (d, y, θ) to $\xi \sim \mathcal{N}(0, I)$ is:

$$\begin{bmatrix} \xi_1 \\ \xi_2 \\ \xi_3 \end{bmatrix} = \begin{bmatrix} T_d(d) \\ T_{y|d}(d, y) \\ T_{\theta|y,d}(d, y, \theta) \end{bmatrix} = \begin{bmatrix} \Phi^{-1}(F(d)) \\ \Phi^{-1}(F(y|d)) \\ \Phi^{-1}(F(\theta|y, d)) \end{bmatrix}$$

KR map of posterior given realizations d^* and y^* is:

$$T_{\theta|y^*, d^*}(\theta) = \Phi^{-1}(F(\theta|y^*, d^*))$$

This is precisely $T_{\theta|y,d}$ conditioned on d^* and y^* : $T_{\theta|y,d}(d^*, y^*, \theta)$!

Map construction [Parno 15]

- Approximate using linear parameterization $\xi = T(z; \gamma)$
(we use a linear architecture of monomial polynomial basis functions)
- Construct $T(z; \gamma)$ by minimizing KL using M samples from target

For i -th dimension of the multivariate map:

$$\min_{\gamma_i} \sum_{m=1}^M \left[0.5 T_i^2(z^{(m)}; \gamma_i) - \ln \left. \frac{\partial T_i(z; \gamma_i)}{\partial z_i} \right|_{z^{(m)}} \right]$$

$$s.t. \quad \left. \frac{\partial T_i(z; \gamma_i)}{\partial z_i} \right|_{z^{(m)}} \geq \lambda_{min} > 0, m = 1, \dots, M$$

Numerically attractive properties:

- Dimensions are separable
- Convex optimization
- Model-independent (only uses samples)

A good joint map also implies good posterior maps

Question: are the resulting posterior maps accurate?

Theorem: ([Huan 15])

- If we construct a joint map that is *optimal* in an *approximation* subspace
- then the resulting posterior maps from conditioning on d_k and y_k are also optimal *on average*
- with respect to the probability measure of d_k and y_k

A good joint map also implies good posterior maps

Theorem

Let the optimal joint map be

$$T_{1:n}^* = \operatorname{argmin}_{T_{1:n} \in \mathcal{T}_{1:n}} D_{KL} \left(f_{1:n}(\cdot) || \tilde{f}_{1:n}(\cdot; T_{1:n}) \right)$$

where $\tilde{f}_{1:n}$ is the target density induced by candidate map $T_{1:n}$. Then, for any $0 < j \leq n$, the dimension-truncated “head” map is also the optimal map for dimensions $(1:j)$, in the sense

$$T_{1:j}^* = \operatorname{argmin}_{T_{1:j} \in \mathcal{T}_{1:j}} D_{KL} \left(f_{1:j}(\cdot) || \tilde{f}_{1:j}(\cdot; T_{1:j}) \right),$$

where $\mathcal{T}_{1:j} \subseteq \mathcal{T}_{1:n}$ is its first j -dimensional truncation.

Corollary

For each $j = 1, \dots, n$, the component map is optimal in an expected sense:

$$T_j^* = \operatorname{argmin}_{T_j \in \mathcal{T}_j} \mathbb{E}_{\mathbf{z}_{1:(j-1)}} \left[D_{KL} \left(f_{j|1:(j-1)}(\cdot | \mathbf{z}_{1:(j-1)}) || \tilde{f}_{j|1:(j-1)}(\cdot | \mathbf{z}_{1:(j-1)}; T_j) \right) \right].$$

