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# *Optimization and Uncertainty Quantification*

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# Outline

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- Inverse UQ – Bayesian methods
  - Statistical Inverse Problem
  - Optimal Experimental Design
- Forward UQ with Polynomial Chaos Expansions
  - Optimal Sampling
  - Compressive Sensing
  - PC Extrema Estimation

## 3 UQ Utility in Optimization

- Optimization under Uncertainty

# Introduction

- Explore connections between Optimization and Uncertainty Quantification
- Optimization problems in UQ
  - Inverse UQ
    - Bayesian methods
    - Statistical inverse problem
    - Experimental design
  - Forward UQ
    - Polynomial Chaos (PC) methods
    - Sampling ... quadrature, response surface fitting
- UQ problems in optimization
  - Forward model surrogate construction
  - Estimation of moments/probabilities in optimization under uncertainty

# Statistical Inverse Problems

- Estimation of model parameters/inputs given (noisy) data on model output observables
  - with quantified uncertainty in inferred parameters
- Conventional deterministic context:
  - Model  $y_m = f(x; \lambda)$ , data  $y$
  - Least squares fitting, minimizing residual  $\|f(x; \lambda) - y\|$
  - Regularization with suitable norms
  - End result is  $x_{\text{BestFit}}$
- Statistical Bayesian context:
  - Use Bayes rule to infer parameter  $\lambda$ 
    - Combine prior information with learning from data
  - Information on  $\lambda$  is in terms of a posterior density conditioned on the data  $p(\lambda|y)$

# Bayes formula for Parameter Inference

- Data Model (fit model + noise model):  $y = f(\lambda) * g(\epsilon)$
- Bayes Formula:

$$p(\lambda, y) = p(\lambda|y)p(y) = p(y|\lambda)p(\lambda)$$

$$\underset{\text{Posterior}}{p(\lambda|y)} = \frac{\overset{\text{Likelihood}}{p(y|\lambda)} \overset{\text{Prior}}{p(\lambda)}}{\underset{\text{Evidence}}{p(y)}}$$

- Prior: knowledge of  $\lambda$  prior to data
- Likelihood: forward model and measurement noise
- Posterior: combines information from prior and data
- Evidence: normalizing constant for present context

# Exploring the Posterior

- Given any sample  $\lambda$ , the un-normalized posterior probability can be easily computed

$$p(\lambda|y) \propto p(y|\lambda)p(\lambda)$$

- Explore posterior w/ Markov Chain Monte Carlo (MCMC)
  - Metropolis-Hastings algorithm:
    - Random walk with proposal PDF & rejection rules
  - Computationally intensive,  $\mathcal{O}(10^5)$  samples
  - Each sample: evaluation of the forward model
    - Surrogate models
- Evaluate moments/marginals from the MCMC statistics

# Optimization and MCMC

- MCMC needs to get enough good  $\lambda$ -samples to describe the posterior density well
- Frequently, the focus is on a given mode/peak
  - although generally multimodal
- In order to get good samples of a particular peak, the random walk needs to be directed to the vicinity of the peak as efficiently as possible
- The structure of the proposal distribution and the random walk algorithm are crucial
- Generally, this is about climbing the posterior density towards its peak at the Maximum A-Posteriori (MAP) parameter value, employing a random walk
- Gradient & Hessian information is useful in this regard

# Optimal Experimental Design – Stochastic Optimization

- Setup:
  - Choose experimental design  $x$
  - Collect data  $y$  to estimate parameters  $\theta$
- Challenge:
  - Choose an **optimal** design  $x^*$  that maximizes the **expected information gain** from the experiment
- Bayesian formulation:

$$p(\theta|y, x) = p(y|\theta, x)p(\theta|x) / p(y|x)$$

$$D(y, x) = D_{\text{KL}}(p(\cdot|y, x) \| p(\cdot|x)) \equiv \int p(\theta|y, x) \ln \frac{p(\theta|y, x)}{p(\theta|x)} d\theta$$

$$U(x) = \mathbf{E}_{p(\cdot|x)}[D(y, x)] \equiv \int D(y, x)p(y|x)dy$$

$$x^* = \underset{x \in \mathcal{D}}{\operatorname{argmax}} U(x)$$

- A stochastic optimization problem
  - noisy random-sampling estimation of integrals for  $U(x)$



