

SAND2014-18343PE

High-Dimensionality Challenges in Uncertainty Quantification

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ICERM Workshop on Approximation, Integration, and Optimization,
Brown University, Providence, RI
Sep. 29 – Oct. 3, 2014

Acknowledgement

B.J. Debusschere, R.D. Berry, K. Sargsyan, C. Safta,
K. Chowdhary, M. Khalil — Sandia National Laboratories, CA

R.G. Ghanem — U. South. California, Los Angeles, CA

O.M. Knio — Duke Univ., Durham, NC

O.P. Le Maître — CNRS, Paris, France

Y.M. Marzouk — Mass. Inst. of Tech., Cambridge, MA

This work was supported by:

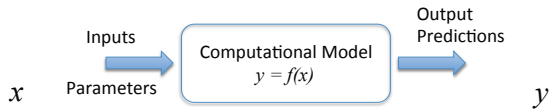
- DOE Office of Science (SC)
- DOE Office of Advanced Scientific Computing Research (ASCR), Scientific Discovery through Advanced Computing (SciDAC)
- DOE ASCR Applied Mathematics program.
- DOE Office of Basic Energy Sciences, Div. of Chem. Sci., Geosci., & Biosci.
- DOE Office of Biological & Environmental Research
- DOE/SC National Energy Research Scientific Computing Center (NERSC) under Contract No. DE-AC02-05CH11231

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Outline

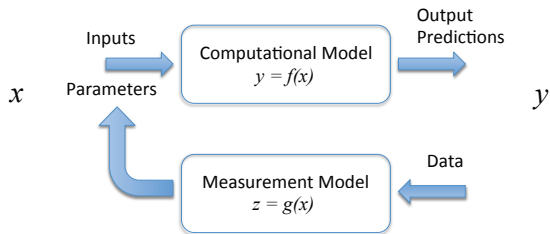
- 1 Introduction
- 2 Random Fields
- 3 Integration in Forward UQ
- 4 Sparsity
- 5 Closure

Uncertainty Quantification and Computational Science



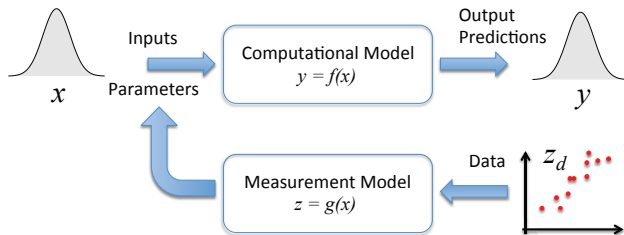
Forward problem

Uncertainty Quantification and Computational Science



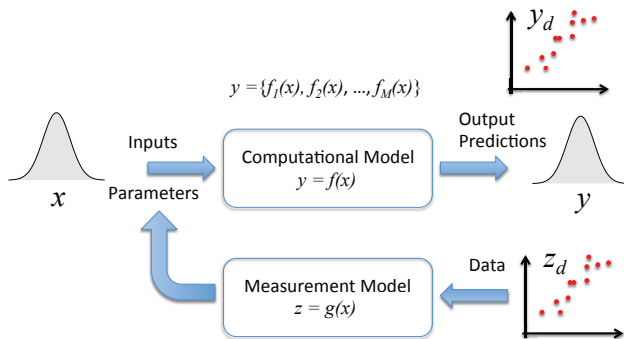
Inverse & Forward problems

Uncertainty Quantification and Computational Science



Inverse & Forward UQ

Uncertainty Quantification and Computational Science



Inverse & Forward UQ

Model validation & comparison, Hypothesis testing

Relevance of hi-dimensionality in UQ

- Large number of degrees of freedom in uncertain model inputs/outputs
 - Random fields – material properties, initial/boundary conditions
 - Model parameters – reaction rate constants, physical constants, phenomenological models
 - State variables – model outputs
- Consequences
 - Joint densities
 - Functional representations of random variables/fields
 - Global sensitivity analysis
 - Forward and Inverse UQ

Probabilistic Forward UQ

—

$$y = f(x)$$

Represent uncertain quantities using probability theory

- Random sampling, MC, QMC
 - Generate random samples $\{x^i\}_{i=1}^N$ from the PDF of $x, p(x)$
 - Bin the corresponding $\{y^i\}$ to construct $p(y)$
 - Not feasible for computationally expensive $f(x)$
 - slow convergence of MC/QMC methods
 - ⇒ very large N required for reliable estimates
- Build a cheap surrogate for $f(x)$, then use MC
 - Collocation – interpolants
 - Regression – fitting
- Galerkin methods
 - Polynomial Chaos (PC)
 - Intrusive and non-intrusive PC methods

Probabilistic Forward UQ & Polynomial Chaos Representation of Random Variables

With $y = f(x)$, x a random variable, estimate the RV y

- Can describe a RV in terms of its
 - density, moments, characteristic function, or
 - as a function on a probability space
- Constraining the analysis to RVs with finite variance
 - ⇒ Represent RV as a spectral expansion in terms of orthogonal functions of standard RVs
 - Polynomial Chaos Expansion
- Enables the use of available functional analysis methods for forward UQ

Polynomial Chaos Expansion (PCE)

- Model uncertain quantities as random variables (RVs)
- Given a *germ* $\xi(\omega) = \{\xi_1, \dots, \xi_n\}$ – a set of *i.i.d.* RVs
 - where $p(\xi)$ is uniquely determined by its moments

Any random variable $Z \in L^2(\Omega, \mathfrak{G}(\xi), P)$ can be written as a PCE:

$$Z(\omega) = Z(\xi) \simeq \sum_{k=0}^P z_k \Psi_k(\xi(\omega))$$

- z_k are mode strengths
- $\Psi_k()$ are multivariate functions orthogonal w.r.t. $p(\xi)$

With dimension n and order p : $P + 1 = \frac{(n + p)!}{n!p!}$

Orthogonality

By construction, the functions $\Psi_k()$ are orthogonal with respect to the density of ξ

$$z_k = \frac{\langle Z\Psi_k \rangle}{\langle \Psi_k^2 \rangle} = \frac{1}{\langle \Psi_k^2 \rangle} \int Z(\lambda(\xi)) \Psi_k(\xi) p_\xi(\xi) d\xi$$

Examples:

- Hermite polynomials with Gaussian basis
- Legendre polynomials with Uniform basis, ...
- Global versus Local PC methods
 - Adaptive domain decomposition of the support of ξ

Random Fields

- A random field (RF) $Z(x, \omega)$ is a function on a product space $D \times \Omega$
 - a much more complex object than a RV
 - an RV at an $x \in D$
 - an infinite dimensional object
- However, in many physical systems, uncertain field quantities, described by RFs, have an underlying *smoothness* due to correlations
- Can be represented with a small no. of stochastic degrees of freedom
- Optimal representation – second-order statistics
 - Karhunen-Loeve expansion

Random Fields – KLE

- Karhunen-Loeve Expansion (KLE) for a RF with a continuous covariance function

$$Z(x, \omega) = \mu(x) + \sum_{i=1}^{\infty} \sqrt{\lambda_i} \zeta_i(\omega) \phi_i(x)$$

- $\mu(x)$ is the mean of $Z(x, \omega)$ at x
- λ_i and $\phi_i(x)$ are the eigenvalues and eigenfunctions of the covariance

$$C(x_1, x_2) = \langle [Z(x_1, \omega) - \mu(x_1)][Z(x_2, \omega) - \mu(x_2)] \rangle$$