Extension to multiple experiments

$k = 0$	1	2	...
$\xi_{d_0} = T_{d_0}(d_0)$	$\xi_{d_0} = T_{d_0}(d_0)$	$\xi_{d_0} = T_{d_0}(d_0)$	
$\xi_{y_0} = T_{y_0}(d_0, y_0)$	$\xi_{y_0} = T_{y_0}(d_0, y_0)$	$\xi_{y_0} = T_{y_0}(d_0, y_0)$	
$\xi_{\theta_0} = T_{\theta_0}(d_0, y_0, \theta)$	$\xi_{d_1} = T_{d_1}(d_0, y_0, d_1)$	$\xi_{d_1} = T_{d_1}(d_0, y_0, d_1)$	
	$\xi_{y_1} = T_{y_1}(d_0, y_0, d_1, y_1)$	$\xi_{y_1} = T_{y_1}(d_0, y_0, d_1, y_1)$	
	$\xi_{\theta_1} = T_{\theta_1}(d_0, y_0, d_1, y_1, \theta)$	$\xi_{d_2} = T_{d_2}(d_0, y_0, d_1, y_1, d_2)$	
		$\xi_{y_2} = T_{y_2}(d_0, y_0, d_1, y_1, d_2, y_2)$	
		$\xi_{\theta_2} = T_{\theta_2}(d_0, y_0, d_1, y_1, d_2, y_2, \theta)$	

- Posterior map after Bayesian inference on $k + 1$ experiments is the n_θ -dimensional $T_{\theta_k | d_0^*, y_0^*, \dots, d_k^*, y_k^*}(\theta)$
- Components grouped by the red rectangular boxes are identical; concatenate unique parts and construct overall map in one shot
- Map constructed using samples of trajectory simulation (exploration and exploitation)

Final algorithm for sOED

- ① **Set parameters:** Select VFA features, exploration measure, L , R_0 , R , T
- ② **Initial exploration:** Simulate R_0 exploration trajectories, without inference
- ③ **Make exploration joint map:** Make T_{explore} from these samples
- ④ For $\ell = 1, \dots, L$
 - (a) **Exploration:** Simulate R exploration trajectories, with inference using T_{explore} , store states visited $\mathcal{X}_{k,\text{explore}}^\ell = \{x_k^r\}_{r=1}^R$
 - (b) **Exploitation:** If $\ell > 1$, simulate T exploitation trajectories by evaluating $d_k = \operatorname{argmax}_{d'_k} \mathbb{E}_{y_k|x_k, d'_k} \left[g_k(x_k, y_k, d'_k) + \tilde{J}_{k+1}^{\ell-1}(\mathcal{F}_k(x_k, y_k, d'_k)) \right]$, with inference using T_{explore} , store states visited $\mathcal{X}_{k,\text{exploit}}^\ell = \{x_k^t\}_{t=1}^T$
 - (c) **Approximate value iteration:** Update \tilde{J}_k^ℓ functions from new regression points $x_k^{rt} \in \{\mathcal{X}_{k,\text{explore}}^\ell \cup \mathcal{X}_{k,\text{exploit}}^\ell\}$ by evaluating $t_k^{rt}(x_k^{rt}) = \max_{d'_k} \mathbb{E}_{y_k|x_k, d'_k} \left[g_k(x_k^{rt}, y_k, d'_k) + \tilde{J}_{k+1}^\ell(\mathcal{F}_k(x_k^{rt}, y_k, d'_k)) \right]$ or terminal reward at all regression points, with inference using T_{explore}
- ⑤ **Extract final policy parameterization:** $\tilde{J}_k^L, \forall k$

Linear-Gaussian problem: problem settings

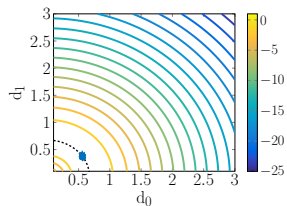
$$y_k = \theta d_k + \epsilon_k$$

- $\theta \sim \mathcal{N}(0, 3^2)$ $d_k \in [0.1, 3]$
 $\epsilon_k \stackrel{\text{iid}}{\sim} \mathcal{N}(0, 1^2)$ $g_k = 0$
- Conjugate family, all posteriors are Gaussian
- $N = 2$ experiments
- Additional terminal reward component (target variance)