# Probabilistic Forward UQ & Polynomial Chaos Representation of Random Variables

With  $y = f(x)$ ,  $x$  a random variable, estimate the RV  $y$

- Can describe a RV in terms of its
  - density, moments, characteristic function, or
  - as a function on a probability space
- Constraining the analysis to RVs with finite variance
  - ⇒ Represent RV as a spectral expansion in terms of orthogonal functions of standard RVs
    - Polynomial Chaos Expansion
- Enables the use of available functional analysis methods for forward UQ

# Polynomial Chaos Expansion (PCE)

- Model uncertain quantities as random variables (RVs)
- Given a *germ*  $\xi(\omega) = \{\xi_1, \dots, \xi_n\}$  – a set of *i.i.d.* RVs
  - where  $p(\xi)$  is uniquely determined by its moments

Any RV in  $L^2(\Omega, \mathfrak{S}(\xi), P)$  can be written as a PCE:

$$u(\mathbf{x}, t, \omega) = f(\mathbf{x}, t, \xi) \simeq \sum_{k=0}^P u_k(\mathbf{x}, t) \Psi_k(\xi(\omega))$$

- $u_k(\mathbf{x}, t)$  are mode strengths
- $\Psi_k()$  are multivariate functions orthogonal w.r.t.  $p(\xi)$

With dimension  $n$  and order  $p$ :  $P + 1 = \frac{(n + p)!}{n!p!}$

# Orthogonality

By construction, the functions  $\Psi_k()$  are orthogonal with respect to the density of  $\xi$

$$u_k(\mathbf{x}, t) = \frac{\langle u \Psi_k \rangle}{\langle \Psi_k^2 \rangle} = \frac{1}{\langle \Psi_k^2 \rangle} \int u(\mathbf{x}, t; \lambda(\xi)) \Psi_k(\xi) p_\xi(\xi) d\xi$$

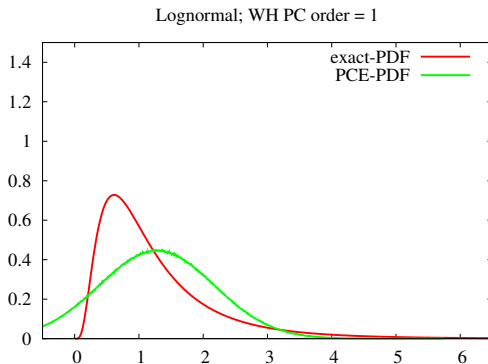
## Examples:

- Hermite polynomials with Gaussian basis
- Legendre polynomials with Uniform basis, ...
- Global versus Local PC methods
  - Adaptive domain decomposition of the support of  $\xi$

# PC Illustration: WH PCE for a Lognormal RV

- Wiener-Hermite PCE constructed for a Lognormal RV
- PCE-sampled PDF superposed on true PDF
- Order = 1

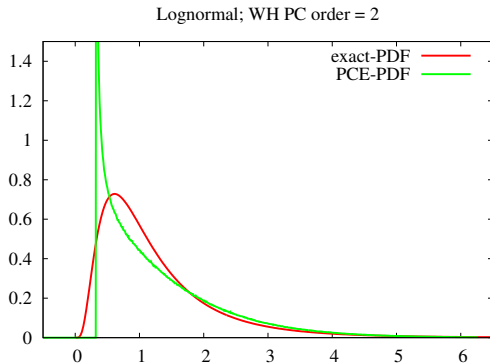
$$\begin{aligned}u &= \sum_{k=0}^P u_k \Psi_k(\xi) \\ &= u_0 + u_1 \xi\end{aligned}$$