- The ζ_i are uncorrelated zero-mean unit-variance RVs

$$\zeta_i(\omega) = \frac{1}{\sqrt{\lambda_i}} \int_D Z(x, \omega) \phi_i(x) dx$$

Rosenblatt transform for Multi-D RVs

- Rosenblatt transformation maps any set of jointly distributed random variables $(\zeta_1, \dots, \zeta_d)$ to uniform *i.i.d.* RVs $\eta_i \stackrel{iid}{\sim} U(0, 1)$, $i = 1, \dots, d$ (Rosenblatt, 1952).
- Relies on conditional CDFs $F_{k|k-1, \dots, 1}(\zeta_k | \zeta_{k-1}, \dots, \zeta_1)$

$$\begin{aligned}
 \eta_1 &= F_1(\zeta_1) \\
 \eta_2 &= F_{2|1}(\zeta_2 | \zeta_1) \\
 &\vdots \\
 \eta_d &= F_{d|d-1, \dots, 1}(\zeta_d | \zeta_{d-1}, \dots, \zeta_1)
 \end{aligned}$$

- A multi-D generalization of 1D CDF mapping.
- Conditional CDFs harder to evaluate in high dimensions

Rosenblatt transformation with samples – KDE

- Given samples $\{\zeta^i\}_{i=1}^N$ of the random vector $\zeta = (\zeta_1, \dots, \zeta_d)$
- Kernel Density Estimation (KDE) is useful to compute the conditional CDFs

$$\begin{aligned}
 \eta_k &= F_{k|k-1,\dots,1}(\zeta_k | \zeta_{k-1}, \dots, \zeta_1) \\
 &= \int_{-\infty}^{\zeta_k} p_{k|k-1,\dots,1}(\zeta'_k | \zeta_{k-1}, \dots, \zeta_1) d\zeta'_k \\
 &= \int_{-\infty}^{\zeta_k} \frac{p_{k,\dots,1}(\zeta'_k, \zeta_{k-1}, \dots, \zeta_1)}{p_{k-1,\dots,1}(\zeta_{k-1}, \dots, \zeta_1)} d\zeta'_k
 \end{aligned}$$

Rosenblatt transformation with samples – KDE

$$\begin{aligned}
 \eta_k &\approx \frac{1}{h\sqrt{2\pi}} \int_{-\infty}^{\zeta_k} \frac{\sum_{i=1}^N \exp\left(-\frac{(\zeta_1 - \zeta_1^i)^2 + \dots + (\zeta'_k - \zeta_k^i)^2}{2h^2}\right)}{\sum_{i=1}^N \exp\left(-\frac{(\zeta_1 - \zeta_1^i)^2 + \dots + (\zeta_{k-1} - \zeta_{k-1}^i)^2}{2h^2}\right)} d\zeta'_k \\
 &= \int_{-\infty}^{\zeta_k} \frac{\sum_{i=1}^N \exp\left(-\frac{(\zeta_1 - \zeta_1^i)^2 + \dots + (\zeta_{k-1} - \zeta_{k-1}^i)^2}{2h^2}\right) \times \frac{1}{h\sqrt{2\pi}} \exp\left(-\frac{(\zeta'_k - \zeta_k^i)^2}{2h^2}\right)}{\sum_{i=1}^N \exp\left(-\frac{(\zeta_1 - \zeta_1^i)^2 + \dots + (\zeta_{k-1} - \zeta_{k-1}^i)^2}{2h^2}\right)} d\zeta'_k \\
 &= \frac{\sum_{i=1}^N \exp\left(-\frac{(\zeta_1 - \zeta_1^i)^2 + \dots + (\zeta_{k-1} - \zeta_{k-1}^i)^2}{2h^2}\right) \times \Phi\left(\frac{\zeta_k - \zeta_k^i}{h}\right)}{\sum_{i=1}^N \exp\left(-\frac{(\zeta_1 - \zeta_1^i)^2 + \dots + (\zeta_{k-1} - \zeta_{k-1}^i)^2}{2h^2}\right)}
 \end{aligned}$$

RF KLE PCE

- Define $\xi_i \stackrel{iid}{\sim} N(0, 1)$, $i = 1, \dots, d$, where $\eta_i = \Phi(\xi_i)$
- The Rosenblatt transform is then

$$\Phi(\xi_k) = F_{k|k-1, \dots, 1}(\zeta_k | \zeta_{k-1}, \dots, \zeta_1), \quad k = 1, \dots, d$$

- Using the inverse Rosenblatt transform, we have

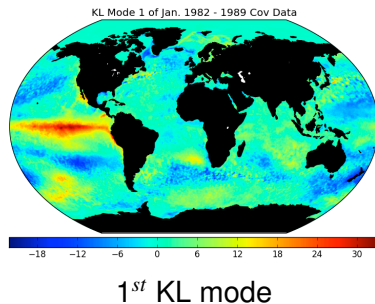
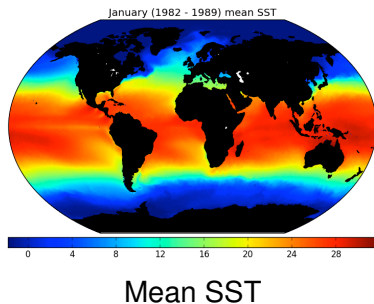
$$\zeta_k = F_{k|k-1, \dots, 1}^{-1}(\Phi(\xi_k) | \xi_{k-1}, \dots, \xi_1) = f_k(\xi_1, \dots, \xi_k)$$

which allows the construction of a WH PCE for each ζ_k and the representation of the RF $Z(x, \omega)$ as a WH PCE

$$Z(x, \omega) = \sum_{j=0}^{\infty} Z_j(x) \Psi_j(\boldsymbol{\xi})$$

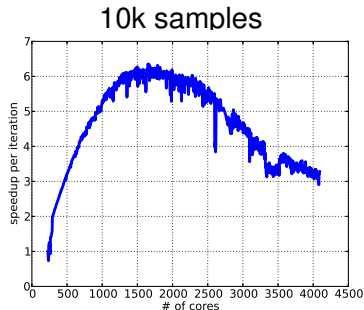
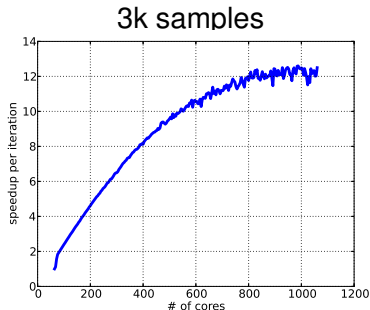
Random Fields – large scale NOAA data – SVD

- KLE for uncertain Sea Surface Temperature (SST)
 - 1/4-degree spatial resolution data
- 10^6 -dimensional random field encompassing spatial and temporal uncertainty in SST data
- SVD using Trilinos / parallelized block Krylov Schur solver



SVD on NOAA SST data – Scaling

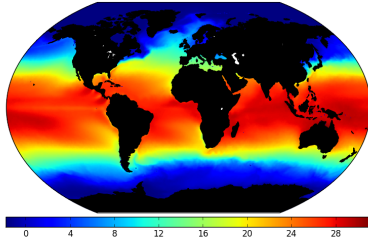
- Parallel SVD: use Trilinos eigensolver library Anasazi
- Strong scaling study – fixed problem size
 - 3-10k samples of a 10^6 -dimensional data set