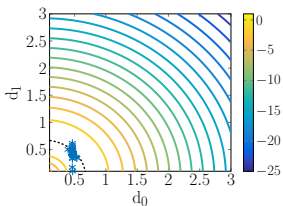
$$g_N = D_{KL}(f(x_{N,b}) || f(x_{0,b})) - 2(\ln \sigma_N^2 - \ln 2)^2$$

- Compare state representations:
 - **analytic** (mean and variance)
 - PDF on a **grid** (50 grid points)
 - **map** (total order polynomial degree 3, made with 10^5 samples)

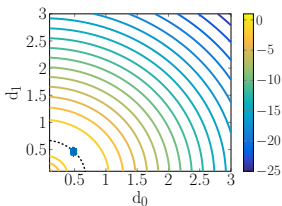
Linear-Gaussian problem: d_0 , d_1 on exact expected utility



analytic



grid



map

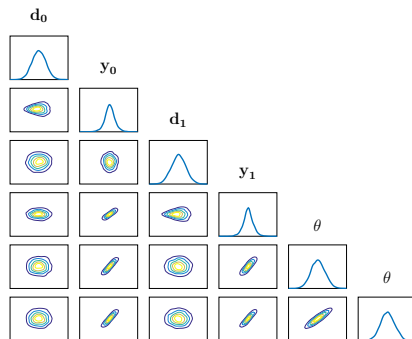
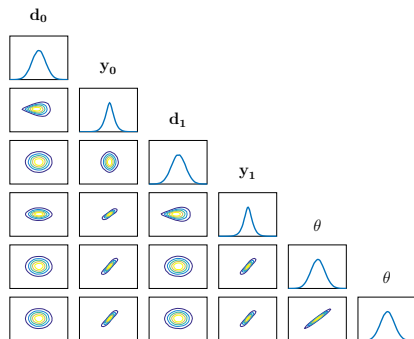
Linear-Gaussian problem: expected total rewards

Expected rewards from 1000 simulated trajectories:

exact	analytic	grid	map
0.7833	0.79	0.76	0.79

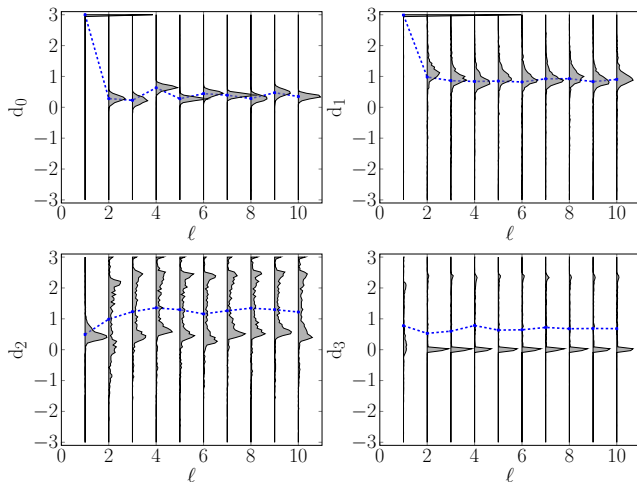
- All values have standard error ± 0.02
- Optimal policy for linear-Gaussian problem with constant noise variance is state-independent and non-unique (“exchangeability” between d_0 and d_1)
- Excellent agreement between analytic, grid, and map methods
- In comparison, exploration policy expected reward is -8.5

Linear-Gaussian problem: joint map

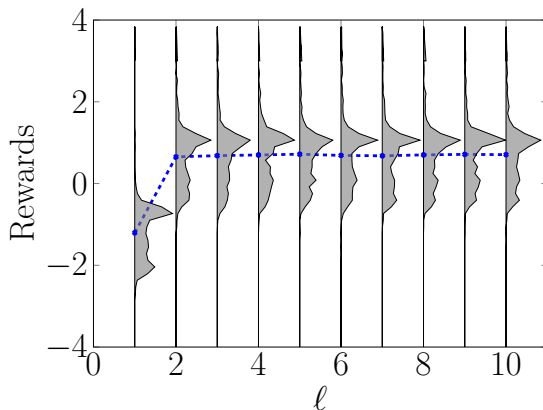


**Joint distribution even for a linear-Gaussian problem
is generally not Gaussian!**

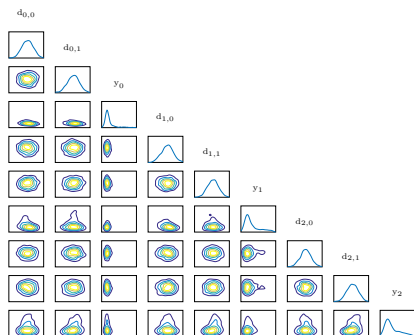
1D source inversion problem: effects of ℓ



