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# *Parameter Estimation with Partial Information*

H. Najm,  
R. Berry, K. Chowdhary, C. Safta,  
K. Sargsyan, & B. Debusschere

**hnnajm@sandia.gov**

Sandia National Laboratories,  
Livermore, CA

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# Outline

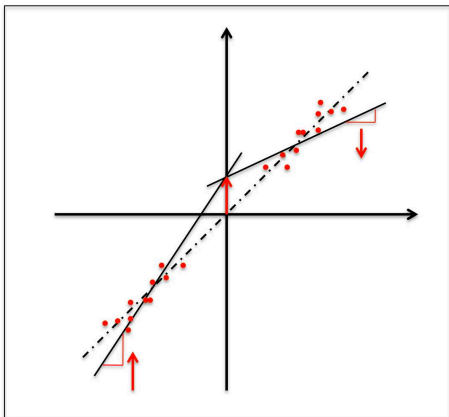
- 1 Introduction
- 2 Inference with Partial Data
- 3 Data Free Inference
- 4 Processed Data Products
- 5 Closure

# Uncertainty in Physical Model Inputs

- Probabilistic UQ requires specification of uncertain inputs
- Require joint PDF on input space
- PDF can be found given data
  - Data is rarely published
- Typically such PDFs are not available from the literature
  - What is typically available are summaries, e.g. nominals and bounds on data, parameters, or fitted model
  - Processed data products
- Uncertainty in computational predictions can depend strongly on detailed structure of the parametric PDF
- Need a procedure to reconstruct a PDF consistent with available information in the absence of the raw data
  - Published summaries
  - Fitting model
  - Other known experimental details

# Remark – Subjective Nature of Uncertainty

- Correlation in uncertain parameters depends strongly on experimental details
- There is no intrinsic objective correlation
- Nominal values are more robust



Data range determines correlation in uncertain slope & intercept, for the same underlying truth

# Bayes formula for Parameter Inference

- Data Model (fit model + noise):  $y = f(\lambda) + \epsilon$
- Bayes Formula:

$$p(\lambda, y) = p(\lambda|y)p(y) = p(y|\lambda)p(\lambda)$$

$$\begin{array}{c}
 \text{Likelihood} \quad \text{Prior} \\
 p(\lambda|y) = \frac{p(y|\lambda) p(\lambda)}{p(y)} \\
 \text{Posterior} \qquad \qquad \qquad \text{Evidence}
 \end{array}$$

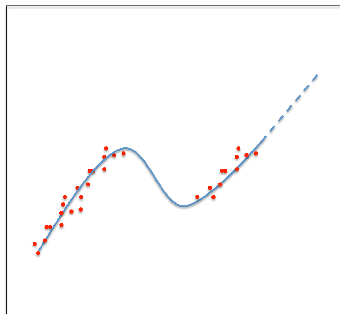
- Prior: knowledge of  $\lambda$  prior to data
- Likelihood: forward model and measurement noise
- Posterior: combines information from prior and data
- Evidence: normalizing constant for present context

# Inference with Partial/Missing Data – Scenarios

- Partial data: some of the data is missing
  - *e.g.* surveys, census, *etc*
  - ⇒ Imputation methods
- All the data is missing but some information is known
  - Moments/statistics, or summaries
    - on missing data
    - on parameters fitted to missing data
    - on outputs of models fitted to missing data
  - Physical constraints
  - Details about the experiment
  - ⇒ Maximum Entropy principle

# Dealing with Missing Data – Imputation

- Use information
  - from existing data
  - from prior knowledge about the model/systemto estimate the missing data
- Interpolation is less risky than extrapolation
  - still risky
- Numerous methods for interpolation and extrapolation
- Assumptions about the structure of the model in the no-data regions