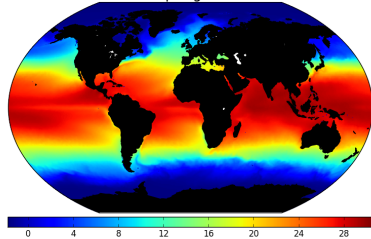
Strong-scaling speedup of the SVD solver
NERSC-Hopper

Mean SST for different seasons, 1982 - 1989

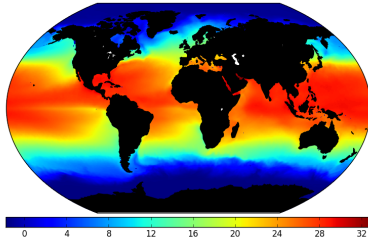
Mean for Winter 1982 - 1989



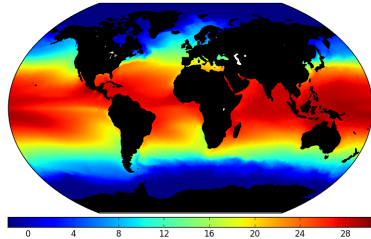
Mean for Spring 1982 - 1989



Mean for Summer 1982 - 1989

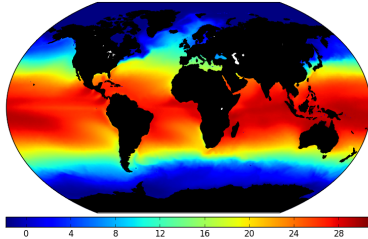


Mean for Fall 1982 - 1989

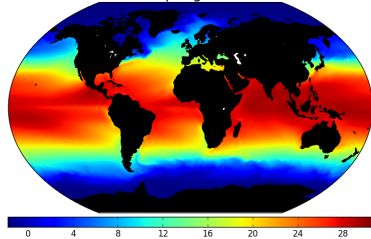


Mean SST for different seasons, 1990 - 1999

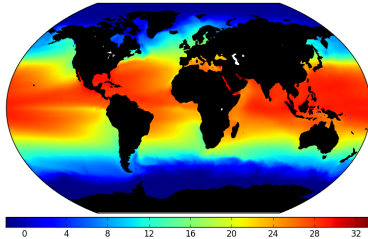
Mean for Winter 1990 - 1999



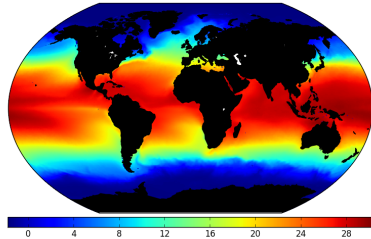
Mean for Spring 1990 - 1999



Mean for Summer 1990 - 1999

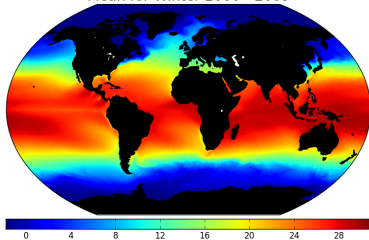


Mean for Fall 1990 - 1999

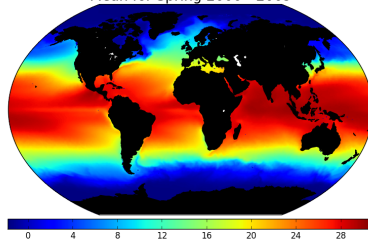


Mean SST for different seasons, 2000 - 2009

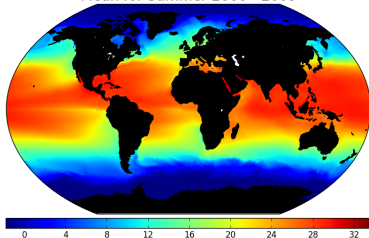
Mean for Winter 2000 - 2009



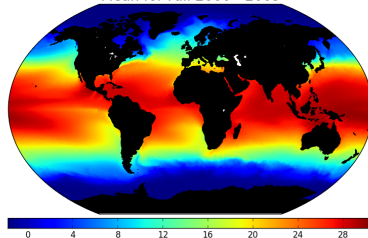
Mean for Spring 2000 - 2009



Mean for Summer 2000 - 2009

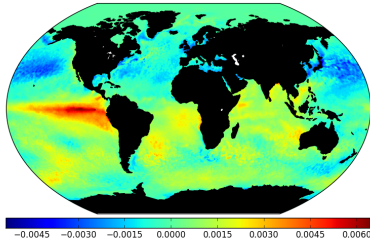


Mean for Fall 2000 - 2009

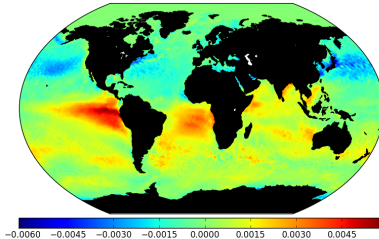


1st principal components for winter over decades

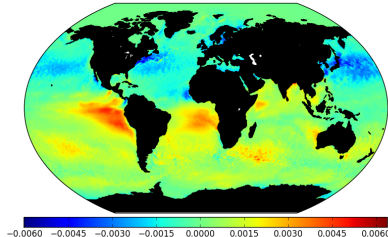
KL mode 1 for Winter 1982 - 1989



KL mode 1 for Winter 1990 - 1999

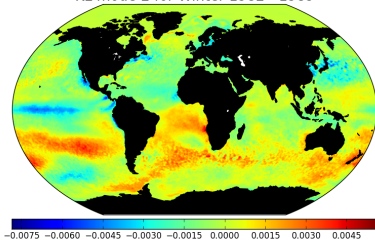


KL mode 1 for Winter 2000 - 2009

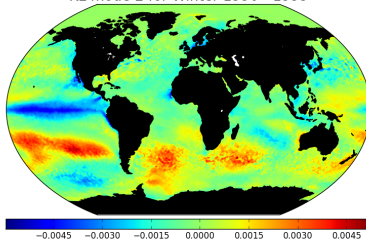


2nd principal components for winter over decades

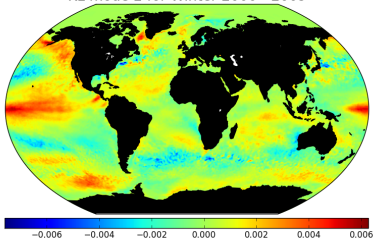
KL mode 2 for Winter 1982 - 1989



KL mode 2 for Winter 1990 - 1999

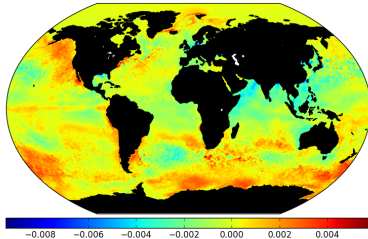


KL mode 2 for Winter 2000 - 2009

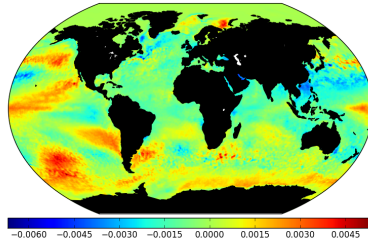


3rd principal components for winter over decades

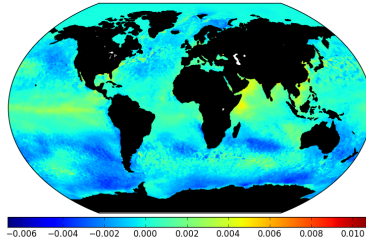
KL mode 3 for Winter 1982 - 1989



KL mode 3 for Winter 1990 - 1999

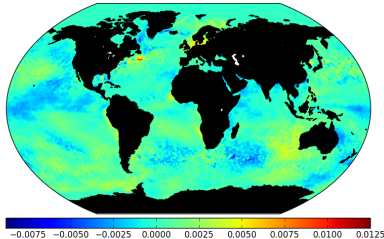


KL mode 3 for Winter 2000 - 2009

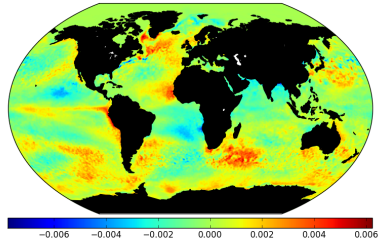


4th principal components for winter over decades

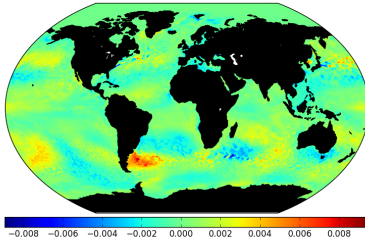
KL mode 4 for Winter 1982 - 1989



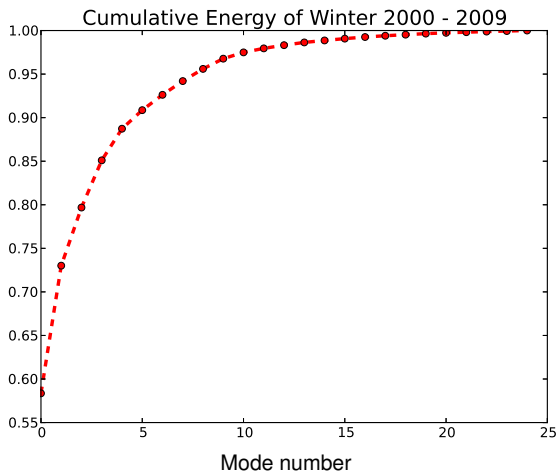
KL mode 4 for Winter 1990 - 1999