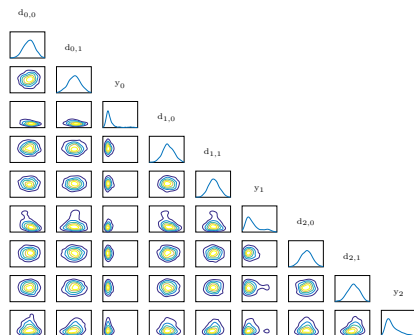
1D source inversion problem: effects of ℓ



2D source inversion problem: joint map (partial)



original samples



samples from map

Joint map does a good job in capturing some highly non-Gaussian, possibly multi-modal, behavior!

Overview

Challenges:

High number of uncertain model parameters (stochastic dimension), requiring many expensive flow solves

Would like to...

- use few runs
- incorporate less expensive simulations from low-fidelity models

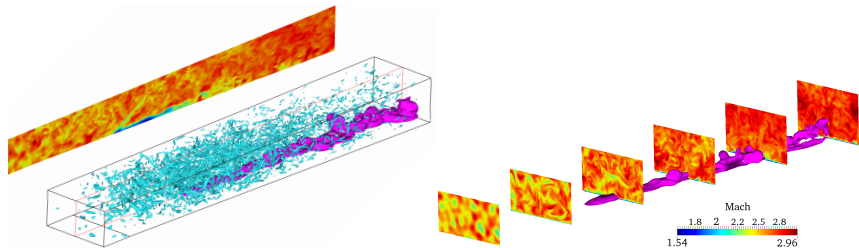
Objectives:

- identify influential uncertain parameters via global sensitivity analysis (GSA)
- characterize uncertainty due to model error resulted from using low-fidelity models

RAPTOR solver

RAPTOR: code by Oefelein *et al.* at Sandia [\[Oefelein 06\]](#)

- Fully-coupled, compressible conservation equations
- High Reynolds number, high-pressure, wide range of Mach number
- Real-fluid equation of state
- Detailed thermodynamics, transport and chemistry
- Non-dissipative, discretely conservative, staggered finite-volume
- Complex geometry treatment



Fuel component—purple

Turbulence—blue

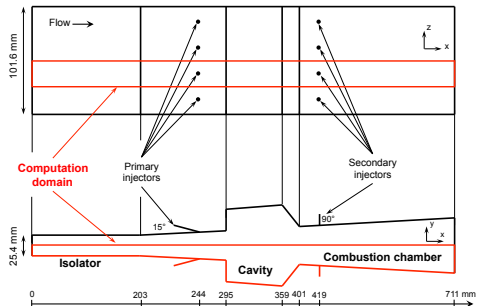
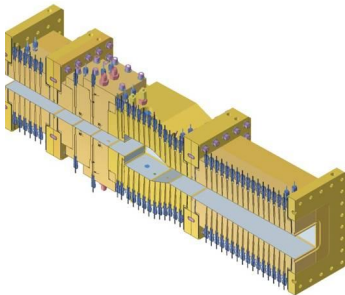
Mach number—cutting planes

Computational domain of initial unit problem

Final: simulation of full combustor domain, match experimental setup
—HIFiRE Direct Connect Rig (HDCR) [bottom-left figure]

Initial unit problem: primary injection section

- omit cavity
- no combustion
- focus on interaction of fuel jet and supersonic air crossflow