# Inference with Partial Data – Imputation

## Bayesian Multiple Imputation (Rubin, 1987)

- Use the posterior predictive conditioned on *observed* data to generate replicates of the missing data  $z_{\text{mis}}$
- For each full data set,  $(z_{\text{obs}}, z_{\text{mis}})$  apply Bayesian inference to get a posterior density  $p(\beta|z_{\text{obs}}, z_{\text{mis}})$
- Marginalize over the missing data to get the observed data posterior  $p(\beta|z_{\text{obs}})$

$$p(\beta|z_{\text{obs}}) = \int q(\beta|z_{\text{obs}}, z_{\text{mis}})f(z_{\text{mis}}|z_{\text{obs}}) dz_{\text{mis}}$$

# No Data – Inference based on Constraints

Maximum Entropy (MaxEnt) Principle:

Maximize ignorance while satisfying given constraints  $\mathcal{C}$

Given data  $z$ , use Bayes rule to get

$$q(\beta|z) = \frac{q(z|\beta)q(\beta)}{q(z)}$$

When data  $z$  is unavailable, but its distribution  $w(z|\mathcal{C})$  can be estimated, then, given a prior  $q(z, \beta)$ , maximizing relative entropy  $\mathcal{E}(p, q)$  leads to the joint MaxEnt posterior

$$p(z, \beta|\mathcal{C}) = q(\beta|z)w(z|\mathcal{C})$$

Marginalizing over  $z$ :  $p(\beta|\mathcal{C}) = \int q(\beta|z)w(z|\mathcal{C})dz$

# Data Free Inference (DFI)

(Berry *et al.*, JCP 2012)

- **Intuition:** In the absence of data,
  - the structure of the fit model
  - nominals and bounds
  - physical constraints
  - other known experimental detailsimplicitly inform the correlation between the parameters
- **Goal:** Make this information *explicit* in the joint PDF
- **DFI:** discover a consensus joint PDF on the parameters consistent with given information
  - Maximum entropy principle
  - Approximate Bayesian Computation (ABC) methods

# DFI Algorithm in General

Given fitted model  $f(x, \beta)$  and statistics from missing data  $z^*$   
 $z^* \Rightarrow p(\beta|z^*) \Rightarrow \mathcal{S}(G(p(\beta|z^*))) \Rightarrow \mathcal{S}(z^*) := \mathcal{S}^*$

- MCMC chain on data space
  - Each chain step  $\Rightarrow$  data  $z$
  - Likelihood  $\pi_\epsilon(z) = K_\epsilon(|\mathcal{S}(z) - \mathcal{S}^*|)$ 
    - $\epsilon \rightarrow 0 \Rightarrow \{\pi_\epsilon(z) \neq 0 \Leftrightarrow \mathcal{S}(z) \rightarrow \mathcal{S}^*\}$
    - $\mathcal{S}(z)$  evaluated from the inner MCMC chain on the parameter space
- MCMC chain on parameter space
  - Estimate posterior  $p(\beta|z)$
  - Evaluate statistic  $\mathcal{S}(z)$
- Pooling/marginalizing consistent posteriors  $p(\beta|z)$

# Opinion Pooling – Linear

Consider  $M$  experts, each providing an opinion (density)  $p_i(\beta)$ .  
Seek a consensus pooled density  $p_0$

$$p_0 = T(p_1, \dots, p_M)$$

Linear average pooling

$$p_0(\beta) = \int_{\Delta} p(\beta|z)w(z)dz$$

– Marginalization before or after pooling is equivalent

With  $z$  sampled from  $w(z)$ ,

$$p_0(\beta|z_1, \dots, z_M) = \frac{1}{M} \sum_{i=1}^M p(\beta|z_i)$$

# Opinion Pooling – Logarithmic

Logarithmic average pooling

$$\ln p_0(\beta) = C + \int_{\Delta} \ln p(\beta|z)w(z)dz$$

- Externally Bayesian; commutes with Bayesian updating
- Does not satisfy the marginalization property