# PC Illustration: WH PCE for a Lognormal RV

- Wiener-Hermite PCE constructed for a Lognormal RV
- PCE-sampled PDF superposed on true PDF
- Order = 2

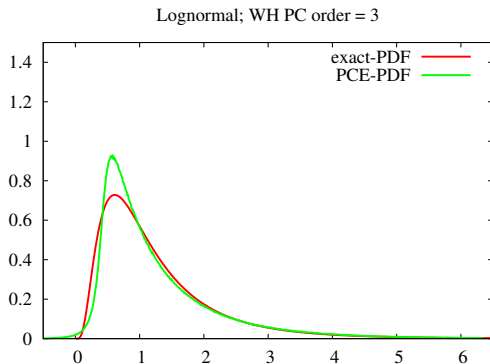
$$\begin{aligned}u &= \sum_{k=0}^P u_k \Psi_k(\xi) \\ &= u_0 + u_1 \xi + u_2 (\xi^2 - 1)\end{aligned}$$



# PC Illustration: WH PCE for a Lognormal RV

- Wiener-Hermite PCE constructed for a Lognormal RV
- PCE-sampled PDF superposed on true PDF
- Order = 3

$$\begin{aligned}u &= \sum_{k=0}^P u_k \Psi_k(\xi) \\ &= u_0 + u_1 \xi + u_2 (\xi^2 - 1) + u_3 (\xi^3 - 3\xi)\end{aligned}$$

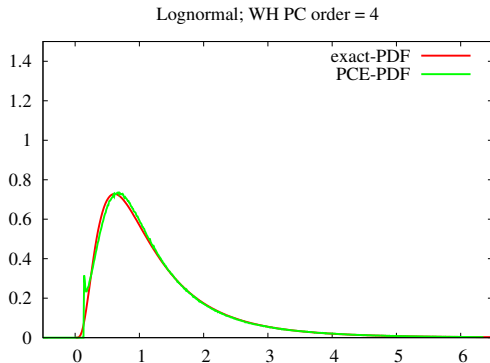


# PC Illustration: WH PCE for a Lognormal RV

- Wiener-Hermite PCE constructed for a Lognormal RV
- PCE-sampled PDF superposed on true PDF
- Order = 4

$$u = \sum_{k=0}^P u_k \Psi_k(\xi)$$

$$= u_0 + u_1 \xi + u_2 (\xi^2 - 1) + u_3 (\xi^3 - 3\xi) + u_4 (\xi^4 - 6\xi^2 + 3)$$

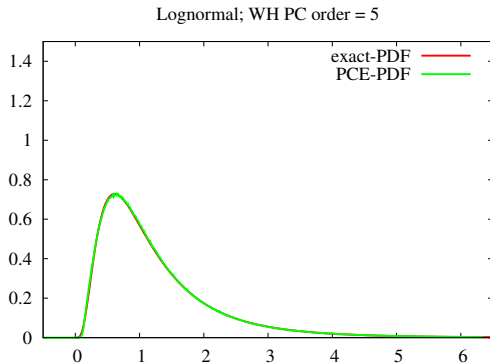


# PC Illustration: WH PCE for a Lognormal RV

- Wiener-Hermite PCE constructed for a Lognormal RV
- PCE-sampled PDF superposed on true PDF
- Order = 5

$$u = \sum_{k=0}^P u_k \Psi_k(\xi)$$

$$\begin{aligned} &= u_0 + u_1 \xi + u_2(\xi^2 - 1) + u_3(\xi^3 - 3\xi) + u_4(\xi^4 - 6\xi^2 + 3) \\ &\quad + u_5(\xi^5 - 10\xi^3 + 15\xi) \end{aligned}$$





# Essential Use of PC in UQ

## Strategy:

- Represent model parameters/solution as random variables
- Construct PCEs for uncertain parameters
- Evaluate PCEs for model outputs

## Advantages:

- Computational efficiency
- Utility
  - Moments:  $E(u) = u_0$ ,  $\text{var}(u) = \sum_{k=1}^P u_k^2 \langle \Psi_k^2 \rangle$ ,  $\dots$
  - Global Sensitivities – fractional variances, Sobol' indices
  - Surrogate for forward model