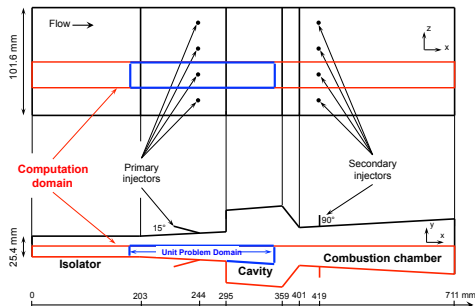
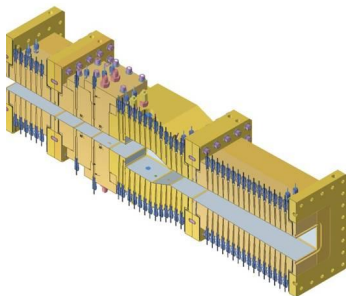


Computational domain of initial unit problem

Final: simulation of full combustor domain, match experimental setup
—HIFiRE Direct Connect Rig (HDCR) [bottom-left figure]

Initial unit problem: primary injection section

- omit cavity
- no combustion
- focus on interaction of fuel jet and supersonic air crossflow



Embedded representation of model error

Motivation: different models are available

⇒ to use them, need to quantify the error due to model structure

Traditional additive form: [Kennedy 01]

$$q_i(s) = \rho_i f_i(s, \lambda) + \delta_i(s)$$

- Flexible for fitting model discrepancy
- Predictions do not obey governing equations
- Difficult to distinguish uncertainty contributions between model error and measurement noise
- $\delta_i(s)$ not transferable for prediction of Qols outside calibration set

Embedding approach: [Sargsyan 15]

$$q_i(s) = f_i(s, \lambda + \delta_i(s, \alpha_i, \xi_i))$$

⇒ physically-meaningful predictions that satisfy governing equations

Bayesian calibration of model error term

Represent the discrepancy term δ using a PCE:

$$\lambda + \delta(\alpha, \xi) = \lambda + \sum_{\beta \neq 0} \alpha_{\beta} \Psi_{\beta}(\xi)$$

Calibrate by inferring all parameters $\tilde{\alpha} \equiv (\lambda, \alpha)$ via

Bayesian inference: $\overbrace{p(\tilde{\alpha}|D)}^{\text{posterior}} \propto \overbrace{p(D|\tilde{\alpha})}^{\text{likelihood}} \overbrace{p(\tilde{\alpha})}^{\text{prior}}$

Posterior explored via adaptive Markov chain Monte Carlo (MCMC)

Through use of surrogate model and likelihood approximation, we can **attribute predictive variance** to different sources:

$$\text{Var}[q_k] = \underbrace{\mathbb{E}_{\tilde{\alpha}} [\sigma_k^2(\lambda, \alpha)]}_{\text{model error}} + \underbrace{\text{Var}_{\tilde{\alpha}} [\mu_k(\lambda, \alpha)]}_{\text{posterior uncertainty}} + \underbrace{\sigma_{k,\text{LOO}}^2}_{\text{surrogate error}}$$

Likelihood approximation, and breakdown of variance

MCMC requires likelihood evaluation $p(D|\tilde{\alpha})$, no analytical form

Enable tractable likelihood evaluation via two approximations:

1. Polynomial surrogate for Qols, built using regression

$$q_k = f_k(\lambda + \delta(\alpha, \xi)) = \hat{f}_k(\lambda + \delta(\alpha, \xi)) + \epsilon_k$$

2. Gauss-marginal approximation to likelihood form

$$p(D|\tilde{\alpha}) \approx L_G(\tilde{\alpha}) = \frac{1}{(2\pi)^{\frac{N}{2}}} \prod_{k=1}^N \frac{1}{\sigma_k(\tilde{\alpha})} \exp \left[-\frac{(\mu_i(\tilde{\alpha}) - g_k)^2}{2\sigma_k^2(\tilde{\alpha})} \right]$$

$$\mu_k(\tilde{\alpha}) \approx \hat{f}_{k,0}(\tilde{\alpha}) \quad \text{and} \quad \sigma_k^2(\tilde{\alpha}) \approx \sum_{\beta \neq 0} \hat{f}_{k,\beta}^2(\tilde{\alpha})$$