With  $z$  sampled from the density  $w(z)$

$$p_0(\beta|z_1, \dots, z_M) = \left[ \prod_{i=1}^M p(\beta|z_i) \right]^{1/M}$$

# DFI Demonstration – Plan

Context: Chemical ignition in a homogeneous CH<sub>4</sub>-Air system

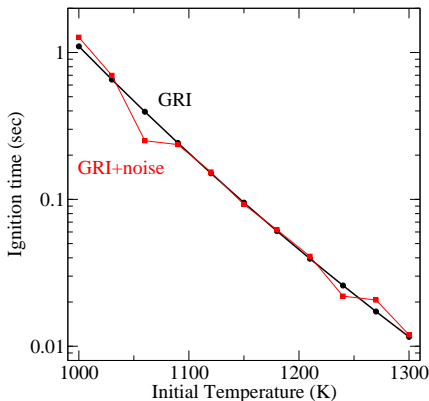
- Use simulated noisy data to infer a joint posterior PDF on parameters of a chemical model
- Discard data, retaining summary information
- Reconstruct the posterior using available information
- Compare the “true” and DFI posteriors

# Generate ignition "data" using a detailed model+noise

- Ignition using a detailed chemical model for methane-air chemistry
- Ignition time versus Initial Temperature
- Multiplicative noise error model
- 11 data points:

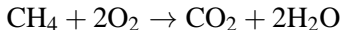
$$\tau_i^d = \tau^{\text{GRI}}(T_i^o) (1 + \sigma \epsilon_i)$$

$$\epsilon \sim N(0, 1)$$



# Fitting with a simple chemical model

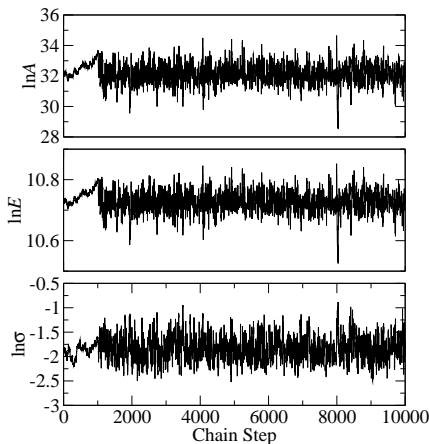
- Fit a global single-step irreversible chemical model



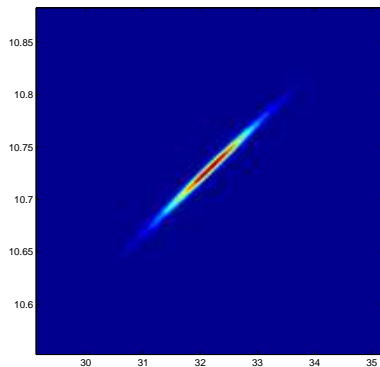
$$\mathfrak{R} = [\text{CH}_4][\text{O}_2]k_f$$

$$k_f = A \exp(-E/R^oT)$$

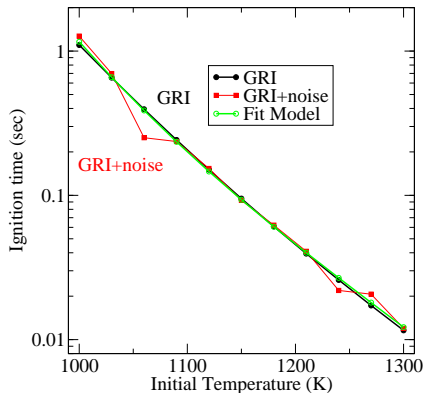
- Infer 3-D parameter vector  $(\ln A, \ln E, \ln \sigma)$
- Good mixing with adaptive MCMC when start at MLE



# Bayesian Inference Posterior and Nominal Prediction



Marginal joint posterior on  $(\ln A, \ln E)$  exhibits strong correlation



Nominal fit model is consistent with the true model

# Data Free Inference Challenge

Discarding initial data, reconstruct marginal  $(\ln A, \ln E)$  posterior using the following information

- Form of fit model
- Range of initial temperature
- Nominal fit parameter values of  $\ln A$  and  $\ln E$
- Marginal 5% and 95% quantiles on  $\ln A$  and  $\ln E$