## Requirement:

- RVs in  $L^2$ , i.e. with finite variance, on  $(\Omega, \mathfrak{G}(\xi), P)$

# Intrusive PC UQ: A direct *non-sampling* method

- Given model equations:  $\mathcal{M}(u(\mathbf{x}, t); \lambda) = 0$
- Express uncertain parameters/variables using PCEs

$$u = \sum_{k=0}^P u_k \Psi_k; \quad \lambda = \sum_{k=0}^P \lambda_k \Psi_k$$

- Substitute in model equations; apply Galerkin projection
- New set of equations:  $\mathcal{G}(U(\mathbf{x}, t), \Lambda) = 0$ 
  - with  $U = [u_0, \dots, u_P]^T$ ,  $\Lambda = [\lambda_0, \dots, \lambda_P]^T$
- Solving this deterministic system once provides the full specification of uncertain model outputs

# Non-intrusive PC UQ

- *Sampling*-based
- Relies on black-box utilization of the computational model
- Evaluate projection integrals *numerically*
- For any quantity of interest  $\phi(\mathbf{x}, t; \lambda) = \sum_{k=0}^P \phi_k(\mathbf{x}, t) \Psi_k(\boldsymbol{\xi})$

$$\phi_k(\mathbf{x}, t) = \frac{1}{\langle \Psi_k^2 \rangle} \int \phi(\mathbf{x}, t; \lambda(\boldsymbol{\xi})) \Psi_k(\boldsymbol{\xi}) p_{\boldsymbol{\xi}}(\boldsymbol{\xi}) d\boldsymbol{\xi}, \quad k = 0, \dots, P$$

- Integrals can be evaluated using
  - A variety of (Quasi) Monte Carlo methods
    - Slow convergence;  $\sim$  indep. of dimensionality
  - Quadrature/Sparse-Quadrature methods
    - Fast convergence; depends on dimensionality

# Optimal Sparse Quadrature – forward UQ

Integration problem, with  $x \in D \subset \mathbb{R}^N$ :

$$I = \int_D f(x) \, dx \approx \hat{I} = \sum_{j=1}^M w_j f(x_j)$$

- Optimization problem
  - Minimize number of (sparse) quadrature points,  $M$
  - Optimize their locations and weights,  $\{w_j, x_j\}_{j=1}^M$
  - For a requisite integration accuracy,  $\|I - \hat{I}\| < \epsilon$
- Regular domains – hypercubes Sinsbeck & Nowak, IJUQ 2015
- Arbitrary domains Ryu & Boyd, Found. Comput. Math. 2015
  - By construction
  - $f(x)$  model failure – unrealistic conditions
  - $f(x)$  code failure – numerical stability / machine faults

# Greedy Sampling Algorithms

- Find the optimal location for the next evaluation  $f(x_k)$ 
  - given existing samples  $x_j, j = 1, \dots, k-1$
- Maximize expected reduction in error
  - given one additional sample/batch-of-samples
- Adaptive multilevel/hierarchical sparse quadrature
  - Selective evaluation of corner samples
- Non-isotropic sparse quadrature
  - Dimension-adaptive sampling

# Interpolant/Regression Surrogates – Optimal Design

- PCE or other surrogate functions built via
  - Interpolation
  - Least-squares regression – noisy forward models
    - intrinsic noise
    - discretization errors
    - sample-averaging noise
    - sparse samples
- The optimal set of points – design
  - Minimize oscillations – particularly in Hi-D
  - Minimize cross-validation fit errors
- Recent work (Narayan *et al.*)
  - Leja sequences for optimal interpolation in Hi-D
  - Optimal random sampling of design points for weighted least-squares regression

# PC coefficients via sparse regression

PCE:

$$y = f(x) \simeq \sum_{k=0}^{K-1} c_k \Psi_k(x)$$

with  $x \in \mathbb{R}^n$ ,  $\Psi_k$  max order  $p$ , and  $K = (p+n)!/p!/n!$

- $N$  samples  $(x_1, y_1), \dots, (x_N, y_N)$
- Estimate  $K$  terms  $c_0, \dots, c_{K-1}$ , s.t.