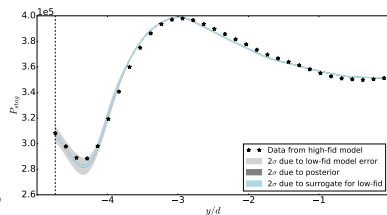
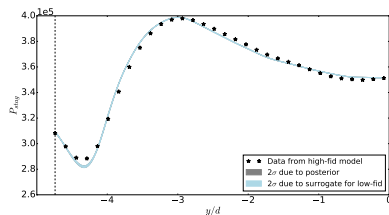
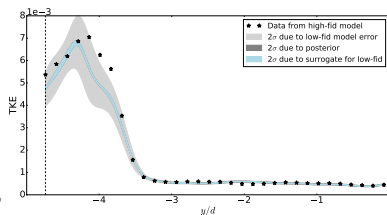
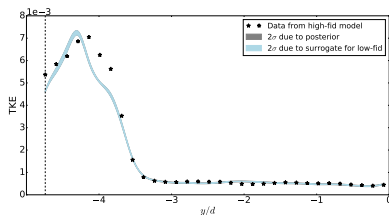
Also **enables attribution of predictive variance** to different sources:

$$\text{Var}[q_k] = \underbrace{\mathbb{E}_{\tilde{\alpha}} [\sigma_k^2(\lambda, \alpha)]}_{\text{model error}} + \underbrace{\text{Var}_{\tilde{\alpha}} [\mu_k(\lambda, \alpha)]}_{\text{posterior uncertainty}} + \underbrace{\sigma_{k,\text{LOO}}^2}_{\text{surrogate error}}$$

Dynamic-vs-Static Smagorinsky turbulence model

Calibrate static Smagorinsky model with dynamic treatment simulations

- Calibrate using TKE profile, $\lambda = C_R$



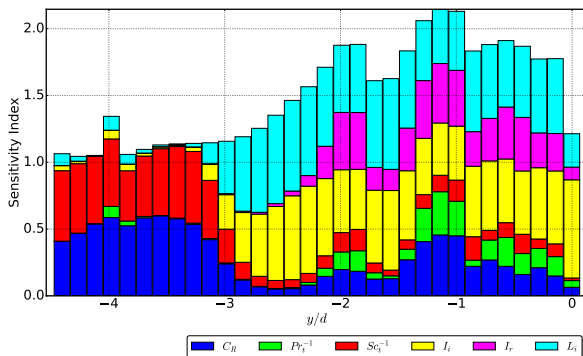
No model error treatment

Embedded model error treatment

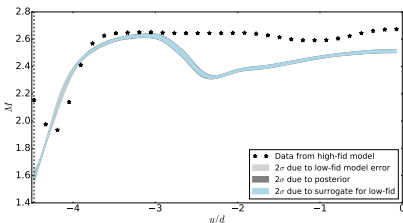
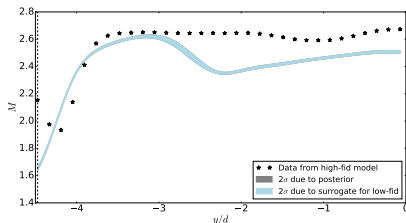
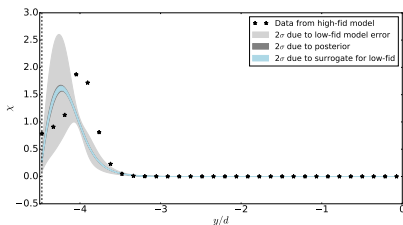
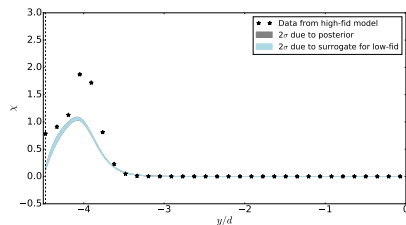
2D-vs-3D: choice of embedding parameters