KL mode 4 for Winter 2000 - 2009



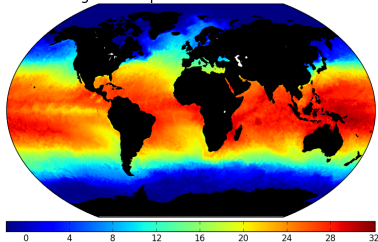
SST RF reconstruction, winter-2000s



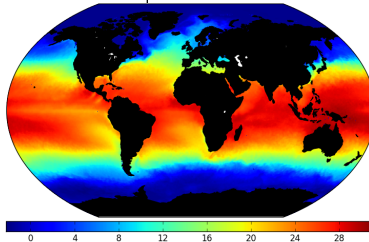
Truncated KLE with ~ 25 modes captures most of the energy

SST RF reconstruction, winter-2000s – 25 modes

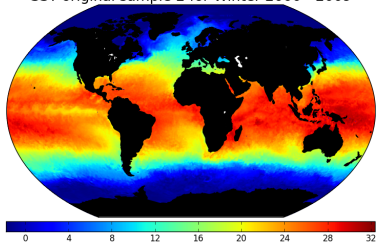
SST original sample 1 for Winter 2000 - 2009



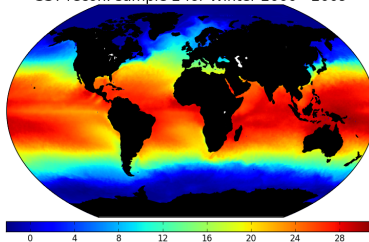
SST recon. sample 1 for Winter 2000 - 2009



SST original sample 2 for Winter 2000 - 2009



SST recon. sample 2 for Winter 2000 - 2009



Random Fields – sparse data

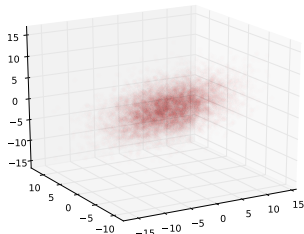
- Developed a Bayesian procedure for KLE construction given sparse data
 - Bayesian Principal Component Analysis (BPCA)/BKLE
- Address challenges arising due to
 - approximate knowledge of the covariance matrix
 - lack of positive definiteness of sample covariance matrix
- BPCA framework explores the space of orthonormal vectors, seeking those that best explain the data
 - Likelihood density $p(\Phi)$ is peaked at

$$\Phi^* = \operatorname{argmin}_{\Phi \in V_k(\mathbb{R}^d)} \sum_{i=1}^n \|x^i - P_{\Phi} x^i\|^2$$

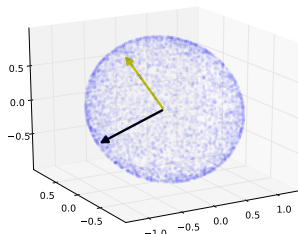
where $V_k(\mathbb{R}^d)$ is the space of k orthonormal d -dimensional vectors

- Resulting KLE incorporates uncertainty due to small number of samples

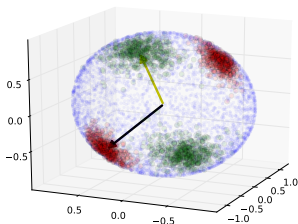
BPCA Example – Data from a 3D MVN



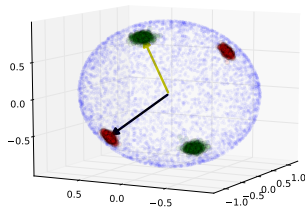
Samples of random variables from a 3D Multivariate Normal (MVN) distribution



First two principal components.
Black is the vector with maximum variance

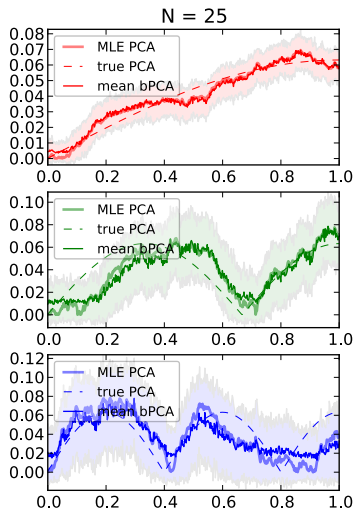


Samples from $p(\Phi)$ using 100 samples, x^i



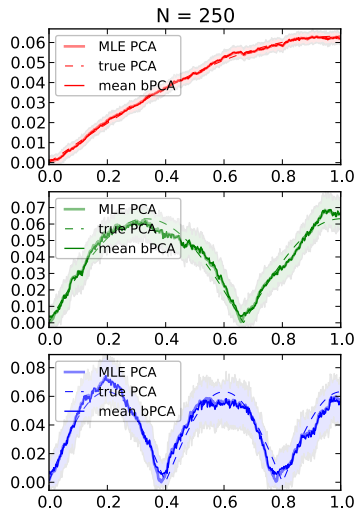
Samples from $p(\Phi)$ using 300 samples, x^i

BPCA Example – Brownian motion – 25 samples



- 500-dimensional Brownian motion stochastic process.
- Using only 25 samples, we compute samples from $p(\Phi)$ and plot the first three principal components.
- The dark solid lines represent the principal components and the shaded region represents error bars based on samples using the Bayesian PCA approach.