$$\min ||\mathbf{y} - \mathbf{A}\mathbf{c}||_2^2$$

where  $\mathbf{y} \in \mathbb{R}^N$ ,  $\mathbf{c} \in \mathbb{R}^K$ ,  $\mathbf{A}_{ik} = \Psi_k(x_i)$ ,  $\mathbf{A} \in \mathbb{R}^{N \times K}$

With  $N \ll K \Rightarrow$  under-determined

- Need some form of regularization

# Regularization – Compressive Sensing (CS)

- $\ell_2$ -norm — Tikhonov regularization; Ridge regression:

$$\min \{ \|\mathbf{y} - \mathbf{A}\mathbf{c}\|_2^2 + \|\mathbf{c}\|_2^2 \}$$

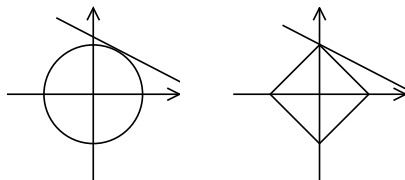
- $\ell_1$ -norm — Compressive Sensing; LASSO; basis pursuit

$$\min \{ \|\mathbf{y} - \mathbf{A}\mathbf{c}\|_2^2 + \|\mathbf{c}\|_1 \}$$

$$\min \{ \|\mathbf{y} - \mathbf{A}\mathbf{c}\|_2^2 \} \quad \text{subject to } \|\mathbf{c}\|_1 \leq \epsilon$$

$$\min \{ \|\mathbf{c}\|_1 \} \quad \text{subject to } \|\mathbf{y} - \mathbf{A}\mathbf{c}\|_2^2 \leq \epsilon$$

⇒ discovery of sparse signals





# Bayesian Regression

- Bayes formula

$$p(\mathbf{c}|D) \propto p(D|\mathbf{c})\pi(\mathbf{c})$$

- Bayesian regression: prior as a regularizer, *e.g.*

- Log Likelihood  $\Leftrightarrow \|\mathbf{y} - \mathbf{A}\mathbf{c}\|_2^2$

- Log Prior  $\Leftrightarrow \|\mathbf{c}\|_p^p$

- Laplace sparsity priors  $\pi(c_k|\alpha) = \frac{1}{2\alpha}e^{-|c_k|/\alpha}$

- LASSO (Tibshirani 1996) ... formally:

$$\min \{ \|\mathbf{y} - \mathbf{A}\mathbf{c}\|_2^2 + \lambda \|\mathbf{c}\|_1 \}$$

Solution  $\sim$  the posterior mode of  $\mathbf{c}$  in the Bayesian model

$$\mathbf{y} \sim \mathcal{N}(\mathbf{A}\mathbf{c}, I_N), \quad c_k \sim \frac{1}{2\alpha}e^{-|c_k|/\alpha}$$

- Bayesian LASSO (Park & Casella 2008)
- Bayesian compressive sensing (Ji 2008)

# PC Extrema Estimation – Global Optimization

- Often it is important to establish PCE positivity

$$\begin{aligned}u_{\text{PC}}(\boldsymbol{\xi}) &\equiv \sum_{k=1}^P u_k \Psi_k(\boldsymbol{\xi}) \\ u_{\min} &\equiv \min_{\boldsymbol{\xi} \in \Xi} u_{\text{PC}}(\boldsymbol{\xi}) > 0\end{aligned}$$

- A global optimization problem
- Nonlinear
- High-dimensional

# UQ Utility in Optimization

## Deterministic optimization problems

- Model surrogates constructed using forward UQ
  - A wide range of UQ methods for efficient surrogate construction in hi-D
  - Surrogates can be built over deterministic spaces employing uniform RVs
  - Readily available surrogate gradient/hessian information