Further, for now, presume

- Multiplicative Gaussian errors
- $N = 8$  data points

# DFI Algorithm Structure

- Explore the space of hypothetical data sets
    - MCMC chain on the data
    - Each state defines a data set
  - For each data set:
    - MCMC chain on the parameters
    - Evaluate statistics on resulting posterior
    - Accept data set if posterior is consistent with given information
  - Evaluate pooled posterior from all acceptable posteriors
- Logarithmic pooling:

$$p(\lambda|y) = \left[ \prod_{i=1}^K p(\lambda|y_i) \right]^{1/K}$$

# Computational issues

- Outer Chain:
  - $N$  data points  $\Rightarrow$  chain is  $(2N + 1)$ -dimensional
    - Single-site update – easy to tune in hi-D
  - Markov chain on a manifold in data space
    - Ideally, use geometric MCMC methods
    - No sense of good mixing – no peak
    - Pursue convergence with # of samples
- Inner Chain:
  - Highly peaked posterior ridge
    - Adaptive MCMC for parameter estimation
    - Start at nominal model
    - Ideal proposal distrib. uses Hessian of Likelihood
- Computationally challenging:
  - Use a Polynomial Chaos surrogate for forward model
  - Run multiple outer chains in parallel, and aggregate resulting acceptable data sets

# Data Chain Likelihood Function

- With  $\beta_0$  being the given nominal parameter vector

$$L(z, \ln \sigma; \beta_0) = F_\delta(z) \frac{p(z, \ln \sigma | \beta_0)}{\max_{z, \ln \sigma} [p(z, \ln \sigma | \beta_0)]}$$

- Consistency check: Nominals

$$p(z, \ln \sigma | \beta_0) \propto p(z | \beta_0, \ln \sigma) \pi(\beta_0, \ln \sigma)$$

– parameter likelihood  $\times$  prior; inner chain

- Consistency check: Bounds

- Evaluate the marginal parameter posterior  $p(\beta|z)$  quantiles
- $F_\delta(z)$  is a Kullback-Leibler density that is maximal when the posterior quantiles are consistent with the given 5%, 95% quantiles on  $\beta$

# Bounds Check employing the Kullback-Leibler Density

- $[p, r, q]$ : trinomial density with probabilities  $p + r + q = 1$
- KL density measuring mismatch between two trinomials:

$$f_{\delta}([p, r, q] \mid [0.05, 0.90, 0.05]) = \exp \left\{ -\delta \left( p \ln \frac{p}{0.05} + r \ln \frac{r}{0.90} + q \ln \frac{q}{0.05} \right) \right\}$$

Then, with the measured  $p(\beta|z)$  quantiles

$$p_1(z) = P[\ln A < (\ln A)_{0.05} \mid z] \quad q_1(z) = P[\ln A > (\ln A)_{0.95} \mid z]$$

$$p_2(z) = P[\ln E < (\ln E)_{0.05} \mid z] \quad q_2(z) = P[\ln E > (\ln E)_{0.95} \mid z]$$

we define

$$F_{\delta}(z) \propto \prod_{i=1}^2 f_{\delta} \left( [p_i(z), 1 - p_i(z) - q_i(z), q_i(z)] \mid [0.05, 0.90, 0.05] \right)$$

# Inner/Parameter Chain Likelihood

Likelihood function employs, by construction, an *i.i.d.* multiplicative gaussian noise model for the dependent data  $\tau$