BPCA Example – Brownian motion – 250 samples



- Using 250 samples
- Modes are evaluated with improved accuracy
- Lower uncertainty

Essential Use of PC in UQ

Strategy:

- Represent model parameters/solution as random variables
- Construct PCEs for uncertain parameters
- Evaluate PCEs for model outputs

Advantages:

- Computational efficiency
- Utility
 - Moments: $E(u) = u_0$, $\text{var}(u) = \sum_{k=1}^P u_k^2 \langle \Psi_k^2 \rangle$, $\dots$
 - Global Sensitivities – fractional variances, Sobol' indices
 - Surrogate for forward model

Requirement:

- RVs in L^2 , *i.e.* with finite variance, on $(\Omega, \mathfrak{G}(\xi), P)$

Intrusive PC UQ: A direct *non-sampling* method

- Given model equations: $\mathcal{M}(u(\mathbf{x}, t); \lambda) = 0$
- Express uncertain parameters/variables using PCEs

$$u = \sum_{k=0}^P u_k \Psi_k; \quad \lambda = \sum_{k=0}^P \lambda_k \Psi_k$$

- Substitute in model equations; apply Galerkin projection
- New set of equations: $\mathcal{G}(U(\mathbf{x}, t), \Lambda) = 0$
 - with $U = [u_0, \dots, u_P]^T$, $\Lambda = [\lambda_0, \dots, \lambda_P]^T$
- Solving this deterministic system once provides the full specification of uncertain model outputs

Non-intrusive PC UQ

- *Sampling*-based
- Relies on black-box utilization of the computational model
- Evaluate projection integrals *numerically*
- For any quantity of interest $\phi(\mathbf{x}, t; \lambda) = \sum_{k=0}^P \phi_k(\mathbf{x}, t) \Psi_k(\boldsymbol{\xi})$

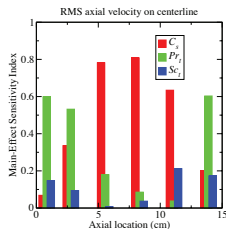
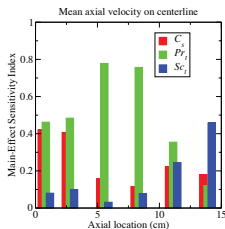
$$\phi_k(\mathbf{x}, t) = \frac{1}{\langle \Psi_k^2 \rangle} \int \phi(\mathbf{x}, t; \lambda(\boldsymbol{\xi})) \Psi_k(\boldsymbol{\xi}) p_{\boldsymbol{\xi}}(\boldsymbol{\xi}) d\boldsymbol{\xi}, \quad k = 0, \dots, P$$

- Integrals can be evaluated using
 - A variety of (Quasi) Monte Carlo methods
 - Slow convergence; $\sim$ indep. of dimensionality
 - Quadrature/Sparse-Quadrature methods
 - Fast convergence; depends on dimensionality

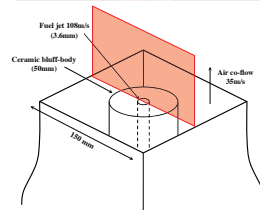
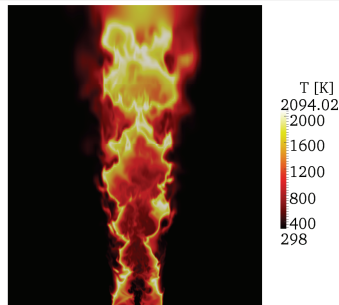
UQ in LES computations: turbulent bluff-body flame

with M. Khalil, G. Lacaze, & J. Oefelein, Sandia Nat. Labs

- $\text{CH}_4\text{-H}_2$ jet, air coflow, 3D flow
- $\text{Re}=9500$, LES subgrid modeling
- 12×10^6 mesh cells, 1024 cores
- 3 days run time, 2×10^5 time steps
- 3 uncertain parameters (C_s , Pr_t , Sc_t)
- 2nd-order PC, 25 sparse-quad. pts



Main-Effect Sensitivity Indices

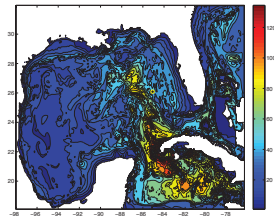
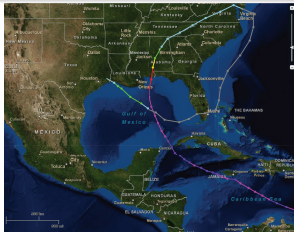


J. Oefelein & G. Lacaze, SNL

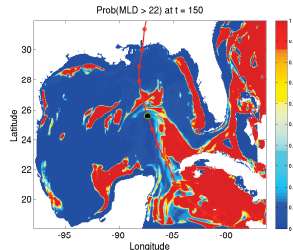
UQ in Ocean Modeling – Gulf of Mexico

A. Alexanderian, J. Winokur, I. Sraj, O.M. Knio, Duke Univ.

A. Srinivasan, M. Iskandarani, Univ. Miami; W.C. Thacker, NOAA



- Hurricane Ivan, Sep. 2004
- HYCOM ocean model (hycom.org)
- Predicted Mixed Layer Depth (MLD)
- Four uncertain parameters, *i.i.d.* U
 - subgrid mixing & wind drag params
- 385 sparse quadrature samples



(Alexanderian *et al.*, Winokur *et al.*, *Comput. Geosci.*, 2012, 2013)

PC and High-Dimensionality

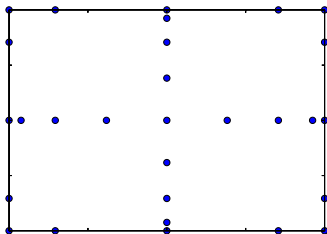
Dimensionality n of the PC basis: $\xi = \{\xi_1, \dots, \xi_n\}$

- $n \approx$ number of uncertain parameters
- $P + 1 = (n + p)!/n!p!$ grows fast with n

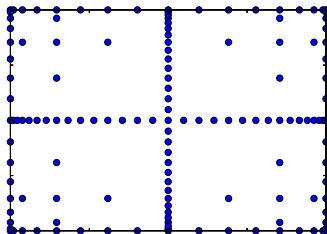
Impacts:

- Size of intrusive PC system
- Hi-D projection integrals $\Rightarrow$ large # non-intrusive samples
 - Sparse quadrature methods

Clenshaw-Curtis sparse grid, Level = 3



Clenshaw-Curtis sparse grid, Level = 5



PC Sparse Quadrature in hiD – Climate land model

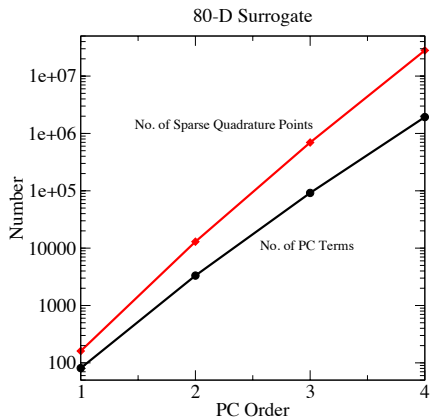
Full quadrature: $N = (N_{\text{ID}})^n$

Sparse Quadrature

- Wide range of methods
- Nested & hierarchical
- Clenshaw-Curtis:
 $N = \mathcal{O}(n^p)$
- Adaptive – greedy algorithms