## Optimization under uncertainty

- Stochastic optimization
- Distributionally Robust optimization
- Robust optimization

# Stochastic Optimization

e.g. stochastic objective

$$\begin{array}{ll} \underset{x \in \mathcal{X}}{\text{minimize}} & \mathbf{E}_{p_{\Theta}}[u(x, \theta)] \\ \text{subject to} & c(x) > 0 \end{array}$$

or penalize variability, and chance constraint

$$\begin{array}{ll} \underset{x \in \mathcal{X}}{\text{minimize}} & \mathbf{E}_{p_{\Theta}}[u(x, \theta) - \gamma \text{Var}(u(x, \theta))] \\ \text{subject to} & \mathbf{P}[c(x, \theta) > 0] > 1 - \alpha \end{array}$$

or minimize conditional value at risk (CVaR)

$$\begin{array}{ll} \underset{x \in \mathcal{X}}{\text{minimize}} & \mathbf{E}_{p_{\Theta}}[u(x, \theta) | u(x, \theta) > u_0] \\ \text{subject to} & \mathbf{P}[u(x, \theta) < u_0] = 1 - \alpha \end{array}$$

# Stochastic Optimization – SAA

$$\underset{x \in \mathcal{X}}{\text{minimize}} \quad \{ U(x) \equiv \mathbf{E}_{p_{\Theta}}[u(x, \theta)] \}$$

where  $\Theta \sim p_{\Theta}(\theta)$ ,  $\theta \in \mathbb{R}^n$ ,  $\mathcal{X} \subset \mathbb{R}^N$

- Presumes knowledge of  $p_{\Theta}(\cdot)$
- Typically relies on sample averaged approximation (SAA)

$$U(x) \approx \hat{U}(x) := \frac{1}{K} \sum_{k=1}^K u(x, \theta_k)$$

- Accurate Monte Carlo estimation requires large  $K$
- $\hat{U}(x)$  is a noisy estimator of  $U(x)$ 
  - Gradients of  $\hat{U}(x)$  challenging to estimate

# Stochastic Optimization with PCE

$$\begin{aligned}\theta_{\text{PC}}(\xi) &= \sum_k \alpha_k \Psi_k(\xi), & u_{\text{PC}}(x, \xi) &= \sum_k u_k(x) \Psi_k(\xi) \\ \hat{U}(x) &= \mathbf{E}_{p_\xi}[u_{\text{PC}}(x, \xi)] = u_0(x)\end{aligned}$$

where

$$u_0(x) = \int u p_\xi d\xi = \sum_i w_i u(x, \theta(\xi_i))$$

is estimated using forward UQ methods

- perhaps intrusively, if  $u(x, \theta)$  is relatively simple
- otherwise non-intrusive, *e.g.* sparse quadrature

Computational efficiency relative to Monte Carlo depends on

- the dimensionality of  $\theta$
- the  $\theta$ -smoothness of  $u(x, \theta)$

# Gradients over the design space

- Estimation of  $\frac{du_0}{dx}$  requires
  - A functional representation of  $u_0(x)$  to be differentiated, or
  - A hi-resolution estimation of  $u_{PC}(x_i, \xi)$ ,  $i = 1, \dots, I_{\text{mesh}}$ , or
  - A PCE for  $\frac{du}{dx}(x, \xi)$ , and hence gradients of the obj. func.
- Alternatively, the PCE can be built over  $(x, \xi)$  [Eldred, IUQ 2011](#)

$$u_{PC}(x, \xi) = \sum_k u_k \Psi_k(x, \xi)$$

- Functional representation of  $u_{PC}(x, \xi)$  over  $x$  is built-in
- Easy access to gradients/hessians over  $x$
- But a higher dimensional forward UQ problem

# PCE in Power Grid Stochastic Optimization

## Scenario Generation – Random Field (RF) Inputs

- Power-grid optimization involves uncertainties
  - in both loads and alternative energy sources
  - The largest uncertainties are in wind and solar generation
  - Being uncertain functions of time, these are RFs
- The Karhunen-Loeve expansion (KLE) provides an optimal representation of RFs, capturing both mean & covariance

$$W(t, \omega) = \mu(t) + \sum_{i=1}^{\infty} \sqrt{\lambda_i} \eta_i(\omega) \phi_i(t)$$