$$\frac{\tau}{\tau^m(T^o, \beta)} - 1 \sim N(0, \sigma^2)$$

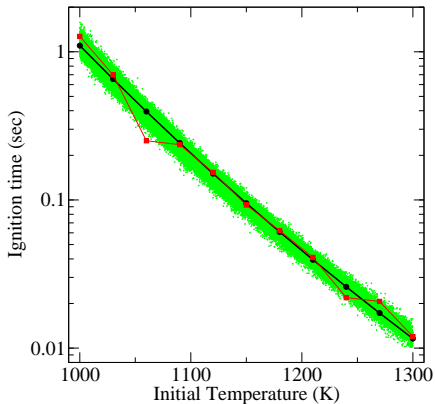
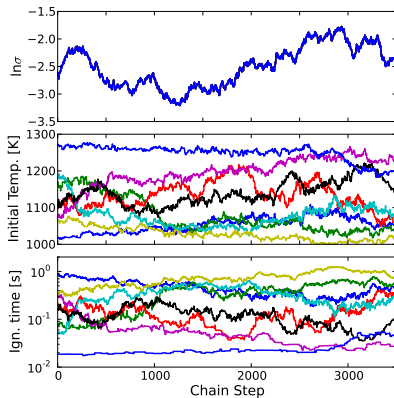
and an *i.i.d.* uniform model for the independent data  $T^o$ .

$$T^o \sim U(T_{\min}^o, T_{\max}^o)$$

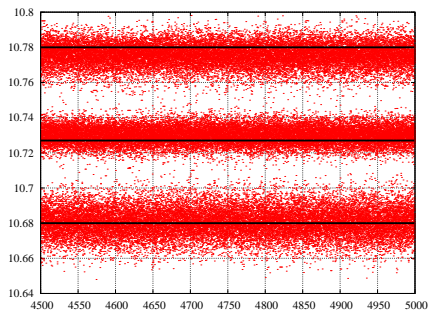
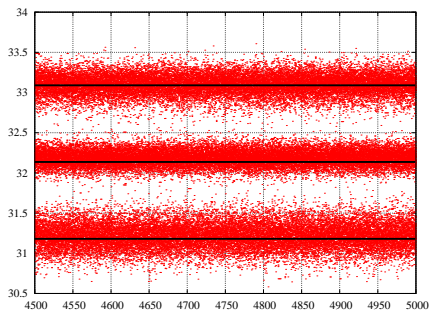
Giving the likelihood for  $z = (\tau_1^d, \dots, \tau_N^d, T_1^o, \dots, T_N^o)$ :

$$p(z | \beta, \ln \sigma) \propto \frac{1}{(2\pi\sigma^2)^{N/2}} \frac{1}{\prod_{i=1}^N \tau^m(T_i^o, \beta)} \exp \left\{ -\frac{1}{2\sigma^2} \sum_{i=1}^N \left[ \frac{\tau_i^d}{\tau^m(T_i^o, \beta)} - 1 \right]^2 \right\}$$

## Short sample from outer/data chain



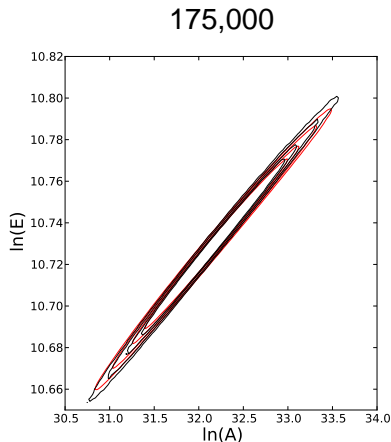
# Scatter of Nominals and 0.05/0.95 Quantiles



- Parameter posterior statistics are in general agreement with the desired nominal/quantile summaries
- Amplitude of scatter can be controlled by choice of weights in the likelihood function of the data chain

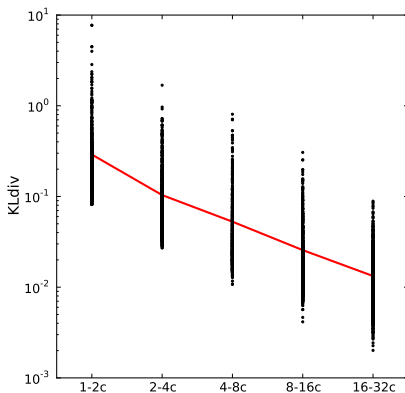
# Pooled DFI Posterior vs. Reference Posterior

- 50 data chains run in parallel  
⇒ Pooled DFI posterior
- Structure of 2D pooled posterior is robust
  - Irrespective of the number of chains
- Very close similarity to the reference posterior
- Available information is essentially a sufficient statistic