Number of points can still be excessive in hi-D

- Large no. of terms
- Reduction/sparsity



Dimensionality Reduction

- Sensitivity analysis
 - Local vs. Global
 - Once-at-a-time; Random sampling; Response surface
- Known correlations in the input space
- Choice of basis
 - Karhunen-Loève Expansion
 - Proper Generalized Decomposition
 - Adaptive sparse representations
- Low dimensional manifolds in high dimensional space
 - Intrinsic dimension
 - PCA; Isometric map; Diffusion map
- HDMR methods (ANOVA)
 - cut-HDMR / Anchored-ANOVA
- Sparsification:
 - Compressive Sensing (CS), LASSO
 - Bayesian Compressive Sensing (BCS)

PC coefficients via sparse regression

PCE:

$$y = f(x) \simeq \sum_{k=0}^{K-1} c_k \Psi_k(x)$$

with $x \in \mathbb{R}^n$, Ψ_k max order p , and $K = (p+n)!/p!/n!$

- N samples $(x_1, y_1), \dots, (x_N, y_N)$
- Estimate K terms $c_0, \dots, c_{K-1}$, s.t.

$$\min ||\mathbf{y} - \mathbf{A}\mathbf{c}||_2^2$$

where $\mathbf{y} \in \mathbb{R}^N$, $\mathbf{c} \in \mathbb{R}^K$, $\mathbf{A}_{ik} = \Psi_k(x_i)$, $\mathbf{A} \in \mathbb{R}^{N \times K}$

With $N \ll K \Rightarrow$ under-determined

- Need some form of regularization

Regularization – Compressive Sensing (CS)

- ℓ_2 -norm — Tikhonov regularization; Ridge regression:

$$\min \{ \|\mathbf{y} - \mathbf{A}\mathbf{c}\|_2^2 + \|\mathbf{c}\|_2^2 \}$$

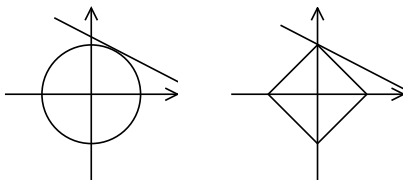
- ℓ_1 -norm — Compressive Sensing; LASSO; basis pursuit

$$\min \{ \|\mathbf{y} - \mathbf{A}\mathbf{c}\|_2^2 + \|\mathbf{c}\|_1 \}$$

$$\min \{ \|\mathbf{y} - \mathbf{A}\mathbf{c}\|_2^2 \} \quad \text{subject to } \|\mathbf{c}\|_1 \leq \epsilon$$

$$\min \{ \|\mathbf{c}\|_1 \} \quad \text{subject to } \|\mathbf{y} - \mathbf{A}\mathbf{c}\|_2^2 \leq \epsilon$$

⇒ discovery of sparse signals



Bayesian Regression

- Bayes formula

$$p(\mathbf{c}|D) \propto p(D|\mathbf{c})\pi(\mathbf{c})$$

- Bayesian regression: prior as a regularizer, *e.g.*
 - Log Likelihood $\Leftrightarrow \|\mathbf{y} - \mathbf{A}\mathbf{c}\|_2^2$
 - Log Prior $\Leftrightarrow \|\mathbf{c}\|_p^p$
- Laplace sparsity priors $\pi(c_k|\alpha) = \frac{1}{2\alpha}e^{-|c_k|/\alpha}$
- LASSO (Tibshirani 1996) ... formally:

$$\min \{ \|\mathbf{y} - \mathbf{A}\mathbf{c}\|_2^2 + \lambda \|\mathbf{c}\|_1 \}$$

Solution $\sim$ the posterior mode of $\mathbf{c}$ in the Bayesian model

$$y \sim \mathcal{N}(\mathbf{A}\mathbf{c}, I_N), \quad c_k \sim \frac{1}{2\alpha}e^{-|c_k|/\alpha}$$

- Bayesian LASSO (Park & Casella 2008)
- Bayesian compressive sensing (Ji 2008)

Bayesian Compressive Sensing (BCS)

- Dimensionality reduction using hierarchical priors

$$p(c_k|\sigma_k^2) = \frac{1}{\sqrt{2\pi}\sigma_k} e^{-\frac{c_k^2}{2\sigma_k^2}} \quad p(\sigma_k^2|\alpha) = \frac{\alpha}{2} e^{-\frac{\alpha\sigma_k^2}{2}}$$

- Effectively, one obtains Laplace *sparsity* prior

$$p(\mathbf{c}|\alpha) = \int \prod_{k=0}^{K-1} p(c_k|\sigma_k^2) p(\sigma_k^2|\alpha) d\sigma_k^2 = \prod_{k=0}^{K-1} \frac{\sqrt{\alpha}}{2} e^{-\sqrt{\alpha}|c_k|}$$

- The parameter α can be further modeled hierarchically, or fixed.
- Evidence maximization dictates values for σ_k^2 , α , σ^2 and allows exact Bayesian solution

$$\mathbf{c} \sim \mathcal{MVN}(\boldsymbol{\mu}, \boldsymbol{\Sigma})$$

with

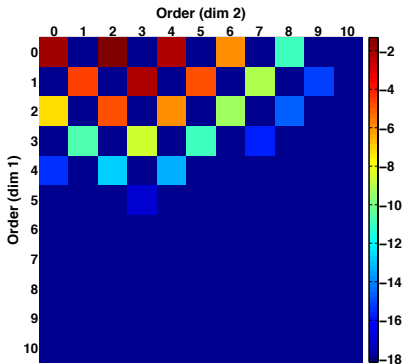
$$\boldsymbol{\mu} = \sigma^{-2} \boldsymbol{\Sigma} \mathbf{P}^T \mathbf{u} \quad \boldsymbol{\Sigma} = \sigma^2 (\mathbf{P}^T \mathbf{P} + \text{diag}(\sigma^2/\sigma_k^2))^{-1}$$

- KEY:** Some $\sigma_k^2 \rightarrow 0$, hence the corresponding basis terms are dropped.

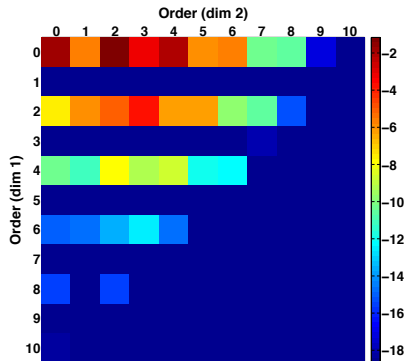
[Tipping, 2001, Ji *et al.*, 2008; Babacan *et al.*, 2010]

BCS removes unnecessary basis terms

$$f(x, y) = \cos(x + 4y)$$



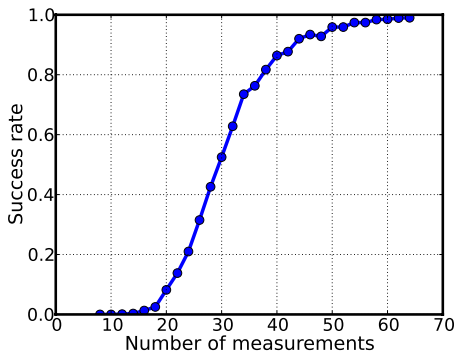
$$f(x, y) = \cos(x^2 + 4y)$$



The square (i, j) represents the (log) spectral coefficient for the basis term $\psi_i(x)\psi_j(y)$.