Calibrate 2D model using 3D model simulations (using χ profile)

- $\lambda = (C_R, Pr_t^{-1}, Sc_t^{-1}, I_i, I_r, L_i)$
- Choose to embed δ (1st-order PCE) in C_R and Sc_t^{-1} based on GSA



2D-vs-3D: predictive quantities



No model error treatment

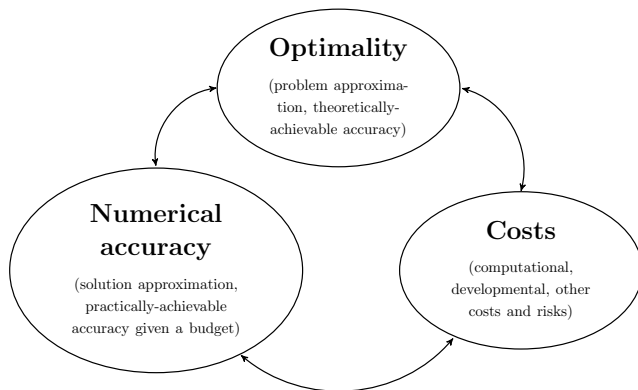
Embedded model error treatment

Ideas for future work

- **Goal-oriented adaptive numerical methods:** automated refinements of component numerical approximations; quantifiable and meaningful error bounds
- **UQ for high-dimensional and expensive models:** model reduction and surrogate modeling, dimension reduction, sparse representations
- **Multi-model management in experimental design and inference:** leverage and combine existing models of different fidelities and resolutions, improve through selective experimental design and data acquisition
- **Model error with automated learning and improvement:** quantify and learn model error through experimental design, improve models through adding sub-components while preserving governing equations and physical principles
- **Practical experimental design formulation choices:** combine low-cost sub-optimal design frameworks via value of feedback and coordination

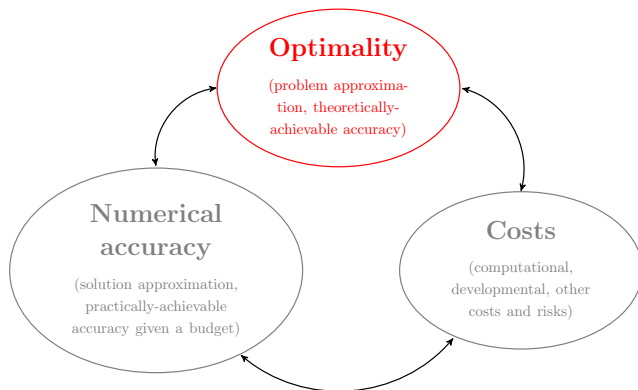
Future work: framework and methodology choices

What experimental design framework and methods to use?

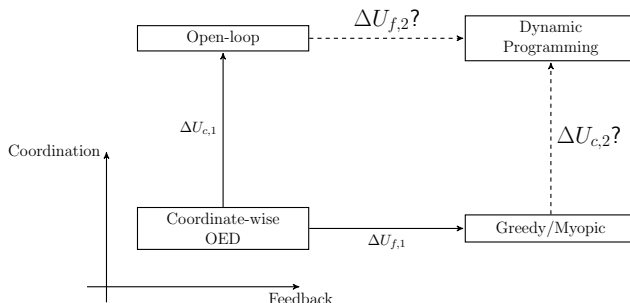


Future work: framework and methodology choices

What experimental design framework and methods to use?



Future work: value of feedback and coordination



Estimate “value of feedback” and “value of coordination”: $\Delta U_{c,2}$, $\Delta U_{f,2}$

- Heuristics: use $\Delta U_{c,1}$ and $\Delta U_{f,1}$ as guesses (cheaper to obtain)
- Theoretical: stochastic bounds $U_{\text{openloop}}, U_{\text{greedy}} \leq U^* \leq U_{\text{hindsight}}$

Ultimately: find appropriate degrees and combinations of feedback and coordination for different subsets of experiments

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