# Empirical Convergence of 3D Pooled Posteriors

- Kullback-Leibler divergence between posteriors of successively increased data volumes
- Combinatorial choices of chains pooled at each stage
  - statistical scatter of KLdiv
- Overall convergence evident  $\propto 1/N$



# Processed Data Products

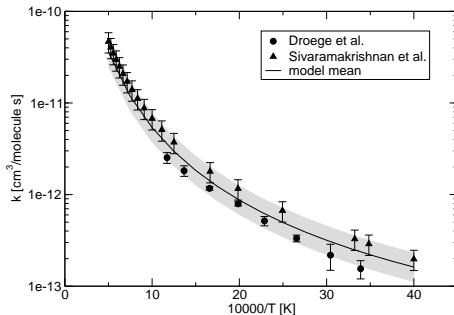
Frequently, we are given published “data” in the form of nominal values and error bars

- Fitting of  $y = f(x; \beta)$

- Data:

$$(x_\ell, y_\ell^o \pm \alpha_\ell), \quad \ell = 1, \dots, L$$

$\pm \alpha_\ell$  are e.g.  $\pm 2\sigma$  error bars



Two interpretations of the  $(y_\ell^o, \alpha_\ell)$  as nominals and bounds

- on the **data**:  $N$  observed random  $y_\ell$  at each  $x_\ell$
- on the **fitted model**: uncertain  $y_\ell = f(x_\ell, \beta)$

# Processed Data Products – Demo

Demonstrate DFI on a simple model problem:

- Generate synthetic noisy data from a double exponential decay model

$$s = 2e^{-2t} + .5e^{-.5t}$$
$$y = s + (\sigma_1 s + \sigma_2)\eta$$

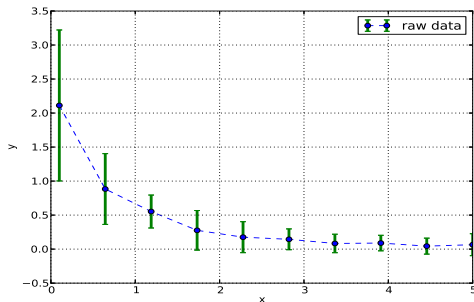
Retain only error bars at different times

- Fit using DFI with a single exponential decay model

$$f = a \exp(-kt)$$
$$y = f + (\sigma_1 f + \sigma_2)\eta$$

# DFI with Processed Data Products, Case I

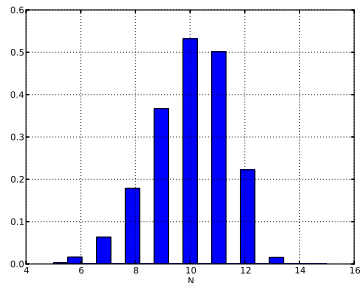
- Given means and variances on the raw data.



- Using MaxEnt:  $y_\ell \sim N(y_\ell^o, \sigma_\ell^2)$ ,  $\ell = 1, \dots, L$
- Generate consistent data samples

# Unknown Dimensionality: Marginalizing over $N$

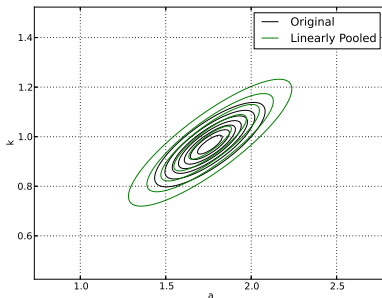
- Generate consistent data sets using MaxEnt
- The number of samples  $N$  is unknown
- Place a density on the number of samples  $N$ , and sample it along with data  $z$