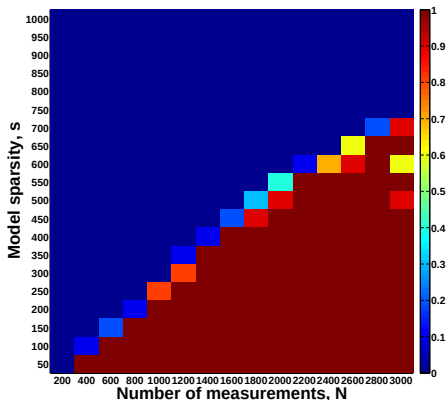
Success rate grows with more data and 'sparser' model

- Test function: $f(\mathbf{x}) = \sum_{k=0}^{K-1} c_k \Psi_k(\mathbf{x})$
- Only S coefficients c_k are non-zero, with $S < N < K$

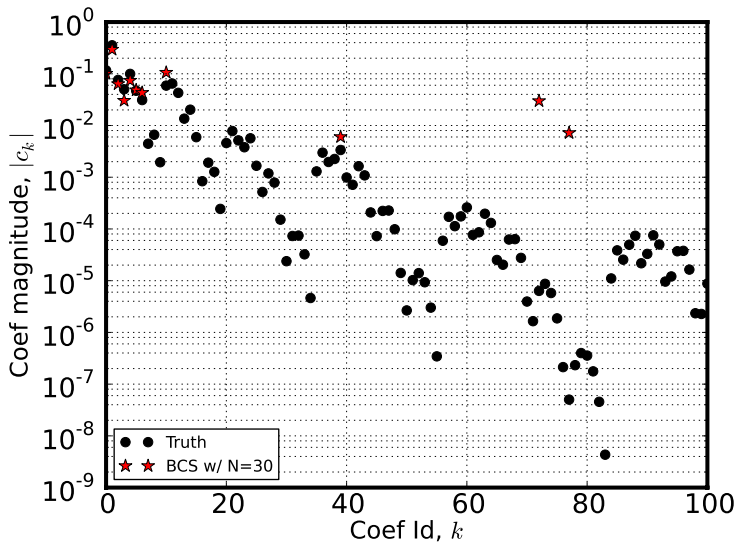


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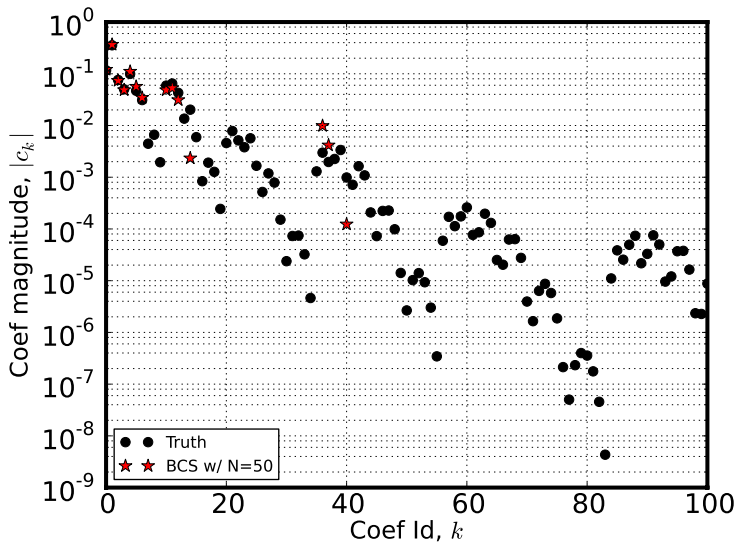
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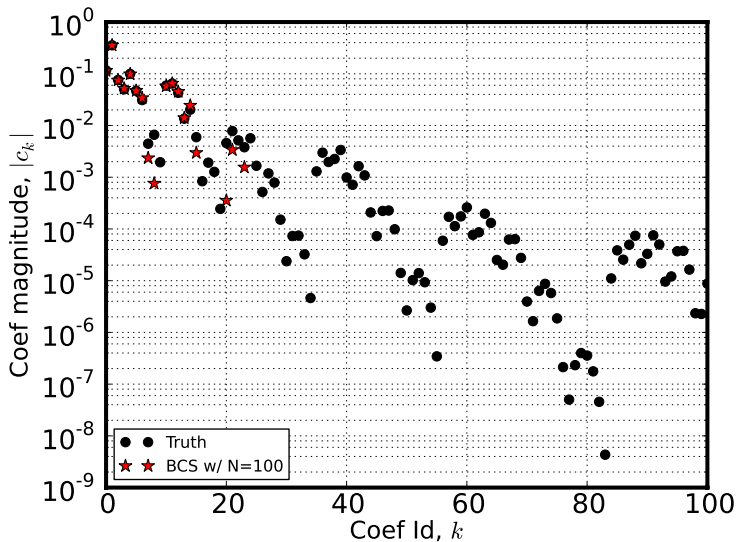
BCS recovers true coefficients with more data



BCS recovers true coefficients with more data



BCS recovers true coefficients with more data



Weighted Bayesian Compressive Sensing

- Dimensionality reduction using hierarchical priors

$$p(c_k|\sigma_k^2) = \frac{1}{\sqrt{2\pi}\sigma_k} e^{-\frac{c_k^2}{2\sigma_k^2}} \quad p(\sigma_k^2|\alpha_k) = \frac{\alpha_k}{2} e^{-\frac{\alpha_k \sigma_k^2}{2}}$$

- Effectively, one obtains Laplace *sparsity* prior

$$p(\mathbf{c}|\alpha_k) = \int \prod_{k=0}^{K-1} p(c_k|\sigma_k^2) p(\sigma_k^2|\alpha_k) d\sigma_k^2 = \prod_{k=0}^{K-1} \frac{\sqrt{\alpha_k}}{2} e^{-\sqrt{\alpha_k}|c_k|}$$

- The parameter α_k can be further modeled hierarchically, or fixed.
- Evidence maximization dictates values for σ_k^2 , α_k , σ^2 and allows exact Bayesian solution

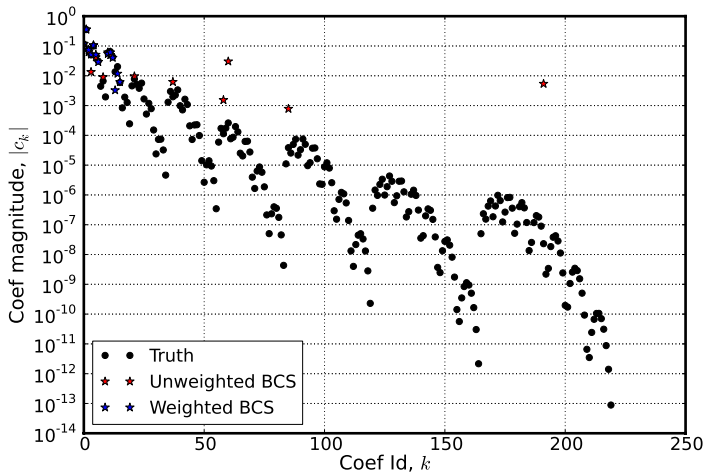
$$\mathbf{c} \sim \mathcal{MVN}(\boldsymbol{\mu}, \boldsymbol{\Sigma})$$

with

$$\boldsymbol{\mu} = \sigma^{-2} \boldsymbol{\Sigma} \mathbf{P}^T \mathbf{u} \quad \boldsymbol{\Sigma} = \sigma^2 (\mathbf{P}^T \mathbf{P} + \text{diag}(\sigma^2/\sigma_k^2))^{-1}$$

- KEY:** Some $\sigma_k^2 \rightarrow 0$, hence the corresponding basis terms are dropped.