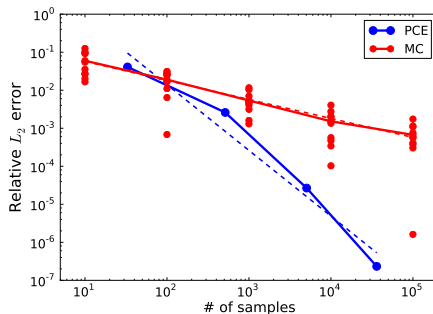
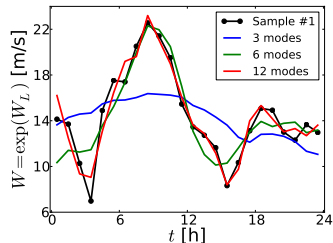
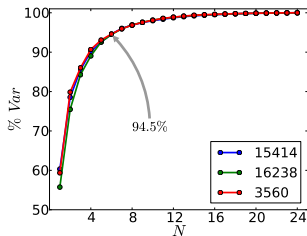
- $\mu(t)$  is the mean of  $W(t, \omega)$  at  $t$
- $\lambda_i$  and  $\phi_i(t)$  are the eigenvalues and eigenfunctions of the covariance  $C(t_1, t_2) = \langle [W(t_1, \omega) - \mu(t_1)][W(t_2, \omega) - \mu(t_2)] \rangle$
- The  $\eta_i$  are uncorrelated zero-mean unit-variance RVs



# PCE in Power Grid Stochastic Optimization

- Consider the Economic Dispatch problem
  - Given a set of generators online
  - Find optimal expected power generation schedules over the next 24 hr
  - Feasibility and operational constraints
- IEEE 118 bus system – 54 generators, 64 loads
- 3 generators replaced by wind farms
- wind data from two sites in Wyoming and one in California
- KLE  $\Rightarrow$  16-dimensional forward UQ problem
- Minimum cost  $Q(x, W(t, \omega)) \approx Q_{\text{PC}}(x, \eta(\xi)) = \sum_k q_k(x) \Psi_k(\xi)$ 
  - Estimate PC coefficients using sparse quadrature
  - Expectation  $q_0(x)$

# Scenario Generation – Random Field (RF) Inputs



# Distributionally Robust Optimization (DRO)

- Presumes imperfect knowledge of  $p_{\Theta}(\cdot)$
- Consider  $p_{\Theta} \in \mathcal{D}$ ,

$$\underset{x \in \mathcal{X}}{\text{minimize}} \quad \left\{ \max_{p_{\Theta} \in \mathcal{D}} \mathbf{E}_{p_{\Theta}}[u(x, \theta)] \right\}$$

- Implementation: define ambiguity set  $\mathcal{D}$ , e.g.
  - given presumed  $\mathcal{S} \supset \text{supp}(p_{\Theta})$  & moments of  $\Theta$ :  $(\mu_0, \Sigma_0)$ 
    - Allow uncertainty in moments [Delage & Ye, OR 2010](#)
  - given max KL-divergence between  $p_{\Theta}$  and a nominal  $p_0$   
[Hu & Hong, 2013](#)
- Utility of PC methods
  - Moment constraints accessible with PCE [Eldred, IJUQ 2011](#)
- Connections to "Optimal UQ" [Owhadi, SIAM Review 2013](#)
- Possible role for Maximum Entropy methods?

# Robust optimization

- Set-based approach – support of uncertain parameter PDF
- Protect against worse case scenario in the set
- Learn set based on samples/historical-realizations of the uncertain-parameters as RVs
- Topics:
  - Role for PDF quantiles?
  - PCE extrema
  - Data-analysis/classification for establishing set boundaries
  - PDF tail behavior – Extreme Value Theory

# The End