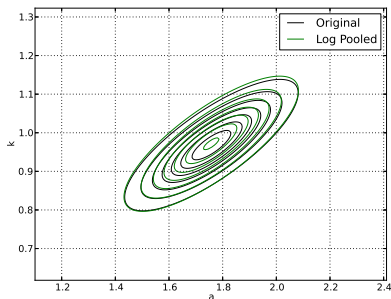
Presumed Beta  
density on  $N$

# Linear and Logarithmic Pooling give Similar Results

## Linear



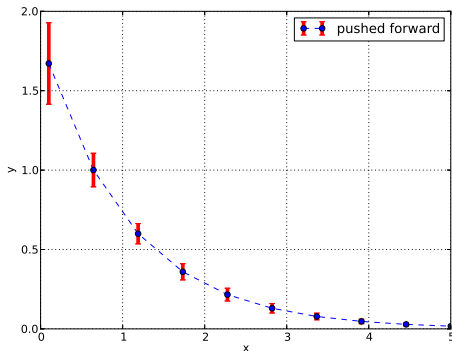
## Logarithmic



- Pooled posteriors on  $(a, k)$ , marginalized over error model hyperparameters  $\sigma_1, \sigma_2$  and  $N$
- Posteriors from individual data sets are largely similar to the reference posterior, hence the similar pooling results

# DFI with Processed Data Products, Case II

- Given means & variances on the uncertain fitted model output



- Consider these as statistics on the pushed forward posterior
- Use these statistics to constrain proposed data sets

# DFI Implementation

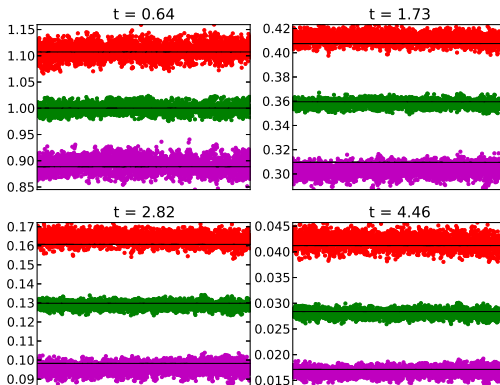
Given required statistics  $\mathcal{S}^*$

Missing data  $z^*$

$$z^* \Rightarrow p(\beta|z^*) \Rightarrow p(f(x, \beta)|z^*) \Rightarrow \mathcal{S}(z^*) := \mathcal{S}^*$$

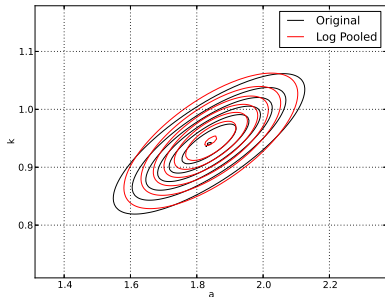
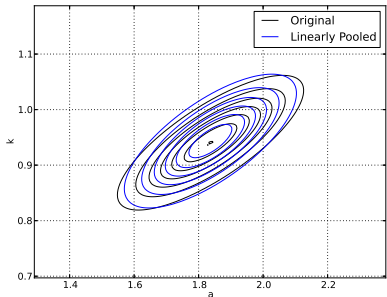
- MCMC chain on data space
  - Each chain step  $\Rightarrow$  data  $z$
  - Likelihood  $\pi_\delta(z) = f_\delta(\mathcal{S}(z) - \mathcal{S}^*)$ 
    - $\mathcal{S}(z)$  evaluated from the inner MCMC chain on the parameter space
- MCMC chain on parameter space
  - Estimate posterior  $p(\beta|z)$
  - Estimate pushed forward posterior  $p(f(x, \beta)|z)$
  - Evaluate statistic  $\mathcal{S}(z)$
- Pooling

# PFP Quantiles



- Consistent pushed forward posterior quantiles clustered in the vicinity of reference values

# Linear and Logarithmic Pooling give Similar Results



- Posteriors from individual data sets are largely similar to the reference posterior, hence the similar pooling results

# Closure

Outlined several missing data contexts

- Partial data – Imputation
- No data
  - Maximum Entropy principle
  - Approximate Bayesian Computation
  - DFI
- Demonstrations
  - Constraints on fitted parameters of interest
  - Processed data products
    - Constraints on raw data
    - Constraints on fitted model uncertain predictions