WBCS recovers true coefficients better



$$f(\mathbf{x}) = x_0 \cos \left(e + \sum_{i=1}^9 x_i / i \right)$$

Iteratively reweighting Compressive Sensing

[Candes *et al.*, 2007]Sparsest solution: $\min \|c\|_0$ such that $u \approx Pc$ Compressive sensing: $\min \|c\|_1$ such that $u \approx Pc$ Weighted compressive sensing: $\min \|Wc\|_1$ such that $u \approx Pc$ For sparse signals, $u = Pc$, with $\|c\|_0 = S < K$, ideal weights are

$$W = \text{diag} \left(\frac{1}{|c_k|} \right) \quad [\text{i.e., } W_{kk} = +\infty \text{ if } c_k = 0]$$

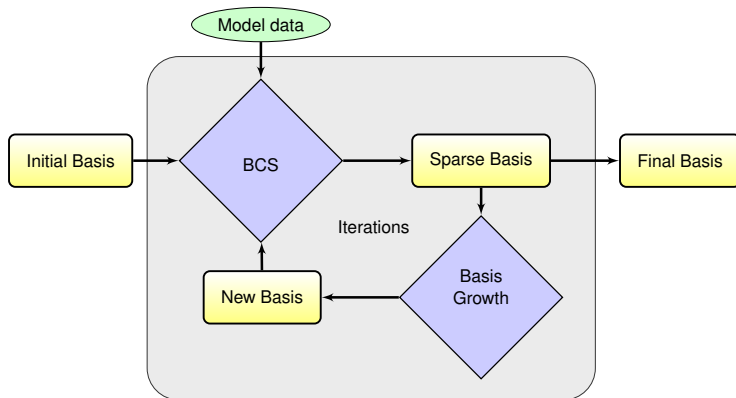
In practice, the true signal coefficients are not known, so ...

 $\Rightarrow$ Iterative re-weighting

$$W^{(i+1)} = \text{diag} \left(\frac{1}{|c_k^{(i)}| + \epsilon} \right) \quad [\epsilon \ll 1 \text{ for stability}]$$

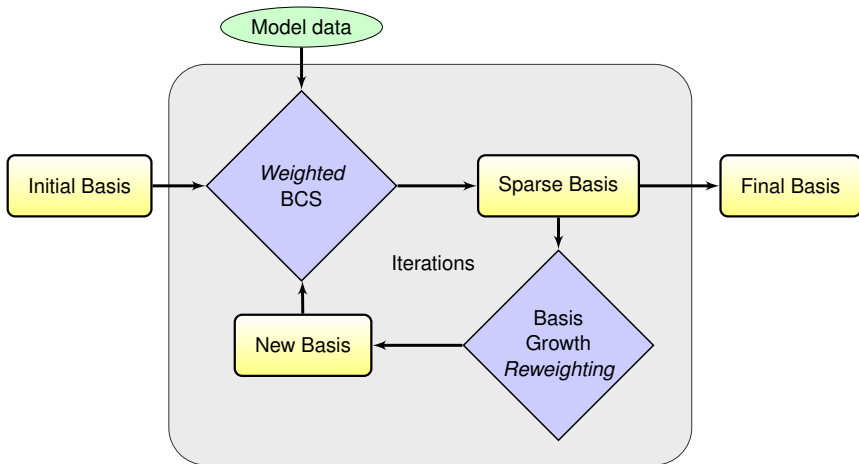
Iterative Bayesian Compressive Sensing (iBCS)

Iterative BCS: Implement an iterative procedure that allows increasing the order for the relevant basis terms while maintaining the dimensionality reduction. [Sargsyan *et al.* 2014]



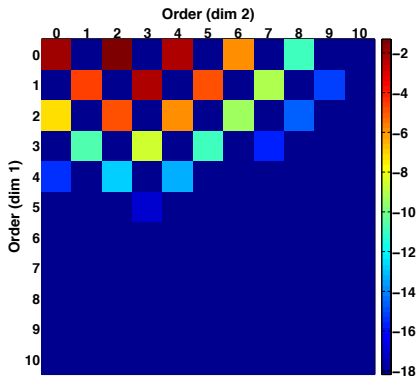
iBCS with reweighting

- Combine basis growth and reweighting!

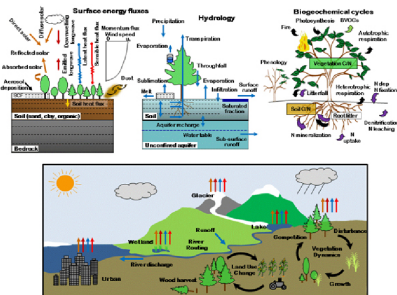


Basis set growth

$$f(x, y) = \cos(x + 4y)$$



Application of Interest: Community Land Model

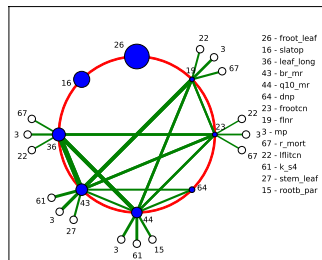
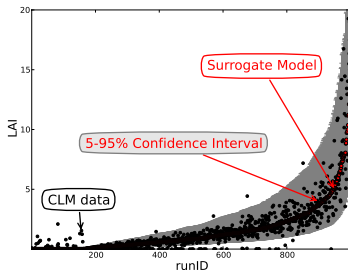


<http://www.cesm.ucar.edu/models/clm/>

- Nested computational grid hierarchy
- A single-site, 1000-yr simulation takes ~ 10 hrs on 1 CPU
- Involves ~ 70 input parameters; some dependent
- Non-smooth input-output relationship

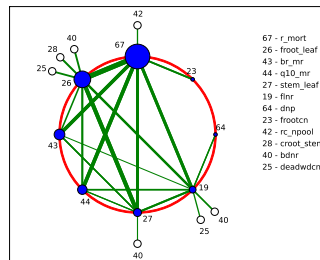
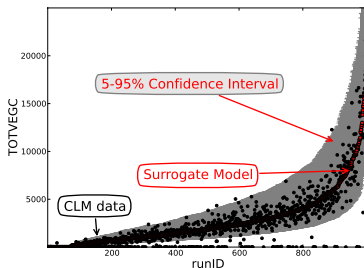
Sparse PC surrogate for the Community Land Model

- Main effect sensitivities : rank input parameters
- Joint sensitivities : most influential input couplings
- About 200 polynomial basis terms in the 70-dimensional space
- Sparse PC will further be used for
 - sampling in a reduced space
 - parameter calibration against experimental data



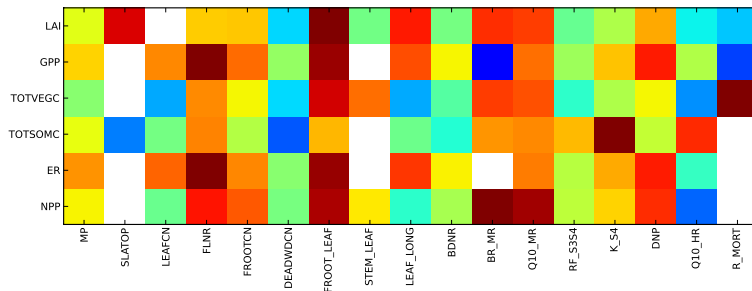
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Closure

- Probabilistic UQ framework
- Hi-dimensionality challenges
- Spectral PC representation of random variables
- Optimal KLE representation of random fields
- Dimensionality reduction – sparsity
- Bayesian compressive sensing